

# Exploring Best Practices in the Design of Experiments

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**University of Minnesota**  
**Carlson School of Management**  
**and**  
**School of Statistics**



# Joint work with Brad Jones

## SAS/JMP



# Overview

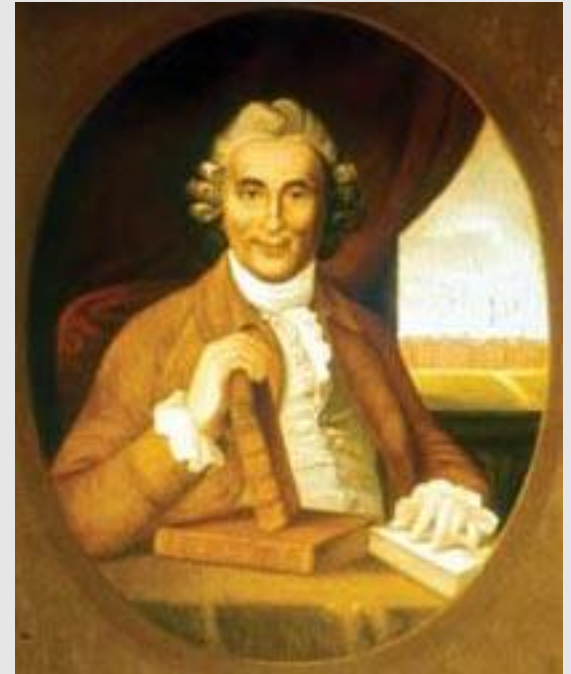
- Where have we been?
- Where are we now?
- Emerging paradigm: problem-driven design
- Six Examples
- What's driving the software?
- Definitive screening and an example
- The future?

# Where have we been?



# Where have we been?

- James Lind, 1753: “A Treatise on Scurvy”
- First (published) one-way layout



## From “A Treatise...”

“On the 20th May, 1747, I took twelve patients in the scurvy on board the Salisbury at sea. Their cases were as similar as I could have them. They all in general had putrid gums, the spots and lassitude, with weakness of their knees.

...

- Two of these were ordered each a quart of cyder a day...
- Two others took twenty five gutts of elixir vitriol three times a day upon an empty stomach...
- Two others took two spoonfuls of vinegar three times a day upon an empty stomach...
- Two of the worst patients, with the tendons in the ham rigid (a symptom none the rest had) were put under a course of sea water...
- Two others had each two oranges and one lemon given them every day...
- The two remaining patients took the bigness of a nutmeg three times a day...

The consequence was that the most sudden and visible good effects were perceived from the use of the oranges and lemons”

# Where have we been?

- Gergonne: 1815
- Designs for polynomial regression, response surface design



- S. C. Peirce: 1870s
- Randomization



- K. Smith, 1918: *Biometrika*, Optimal design for polynomial regression

# R. A. Fisher put it all together

Fisher, 1920s:

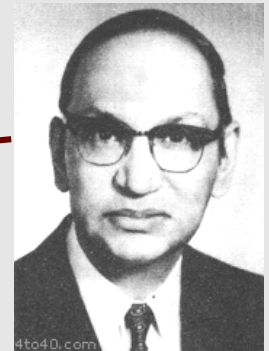
- Randomization as mathematical basis for analysis
- Local control and blocking
- Replication
- Factorial designs
- Split plot designs
- Confounding
- ANOVA
- F, t distributions
- Etc., etc.





# 1920's-1950's: Orthogonality is the driving principle

- Fisher, Yates: need for ease of computation, independence of effects
- R. C. Bose, C.R. Rao, and Indian School: Combinatorics, BIBDs, PBIBDs
- Finney, 1945: Fractional replication
- Plackett and Burman, 1946



# ...culminating in the $2^{k-p}$ System

VOL. 3, No. 3

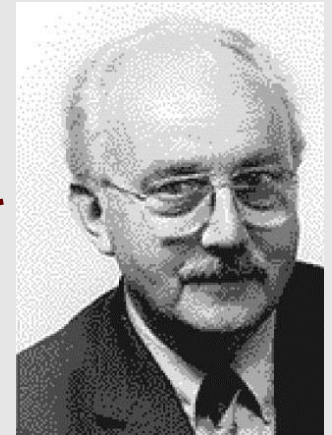
TECHNOMETRICS

AUGUST, 1961


## The $2^{k-p}$ Fractional Factorial Designs\* Part I.

G. E. P. BOX AND J. S. HUNTER

*Statistics Department, University of Wisconsin and Mathematics Research Center,  
University of Wisconsin*



# 1950s: Baby steps away from orthogonality




Journal of the Royal Statistical Society  
SERIES B (METHODOLOGICAL)  
Vol. XIII, No. 1, 1951

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ON THE EXPERIMENTAL ATTAINMENT OF OPTIMUM CONDITIONS


By G. E. P. BOX and K. B. WILSON  
*Imperial Chemical Industries, Dyestuffs Division Headq  
Blackley, Manchester*

A BASIS FOR THE SELECTION OF A RESPONSE  
SURFACE DESIGN\*



G. E. P. Box  
*Princeton University*

AND



NORMAN R. DRAPER  
*University of North Carolina and Imperial Chemical Industries*

Also Box and Lucas, 1959, Nonlinear Design



# Optimal Design—Pioneers

- Kiefer and Wolfowitz (1959)

- Fedorov (1969)

- Wynn (1969)



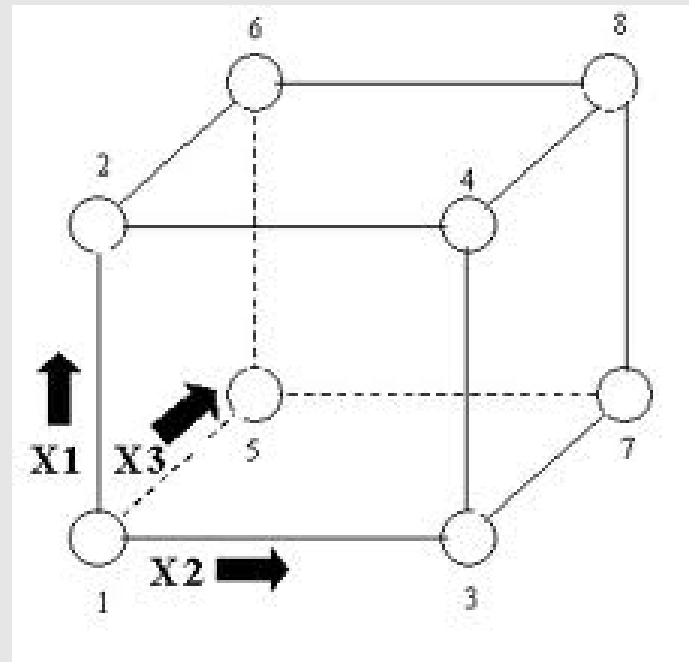
# Shift to problem-driven design (70s, 80s 90s, 00s)

- Toby Mitchell---Detmax
- Constrained mixture designs—John Cornell, Ron Snee, Greg Piepel
- Design augmentation
- Optimal blocking (Cook, Nachtsheim '87, Donev, Atkinson '87)
- Model robust design (Lauter, 1974)
- Bayesian design (Chaloner, '80, DuMouchel and Jones, '94)
- Optimal split-plot designs, Goos, '2002

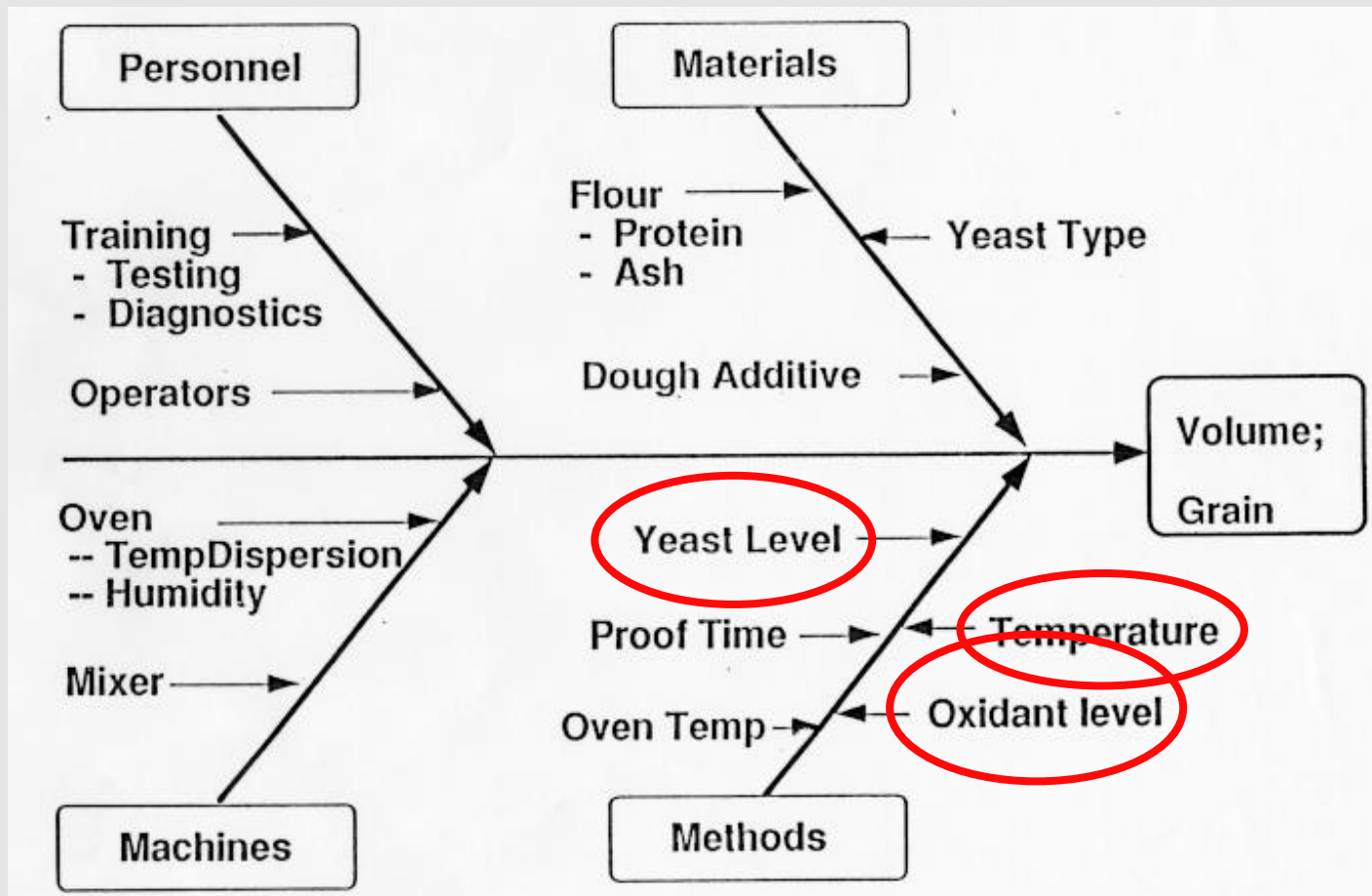


# Where are we now?

# What is a Factorial Experiment?



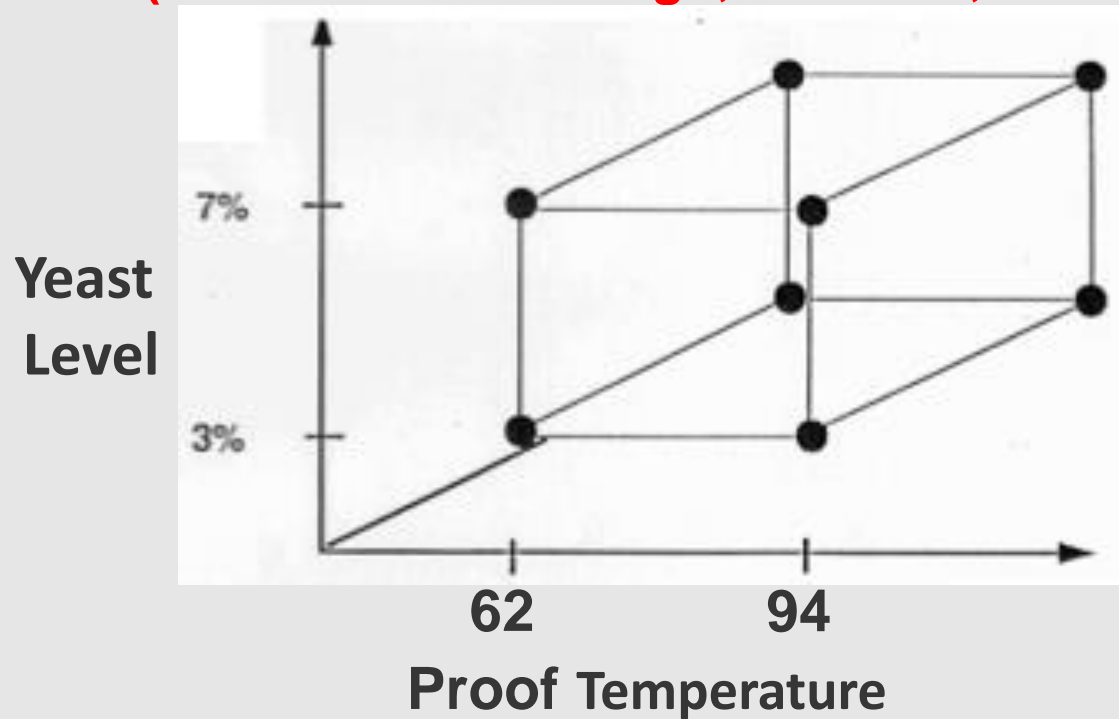
# Bread Baking Optimization: Which Factors?



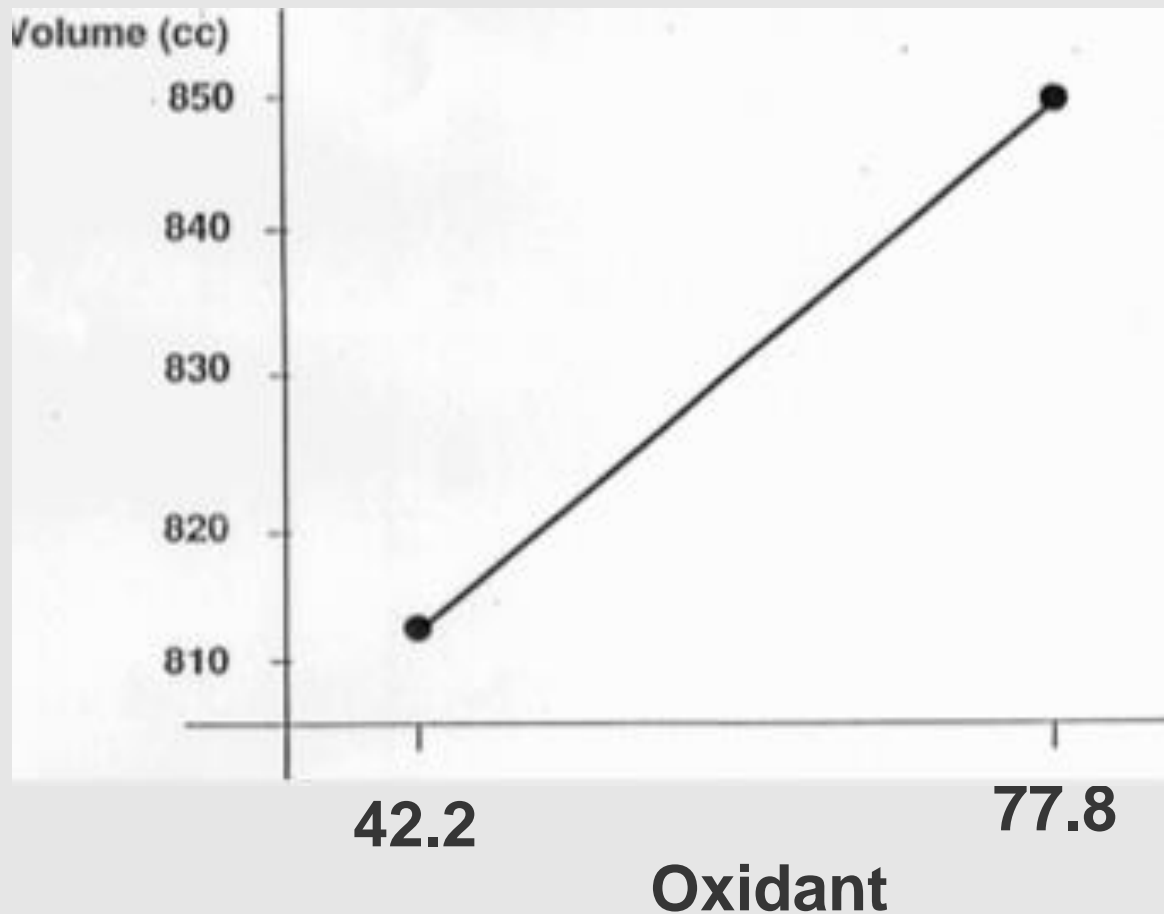


# Factorial Design Approach

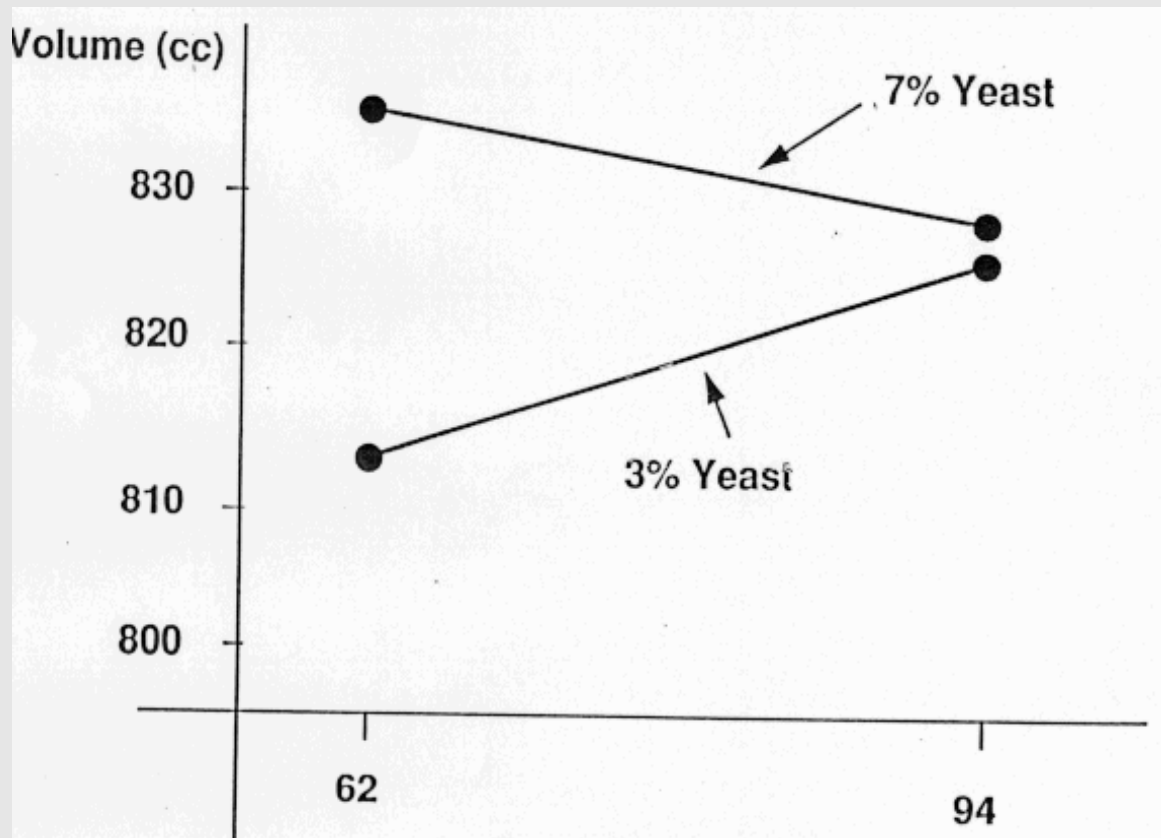
**One Test at Each Corner (n=8)  
(2x2x2 factorial design, 3 factors, 2 levels)**



# Analysis: Main Effects Plot



## Bonus: We can identify interactions



Proof Temperature

# JMP Demo: Bread Baking Experiment

# Paradigms for DOE

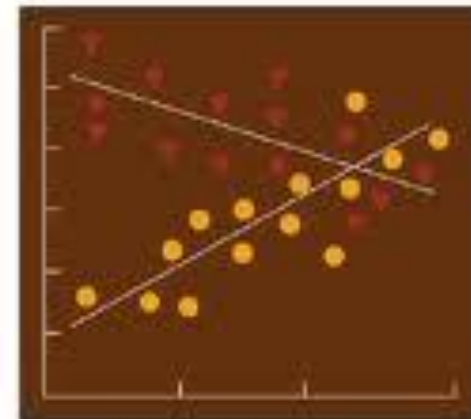
## Classical:

- Fisher paradigm
- Textbook paradigm
- duPont paradigm

## Optimal:

- Problem-driven paradigm

APPLIED LINEAR  
STATISTICAL MODELS  
FIFTH EDITION



Kutner  
Nachtsheim  
Neter

# The Classical Approach to DOE

1. State the objectives and your constraints
2. Find a design from a book or a table or a software program that most closely matches your objectives (requires DOE expertise)
3. Implement the design
4. Analyze results

# Classical design paradigms pros, cons

## Pros

1. Designs are tried & true
2. Easy analysis

# Classical design paradigms pros, cons

## Pros

1. Designs are tried & true
2. Easy analysis

## Cons

1. Requires much expertise
2. Rigidity: designs often do not match the problem



# What DOE Expertise is Needed?

- **Full factorial designs**

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- **Interactions**

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- $2^{k-p}$  fractional factorial designs

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- Interactions
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- $2^{k-p}$  fractional factorial designs
- Confounding schemes and aliasing
- **Resolution**

# What DOE Expertise is Needed?

- Full factorial designs
- Interactions
- $2^k$  factorial designs
- $2^{k-p}$  fractional factorial designs
- Confounding schemes and aliasing
- Resolution
- Screening designs (Plackett Burman, Res III FF)

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**OK, What else?**



# What DOE Expertise (continued)?

- **Response surface experiments**

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- **Blocked experiments**

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- Response surface experiments
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- **Nested designs**

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- **Repeated measures and split plot designs**

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- Repeated measures and split plot designs
- **Incomplete block designs**

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- Incomplete block designs
- Mixture experiments and mixture models

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- **Supersaturated designs**

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- Response surface experiments
- Blocked experiments
- Nested designs
- Repeated measures and split plot designs
- Incomplete block designs
- Mixture experiments and mixture models
- Supersaturated designs





# When might classical designs fail to match problem?

- **Multi-level designs: some factors at 2 levels, some at 3, some at four**

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- **Some factors are qualitative, some not**

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- **My block sizes cannot be the same.**

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- **I want to augment an existing design.**

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- Some factors are qualitative, some not
- I have a mixture experiment---factor levels are percentages that must add to one
- My block sizes cannot be the same.
- I want to augment an existing design.
- **I have a special model, etc., etc., etc.**

# New Paradigm: Problem-driven design



# Problem-Driven Paradigm

- With good software, approach easy to use and to teach.
- One, unified approach to (nearly) all design problems:
  1. Describe your responses
  2. Describe your factors
  3. Describe your objectives (model terms)
  4. Describe your constraints and budget ( $n$ )
  5. Press button to create the design
  6. Check diagnostics, do sensitivity analysis
  7. Manually alter as with classical design

# Problem-driven paradigm pros, cons:

## Pros

1. Design tailored to problem
2. **One unified approach to nearly all problems**

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**Fading fast**



# Six Illustrations with JMP

# Example 1: Extracting food solids from peanuts in solution, n = 16

	Factor	Low	High
1	Water pH level	6.95	8.0
2	Water temp	20C	60C
3	Extraction time	15	40
4	Water-Peanuts Ratio	5	9
5	Agitation speed	5,000	10,000
6	Hydrolyzed?	N	Y
7	Presoaking?	N	Y

# Classical Approach: $2^{7-3}$ Resolution IV FF

## Problem:

- Seven two-level factors in 16 runs.
- Main effects are critical
- Two-factor interactions might be present, might not.
- Standard approach: Resolution IV fractional factorial design

**Well OK, but how did I happen to know this?**

# Knowledge Requirements:

- Main effects, interactions (model terms)
- Confounding schemes, aliasing
- Resolution
- Fractional factorials
- Plackett-Burman designs
- Blocking (maybe)



# Contrast: Custom Design Approach

## Knowledge Requirements:

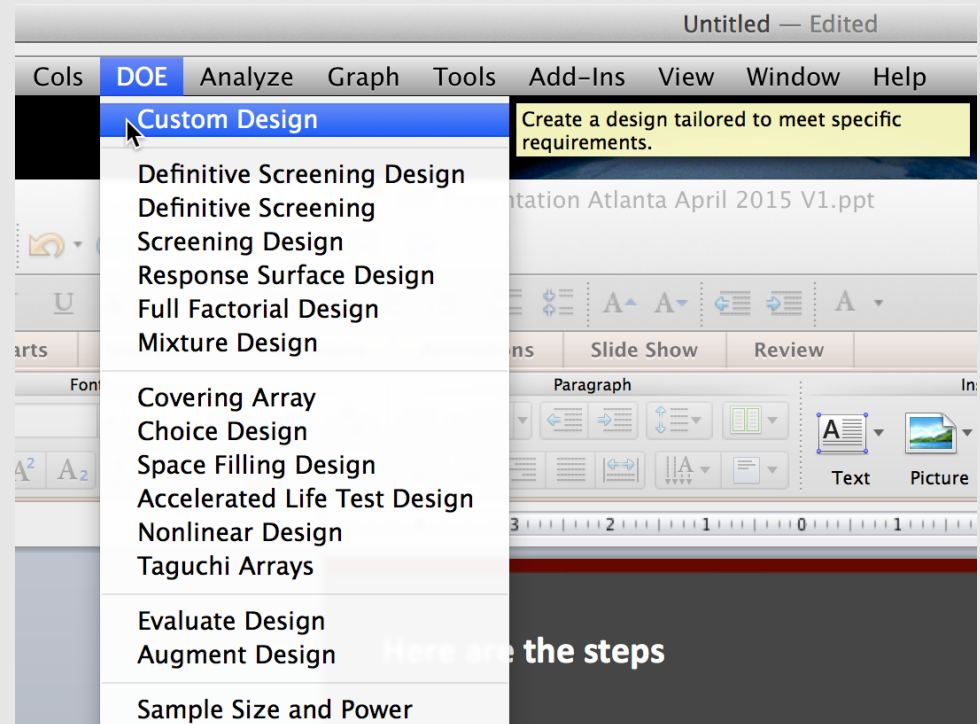
1. Main effects, interactions (model terms)
2. Knowledge of where the “button” is.

## Here's how it's done:

- Specify the main effects model
- Give the sample size as usual
- Push the button

# Here are the steps (Step 1)

## 1. DOE >> Custom Design



# Here are the steps (Steps 2 and 3)

1. DOE >> Custom Design
2. Give the response
3. Give the factors levels (or ranges)

▼ Custom Design

▼ Responses

Add Response ▼ Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Peanut Solids <i>optional item</i>	Maximize	.	.	.

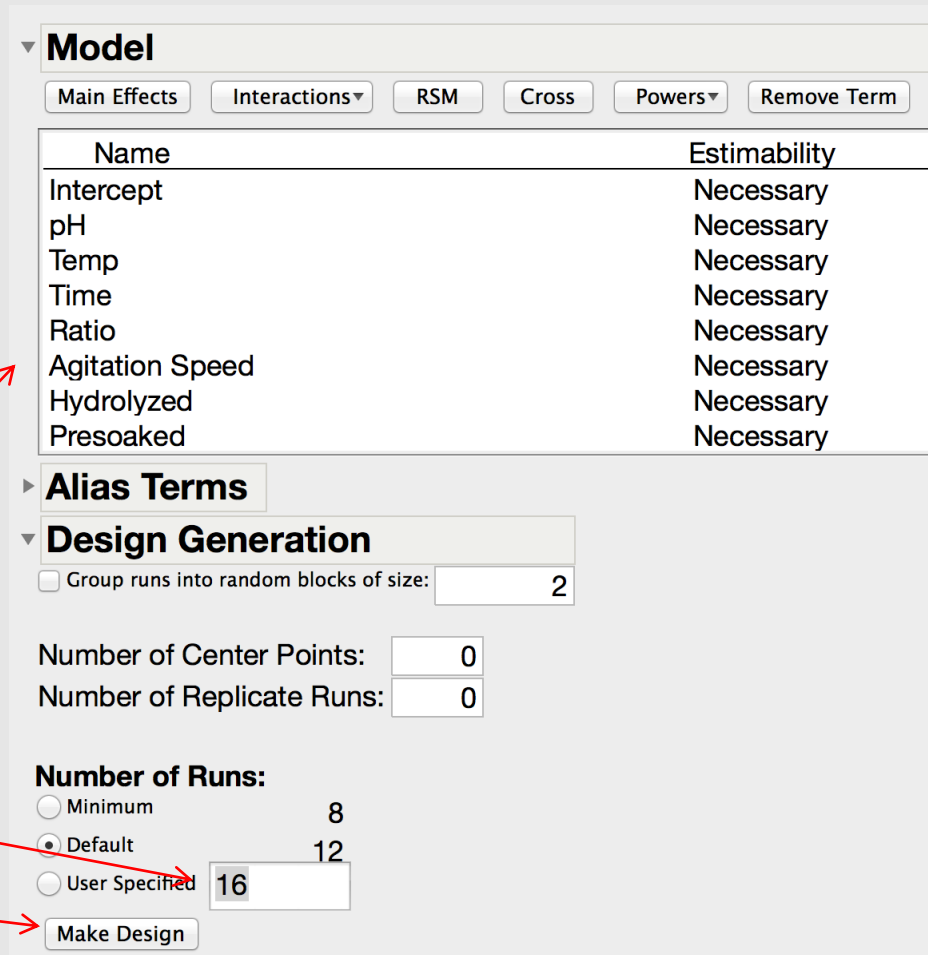
▼ Factors

Add Factor ▼ Remove Add N Factors 1

Name	Role	Changes	Values	
▲ pH	Continuous	Easy	6.95	8
▲ Temp	Continuous	Easy	20	60
▲ Time	Continuous	Easy	15	40
▲ Ratio	Continuous	Easy	5	9
▲ Agitation Speed	Continuous	Easy	5000	10000
▼ Hydrolyzed	Categorical	Easy	No	Yes
▼ Presoaked	Categorical	Easy	No	Yes

# Here are the steps (Steps 4, 5 and 6)

1. DOE >> Custom Design
2. Give the response
3. Give the factors and levels
4. JMP specifies the model
5. Give sample size
6. Press the button



The screenshot shows the JMP Model Builder interface. Red arrows from the numbered steps point to the following elements:

- Step 4 points to the **Model** section.
- Step 5 points to the **Number of Runs** section.
- Step 6 points to the **Make Design** button.

**Model**

Main Effects Interactions RSM Cross Powers Remove Term

Name	Estimability
Intercept	Necessary
pH	Necessary
Temp	Necessary
Time	Necessary
Ratio	Necessary
Agitation Speed	Necessary
Hydrolyzed	Necessary
Presoaked	Necessary

**Alias Terms**

**Design Generation**

☐ Group runs into random blocks of size: 2

Number of Center Points: 0

Number of Replicate Runs: 0

**Number of Runs:**

☐ Minimum 8

☒ Default 12

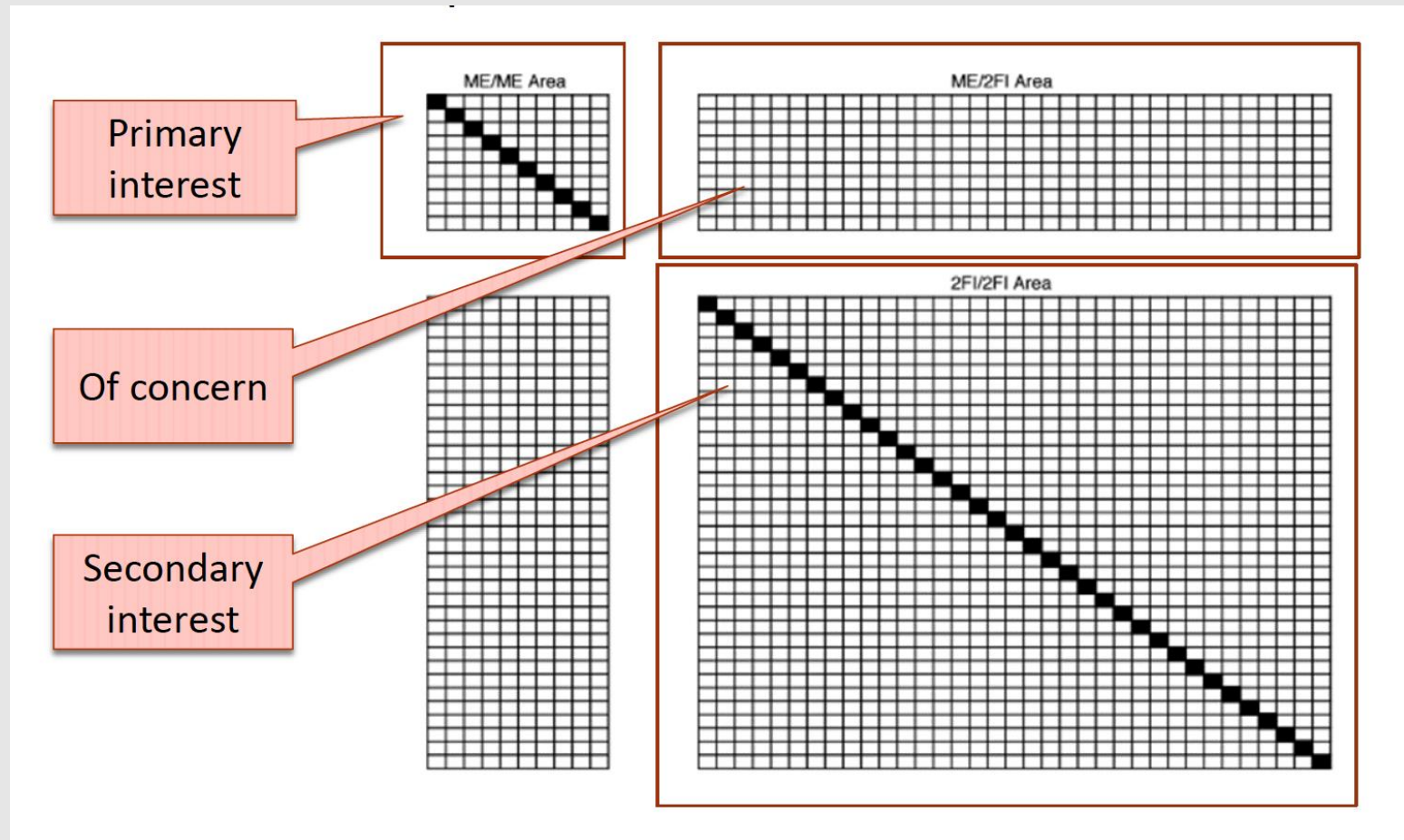
☐ User Specified 16

Make Design

# The result: Orthogonal, D-optimal 16-Run Design

	pH	Temp	Time	Ratio	Agitation Speed	Hydrolyzed	Presoaked	Peanut Solids
1	6.95	60	15	5	5000	Yes	No	•
2	6.95	60	40	5	10000	No	No	•
3	8	20	15	5	5000	Yes	Yes	•
4	6.95	20	40	5	10000	Yes	No	•
5	6.95	60	40	9	5000	No	Yes	•
6	6.95	20	40	9	5000	Yes	Yes	•
7	6.95	20	15	9	10000	No	Yes	•
8	8	60	40	9	5000	Yes	No	•
9	8	60	15	5	10000	Yes	Yes	•
10	8	20	15	9	5000	No	No	•
11	8	20	40	5	10000	No	Yes	•
12	8	60	15	9	10000	No	No	•
13	8	60	40	9	10000	Yes	Yes	•
14	8	20	40	5	5000	No	No	•
15	6.95	20	15	9	10000	Yes	No	•
16	6.95	60	15	5	5000	No	Yes	•

# One way to evaluate the design: Correlation Cell Plot

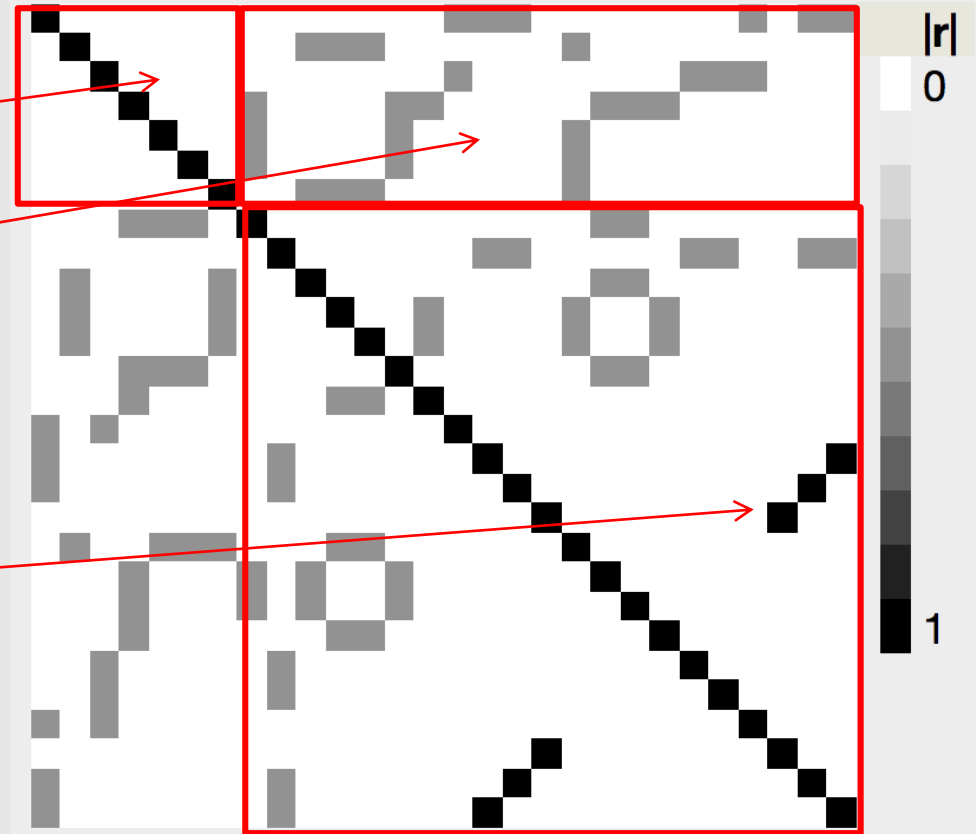


# Our Design

1. Primary interest (Yay!)

2. Of concern  
(Some correlations)

1. Secondary interest  
(Some two-factor  
interactions directly  
confounded, Boo!)





# Can we ask more of our design?

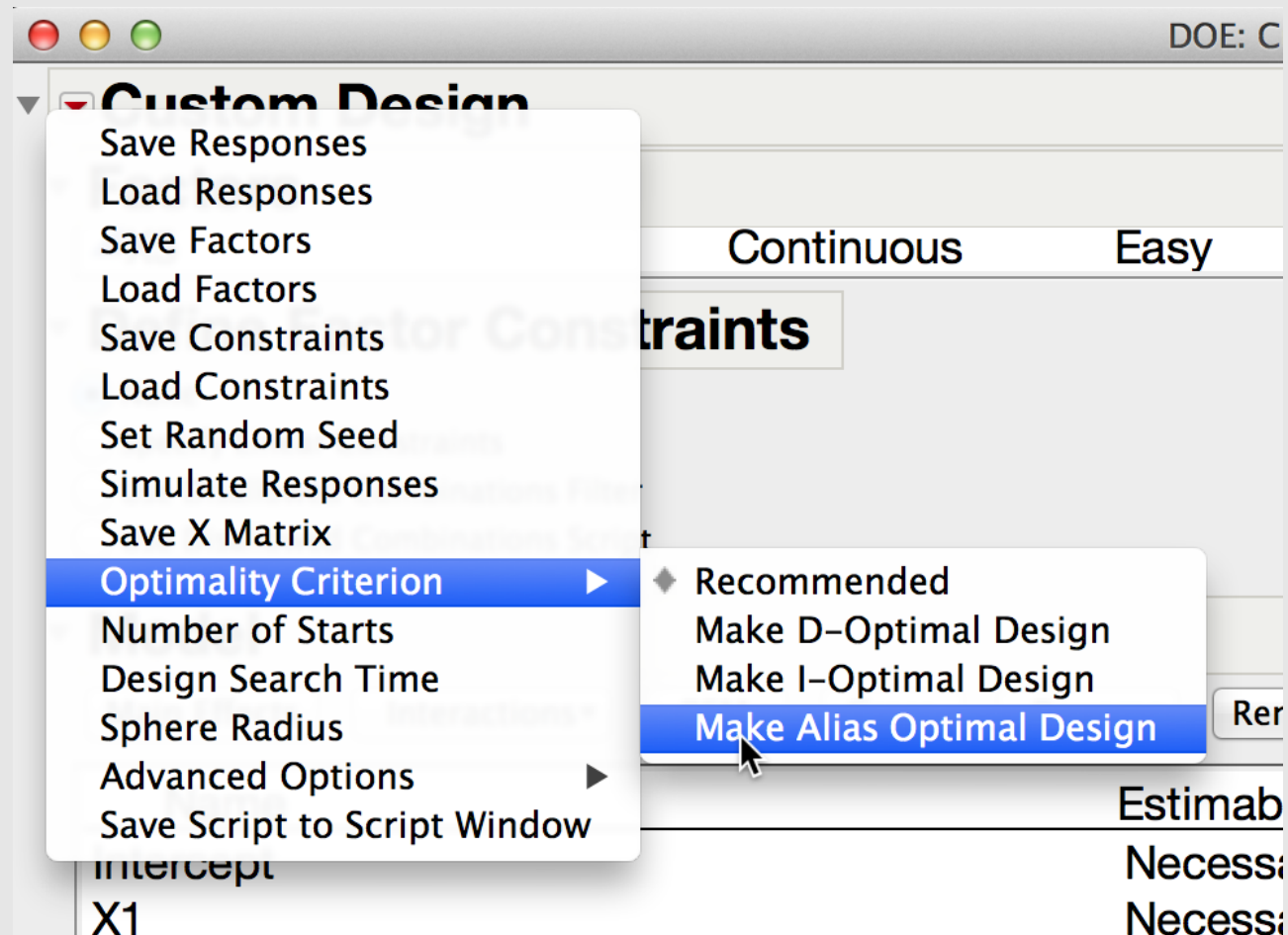
- Let's tell JMP to find a design that minimizes the aliasing (correlations) between main effects and two-factor interactions

# Only need to change the optimality criterion, and re-do

Red triangle >>

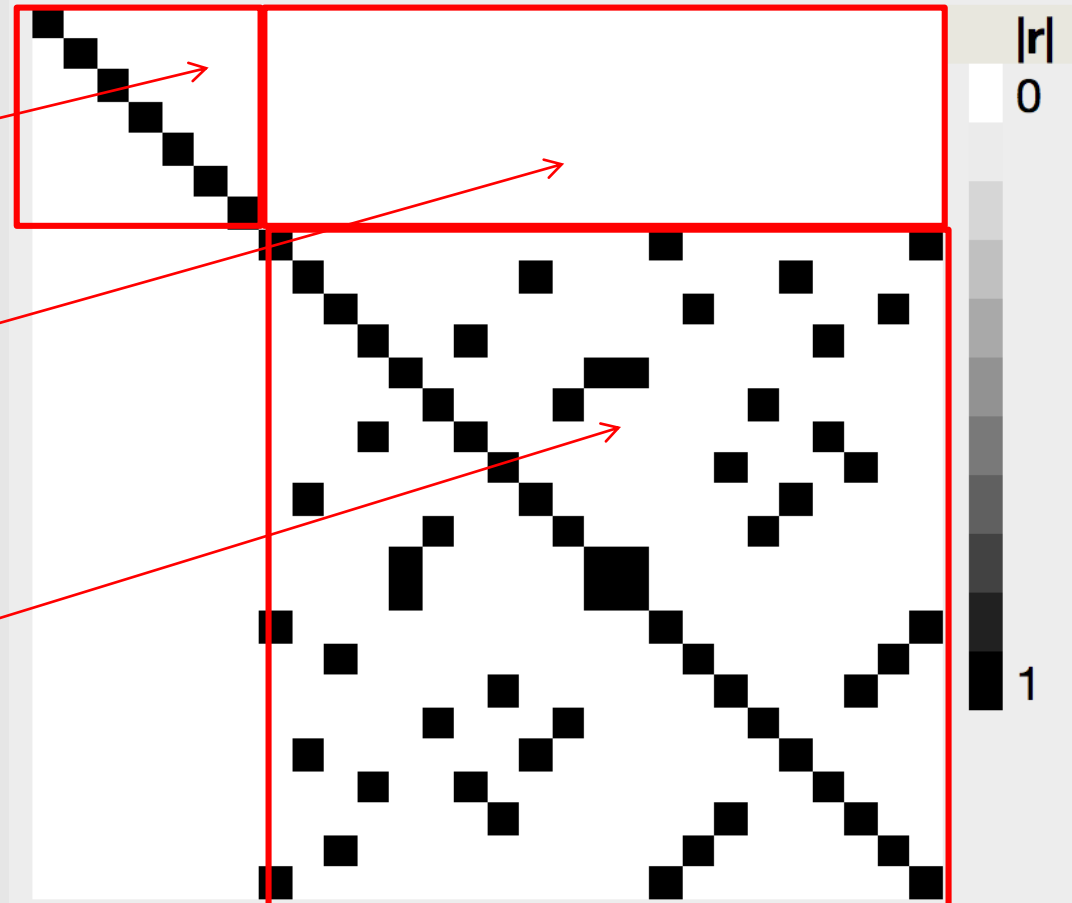
Optimality  
Criterion >>

Make Alias  
Optimal Design



# Result: Resolution IV fractional factorial design

1. Primary interest  
(Yay!)
2. Of concern  
(Yay!)
1. Secondary interest  
(Some two-factor interactions directly confounded)



# How did JMP do that?

## Efficient Designs With Minimal Aliasing

**Bradley JONES**

SAS Institute  
Cary, NC 27513  
and  
Faculty of Applied Economics  
Universiteit Antwerpen  
Prinsstraat 13  
2000 Antwerpen, Belgium  
([Bradley.Jones@jmp.com](mailto:Bradley.Jones@jmp.com))

**Christopher J. NACHTSHEIM**

Operations and Management Science Department  
Carlson School of Management  
University of Minnesota  
Minneapolis, MN 55455  
([Nacht001@umn.edu](mailto:Nacht001@umn.edu))

For some experimenters, a disadvantage of the standard optimal design approach is that it does not consider explicitly the aliasing of specified model terms with terms that are potentially important but are not included in the model. For example, when constructing an optimal design for a first-order model, aliasing of main effects and interactions is not considered. This can lead to designs that are optimal for estimation of the primary effects of interest, yet have undesirable aliasing structures. In this article, we construct exact designs that minimize the squared norm of the alias matrix subject to constraints on design efficiency. We demonstrate use of the method for the construction of screening and response surface designs.

**Technometrics, 2011, Vol 53, pp 62--71**

# JMP Demo: Peanut Solids...Over to Brad

## Contrasts

Term	Contrast		Lenth	Individual	Simultaneous	
			t-Ratio	p-Value	p-Value	Aliases
Ratio	-1.95810		-11.38	<.0001*	0.0004*	Pre-soak?*pH*Water Temp
Pre-soak?	-0.76222		-4.42	0.0041*	0.0396*	
Agitation Speed	0.68295		3.96	0.0064*	0.0598	
Hydrolyze?	-0.41454		-2.41	0.0319*	0.3035	
pH	0.11962		0.69	0.4718	1.0000	Ratio*Pre-soak?*Water Temp
Water Temp	0.00295		0.02	0.9888	1.0000	Ratio*Pre-soak?*pH
Extraction Time	0.00113		0.01	0.9959	1.0000	
Ratio*Ratio	0.01944 *		0.11	0.9192	1.0000	
Ratio*Pre-soak?	0.33411		1.94	0.0697	0.5543	pH*Water Temp
Ratio*Agitation Speed	0.11019		0.64	0.5591	1.0000	
Pre-soak?*Agitation Speed	-0.17147		-0.99	0.3061	0.9963	
Ratio*Hydrolyze?	0.29993 *		1.74	0.0922	0.6848	Agitation Speed*pH
Pre-soak?*Hydrolyze?	-0.06187 *		-0.36	0.7398	1.0000	Agitation Speed*Water Temp
Agitation Speed*Hydrolyze?	-0.29757 *		-1.73	0.0942	0.6955	Ratio*pH, Pre-soak?*Water Temp
Pre-soak?*pH	0.01945		0.11	0.9192	1.0000	Ratio*Water Temp
Pre-soak?*Extraction Time	0.14833 *		0.86	0.3662	0.9997	
Ratio*Pre-soak?*Hydrolyze?	-0.10430 *		-0.61	0.5814	1.0000	Pre-soak?*Agitation Speed*pH, Ratio*Agitation Speed*Water Temp

## Example 2: Maximize recall of TV commercial

### Factors:

**A: Length of commercial  
(10 to 30 seconds)**

**B: Number of repetitions  
(1 to 3)**

**C: Amount of time since last  
viewing (1 to 5 days)**



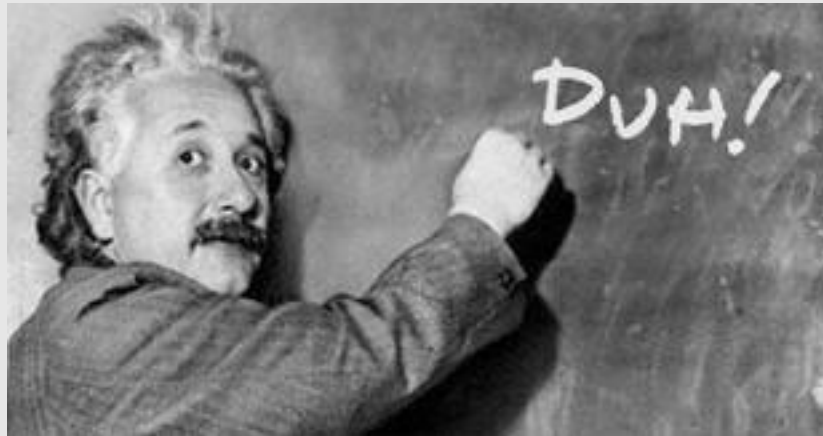
# Classical Design Knowledge Requirements:

- Main effects, interactions, second-order powers (model terms)
- Confounding and resolution (for corners)
- Fractional factorials (for corners)
- CCDs, Box-Behnken designs
- Blocking (sort of)
- Axial points

# Contrast: Problem-Driven Design Approach

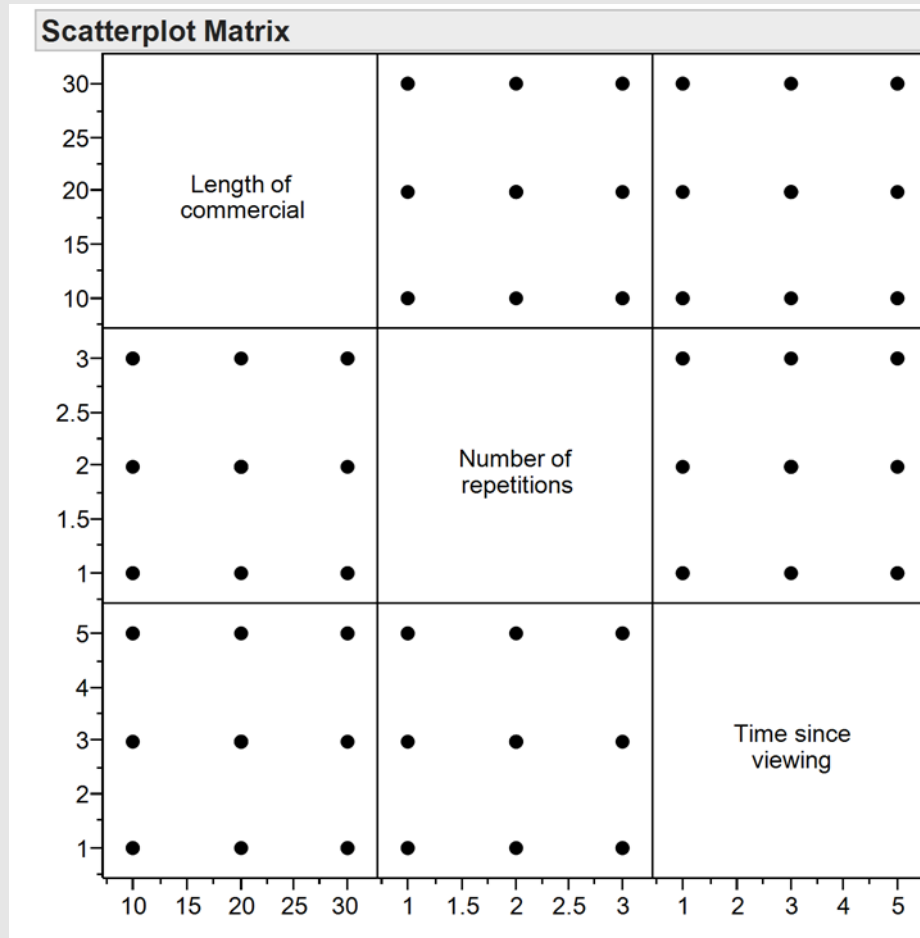
## Knowledge Requirements:

1. Main effects, interactions, quadratic terms  
(RSM button)

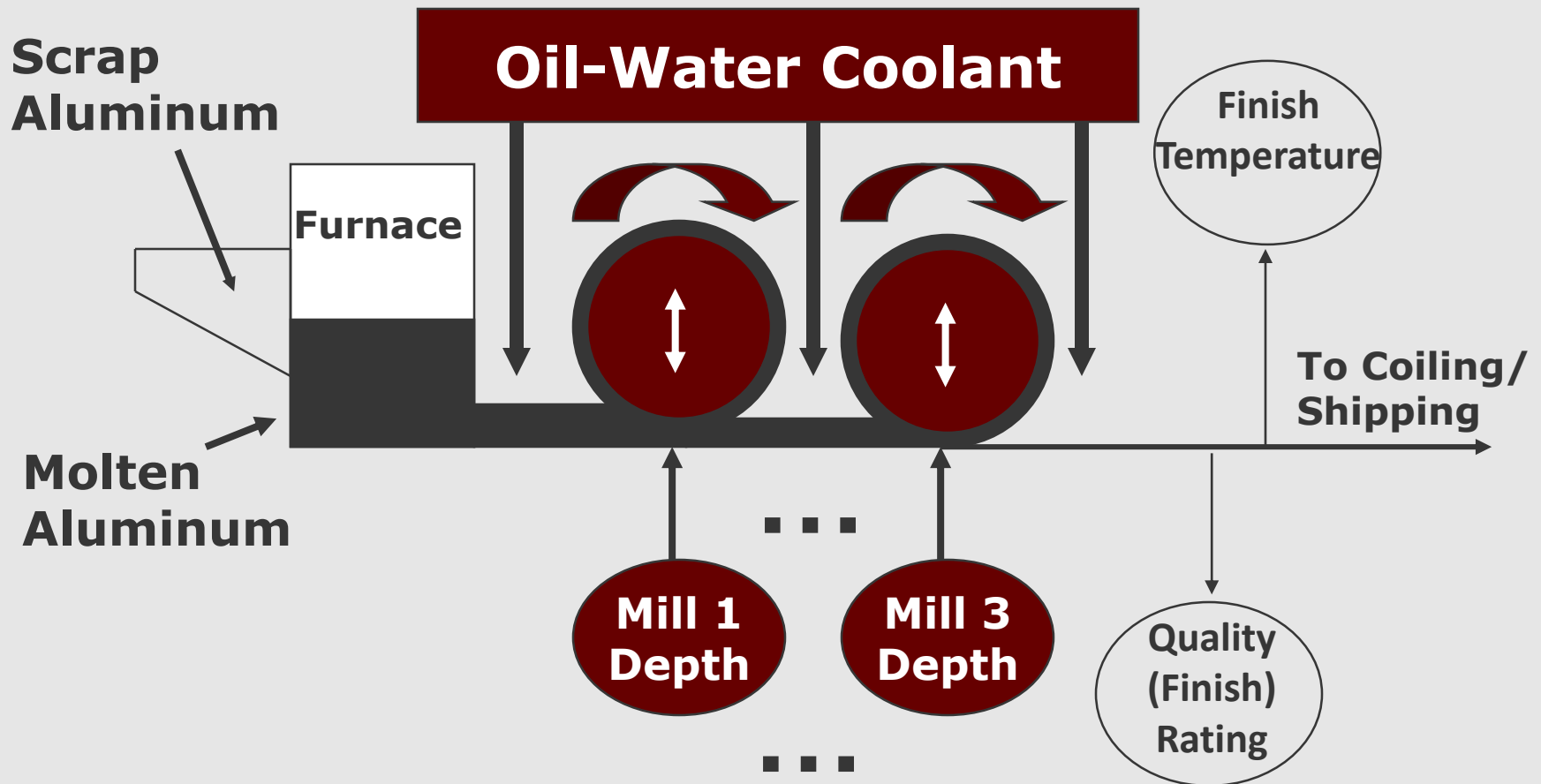




# JMP Demo: Over to Brad



## Example 3: Sheet Aluminum Rolling Mill



# Textbook approach to Aluminum Milling

	Factor	Low	High
1	Mill depth 1	0.1	0.8
2	Mill depth 2	0.1	0.8
3	Mill depth 3	0.1	0.8
4	Spray volume	L	H
5	Spray oil content	L	H

**But there is a problem:**

**We have constrained mill depths (MDs):**

$$MD1 + MD2 + MD3 = 1$$

$$.1 < MD1 < .8$$

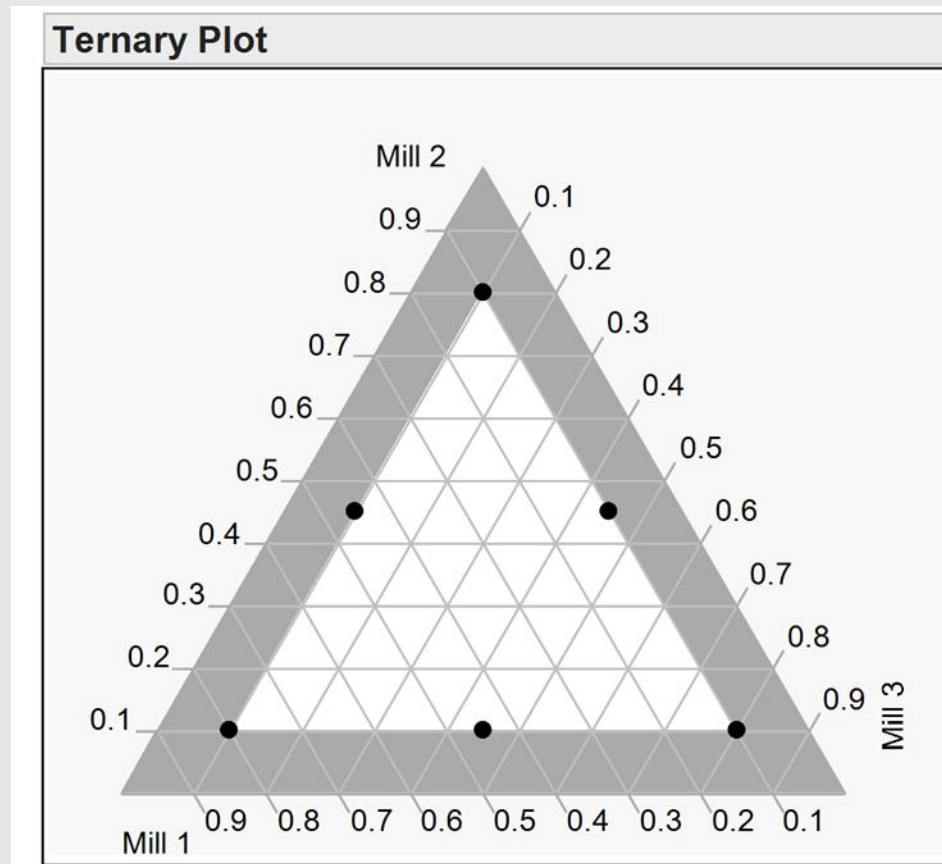
$$.1 < MD2 < .8$$

$$.1 < MD3 < .8$$

**None of the tabled designs work!**



# JMP Demo: Over to Brad



# DOE to the rescue!

- **Result: When mill depth 3 increased, it improved the quality of the surface**



**Scrap reduced from 50% to 10%;  
Company saved!**

# Split Plots: DOE with Easy and Hard to Change Factors



Very common in industrial experiments!

Can trip up even professional statisticians

***“All industrial experiments are split-plot experiments”***

**--Cuthbert Daniel**

# Confession of a Text Book Author

- I have only recently appreciated how effective split plot experiments really are.
- Thorough treatment should be standard in any introductory design course, any text, any statistics package.



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- **They merited a whopping three pages in my text book!**

## Recent Paper---My Conversion

# Split-Plot Designs: What, Why, and How

BRADLEY JONES

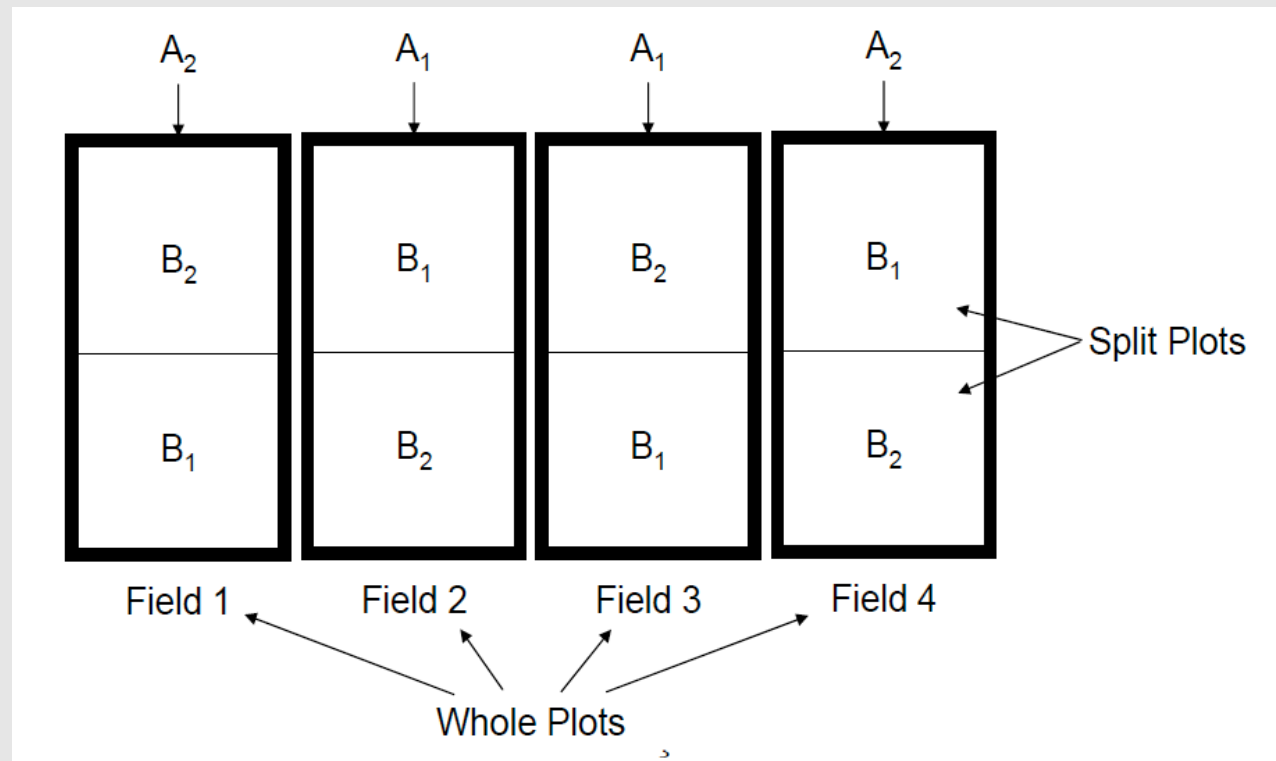
*SAS Institute, Cary, NC 27513*

CHRISTOPHER J. NACHTSHEIM

*Carlson School of Management, University of Minnesota, Minneapolis, MN 55455*

# So What is a Split-Plot Design and Why the Name?

Invented by  
Fisher for  
Agricultural  
Applications



## Example 4: Corrosion resistance of steel bars (Box, Hunter, and Hunter text book)



**DOE with two factors:**

- 1. Furnace (curing) temperatures  
(360°C, 370°C, 380°C)**
- 2. Coating type ( $C_1$ ,  $C_2$ ,  $C_3$ ,  $C_4$ )**

# Easy-to-change and hard-to-change factors

**Problem: Not all factors are created equally!**



**Furnace temperature:**

**Hard to change!**

**Coating type:**

**Easy to change!**

# Actual Corrosion Resistance DOE

Whole-Plots (H2C)		Split-Plots (E2C)			
Furnace Run	Temperature	Coating in randomized order			
1	360	C2	C1	C3	C4
2	380	C1	C3	C4	C2
3	370	C3	C1	C2	C4
4	380	C4	C3	C2	C1
5	370	C4	C1	C3	C2
6	360	C1	C4	C2	C3

**Advantage:  $n = 24$  but only 6 furnace runs!**

# Key Split Plot DOE Characteristic

**Two randomizations:**

**Step 1: Randomly assign temps to furnace runs**

**Step 2: For each furnace run...**

**Randomly assign coatings to four subgroups of bars  
within the furnace run**

# Contrast with ordinary DOE

In usual DOE: just  
one randomization:

3x4 factorial with an  
added replicate  
requires 24 furnace  
runs

(Only one coating  
type per furnace run)

3x4 Factorial				
Design	3x4 Factorial			
Screening				
Model				
Columns (4/0)				
Pattern				
Temperature *				
Coating *				
Furnace Reset?				
Rows				
All rows	24			
Selected	0			
Excluded	0			
Hidden	0			
Labelled	0			

	Pattern	Temperature	Coating
1	22	370	C2
2	13	360	C3
3	24	370	C4
4	31	380	C1
5	11	360	C1
6	32	380	C2
7	34	380	C4
8	34	380	C4
9	14	360	C4
10	31	380	C1
11	24	370	C4
12	23	370	C3
13	33	380	C3
14	13	360	C3
15	21	370	C1
16	21	370	C1
17	22	370	C2
18	23	370	C3
19	33	380	C3
20	12	360	C2
21	12	360	C2
22	32	380	C2
23	11	360	C1
24	14	360	C4



# All-Too-Common Mistake

- Step 1.**            **Create ordinary randomized DOE**
- Step 2.**            **Sort the runs by the hard-to-change factor  
and run in that order  
(DOE is actually now a split plot design)**
- Step 3.**            **Analyze as though a randomized DOE**

# Analysis depends on the randomization

Usual analysis	Factor	p-value	Significant?	
	Temperature	0.003	Yes	Wrong!
	Coatings	0.386	No	Wrong!
	Temp x Coatings	0.852	No	Wrong!

# Analysis depends on the randomization

Usual analysis	Factor	p-value	Significant?	
	Temperature	0.003	Yes	Wrong!
	Coatings	0.386	No	Wrong!
	Temp x Coatings	0.852	No	Wrong!

Correct analysis	Factor	p-value	Significant?	
	Temperature	0.209	No	Right!
	Coatings	0.002	Yes	Right!
	Temp x Coatings	0.024	Yes	Right!

# JMP Demo: Over to Brad



## REML ANOVA Table

Source	Nparm	DF	DFDen	F Ratio	Prob > F
Temperature	2	2	3	2.7548	0.2093
Coating	3	3	9	11.4798	0.0020*
Temperature*Coating	6	6	9	4.3757	0.0241*

## ANOVA Assuming Completely Randomized Design

Source	Nparm	DF	Sum of Squares	F Ratio	Prob > F
Temperature	2	2	26519.250	10.2256	0.0026*
Coating	3	3	4289.125	1.1026	0.3860
Temperature*Coating	6	6	3269.750	0.4203	0.8518

# Take-Aways

1. Split plots are easier to run!
1. Lose info on whole plot main effects, but...
1. Gain info on split-plot factors and split-plot by whole-plot factor interactions
2. Overall efficiency can be better than CRD
1. Teach these methods in workshops on DOE

# What's driving the software?

# What's driving the software?

## Optimal design construction, optimal blocking

- Coordinate exchange algorithm, Meyer and Nachtsheim, *Technometrics*, 1995
- Optimal blocking, Cook, Nachtsheim, *Technometrics*, '1987

## Bayesian D-Optimal Designs

- DuMouchel and Jones, *Technometrics*, 1994
- Jones, Lin, Nachtsheim, *Journal of Statistical Planning and Inference*, 2006

## Optimal Split-Plot and Split-Split Plot Designs

- Goos and Jones, *Biometrika*, 2009



# Optimal designs depend on a model!

OH YEAH?



1. No more so than classical designs.
2. They should!
3. If you're going to lose sleep, see model robustness, Bayesian, multi-objective design literature



# Lots of work in Model Robust and Multi-Objective Design

## Randomly chosen examples

1. Li/Nachtsheim, 2000, Model Robust Factorial Designs
2. DuMouchel/Jones, 1994, Bayesian D-optimal Designs
3. Jones/Nachtsheim, 2011 Minimal Aliasing Designs

# Definitive Screening Designs

# Multi-objective designs: six-factor example, $n = 12$

- Standard choice: Plackett-Burman design
- Plackett-Burman designs have “complex aliasing of the main effects by two-factor interactions.

# Alias Matrix for PB design

Full Model:  $1 + 6 + 15 = 22$  terms

$$Y = X_1\beta_1 + X_2\beta_2 + \epsilon$$

Model we can fit:

$$Y = X_1\beta_1 + \epsilon^*$$

Leads to bias:

$$E(\hat{\beta}_1) = \beta_1 + A\beta_2$$

Alias matrix:

$$A = (X_1'X_1)^{-1}X_1'X_2$$

# Alias Matrix of Plackett-Burman design

Alias Matrix

Effect	A*B	A*C	A*D	A*E	A*F	B*C	B*D	B*E	B*F	C*D	C*E	C*F	D*E	D*F	E*F
Intercept	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
A	0	0	0	0	0	0.333	-0.33	0.333	0.333	0.333	-0.33	0.333	-0.33	0.333	-0.33
B	0	0.333	-0.33	0.333	0.333	0	0	0	0	-0.33	-0.33	-0.33	-0.33	0.333	0.333
C	0.333	0	0.333	-0.33	0.333	0	-0.33	-0.33	-0.33	0	0	0	-0.33	0.333	-0.33
D	-0.33	0.333	0	-0.33	0.333	-0.33	0	-0.33	0.333	0	-0.33	0.333	0	0	-0.33
E	0.333	-0.33	-0.33	0	-0.33	-0.33	-0.33	0	0.333	-0.33	0	-0.33	0	-0.33	0
F	0.333	0.333	0.333	-0.33	0	-0.33	0.333	0.333	0	0.333	-0.33	0	-0.33	0	0

PB (non-regular) design has “complex aliasing”

# If only there were another design with this alias matrix:


Alias Matrix

Effect	A*B	A*C	A*D	A*E	A*F	B*C	B*D	B*E	B*F	C*D	C*E	C*F	D*E	D*F	E*F
Intercept	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
A	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
B	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
C	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
D	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
E	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
F	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0



# Turns out there is:

**Six foldover  
pairs**



Run	A	B	C	D	E	F
1	0	1	-1	-1	-1	-1
2	0	-1	1	1	1	1
3	1	0	-1	1	1	-1
4	-1	0	1	-1	-1	1
5	-1	-1	0	1	-1	-1
6	1	1	0	-1	1	1
7	-1	1	1	0	1	-1
8	1	-1	-1	0	-1	1
9	1	-1	1	-1	0	-1
10	-1	1	-1	1	0	1
11	1	1	1	1	-1	0
12	-1	-1	-1	-1	1	0
13	0	0	0	0	0	0

# Definitive Screening Design for 6 factors

Center point in  
each row

Run	A	B	C	D	E	F
1	0	1	-1	-1	-1	-1
2	0	-1	1	1	1	1
3	1	0	-1	1	1	-1
4	-1	0	1	-1	-1	1
5	-1	-1	0	1	-1	-1
6	1	1	0	-1	1	1
7	-1	1	1	0	1	-1
8	1	-1	-1	0	-1	1
9	1	-1	1	-1	0	-1
10	-1	1	-1	1	0	1
11	1	1	1	1	-1	0
12	-1	-1	-1	-1	1	0
13	0	0	0	0	0	0



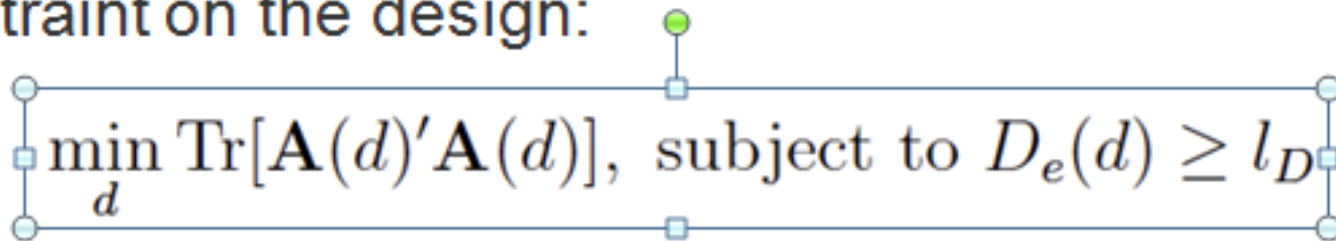
# Definitive Screening Design for 6 factors

Run	A	B	C	D	E	F
1	0	1	-1	-1	-1	-1
2	0	-1	1	1	1	1
3	1	0	-1	1	1	-1
4	-1	0	1	-1	-1	1
5	-1	-1	0	1	-1	-1
6	1	1	0	-1	1	1
7	-1	1	1	0	1	-1
8	1	-1	-1	0	-1	1
9	1	-1	1	-1	0	-1
10	-1	1	-1	1	0	1
11	1	1	1	1	-1	0
12	-1	-1	-1	-1	1	0
13	0	0	0	0	0	0

One overall  
center point

# How did we find this design?\*

Design criterion: Minimize  $\|\mathbf{A}\|^2$  subject to an efficiency constraint on the design:

$$\min_d \text{Tr}[\mathbf{A}(d)' \mathbf{A}(d)], \text{ subject to } D_e(d) \geq l_D$$


where the D-efficiency of the design is:

$$D_e(d) = \left[ \frac{|\mathbf{X}_1(d)' \mathbf{X}_1(d)|}{|\mathbf{X}_1(d_D)' \mathbf{X}_1(d_D)|} \right]^{1/p_1}$$

\*Jones, Nachtsheim, *Technometrics*, 2011

# How did we find this design?\*

- Minimize the average magnitude of the correlations between main effects and two-factor interactions
- Subject to a constraint on the statistical efficiency of the design (e.g., efficiency  $> 90\%$ )

\*Jones, Nachtsheim, *Technometrics*, 2011

# Now generalize this structure for $m$ factors

Table 1: General design structure for  $m$  factors

Foldover Pair	Run ( $i$ )	Factor Levels				
		$x_{i,1}$	$x_{i,2}$	$x_{i,3}$	$\dots$	$x_{i,m}$
1	1	0	$\pm 1$	$\pm 1$	$\dots$	$\pm 1$
	2	0	$\mp 1$	$\mp 1$	$\dots$	$\mp 1$
2	3	$\pm 1$	0	$\pm 1$	$\dots$	$\pm 1$
	4	$\mp 1$	0	$\mp 1$	$\dots$	$\mp 1$
3	5	$\pm 1$	$\pm 1$	0	$\dots$	$\pm 1$
	6	$\mp 1$	$\mp 1$	0	$\dots$	$\mp 1$
$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\vdots$	$\ddots$	$\vdots$
$m$	$2m - 1$	$\pm 1$	$\pm 1$	$\pm 1$	$\dots$	0
	$2m$	$\mp 1$	$\mp 1$	$\mp 1$	$\dots$	0
Centerpoint	$m + 1$	0	0	0	$\dots$	0

Can we find  
great designs  
for any  $m$ ?

# It turns out there is a “Conference Matrix” solution

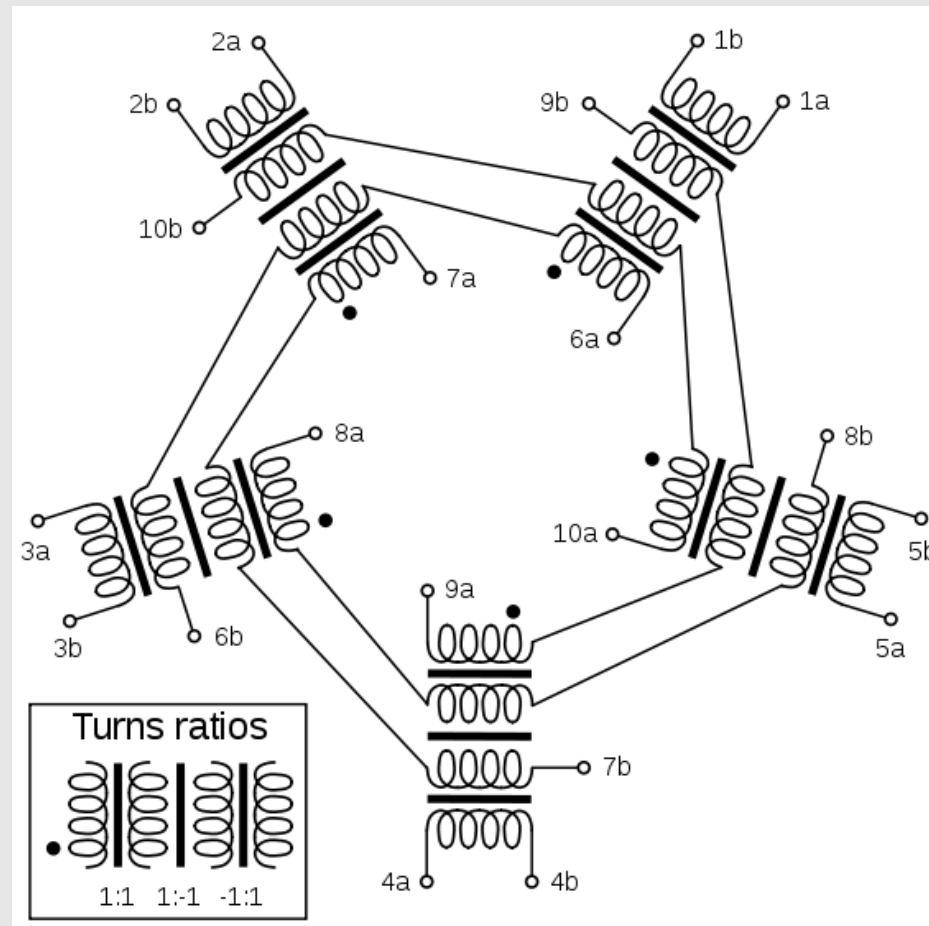
An  $m \times m$  square matrix  $C$  with 0 diagonal and +1 or -1 off diagonal elements such that:

$$C^T C = (m - 1)I_{m \times m}$$

- Arose in connection with a problem in telephony and first described and named by **Vitold Belevitch**
- Belevitch was interested in constructing ideal telephone conference networks from ideal transformers and discovered that such networks were represented by  $C$ .

# Conference Matrices & Telephony

From  
Wikipedia  
article on  
Conference  
Matrices



# Conference Matrix of Order 6

$$\mathbf{C} = \begin{pmatrix} 0 & +1 & +1 & +1 & +1 & +1 \\ +1 & 0 & +1 & -1 & -1 & +1 \\ +1 & +1 & 0 & +1 & -1 & -1 \\ +1 & -1 & +1 & 0 & +1 & -1 \\ +1 & -1 & -1 & +1 & 0 & +1 \\ +1 & +1 & -1 & -1 & +1 & 0 \end{pmatrix}$$

Here the amazing result:

Form the augmented matrix:

$$D = \begin{bmatrix} \hat{e} + C\hat{u} \\ \hat{e} & \hat{u} \\ \hat{e} - C\hat{u} \\ \hat{e} & 0 & \hat{u} \end{bmatrix}$$

...and you get an orthogonal (for main effects)  
**definitive screening design!**



# Design Properties

1. The number of required runs is only one more than twice the number of factors.

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3. Unlike all predecessors, these are three-level designs, so we can estimate curvatures!

# Design Properties

1. The number of required runs is only one more than twice the number of factors.
2. Unlike Plackett-Burman and Resolution III and IV fractional factorial designs, main effects are completely independent of two-factor interactions.
3. Unlike all predecessors, these are three-level designs, so we can estimate curvatures!
4. Designs are capable of estimating all possible response surface models involving three or fewer factors with very high levels of statistical efficiency.

# Impact? A break-through solution for sequestering greenhouse gasses

## 2012 Statistics in Chemistry Award Winner



*Bradley Jones, SAS Institute, JMP Division*



*Scott Allen, Novomer, Inc.*

To Bradley Jones and Scott Allen for outstanding collaborative work in developing a new catalyst for CO<sub>2</sub>-based polymers that sequester CO<sub>2</sub>.

# Impact? First published DSD case study, 2013

Biotechnol Lett

DOI 10.1007/s10529-012-1089-y

ORIGINAL RESEARCH PAPER

## **Efficient biological process characterization by definitive-screening designs: the formaldehyde treatment of a therapeutic protein as a case study**

**Axel Erler • Nuria de Mas • Philip Ramsey •  
Grant Henderson**

## Impact? From the conclusions:

***“Definitive-screening designs were used to efficiently select a model describing the formulation of a protein under clinical development. The ability of the single definitive screening design to identify and model all the active effects obviated the need for further experimentation, reducing the total number of experimental runs required to 17 from the greater than or equal to 70 runs that would have been necessary using the traditional screening/RSM approach.”***

# Awards for Definitive Screening Designs

**“A New Class of Three-Level Screening Designs for Definitive Screening in the Presence of Second-Order Effects”, *Journal of Quality Technology*, Jan., 2011**

- **Brumbaugh Award, 2011**
- **Lloyd S. Nelson Award, 2012**
- **(Led to) Statistics in Chemistry Award, 2012**
- **Much work now in combinatorial aspects**



## Upshot – Definitive Screening Designs

1. Our view: engineers, scientists prefer three levels.
2. Can estimate curvatures
3. Can disentangle interactions
4. Designs project to fully estimable response surface designs in two or three factors, so with effect sparsity, two experiments for the price of one.
5. **We see little or no reason to continue the practice of using  $2^{k-p}$  designs for four or more continuous factors!!**

# BUT!...a limitation

**DSDs can only be used with  
continuous (quantitative) factors**

# BUT!...a limitation

DSDs can only be used with  
continuous (quantitative) factors

**DANG!!!!**

# Not any more:

## Definitive Screening Designs with Added Two-Level Categorical Factors

BRADLEY JONES

*SAS Institute, Cary, NC*

CHRISTOPHER J. NACHTSHEIM

*Carlson School of Management, University of Minnesota, Minneapolis, MN*

- *Journal of Quality Technology*, April 2013

# Example 5: Extracting food solids from peanuts in solution, again, $n = 18$

		Factor	Low	High
<b>5 continuous factors</b>	1	Water pH level	6.95	8.0
	2	Water temp	20C	60C
	3	Extraction time	15	40
	4	Water-Peanuts Ratio	5	9
	5	Agitation speed	5,000	10,000
<b>2 categorical factors</b>	6	Hydrolyzed?	N	Y
	7	Presoaking?	N	Y

# JMP Demo: Peanut Solids as a DSD---Over to Brad

Recall our conclusion: **There are three active interaction effects, but we cannot identify them!**

Using the same number of runs (16, plus 2 center points) the definitive screening design will not only identify the same main effects, but it will:

- Investigate all curvatures (concluding there are none)
- Investigate all interactions and **definitively** identify the three active interactions

## Example 6: A Recent Experiment at In'Tech Industries

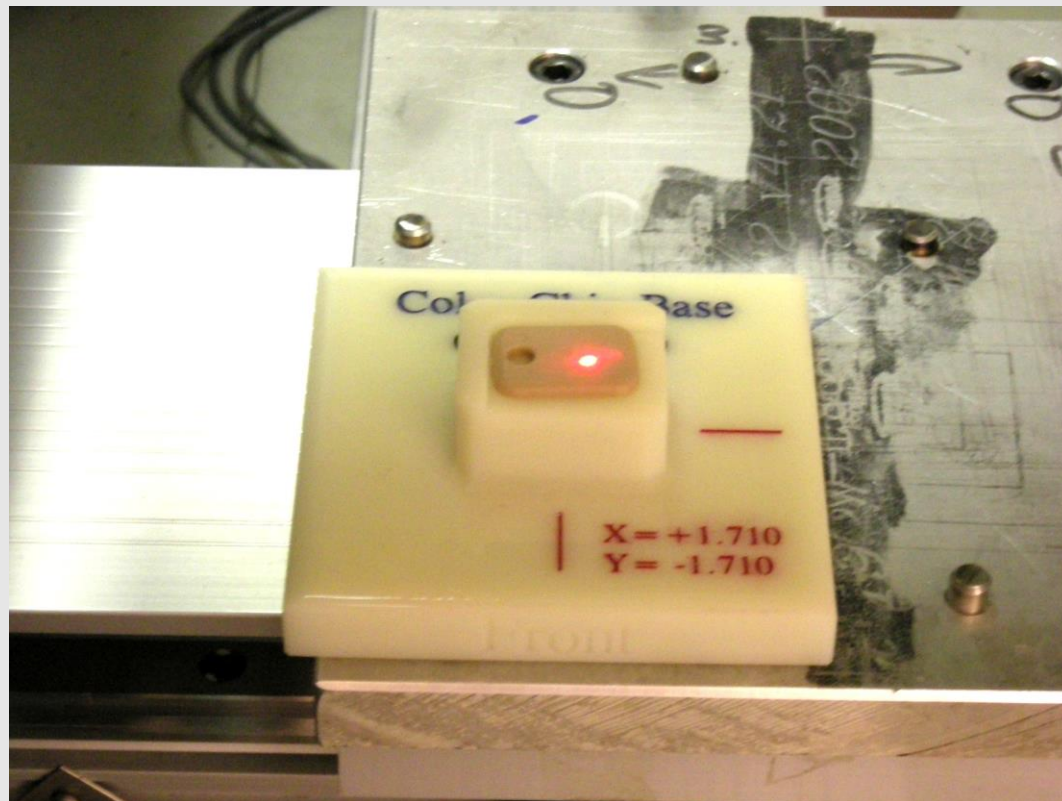


# Problem: Need to laser etch labels on small plastic parts in an “optimal” fashion



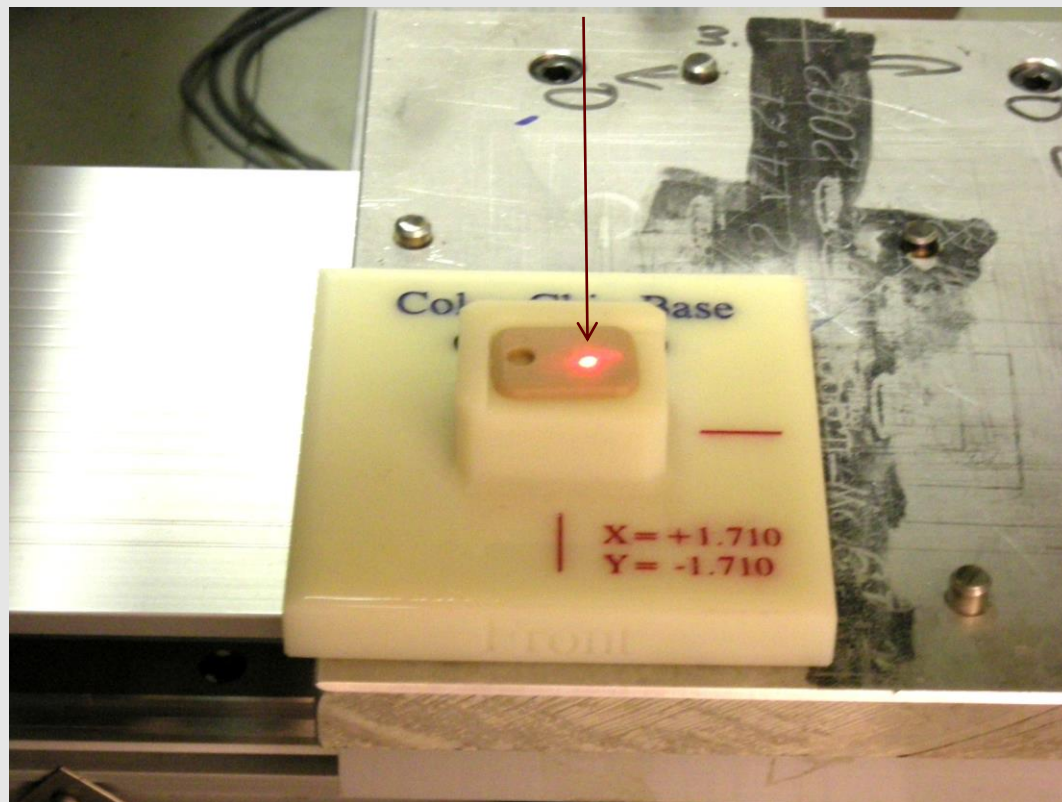


# Laser etching in progress....

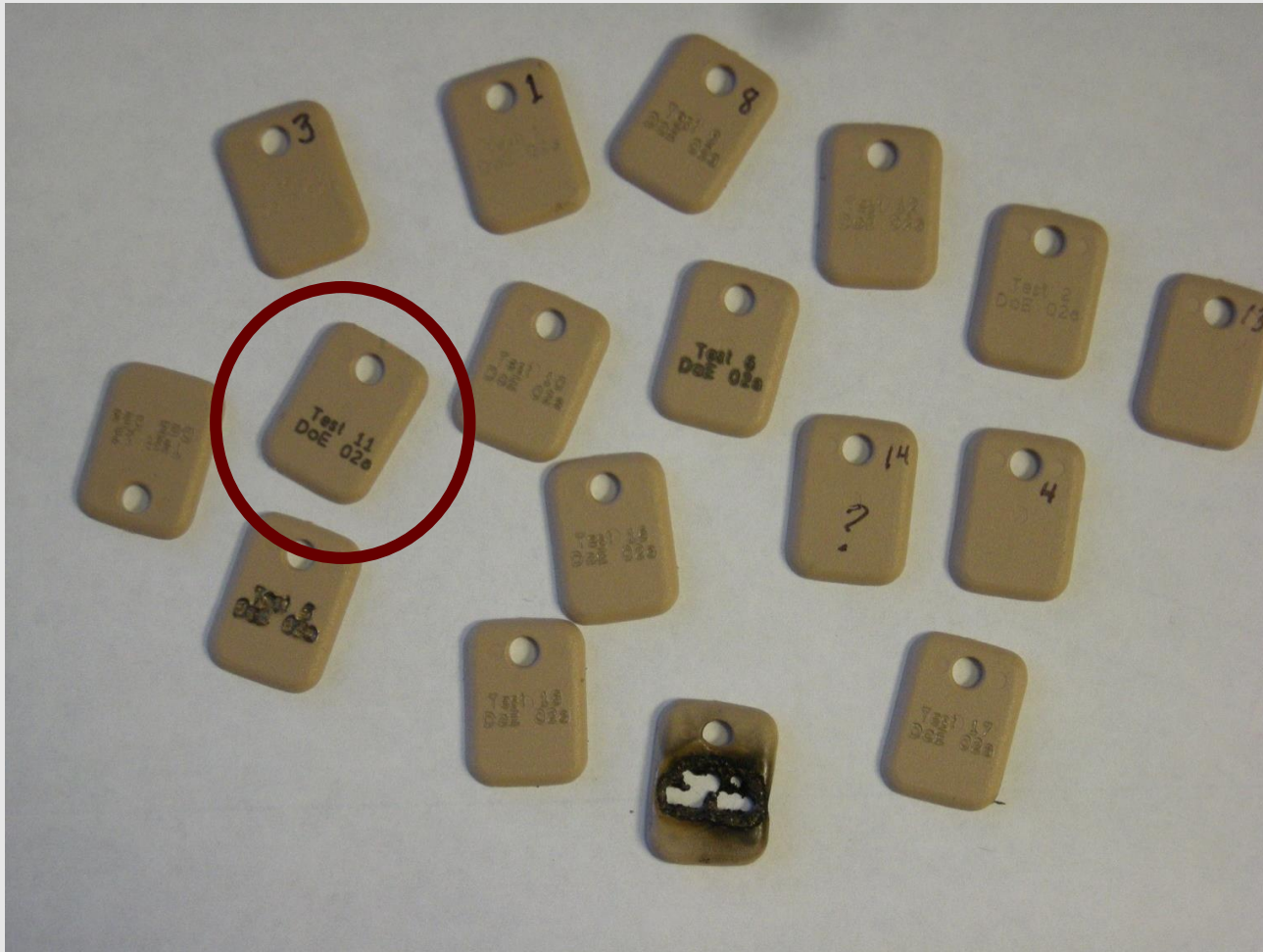


# Laser etching in progress....

## Laser

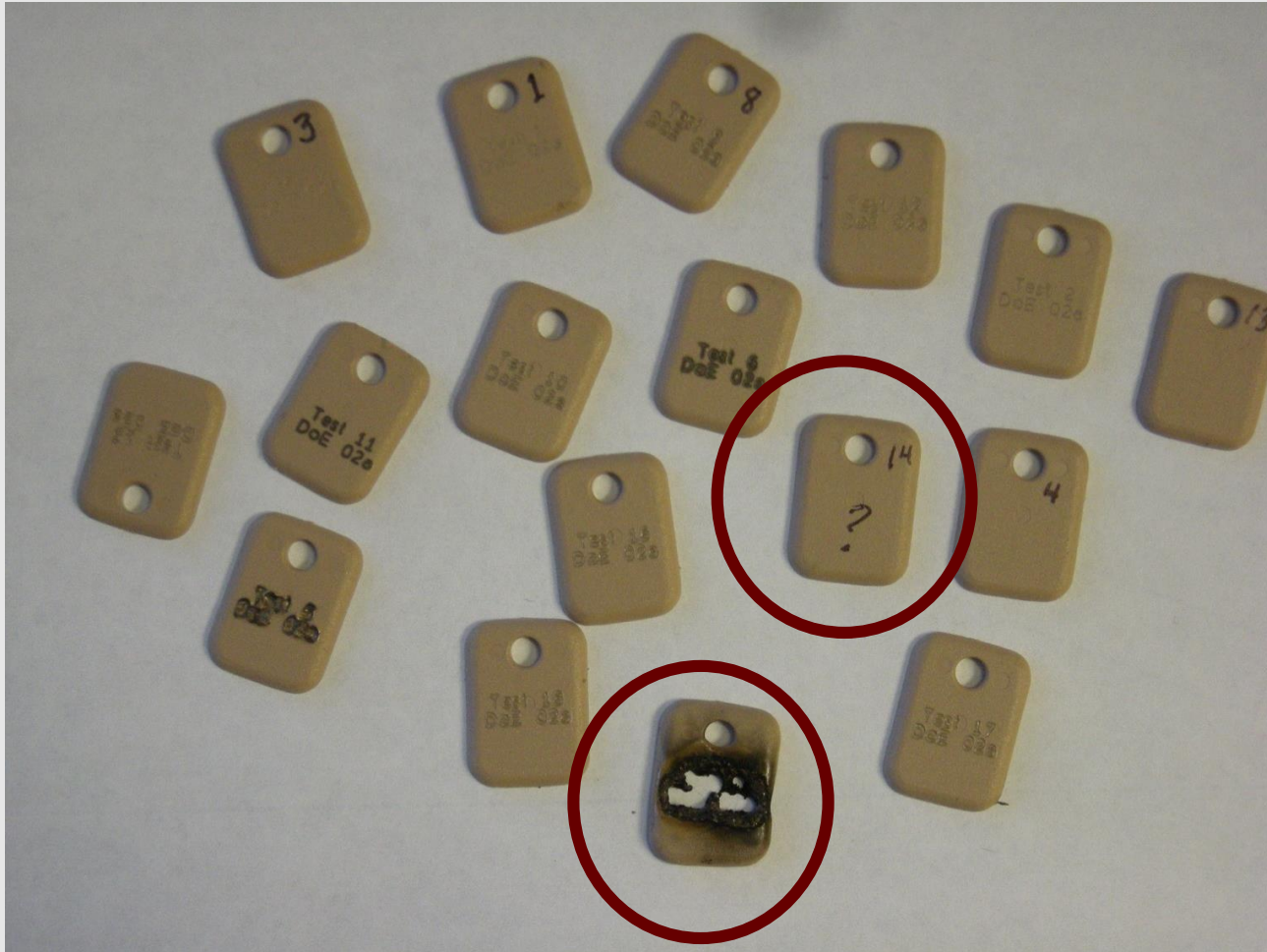


# Initial experience: From easy to read,





# Initial results: ..... to not so easy to read



# Factors and ranges....

Factors	Factor ranges	
	Low level	High level
Mark Speed	8	15
Frequency	1	5
Percent Power	15	55
Repetitions	1	5
Humidity	5%	15%

Blocking factor is operator

# Ultimate objective: optimization

## Experimental plan

**Step 1: Screen (5 factors @ 3 levels - DSD)**

**Step 2: Optimize: RSD in fewer(?) factors**

**Complication: Need to block on operators**

**Conundrum: How to block a definitive screening design?**

# Solution to appear in Technometrics, 2015

## Blocking Schemes for Definitive Screening Designs

BRADLEY JONES

*SAS Institute, Cary, NC*

CHRISTOPHER J. NACHTSHEIM

*Carlson School of Management, University of Minnesota, Minneapolis, MN*

### Abstract

Jones and Nachtsheim (2011) proposed a new class of screening designs called definitive screening designs. As originally presented, these designs are three-level designs for quantitative factors that provide estimates of main effects that are unbiased by any second-order effect and require only one more than twice as many runs as there are factors. Definitive screening designs avoid direct confounding of any pair of second-order effects, and, for designs that have more than five factors, project to efficient response surface designs for any two or three factors. Recently, Jones and Nachtsheim (2013) expanded the applicability of these designs by showing how to include any number of two-level categorical factors. However, methods for blocking definitive screening de-

# Analyzing DSDs



# Simplest idea – Fit the main (linear) effects model

## Advantages:

MEs unbiased – you can believe the coefficient estimates



# Fit the Main Effects Model

## Disadvantages:

Estimate of  $\sigma$  inflated with strong 2FIs or quadratic effects

May make active MEs appear not statistically significant.

You cannot believe the coefficient standard errors.



# Analysis Idea #2 – Use Stepwise Regression

## Advantages:

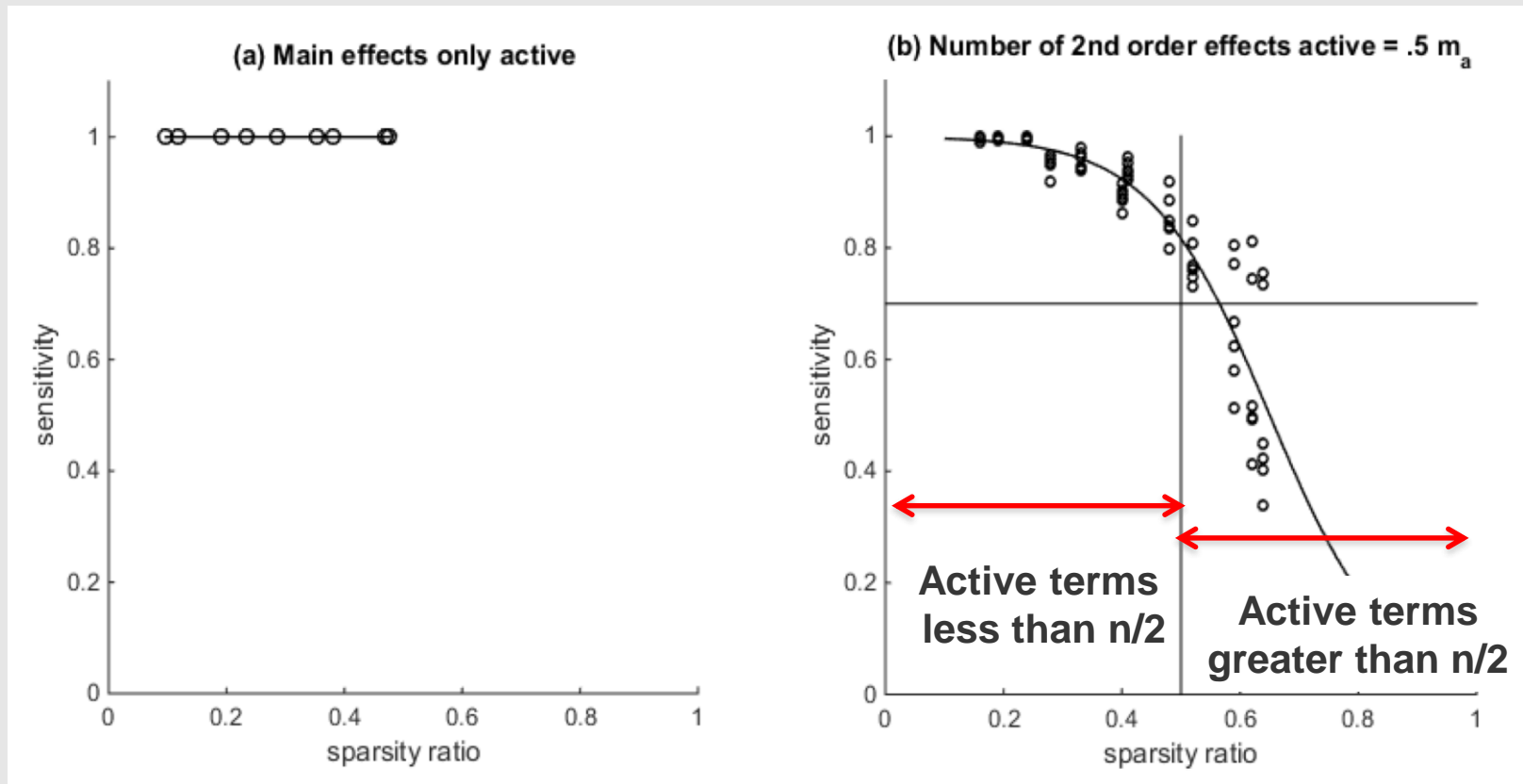
- Easy to do
- Available in most software
- Specify the response surface model
- Use forward variable selection based on the AICc criterion
- Or use the Dantzig Selector



# Limitations of DSDs

1. Stepwise/AICc and Dantzig break down if there are more than about  **$n/2$  active terms** in the model
2. The quadratic effect must be large (2-3 sigma) to have good power.

# Simulation Results\*



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# Fix 1: “Separate” Analyses of First-, Second-Order Effects

Since main effects and 2<sup>nd</sup> order effects are orthogonal to each other we can split the response (Y) into two new responses

- One response for identifying main effects – call it Y\_ME
- One response for identifying the intercept and the 2<sup>nd</sup> order effects – call it Y\_2nd
- And the two columns are orthogonal to each other

## Fix 2: Employ two or more fake factors

- Adding **fake factors** provides a way to estimate variance without repeating center runs!
  - **Fake factors** are orthogonal to the real factors and all the 2<sup>nd</sup> order effects
  - Assuming the 3<sup>rd</sup> and higher order effects are negligible, we can use the fake factor degrees of freedom to create an unbiased estimate of the error variance!
- Use this estimate to identify active main effects

# Analysis Start – Computing the New Responses

1. Fit the main effects model (**no Intercept**) to the real and fake factors and compute the residuals – these residuals are the responses for the 2<sup>nd</sup> order effects (Y2nd).
2. Save the predicted values from the above regression for the main effects response (YME).



## Stage 1 – Identify Active Main Effects

1. Recall that the residuals from fitting the main effects data to the real factors only have 2 degrees of freedom.
2. To estimate  $\sigma^2$ , sum the squared residuals from this fit and divide the result by 2.
3. Using this estimate, do t-tests of each coefficient and retain the active main effects

## Stage 2 – Identify 2<sup>nd</sup> Order Effects Assuming Heredity

- Form all the 2<sup>nd</sup> order terms involving the active main effects
- Using  $Y_{2nd}$ , do forward all subsets regression up to the point where the MSE of the best 2<sup>nd</sup> order model for a given number of terms is not significantly larger than your estimate of  $\sigma^2$  from Stage 1

# Example:

A	B	C	D	E	F	Fake 1	Fake 2	Y	Y_2nd	Y_ME
0	1	1	1	1	1	1	1	94.51	101.04	-6.53
0	-1	-1	-1	-1	-1	-1	-1	107.57	101.04	6.53
1	0	1	1	-1	1	-1	-1	94.36	101.1...	-6.815
-1	0	-1	-1	1	-1	1	1	107.99	101.1...	6.815
1	-1	0	1	1	-1	1	-1	91.80	90.525	1.275
-1	1	0	-1	-1	1	-1	1	89.25	90.525	-1.275
1	-1	-1	0	1	1	-1	1	93.70	94.485	-0.785
-1	1	1	0	-1	-1	1	-1	95.27	94.485	0.785
1	1	-1	-1	0	1	1	-1	89.55	88.71	0.84
-1	-1	1	1	0	-1	-1	1	87.87	88.71	-0.84
1	-1	1	-1	-1	0	1	1	94.58	95.235	-0.655
-1	1	-1	1	1	0	-1	-1	95.89	95.235	0.655
1	1	-1	1	-1	-1	0	1	93.23	89.58	3.65
-1	-1	1	-1	1	1	0	-1	85.93	89.58	-3.65
1	1	1	-1	1	-1	-1	0	98.11	95.815	2.295
-1	-1	-1	1	-1	1	1	0	93.52	95.815	-2.295
0	0	0	0	0	0	0	0	99.75	99.75	0

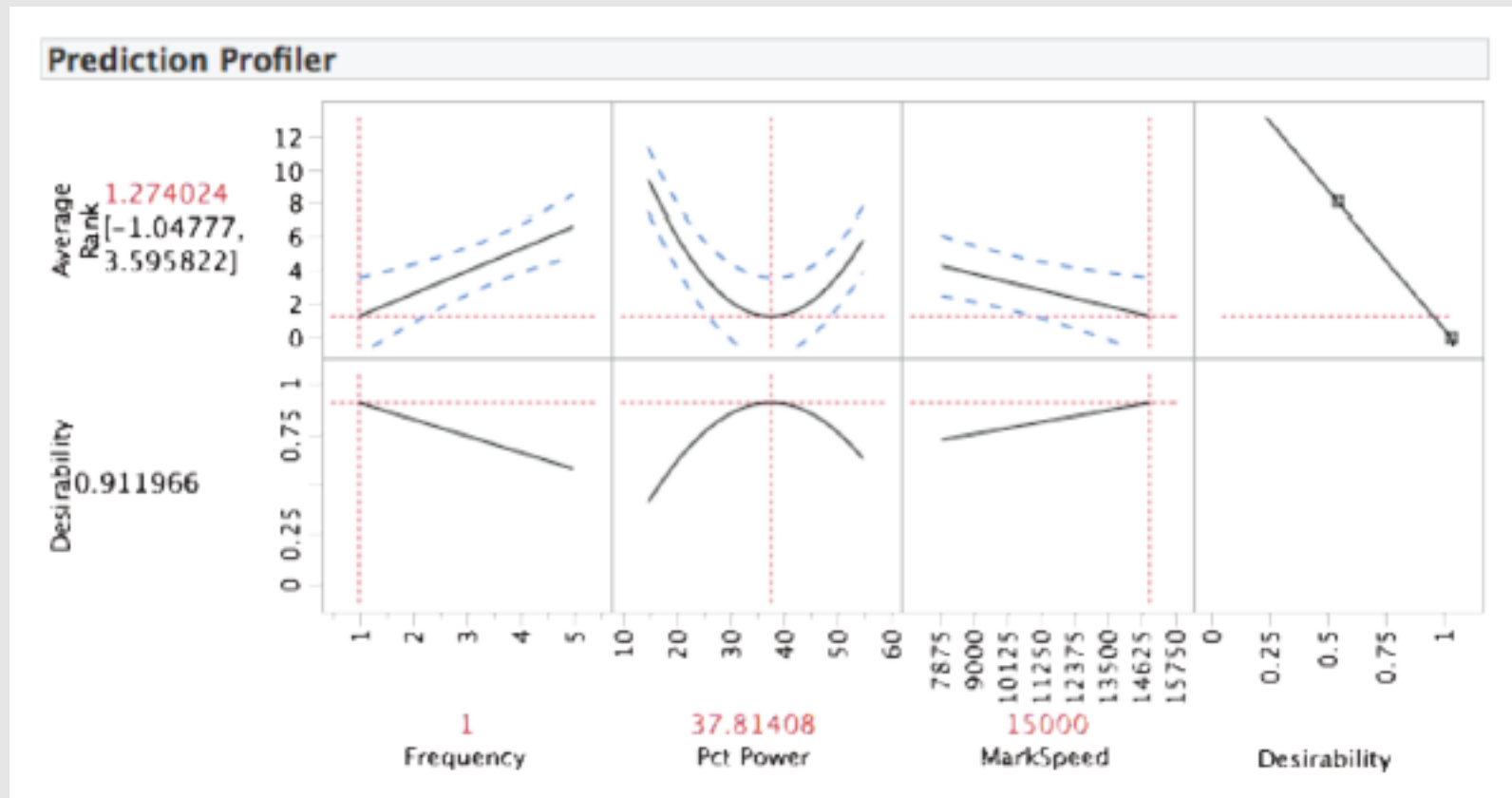
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# Laser Etch Design: 15-run DSD in 3 blocks

Run	MarkSpeed	Frequency	Pct Power	Repetitions	Humidity	Block
1	11500	1	15	1	Low	1
2	8000	3	15	5	High	1
3	11500	3	35	3	Medium	1
4	11500	5	55	5	High	1
5	15000	3	55	1	Low	1
6	15000	1	55	3	High	2
7	8000	5	15	3	Low	2
8	8000	1	35	1	High	2
9	11500	3	35	3	Medium	2
10	15000	5	35	5	Low	2
11	15000	1	15	5	Medium	3
12	8000	1	55	5	Low	3
13	11500	3	35	3	Medium	3
14	15000	5	15	1	High	3
15	8000	5	55	1	Medium	3

**Recall: Response is average rank (smaller is better)**

# Prediction profiler plots---JMP Demo



# Verification run: Clear, easy to read, cheap



# So, what are the current best practices?

Current best practices are:

- **Problem-driven**
- **Three-level, where possible**
- **Multi-objective**
- **Supersaturated**

