

Generalized Linear Mixed Models

What, Why and How

Timeline for GLMMs

- ☐ Mixed model issues appear in literature in 1930s (Yates)
- ☐ Mixed models named as such appear in 1950s
- ☐ Widespread applicability of mixed models not fully appreciated until 1980s
- ☐ Viable mixed model software available in 1990s
- ☐ Generalized linear models appear in literature in 1970s
- ☐ Generalized + mixed linear model literature in 1990s
- ☐ Viable GLMM software appears mid-to-late 2000s
- ☐ Appreciation of widespread applicability of GLMM still a work in progress

Setting for Statistical Models

Classic Statistical
Model Format:
response =
systematic + random/residual

		systematic/explanatory		random/residual	
		Categorical (ANOVA) $\mu + \tau_i$	continuous (regression) $\beta_0 + \beta_1 X$	model effects	Residual error structure (e.g. serial/spatial)
response variable					
Gaussian (normal)					
Non- Gaussian	discrete proportion binomial multinomial				
	continuous proportion beta				
	count Poisson negative Binomial				
	time to event				
	etc....				

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response variable		$\mu + \tau_i$	$\beta_0 + \beta_1 X$	model effects	
Gaussian (normal)		Linear Model classical ANOVA and regression			
Non-Gaussian	discrete proportion binomial multinomial				
	continuous proportion beta				
	count Poisson negative Binomial				
	time to event				
	etc....				

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Gaussian (normal)		Linear Mixed Model (LMM)			
Non-Gaussian	discrete proportion binomial multinomial				
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response variable		$\mu + \tau_i$	$\beta_0 + \beta_1 X$		
Non-Gaussian	Gaussian (normal)	Generalized Linear Mixed Model (GLMM)			
	discrete proportion binomial multinomial				
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	etc....				

Defining Elements of GLMM

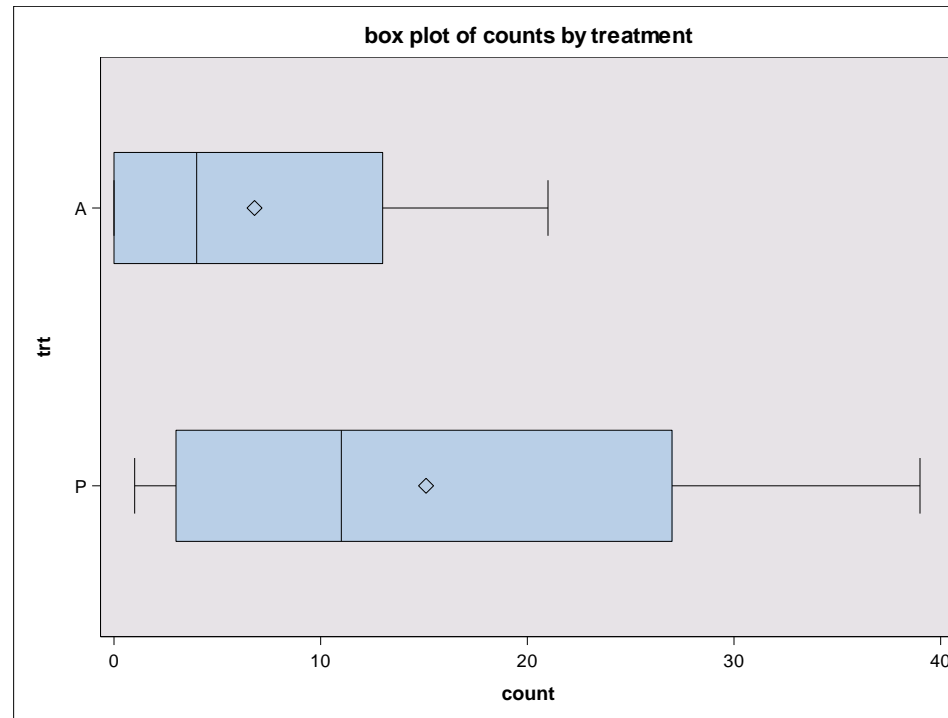
- ❑ Link function = linear predictor
- ❑ $\eta = g(\mu) = X\beta + Zb$
- ❑ β denotes fixed (ANOVA or regression) effects
- ❑ b denotes random model effects, assume $b \sim N(0, G)$
- ❑ y denotes observations
- ❑ $y|b \sim \mathcal{D}(\mu, \Sigma)$
- ❑ linear predictor is mixed model; \mathcal{D} and link function accommodate non-Gaussian data

Why GLMMs - Motivating Examples

- Main issues are target of inference & accuracy of inferential statistics
- Example 1
- Paired Comparison Experiment:
 - a.k.a. Randomized Complete Block Design
 - 10 Pairs / Blocks
 - 2 Treatments – “Treatment 0”, “Treatment 1”
e.g. **“control”** & **“test”**; **“A”** & **“P”**
 - Response: **count**
e.g. number of events / claims / defects = 0,1,2,...

Example 1: The Data

Obs	clinic	trt	count
1	1	A	7
2	1	P	5
3	2	A	1
4	2	P	1
5	3	A	13
6	3	P	15
7	4	A	5
8	4	P	7
9	5	A	3
10	5	P	32
11	6	A	0
12	6	P	1
13	7	A	21
14	7	P	21
15	8	A	0
16	8	P	3
17	9	A	18
18	9	P	39
19	10	A	0
20	10	P	27

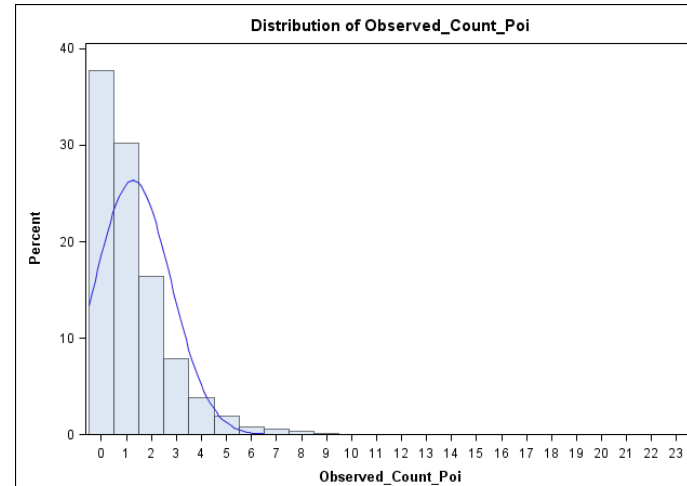


Treatment	Mean	Variance	Median
A	6.80	61.7	7.86
P	15.1	193.9	13.92

Count Distributions?

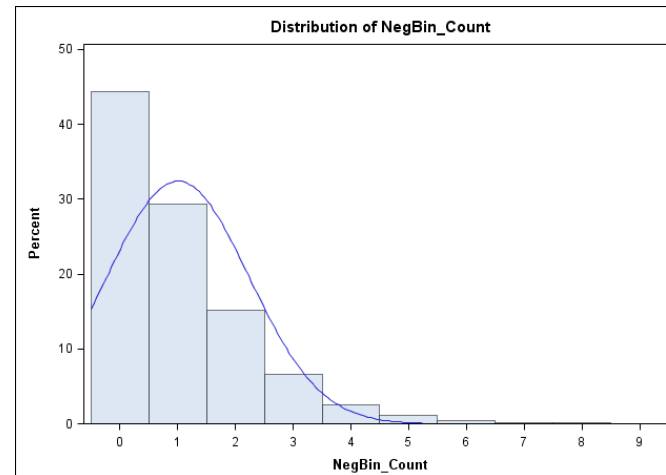
Poisson

$$\lambda = 1$$



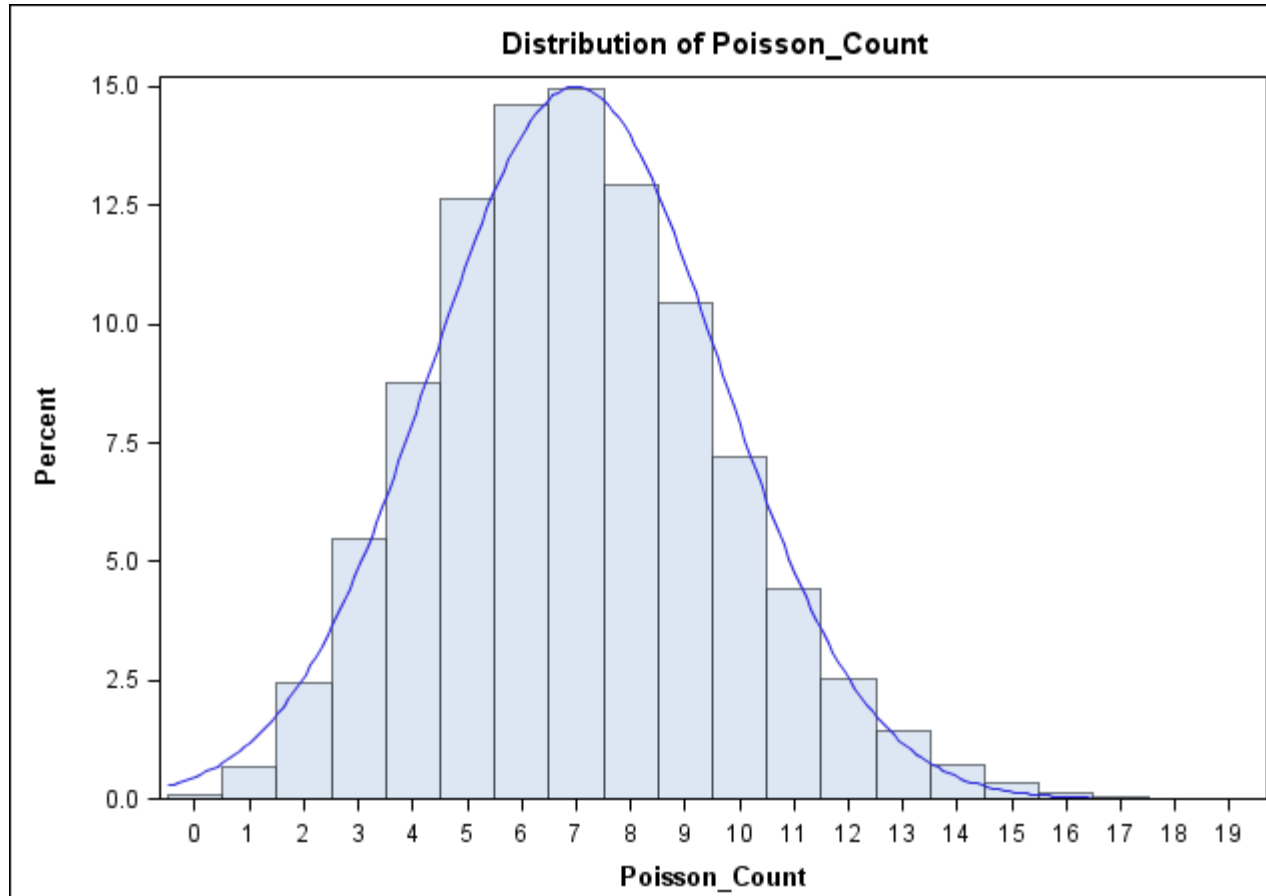
**Negative
Binomial**

$$\lambda = 1; \phi = 0.5$$



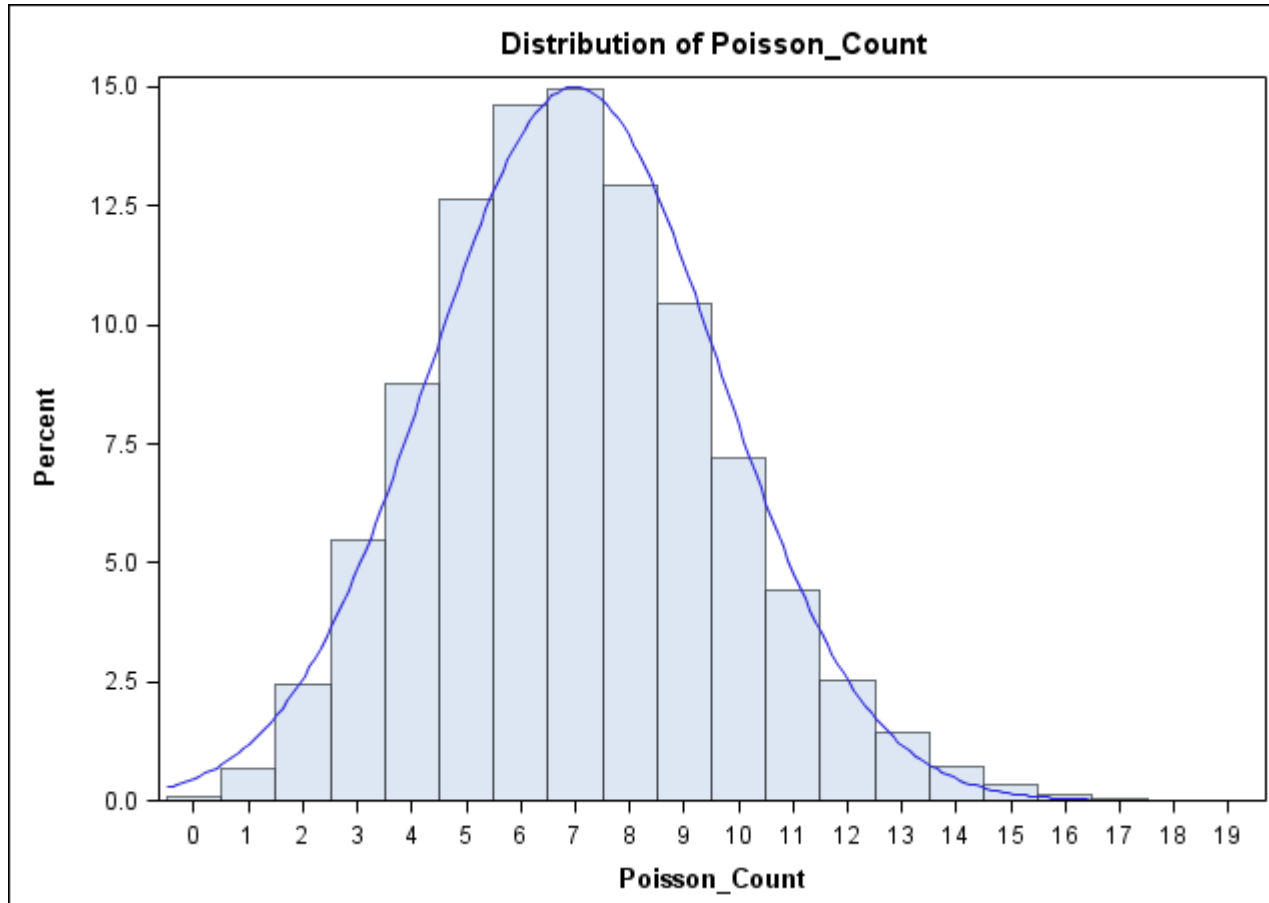
Count Distributions?

Poisson
 $\lambda = 7$



Count Distributions?

Poisson
 $\lambda = 7$



Doesn't this imply
count \sim approx Normal
 \Rightarrow **ANOVA** with **count** as
response variable okay?

Linear Model for RCBD Count Data

- ❑ ANOVA → “general” linear model for RCBD
- ❑ Model: **count = intercept + trt + block + resid**

$$count_{ij} = \mu + \tau_i + b_j + e_{ij}$$

- ❑ You can implement ANOVA using linear mixed model (LMM) software

Example 1 - Analysis 1 – ANOVA

Source	d.f.	MS	F	Pr>F
block	9	180.9		
treatment	1	344.5	4.61	0.0603
error	9	74.7		
Total	19			

Estimates of Treatment Means

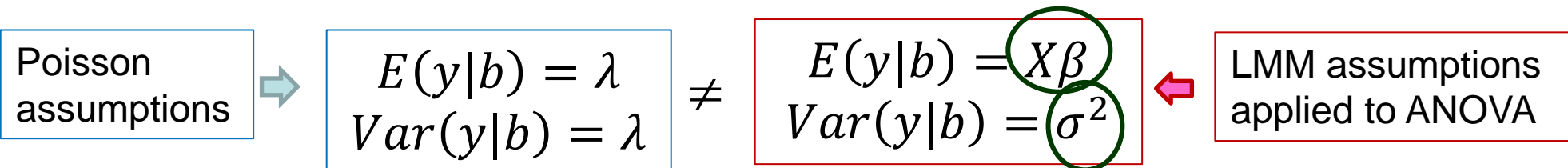
Treatment	Estimate	Std Error
A	6.8	2.73
P	15.1	2.73

??Poisson → mean = variance??

Neg Bin → mean \propto variance

Problems using LMM ($y = X\beta + Zb + e$) for count data

- Assumes $X\hat{\beta}$ estimates $E(y) = \lambda$
- $\hat{\lambda}$ must be ≥ 0
- No guarantee $X\hat{\beta} \geq 0$
e.g. regression provides easy examples
- Logical issues



- with count data, “Residual” has no meaning
- We need a better approach

Possible Models – Count Data

- Normal Approximation
 - $y_{ij} = \mu + \tau_i + b_j + e_{ij}; b_j \text{ iid } N(0, \sigma_b^2); e_{ij} \text{ iid } N(0, \sigma^2)$
 - standard ANOVA model
- Variance Stabilizing Transformation
 - $\log(y_{ij}) = \mu + \tau_i + b_j + e_{ij}$ -- $\log(y_{ij} + 1)$ if there are zeros
 - $\sqrt{y_{ij} + 3/8} = \mu + \tau_i + b_j + e_{ij}$
 - variance stabilizing transformation – standard pre-GLMM
- Poisson Generalized Linear Mixed Model (GLMM)
 - $y_{ij} | b_j \sim \text{Poisson}(\lambda_{ij}); b_j \text{ iid } N(0, \sigma_b^2)$
 - $\eta_{ij} = \log(\lambda_{ij}) = \eta + \tau_i + b_j$ “Naive model”
 - $\eta_{ij} = \log(\lambda_{ij}) = \eta + \tau_i + b_j + u_{ij}; u_{ij} \text{ iid } N(0, \sigma_u^2)$ better – we’ll see why

Determining an Appropriate Model

Types of Blocked Designs

- Multi-location, multi-center, multi-clinic
- Matched pairs
- Before and after on same subject
- Field plots
- etc.

motivating a
better approach

Visualization

Treatment design	
Trt 1	Trt 2

"Experiment" (study) design		
Block	Unit	
1		
2		
3		
4		
....		
10		

Full design		
Block	Unit	
1	trt 1	trt 2
2	trt 2	trt 1
3	trt 2	trt 1
4	trt 1	trt 2
....		
10	trt 2	trt 1

Repurposed ANOVA

Experiment		Treatment		Combined	
Source	d.f.	Source	d.f.	Source	d.f.
block	9			block	9
		trt	$2-1=1$	trt	1
unit(block)	$10 \times (2-1) = 10$	“parallels” (Yates, Fisher, 1935)	18	unit(block) trt a.k.a. “residual” a.k.a “block x trt”	$10-1=9$
Total	19	Total	19	Total	19

Repurposed ANOVA and Sensible Model

sensible model → one-to-one ANOVA effect – model parameter match

LMM Linear predictor for blocked design includes **block** and **treatment** effect
However, LMM also accounts for unit variation. If GLMM mimics LMM literally, **trouble**

combined		model		
Source	d.f.	LMM	naive GLMM	GLMM accounting for unit effect
block	9	b_j	b_j	b_j
treatment	1	τ_j	τ_j	τ_j
unit(block) block x trt “residual”	9	e_{ij} or σ^2	→ overdispersion likely	bt_{ij} (or u_{ij})
total	19			

overdispersion: model fails to adequately account for variation in data

usual consequence: confidence intervals too narrow; type I error rate ↑

Repurposed ANOVA to determine appropriate model

Combined	
Source	d.f.
block	9
trt	1
unit(block) trt a.k.a. “residual” a.k.a “block x trt”	10-1=9
Total	19

$$b_j \sim N(0, \sigma_b^2)$$

$$unit_{ij} \equiv bt_{ij} \sim N(0, \sigma_u^2)$$

$$y_{ij} | b_j, bt_{ij} \sim \text{Poisson}(\lambda_{ij})$$

resulting linear predictor

$$\begin{aligned} \eta_{ij} &= \log(\lambda_{ij}) \\ &= \eta + \tau_i + b_j + bt_{ij} \end{aligned}$$

resulting estimate of λ_i is

$$\exp(\hat{\eta} + \hat{\tau}_i)$$

(called **inverse link**)

Distributions imply Target of Inference

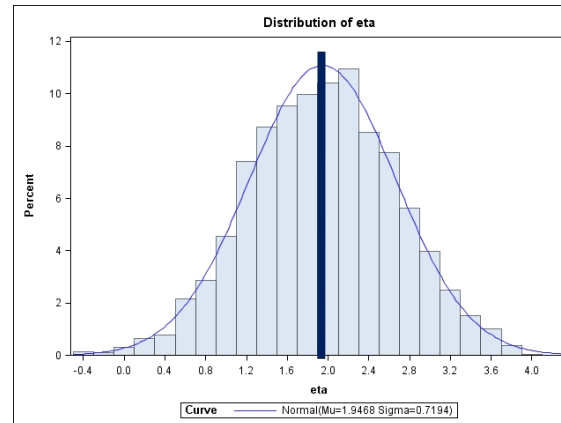
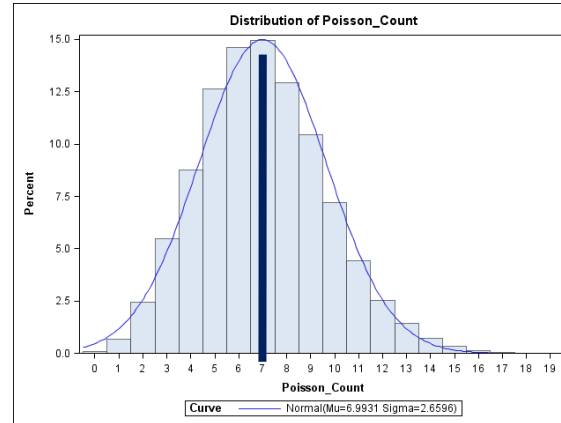
$$Y | b \sim \text{Poisson}(\lambda = 7)$$

$$b \sim NI(0, \sigma_b^2)$$

$$g(\lambda | b) = \eta + \tau_i + b_j + (bt)_{ij}$$

$$g(\lambda | b) = \eta = \log(\lambda | b)$$

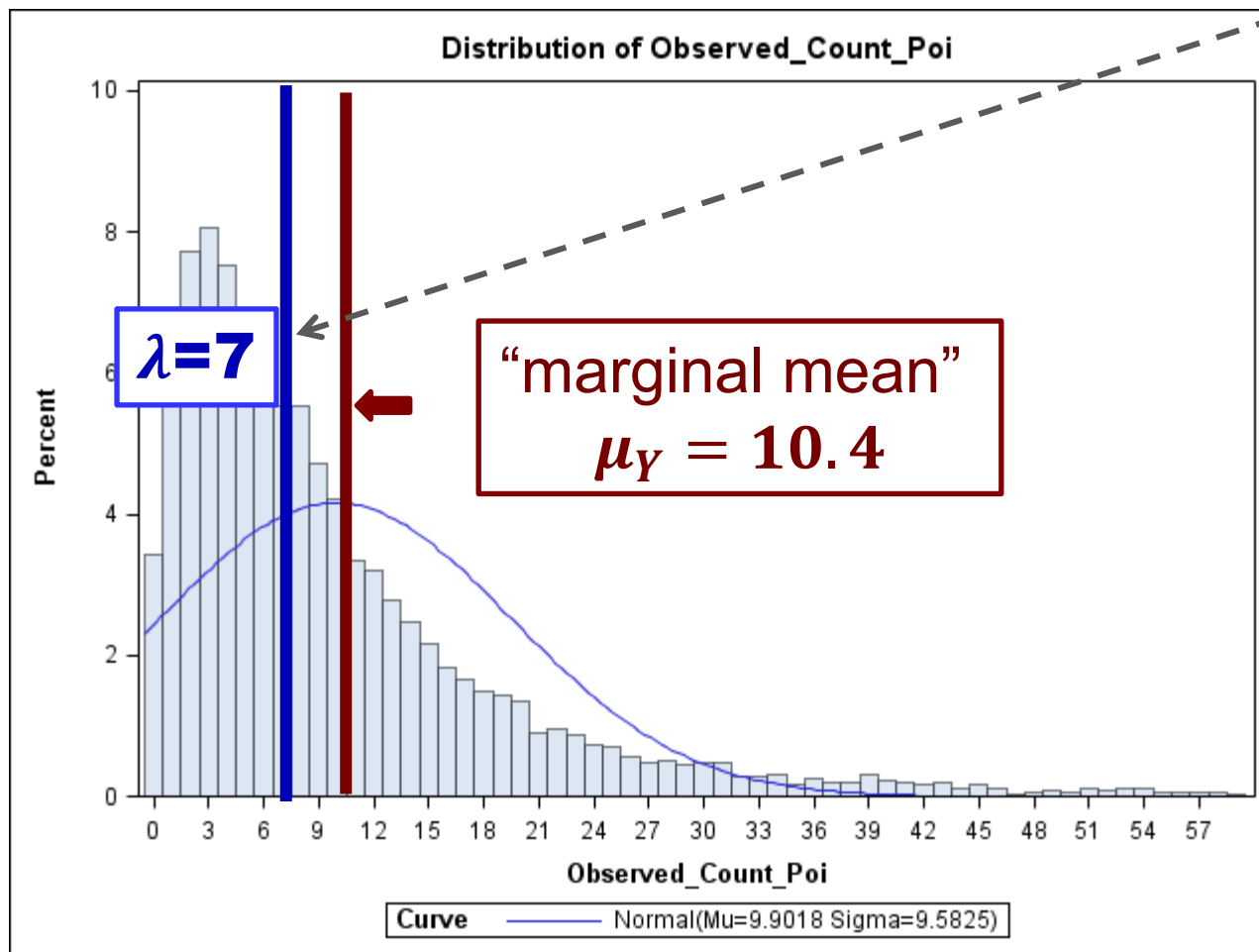
$$\lambda = 7 \Rightarrow \eta \sim N(1.95, \sigma_b^2)$$



These distributions define the GLMM, but you cannot observe either directly

Resulting Distribution of the Observations

$$block \sim N(0, 0.8) \quad y | block \sim \text{Poisson}(\lambda = 7)$$



classical ANOVA
will target μ_Y
overestimates λ

GLMM
will target the true λ

transformations
????
(somewhere in between?)

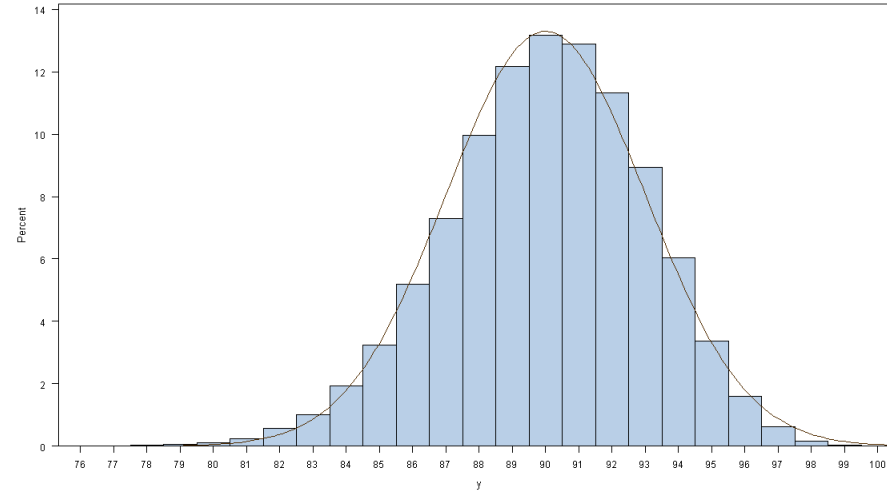
Same Issue Occurs with Binomial Data

- Example 2
- Paired Comparison, a.k.a. blocked design
- Two treatments, ***B*** blocks
- Response variable is Y “successes” out of N observations on given trt-block
- Hence $y_{ij}|b_j \sim \text{Binomial}(N_{ij}, p_{ij})$ where p denotes probability of success

Normal Approximation – Isn't this okay?

$$\pi = 0.90$$

$$N = 100$$



$$N = 100 \quad \text{therefore} \quad Y_i / N \sim \text{approx } N\left(\pi, \frac{\pi(1-\pi)}{N}\right)$$

$$\text{ANOVA via LMM:} \quad Y_i / N = p_i = \mu + \tau_i + b_j + e_i$$

Repurposed ANOVA to determine appropriate binomial GLMM

Combined	
Source	d.f.
block	9
trt	1
unit(block) trt a.k.a. “residual” a.k.a “block x trt”	10-1=9
Total	19

$$b_j \sim N(0, \sigma_b^2)$$

$$unit_{ij} \equiv bt_{ij} \sim N(0, \sigma_u^2)$$

$$y_{ij} | b_j, bt_{ij} \sim \text{Binomial}(N_{ij}, p_{ij})$$

resulting linear predictor

$$\begin{aligned} \eta_{ij} &= \text{logit}(p_{ij}) \\ &= \log\left(\frac{p_{ij}}{1 - p_{ij}}\right) \\ &= \eta + \tau_i + b_j + bt_{ij} \end{aligned}$$

resulting estimate of p_i is

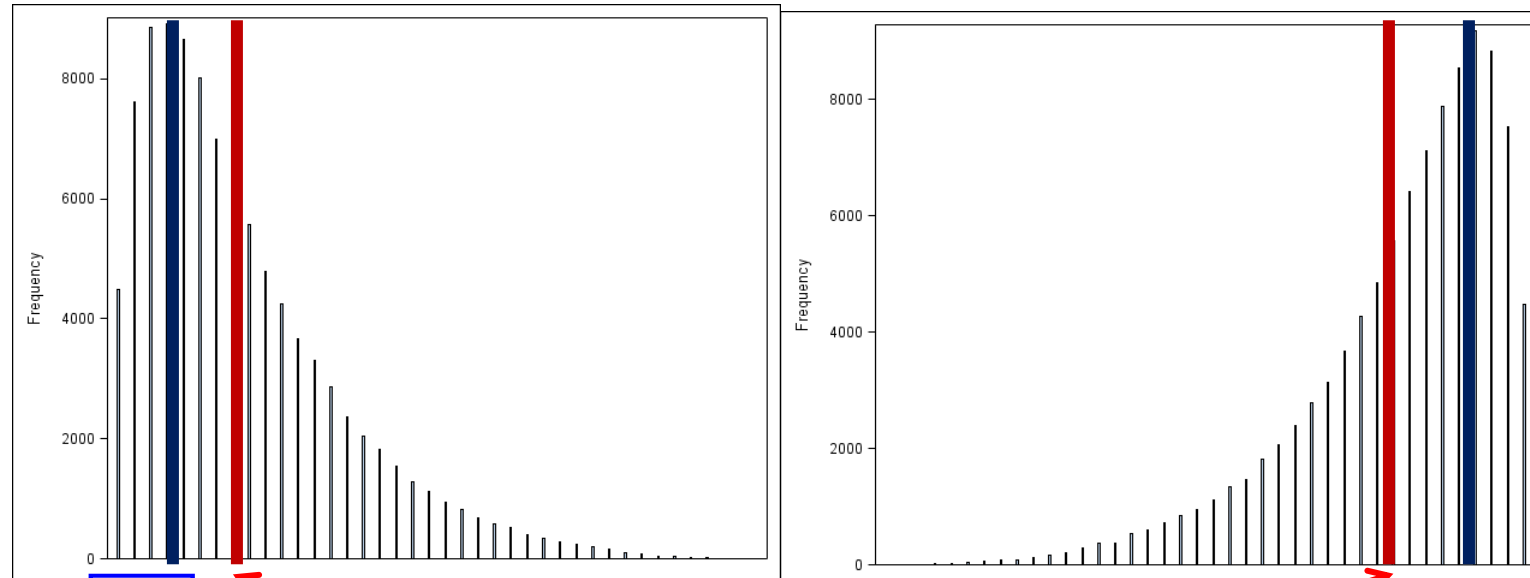
$$\frac{1}{1 + \exp[-(\hat{\eta} + \hat{\tau}_i)]}$$

logistic inverse link

Consequence of Normal Approximation vs. GLMM

If $p < 0.5$

If $p > 0.5$



p_1

normal approx (ANOVA)

estimates marginal μ_Y

binomial GLMM

estimates p_i

p_2

- **IMPORTANT:**
what is your target & why?
- **IF** you get this wrong:
 - inaccurate estimates of p_i & odds-ratio
 - loss of power

List of examples

□ Elizabeth will cover

- **Example 11.5 in *SAS for Mixed Models* (Stroup, et al. 2018).** Multi-center clinical trial with binomial data (from Beitler and Landis, *Biometrics*, 1985)
- **Example 12.3 in *SAS for Mixed Models*.** Manufacturing; data from multiple lots; response variable is number of micro-sites (discrete count). Random coefficient regression

□ Also in *SAS for Mixed Models (2018)*

- **Example 11.6.** Estimate genetic parameters from count data
- **Example 13.3.** Split-plot (multi-level mixed model) with count data
- **Example 13.4.** Repeated measures (longitudinal data) with binomial data

□ Situations calling for GLMMs

- **designs:** blocks; matched pairs; multiple sources, batches, lots, locations, times; multi-level, split-plot; repeated measures; longitudinal – **all call for mixed model**
- If, in addition, you have **non-Gaussian response variable(s), then GLMM**