



Version 15

Design of Experiments Guide

*"The real voyage of discovery consists not in seeking new
landscapes, but in having new eyes."*

Marcel Proust

JMP, A Business Unit of SAS
SAS Campus Drive
Cary, NC 27513

15.1

The correct bibliographic citation for this manual is as follows: SAS Institute Inc. 2020. *JMP® 15 Design of Experiments Guide*. Cary, NC: SAS Institute Inc.

JMP® 15 Design of Experiments Guide

Copyright © 2020, SAS Institute Inc., Cary, NC, USA

All rights reserved. Produced in the United States of America.

U.S. Government License Rights; Restricted Rights: The Software and its documentation is commercial computer software developed at private expense and is provided with RESTRICTED RIGHTS to the United States Government. Use, duplication or disclosure of the Software by the United States Government is subject to the license terms of this Agreement pursuant to, as applicable, FAR 12.212, DFAR 227.7202-1(a), DFAR 227.7202-3(a) and DFAR 227.7202-4 and, to the extent required under U.S. federal law, the minimum restricted rights as set out in FAR 52.227-19 (DEC 2007). If FAR 52.227-19 is applicable, this provision serves as notice under clause (c) thereof and no other notice is required to be affixed to the Software or documentation. The Government's rights in Software and documentation shall be only those set forth in this Agreement.

SAS Institute Inc., SAS Campus Drive, Cary, North Carolina 27513-2414.

September 2019

February 2020

SAS® and all other SAS Institute Inc. product or service names are registered trademarks or trademarks of SAS Institute Inc. in the USA and other countries. ® indicates USA registration.

Other brand and product names are trademarks of their respective companies.

SAS software may be provided with certain third-party software, including but not limited to open-source software, which is licensed under its applicable third-party software license agreement. For license information about third-party software distributed with SAS software, refer to <http://support.sas.com/thirdpartylicenses>.

Get the Most from JMP

Whether you are a first-time or a long-time user, there is always something to learn about JMP.

Visit JMP.com to find the following:

- live and recorded webcasts about how to get started with JMP
- video demos and webcasts of new features and advanced techniques
- details on registering for JMP training
- schedules for seminars being held in your area
- success stories showing how others use JMP
- a blog with tips, tricks, and stories from JMP staff
- a forum to discuss JMP with other users

<https://www.jmp.com/getstarted>

Contents

Design of Experiments Guide

1	Learn about JMP	19
	Documentation and Additional Resources	
	Formatting Conventions	21
	JMP Help	22
	JMP Documentation Library	22
	Additional Resources for Learning JMP	28
	Tutorials	28
	Sample Data Tables	28
	Learn about Statistical and JSL Terms	29
	Learn JMP Tips and Tricks	29
	Tooltips	29
	JMP User Community	30
	Free Online Statistical Thinking Course	30
	New User Welcome Kit	30
	Statistics Knowledge Portal	30
	JMP Training	30
	JMP Books by Users	31
	The JMP Starter Window	31
	Technical Support	31
2	Introduction to DOE	33
	Overview of Design of Experiment Platforms	
3	Starting Out with DOE	39
	Example and Key Concepts	
	Overview of Experimental Design and the DOE Workflow	41
	The Coffee Strength Experiment	42
	Define the Study and Goals	43
	Create the Design	45
	Run the Experiment	51
	Analyze the Data	51
	The DOE Workflow: Describe, Specify, Design	57
	Define Responses and Factors	58

Specify the Model	59
Generate the Design	60
Evaluate the Design	61
Make the Table	61
Principles and Guidelines for Experimental Design	62
Effect Hierarchy	62
Effect Heredity	62
Effect Sparsity	63
Center Points, Replicate Runs, and Testing	63

4 Custom Designs	65
Construct Designs That Meet Your Needs	
Overview of Custom Design	67
Example of a Custom Design	68
Create the Design	70
Analyze the Data	80
Custom Design Window	86
Responses	86
Factors	88
Define Factor Constraints	95
Model	98
Alias Terms	100
Design Generation	101
Design	103
Design Evaluation	103
Output Options	104
Custom Design Options	107
Description of Options	107
Simulate Responses	110
Save X Matrix	112
Number of Starts	115
Design Search Time	115
Set Delta for Power	116
Technical Details for Custom Designs	117
Designs with Randomization Restrictions	117
Covariates with Hard-to-Change Levels	126
Numbers of Whole Plots and Subplots	126
Optimality Criteria	127
D-Efficiency	132
Coordinate-Exchange Algorithm	132

5	Examples of Custom Designs	133
	Perform Experiments That Meet Your Needs	
	Screening Experiments	135
	Design That Estimates Main Effects Only	135
	Design That Estimates All Two-Factor Interactions	138
	Design That Avoids Aliasing of Main Effects and Two-Factor Interactions	141
	Supersaturated Screening Designs	144
	Design for Fixed Blocks	152
	Response Surface Experiments	156
	Response Surface Design	156
	Response Surface Design with Flexible Blocking	163
	Comparison of a D-Optimal and an I-Optimal Response Surface Design	166
	Response Surface Design With Constraints and Categorical Factor	170
	Mixture Experiments	176
	Mixture Design with Nonmixture Factors	176
	Mixture of Mixtures Design	181
	Experiments with Covariates	185
	Design with Fixed Covariates	185
	Design with Hard-to-Change Covariates	188
	Design with a Linear Time Trend	192
	Experiments with Randomization Restrictions	195
	Split-Plot Experiment	195
	Two-Way Split-Plot Experiment	200
	Experiments for Robust Process and Product Design	207
6	Augment Designs	211
	Example Using the Augment Design Platform	213
	Analyze the Augmented Design	215
	Launch the Augment Design Platform	219
	Augment Design Window	220
	Factors	220
	Define Factor Constraints	221
	Augmentation Choices	224
	Augment Design Platform Options	225
	Additional Examples of Augmentation Choices	229
	Replicate a Design	229
	Add Center Points to a Design	231
	Create a Foldover Design	231
	Add Axial Points to a Design	232

Space Filling Augmentation	233
Statistical Details	235
7 Definitive Screening Designs	237
Overview of Definitive Screening Design	239
Examples of Definitive Screening Designs	241
Definitive Screening Design	242
Comparison with a Fractional Factorial Design	244
Definitive Screening Design with Blocking	246
Comparison of a Definitive Screening Design with a Plackett-Burman Design	250
Definitive Screening Design Window	253
Responses	254
Factors	256
Design Options	258
Design	260
Design Evaluation	261
Output Options	261
Definitive Screening Design Options	263
Simulate Responses	265
Technical Details	268
Structure of Definitive Screening Designs	268
Analysis of Experimental Data	270
8 The Fit Definitive Screening Platform	273
Analyze Data from Definitive Screening Experiments	
Overview of the Fit Definitive Screening Platform	275
Identification of Active Effects in DSDs	275
Effective Model Selection for DSDs	275
Example of the Fit Definitive Screening Platform	276
Fit the Model	276
Examine Results	277
Reduce the Model	279
Launch the Fit Definitive Screening Platform	280
Fit Definitive Screening Report	281
Stage 1 - Main Effect Estimates	281
Stage 2 - Even Order Effect Estimates	282
Combined Model Parameter Estimates	283
Main Effects Residual Plots	284
Prediction Profiler	284
Fit Definitive Screening Platform Options	284

Technical Details for the Fit Definitive Screening Platform	285
The Effective Model Selection for DSDs Algorithm	285
9 Screening Designs	289
Overview of Screening Designs	291
Underlying Principles	291
Analysis of Screening Design Results	292
Examples of Screening Designs	292
Compare a Fractional Factorial Design and a Main Effects Screening Design	292
Main Effects Screening Design where No Standard Design Exists	300
Screening Design Window	302
Responses	304
Factors	306
Choose Screening Type	308
Choose from a List of Fractional Factorial Designs	308
Design Type	309
Resolution as a Measure of Confounding	312
Display and Modify Design	313
Main Effects Screening Designs	316
Design Generation	317
Design	317
Design Evaluation	317
Output Options	318
Make Table	319
Screening Design Options	320
Additional Examples of Screening Designs	323
Modify Generating Rules in a Fractional Factorial Design	323
Plackett-Burman Design	329
10 The Fit Two Level Screening Platform	335
Analyze Data from Screening Experiments	
Overview of the Fit Two Level Screening Platform	337
Example of Fit Two Level Screening Platform	337
Launch the Fit Two Level Screening Platform	340
The Screening Report	341
Contrasts	342
Half Normal Plot	342
Make or Run Model	343
Additional Fit Two Level Screening Analysis Examples	343
Analyze a Plackett-Burman Design	343

Effect Heredity Example	344
Analyze a Supersaturated Design	346
Technical Details for the Fit Two Level Screening Platform	348
Order of Effect Entry	348
Fit Two Level Screening as an Orthogonal Rotation	348
Lenth's Pseudo-Standard Error	349
Lenth t-Ratios	349
11 Response Surface Designs	353
Overview of Response Surface Designs	355
Example of a Response Surface Design	357
Construct a Box-Behnken Design	358
Analyze the Experimental Data	359
Explore Optimal Settings	361
Response Surface Design Window	364
Responses	365
Factors	367
Choose a Design	368
Specify Output Options	370
Make Table	371
Response Surface Design Options	372
12 Full Factorial Designs	375
Overview of Full Factorial Design	377
Example of a Full Factorial Design	377
Construct the Design	377
Analyze the Experimental Data	379
Full Factorial Design Window	383
Responses	383
Factors	385
Select Output Options	388
Make Table	389
Full Factorial Design Options	390
13 Mixture Designs	393
Designs for Formulations	
Overview of Mixture Designs	395
Mixture Design Window	395
Responses	396

Factors	398
Ternary Plot Overview	399
Examples of Mixture Design Types	400
Optimal Mixture Design	400
Simplex Centroid Design	402
Simplex Lattice Design	403
Extreme Vertices Design	405
ABCD Design	409
Space Filling Design	410
Fitting Mixture Designs	413
Whole Model Tests and Analysis of Variance Reports	414
Understanding Response Surface Reports	414
A Chemical Mixture Example	414
Create the Design	414
Analyze the Mixture Model	417
The Prediction Profiler	417
The Mixture Profiler	418
14 Taguchi Designs	421
Designs with Noise Factors	
Overview of Taguchi Designs	423
Example of a Taguchi Design	423
Taguchi Design Window	427
Response	428
Factors	428
Choose Inner and Outer Array Designs	429
Display Coded Design	429
Make the Design Table	430
Taguchi Design Options	430
15 Evaluate Designs	433
Explore Properties of Your Design	
Overview of the Evaluate Design Platform	435
Example of Evaluate Design	436
Assessing the Impact of Lost Runs	436
Evaluating Power Relative to a Specified Model	443
Evaluate Design Launch Window	445
Evaluate Design Window	445
Factors	447
Model	447
Alias Terms	448

Design	449
Design Evaluation	449
Power Analysis	449
Prediction Variance Profile	456
Fraction of Design Space Plot	457
Prediction Variance Surface	458
Estimation Efficiency	460
Alias Matrix	462
Color Map on Correlations	463
Design Diagnostics	465
Evaluate Design Options	467
16 Compare Designs	469
Compare and Evaluate Designs Simultaneously	
Overview of the Compare Designs Platform	471
Examples of Comparing Designs	471
Designs of Same Run Size	471
Designs of Different Run Sizes	476
Split Plot Designs with Different Numbers of Whole Plots	480
Launch the Compare Designs Platform	484
Compare Designs Window: Specify Model and Alias Terms	486
Reference Design	486
Factors	486
Model	486
Alias Terms	487
Compare Designs Window: Design Evaluation	487
Power Analysis	488
Prediction Variance Profile	491
Fraction of Design Space Plot	493
Relative Estimation Efficiency	495
Alias Matrix Summary	496
Absolute Correlations	498
Design Diagnostics	501
Compare Designs Options	503
17 Prospective Sample Size and Power	505
Launch the Sample Size and Power Platform	508
One Sample Mean Calculator	509
Examples of the One Sample Mean Calculator	510
One Sample Mean Calculator Fields	514

Statistical Details for the One Sample Mean Calculator	515
Two Sample Means Calculator	516
Example of the Two Sample Means Calculator	516
Two Sample Means Calculator Fields	517
Statistical Details for the Two Sample Means Calculator	518
k Sample Means Calculator	519
Example of the k Sample Means Calculator	519
k Sample Means Calculator Fields	521
Statistical Details for the k Sample Means Calculator	522
One Sample Standard Deviation Calculator	523
Example of the One Sample Standard Deviation Calculator	523
One Sample Standard Deviation Calculator Fields	524
Statistical Details for the One Sample Standard Deviation Calculator	525
One Sample Proportion Calculator	526
Example of the One Sample Proportion Calculator	527
One Sample Proportion Calculator Fields	528
Statistical Details for the One Proportion Calculator	529
Two Sample Proportions Calculator	530
Example of the Two Sample Proportions Calculator	530
Two Sample Proportions Calculator Fields	531
Statistical Details for the Two Proportions Calculator	532
Counts per Unit Calculator	533
Example of the Counts per Unit Calculator	533
Counts per Unit Calculator Fields	534
Statistical Details for the Counts per Unit Calculator	535
Sigma Quality Level Calculator	535
Example of the Sigma Quality Level Calculator	536
Example Computing the Number of Defects	536
Sigma Quality Level Calculator Fields	537
Statistical Details for the Sigma Quality Calculator	537
Reliability Test Plan Calculator	538
Example of the Reliability Test Plan Calculator	538
Reliability Test Plan Calculator Fields	539
Statistical Details for the Reliability Test Plan Calculator	541
Reliability Demonstration Calculator	542
Example of the Reliability Demonstration Calculator	542
Reliability Demonstration Calculator Fields	544
Statistical Details for the Reliability Demonstration Calculator	545

18	Choice Designs	547
	Create a Design for Selecting Preferred Product Profiles	
	Overview of Choice Designs	549
	Choice Design Terminology	549
	Bayesian D-Optimality	550
	Example of a Choice Design	550
	Example of a Choice Design with Analysis	553
	Create a Choice Design for a Pilot Study	553
	Analyze the Pilot Study Data	555
	Design the Final Choice Experiment Using Prior Information	556
	Run the Design and Analyze the Results	558
	Choice Design Window	561
	Attributes	561
	Model	562
	Design Generation	563
	Design	564
	Make Table	566
	Choice Design Options	566
	Technical Details	567
	Bayesian D-Optimality and Design Construction	567
	Utility-Neutral and Local D-Optimal Designs	568
19	MaxDiff Design	569
	Create a Design for Selecting Best and Worst Items	
	Overview of the MaxDiff Design Platform	571
	Example of a MaxDiff Design	571
	MaxDiff Design Launch Window	575
	MaxDiff Window	575
	Design Options Outline	575
	Design Outline	576
	Pairwise Incidence Matrix	576
	Make Table	576
	MaxDiff Options	576
20	Covering Arrays	579
	Detecting Component Interaction Failures	
	Overview of Covering Arrays	581
	Example of a Covering Array with No Factor Level Restrictions	583
	Create the Design	583
	Analyze the Experimental Data	585
	Example of a Covering Array with Factor Level Restrictions	587

Create the Design	588
Analyze the Experimental Data	592
Covering Array Window	594
Factors	594
Restrict Factor Level Combinations	595
Design	598
Metrics	600
Output Options	601
The Covering Array Data Table	601
Covering Array Options	603
Technical Details for Covering Arrays	604
Algorithm for Optimize	604
Formulas for Metrics	604
21 Space-Filling Designs	607
Overview of Space-Filling Designs	609
Space Filling Design Window	610
Responses	610
Factors	612
Define Factor Constraints	614
Space Filling Design Methods	617
Design	618
Design Diagnostics	618
Design Table	618
Space Filling Design Options	618
Sphere-Packing Designs	621
Creating a Sphere-Packing Design	621
Visualizing the Sphere-Packing Design	622
Latin Hypercube Designs	624
Creating a Latin Hypercube Design	624
Visualizing the Latin Hypercube Design	625
Uniform Designs	627
Comparing Sphere-Packing, Latin Hypercube, and Uniform Methods	629
Minimum Potential Designs	630
Maximum Entropy Designs	633
Gaussian Process IMSE Optimal Designs	634
Fast Flexible Filling Designs	636
FFF Optimality Criterion	636
Set Average Cluster Size	637

Constraints	637
Creating and Viewing a Constrained Fast Flexible Filling Design	638
Creating a Space-Filling Design for a Map Shape	640
Example of a Sphere-Packing Design	645
Create the Sphere-Packing Design for the Borehole Data	645
Guidelines for the Analysis of Deterministic Data	646

22 Accelerated Life Test Designs

Designing Experiments for Accelerated Life Tests	
Overview of Accelerated Life Test Designs	653
Example of an Accelerated Life Test Design	653
Obtain Prior Estimates	654
Enter Basic Specifications	655
Enter Prior Information and Remaining Specifications	656
Make the Design	658
Run Simulations	660
Save Design	661
Example of Augmenting an Accelerated Life Test Design	661
Accelerated Life Test Plan Window	666
Specify the Design Structure	667
Specify Acceleration Factors	667
Specify Design Details	668
Review Balanced Design Diagnostics and Update Specifications	671
Create and Assess the Optimal Design	672
Create Design Tables	674
Platform Options	674
Statistical Details for the ALT Design Platform	675

23 Nonlinear Designs

Overview of Nonlinear Designs	681
Examples of Nonlinear Designs	682
Create a Nonlinear Design with No Prior Data	682
Augment a Design Using Prior Data	687
Create a Design for a Binomial Response	691
Nonlinear Design Launch Window	695
Nonlinear Design Window	696
Factors	696
Parameters	697
Design Generation	697
Design	698

Make Table or Augment Table	698
Nonlinear Design Options	698
Statistical Details for Nonlinear Designs	699
Nonlinear Models	699
Radial-Spherical Integration of the Optimality Criterion	700
Finding the Optimal Design	700
24 Balanced Incomplete Block Designs	703
Example of Balanced Incomplete Block Designs	705
Balanced Incomplete Block Design Launch Window	707
Balanced Incomplete Block Design Window	708
Specify Treatments, Blocks, and Make Design	708
Verify Design and Make Table	709
Balanced Incomplete Block Design Options	709
Statistical Details for Balanced Incomplete Block Designs	710
25 Group Orthogonal Supersaturated Designs	713
Create Group Orthogonal Supersaturated Designs	
Overview of Group Orthogonal Supersaturated Designs	715
Example of Group Orthogonal Supersaturated Designs	715
Launch the Group Orthogonal Supersaturated Designs Platform	717
Group Orthogonal Supersaturated Design Window	717
Specify Factors, Structure, and Make Design	718
Verify Design and Make Table	718
Group Orthogonal Supersaturated Design Options	719
Statistical Details for Group Orthogonal Supersaturated Designs	720
26 The Fit Group Orthogonal Supersaturated Platform	721
Analyze Data from Group Orthogonal Supersaturated Designs	
Overview of the Fit Group Orthogonal Supersaturated Platform	723
Example of Fit Group Orthogonal Supersaturated	723
Launch the Fit Group Orthogonal Supersaturated Platform	726
The Fit Group Orthogonal Supersaturated Report	727
Fit Group Orthogonal Supersaturated Platform Options	728
A Column Properties	729
Understanding Column Properties Assigned by DOE	
Adding and Viewing Column Properties	731
Response Limits	732
Response Limits Example	734
Editing Response Limits	737

Design Role	740
Design Role Example	740
Coding	742
Low and High Values	743
Coding Column Property and Center Polynomials	744
Coding Example	744
Assigning Coding	747
Mixture	750
Mixture Example	752
Factor Changes	755
Factor Changes Example	757
Value Order	759
Value Order Example	761
Assigning Value Order	762
Value Labels	763
Value Labels Example	763
RunsPerBlock	765
RunsPerBlock Example	766
ConstraintState	766
ConstraintState Example	767
B Technical Details	769
The Alias Matrix	771
Designs with Hard or Very Hard Factor Changes	771
Designs with If Possible Effects	772
Power Calculations	772
Power for a Single Parameter	772
Power for a Categorical Effect	774
Relative Prediction Variance	775
C References	777
D Technology License Notices	783

Chapter **1**

Learn about JMP

Documentation and Additional Resources

This chapter includes details about JMP documentation, such as book conventions, descriptions of each JMP document, the Help system, and where to find other support.

Contents

Formatting Conventions	21
JMP Help	22
JMP Documentation Library	22
Additional Resources for Learning JMP	28
Tutorials	28
Sample Data Tables	28
Learn about Statistical and JSL Terms	29
Learn JMP Tips and Tricks	29
Tooltips	29
JMP User Community	30
Free Online Statistical Thinking Course	30
New User Welcome Kit	30
Statistics Knowledge Portal	30
JMP Training	30
JMP Books by Users	31
The JMP Starter Window	31
Technical Support	31

Formatting Conventions

The following conventions help you relate written material to information that you see on your screen:

- Sample data table names, column names, pathnames, filenames, file extensions, and folders appear in Helvetica (or sans-serif online) font.
- Code appears in *Lucida Sans Typewriter* (or monospace online) font.
- Code output appears in *Lucida Sans Typewriter* italic (or monospace italic online) font and is indented farther than the preceding code.
- **Helvetica bold** formatting (or bold sans-serif online) indicates items that you select to complete a task:
 - buttons
 - check boxes
 - commands
 - list names that are selectable
 - menus
 - options
 - tab names
 - text boxes
- The following items appear in italics:
 - words or phrases that are important or have definitions specific to JMP
 - book titles
 - variables
- Features that are for JMP Pro only are noted with the JMP Pro icon . For an overview of JMP Pro features, visit <https://www.jmp.com/software/pro>.

Note: Special information and limitations appear within a Note.

Tip: Helpful information appears within a Tip.

JMP Help

JMP Help in the Help menu enables you to search for information about JMP features, statistical methods, and the JMP Scripting Language (or *JSL*). You can open JMP Help in several ways:

- Search and view JMP Help on Windows by selecting the **Help > JMP Help**.
- On Windows, press the F1 key to open the Help system in the default browser.
- Get help on a specific part of a data table or report window. Select the Help tool  from the **Tools** menu and then click anywhere in a data table or report window to see the Help for that area.
- Within a JMP window, click the **Help** button.

Note: The JMP Help is available for users with Internet connections. Users without an Internet connection can search all books in a PDF file by selecting **Help > JMP Documentation Library**. See “[JMP Documentation Library](#)” on page 22 for more information.

JMP Documentation Library

The Help system content is also available in one PDF file called *JMP Documentation Library*. Select **Help > JMP Documentation Library** to open the file. If you prefer searching individual PDF files of each document in the JMP library, download the files from <https://www.jmp.com/documentation>.

The following table describes the purpose and content of each document in the JMP library.

Document Title	Document Purpose	Document Content
<i>Discovering JMP</i>	If you are not familiar with JMP, start here.	Introduces you to JMP and gets you started creating and analyzing data. Also learn how to share your results.
<i>Using JMP</i>	Learn about JMP data tables and how to perform basic operations.	Covers general JMP concepts and features that span across all of JMP, including importing data, modifying columns properties, sorting data, and connecting to SAS.

Document Title	Document Purpose	Document Content
<i>Basic Analysis</i>	Perform basic analysis using this document.	Describes these Analyze menu platforms: • Distribution • Fit Y by X • Tabulate • Text Explorer Covers how to perform bivariate, one-way ANOVA, and contingency analyses through Analyze > Fit Y by X. How to approximate sampling distributions using bootstrapping and how to perform parametric resampling with the Simulate platform are also included.
<i>Essential Graphing</i>	Find the ideal graph for your data.	Describes these Graph menu platforms: • Graph Builder • Scatterplot 3D • Contour Plot • Bubble Plot • Parallel Plot • Cell Plot • Scatterplot Matrix • Ternary Plot • Treemap • Chart • Overlay Plot The book also covers how to create background and custom maps.
<i>Profilers</i>	Learn how to use interactive profiling tools, which enable you to view cross-sections of any response surface.	Covers all profilers listed in the Graph menu. Analyzing noise factors is included along with running simulations using random inputs.

Document Title	Document Purpose	Document Content
<i>Design of Experiments Guide</i>	Learn how to design experiments and determine appropriate sample sizes.	Covers all topics in the DOE menu.
<i>Fitting Linear Models</i>	Learn about Fit Model platform and many of its personalities.	Describes these personalities, all available within the Analyze menu Fit Model platform: • Standard Least Squares • Stepwise • Generalized Regression • Mixed Model • MANOVA • Loglinear Variance • Nominal Logistic • Ordinal Logistic • Generalized Linear Model

Document Title	Document Purpose	Document Content
<i>Predictive and Specialized Modeling</i>	Learn about additional modeling techniques.	Describes these Analyze > Predictive Modeling menu platforms: • Neural • Partition • Bootstrap Forest • Boosted Tree • K Nearest Neighbors • Naive Bayes • Support Vector Machines • Model Comparison • Make Validation Column • Formula Depot Describes these Analyze > Specialized Modeling menu platforms: • Fit Curve • Nonlinear • Functional Data Explorer • Gaussian Process • Time Series • Matched Pairs Describes these Analyze > Screening menu platforms: • Modeling Utilities • Response Screening • Process Screening • Predictor Screening • Association Analysis • Process History Explorer

Document Title	Document Purpose	Document Content
<i>Multivariate Methods</i>	Read about techniques for analyzing several variables simultaneously.	Describes these Analyze > Multivariate Methods menu platforms: • Multivariate • Principal Components • Discriminant • Partial Least Squares • Multiple Correspondence Analysis • Structural Equation Models • Factor Analysis • Multidimensional Scaling • Item Analysis Describes these Analyze > Clustering menu platforms: • Hierarchical Cluster • K Means Cluster • Normal Mixtures • Latent Class Analysis • Cluster Variables
<i>Quality and Process Methods</i>	Read about tools for evaluating and improving processes.	Describes these Analyze > Quality and Process menu platforms: • Control Chart Builder and individual control charts • Measurement Systems Analysis • Variability / Attribute Gauge Charts • Process Capability • Model Driven Multivariate Control Chart • Pareto Plot • Diagram • Manage Spec Limits

Document Title	Document Purpose	Document Content
<i>Reliability and Survival Methods</i>	Learn to evaluate and improve reliability in a product or system and analyze survival data for people and products.	Describes these Analyze > Reliability and Survival menu platforms: • Life Distribution • Fit Life by X • Cumulative Damage • Recurrence Analysis • Degradation • Destructive Degradation • Reliability Forecast • Reliability Growth • Reliability Block Diagram • Repairable Systems Simulation • Survival • Fit Parametric Survival • Fit Proportional Hazards
<i>Consumer Research</i>	Learn about methods for studying consumer preferences and using that insight to create better products and services.	Describes these Analyze > Consumer Research menu platforms: • Categorical • Choice • MaxDiff • Uplift • Multiple Factor Analysis
<i>Scripting Guide</i>	Learn about taking advantage of the powerful JMP Scripting Language (JSL).	Covers a variety of topics, such as writing and debugging scripts, manipulating data tables, constructing display boxes, and creating JMP applications.
<i>JSL Syntax Reference</i>	Read about many JSL functions on functions and their arguments, and messages that you send to objects and display boxes.	Includes syntax, examples, and notes for JSL commands.

Additional Resources for Learning JMP

In addition to reading JMP help, you can also learn about JMP using the following resources:

- [“Tutorials”](#)
- [“Sample Data Tables”](#)
- [“Learn about Statistical and JSL Terms”](#)
- [“Learn JMP Tips and Tricks”](#)
- [“Tooltips”](#)
- [“JMP User Community”](#)
- [“Free Online Statistical Thinking Course”](#)
- [“New User Welcome Kit”](#)
- [“Statistics Knowledge Portal”](#)
- [“JMP Training”](#)
- [“JMP Books by Users”](#)
- [“The JMP Starter Window”](#)

Tutorials

You can access JMP tutorials by selecting **Help > Tutorials**. The first item on the **Tutorials** menu is **Tutorials Directory**. This opens a new window with all the tutorials grouped by category.

If you are not familiar with JMP, start with the **Beginners Tutorial**. It steps you through the JMP interface and explains the basics of using JMP.

The rest of the tutorials help you with specific aspects of JMP, such as designing an experiment and comparing a sample mean to a constant.

Sample Data Tables

All of the examples in the JMP documentation suite use sample data. Select **Help > Sample Data Library** to open the sample data directory.

To view an alphabetized list of sample data tables or view sample data within categories, select **Help > Sample Data**.

Sample data tables are installed in the following directory:

On Windows: C:\Program Files\SAS\JMP\15\Samples\Data

On macOS: \Library\Application Support\JMP\15\Samples\Data

In JMP Pro, sample data is installed in the JMPPRO (rather than JMP) directory.

To view examples using sample data, select **Help > Sample Data** and navigate to the Teaching Resources section. To learn more about the teaching resources, visit <https://jmp.com/tools>.

Learn about Statistical and JSL Terms

The **Help** menu contains the following indexes:

Statistics Index Provides definitions of statistical terms.

Scripting Index Lets you search for information about JSL functions, objects, and display boxes. You can also edit and run sample scripts from the Scripting Index and get help on the commands.

Learn JMP Tips and Tricks

When you first start JMP, you see the Tip of the Day window. This window provides tips for using JMP.

To turn off the Tip of the Day, clear the **Show tips at startup** check box. To view it again, select **Help > Tip of the Day**. Or, you can turn it off using the Preferences window.

Tooltips

JMP provides descriptive tooltips (or *hover labels*) when you place your cursor over items, such as the following:

- Menu or toolbar options
- Labels in graphs
- Text results in the report window (move your cursor in a circle to reveal)
- Files or windows in the Home Window
- Code in the Script Editor

Tip: On Windows, you can hide tooltips in the JMP Preferences. Select **File > Preferences > General** and then deselect **Show menu tips**. This option is not available on macOS.

JMP User Community

The JMP User Community provides a range of options to help you learn more about JMP and connect with other JMP users. The learning library of one-page guides, tutorials, and demos is a good place to start. And you can continue your education by registering for a variety of JMP training courses.

Other resources include a discussion forum, sample data and script file exchange, webcasts, and social networking groups.

To access JMP resources on the website, select **Help > JMP User Community** or visit <https://community.jmp.com>.

Free Online Statistical Thinking Course

Learn practical statistical skills in this free online course on topics such as exploratory data analysis, quality methods, and correlation and regression. The course consists of short videos, demonstrations, exercises, and more. Visit <https://www.jmp.com/statisticalthinking>.

New User Welcome Kit

The New User Welcome Kit is designed to help you quickly get comfortable with the basics of JMP. You'll complete its thirty short demo videos and activities, build your confidence in using the software, and connect with the largest online community of JMP users in the world. Visit <https://www.jmp.com/welcome>.

Statistics Knowledge Portal

The Statistics Knowledge Portal combines concise statistical explanations with illuminating examples and graphics to help visitors establish a firm foundation upon which to build statistical skills. Visit <https://www.jmp.com/skp>.

JMP Training

SAS offers training on a variety of topics led by a seasoned team of JMP experts. Public courses, live web courses, and on-site courses are available. You might also choose the online e-learning subscription to learn at your convenience. Visit <https://www.jmp.com/training>.

JMP Books by Users

Additional books about using JMP that are written by JMP users are available on the JMP website. Visit <https://www.jmp.com/books>.

The JMP Starter Window

The JMP Starter window is a good place to begin if you are not familiar with JMP or data analysis. Options are categorized and described, and you launch them by clicking a button. The JMP Starter window covers many of the options found in the Analyze, Graph, Tables, and File menus. The window also lists JMP Pro features and platforms.

- To open the JMP Starter window, select **View (Window on macOS) > JMP Starter**.
- To display the JMP Starter automatically when you open JMP on Windows, select **File > Preferences > General**, and then select **JMP Starter** from the Initial JMP Window list. On macOS, select **JMP > Preferences > Initial JMP Starter Window**.

Technical Support

JMP technical support is provided by statisticians and engineers educated in SAS and JMP, many of whom have graduate degrees in statistics or other technical disciplines.

Many technical support options are provided at <https://www.jmp.com/support>, including the technical support phone number.

Chapter 2

Introduction to DOE

Overview of Design of Experiment Platforms

The JMP DOE platforms help you to design, evaluate, and analyze experiments. Most of the platforms focus on constructing designs. Other platforms support the design effort. This section provides a quick overview of each of the platforms found under the DOE menu.

Design Construction Platforms

Custom Design Constructs designs that fit a wide variety of settings. Custom designs tend to be more cost effective and flexible than approaches based exclusively on classical designs.

Custom designs accommodate various types of factors, constraints, and disallowed combinations. You can specify which effects are necessary to estimate and which are desirable to estimate, given the number of runs. You can specify a number of runs that matches the budget for your experimental situation. Custom designs also support hard-to-change and very-hard-to-change factors, allowing you to construct split-plot and related designs.

The Custom Design platform constructs many special design types:

- screening
- response surface
- mixture
- random block
- split-plot
- split-split-plot
- two-way split-plot

You can construct classical screening, response surface, and mixture designs using other platforms. However, the Custom Design platform gives you flexibility that is not available in the other platforms. Constructing designs for split-plot situations can be done only using the Custom Design platform.

Definitive Screening Design Constructs screening designs for continuous and two-level categorical factors. Definitive screening designs are useful if you suspect active interactions or curvature. Definitive screening designs enable you to identify the source of

strong nonlinear effects while avoiding complete confounding between any effects up through the second order.

Definitive screening designs are most appropriate for experimentation with four or more factors. Definitive screening designs support grouping runs into blocks. The number of blocks is user-specified.

Screening Design Constructs screening designs for continuous, discrete numeric, and categorical factors with an arbitrary number of levels. When standard designs exist, you have two options:

- Choose from a list of classical screening designs. These designs allow two-level continuous factors or two- or three-level categorical or discrete continuous factors.
- Generate a design that is orthogonal or nearly orthogonal for main effects.
Near-orthogonal designs allow for categorical and discrete numeric factors with any number of levels, as well as two-level continuous factors. These designs focus on estimating main effects in the presence of negligible interactions.

For many screening situations, standard designs are not available. In these situations, you can construct near-orthogonal screening designs.

Response Surface Design Constructs designs that model a quadratic function of continuous factors. To fit the quadratic effects, response surface designs require three settings for each factor. JMP provides response surface designs for up to eight factors.

You can choose from a list of Central Composite or Box-Behnken designs. When appropriate, Central Composite designs that block orthogonally are included in the list. Various modifications to Central Composite designs are supported.

Full Factorial Design Constructs full factorial designs for any number of continuous or categorical factors, both with arbitrarily many levels. A full factorial design has a run at every combination of settings of the factors. Full factorial designs tend to be large. The number of runs equals the product of the numbers of factor levels.

Mixture Design Constructs designs that you use when factors are ingredients in a mixture. In a mixture experiment, a change in the proportion of one ingredient requires that one or more of the remaining ingredients change to maintain the sum. Choose from among several design types, including some classical mixture design approaches: optimal, simplex centroid, simplex lattice, extreme vertices, ABCD, and space filling. For optimal, extreme vertices, and space filling mixture designs, you can specify linear inequality constraints to limit the design space.

Taguchi Arrays Constructs designs that you use for signal-to-noise analysis. The designs are based on Taguchi's inner and outer array approach. Control factor settings constitute the

inner array and noise factor settings form the outer array. The mean and signal-to-noise ratio are the responses of interest.

An alternative to using a Taguchi array is to construct a custom design that includes control factors, noise factors, and control-by-noise interactions. Such designs, called *combined arrays*, are generally more cost-effective and informative than Taguchi arrays.

Choice Design Constructs designs that you can use to compare prospective products. The factors in a choice design are product attributes. The design arranges product profiles, which are combinations of various attributes, in pairs or in groups of three or four. The experiment consists of having respondents indicate which profile in a pair of profiles that they prefer. You can generate a choice design that reflects prior information about the product attributes.

MaxDiff Constructs a design consisting of choice sets that can be presented to respondents as part of a MaxDiff study. Respondents report only the most and least preferred options from among a small set of choices. This forces respondents to rank options in terms of preference, which often results in rankings that are more definitive than rankings obtained using standard preference scales.

JMP[®] PRO Covering Array Constructs combinatorial designs that you can use to test software, networks, and other systems. A strength t covering array has the property that every combination of levels of every t factors appears in at least one run. Covering arrays allow for any number of categorical factors, each with an arbitrary number of levels. Disallowed combinations can be specified.

Space Filling Design Constructs designs for situations where the system of interest is deterministic or near-deterministic. A standard application involves creating a simpler surrogate model of a highly complex deterministic computer simulation model.

In a deterministic system, there is no variation. The goal is to minimize the difference between the fitted model and the true model (bias). Space-filling designs attempt to meet this goal either by spreading the design points out as far from each other as possible or by spacing the points evenly over the design region.

JMP provides seven space-filling design approaches. One of these approaches, the fast flexible filling design, accommodates categorical factors with any number of levels and supports linear constraints.

Accelerated Life Test Design Constructs and augments designs useful for testing products at extreme conditions which are intended to accelerate failure time. Use experimental results to predict reliability under normal operating conditions.

The life distribution can be lognormal or Weibull. Designs can include one or two accelerating factors. If there are two accelerating factors, you can choose to include their interaction. You can specify prior distributions for the acceleration model parameters. D-optimal and two types of I-optimal designs are available.

Nonlinear Design Constructs and augments designs that you use to fit models that are nonlinear in their parameters. You can construct a design using estimates from a model fit to existing data. You can also construct a design by applying prior knowledge if you do not have model-based estimates.

Balanced Incomplete Block Design Constructs design for testing a treatments in b blocks where only k treatments ($k < a$) can be run in any one block

Group Orthogonal Supersaturated Design Constructs supersaturated screening designs. They are appropriate in early stage work when the number of factors to be investigated is larger than the number of feasible runs. A group orthogonal supersaturated design is a special class of two-level supersaturated designs with properties that are desirable for model selection.

Supporting Platforms

Augment Design Adds runs to existing designs in such a way that the resulting design is optimal. Augment Design enables you to conduct experiments in an iterative fashion. You can replicate the design, add center points, create a fold-over design, add axial points, add points to create a space-filling design, or augment the design with a specified number of runs. You can group runs into blocks to distinguish the original runs from the augmented runs. You can add model effects that were not in the original model and specify requirements for these effects.

Fit Definitive Screening Design Analyzes definitive screening designs using a methodology called *Effective Model Selection for DSDs*. This methodology takes advantage of the special structure of definitive screening designs.

Fit Group Orthogonal Supersaturated Analyzes group orthogonal supersaturated designs. This analysis technique takes advantage of the group orthogonal structure of group orthogonal supersaturated designs.

Evaluate Design Provides diagnostics for an existing experimental design. The Evaluate Design platform provides various ways for you to assess the strengths and limitations of your design. The platform can be used with any data table, not only designs created using JMP.

Several diagnostics are provided:

- power analysis
- prediction variance plots
- estimation efficiency for parameters
- the alias matrix, showing the bias structure for model effects
- a color map showing absolute correlations among effects
- design efficiency values

Compare Designs Compares up to four designs to a reference design. Use to explore, evaluate, and compare design performance. Diagnostics show how the designs perform relative to each other and how they perform in an absolute sense.

Sample Size and Power Provides sample size and power calculations for a variety of testing situations: one or more sample means, a standard deviation, one or two proportions, counts per unit (Poisson mean), and sigma quality level. For these options, you specify two of three quantities to compute the third. These three quantities are the difference you want to detect, the sample size, and the power. If you supply only one of these values, a plot of the relationship between the other two values is provided.

You can compute the sample size required for a reliability test plan, where your goal is to estimate failure probabilities. You can also compute the sample size required for a reliability demonstration, where your goal is to demonstrate that a product meets or exceeds a specified standard.

Chapter 3

Starting Out with DOE Example and Key Concepts

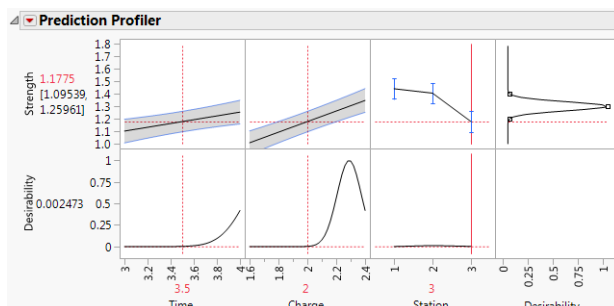
A designed experiment is a controlled set of tests designed to model and explore the relationship between factors and one or more responses. JMP includes a variety of tools that enable you to create efficient experimental designs that work for your situation. In particular, these classes of designs are available:

- The Custom Design platform customizes a design for your unique situation. It constructs designs that accommodate any number of factors of any type and factors that are difficult to change (split plot situations). You control the number of runs.
- The Definitive Screening Design platform constructs an innovative class of screening designs where main effects are not aliased with each other or with two-way interactions. These designs also allow estimation of quadratic terms.
- The Screening Design, Response Surface Design, Full Factorial Design, Mixture Design, and Taguchi Arrays platforms construct traditional experimental designs.
- The Covering Array, Choice Design, Space Filling Design, Accelerated Life Test Design, and Nonlinear Design platforms construct specialized design types.

The Evaluate Design and Augment Design platforms provide tools for evaluating and augmenting existing design. The Sample Size and Power platform addresses sample size and power calculations for specialized situations.

This chapter presents an example that illustrates the JMP approach to DOE. This chapter also discusses the framework for DOE, the workflow that supports design creation, and principles that are fundamental to DOE.

Figure 3.1 Example of a Profiler Plot



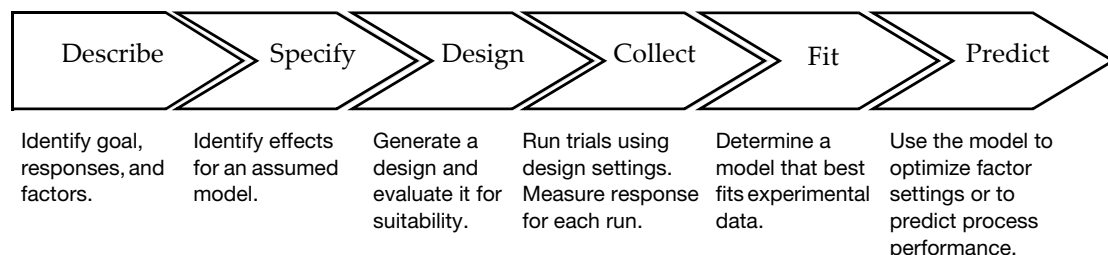
Contents

Overview of Experimental Design and the DOE Workflow	41
The Coffee Strength Experiment.....	42
Define the Study and Goals.....	43
Create the Design	45
Run the Experiment	51
Analyze the Data.....	51
The DOE Workflow: Describe, Specify, Design	57
Define Responses and Factors.....	58
Specify the Model	59
Generate the Design	60
Evaluate the Design	61
Make the Table.....	61
Principles and Guidelines for Experimental Design	62
Effect Hierarchy.....	62
Effect Heredity.....	62
Effect Sparsity	63
Center Points, Replicate Runs, and Testing.....	63

Overview of Experimental Design and the DOE Workflow

A six-step framework provides the structure for designing an experiment, running the experimental trials, and analyzing the results. Sound engineering and process knowledge is critical to all of these steps.

Figure 3.2 Framework for Experimental Design



You perform the first three steps in the DOE platforms. The end result is a design that can be run in your work environment. For a detailed description of the workflow for these three steps, see [“The DOE Workflow: Describe, Specify, Design”](#) on page 57.

Describe Determine the goal of your experiment. Identify responses and factors.

Your goal might be to identify active factors, to find optimal factor settings, or to build a predictive model.

Specify Determine or specify an assumed model that you believe adequately describes the physical situation.

Your *assumed* model is an initial model that ideally contains all the effects that you want to estimate. In some platforms, you can explicitly build the model of interest. In others, the model is implicit in the choices that you make. For example, in the Screening Design platform, you might select a model with a given resolution. The resolution of the design determines which effects are confounded. Confounding of effects potentially leads to ambiguity about which effect is truly active.

Design Generate a design that is consistent with your assumed model. Evaluate this design to understand its strengths and limitations, and to ensure that it provides the information that you need, given your model and goals.

The Design Evaluation or Design Diagnostics outline in the design generation platform give you insight about the properties of your design.

The next step is the data collection phase, where the experiment is run under controlled conditions.

Collect Conduct each of the trials and record the response values.

After you run your experiment, scripts in the generated data table help you fit a model using platforms such as Fit Model and Screening. Depending on your goal, the model can help you identify active effects or find optimal settings.

Fit Fit your assumed model to the experimental data.

Use the JMP modeling platforms to fit and refine your model. In some situations, you might need to augment the design and perform additional runs to resolve model ambiguity.

Predict Use your refined model to address your experimental goals.

Determine which effects are active, find factor levels to optimize responses, or build a predictive model.

Designed experiments are typically used sequentially to construct process knowledge. A design strategy often begins with a screening design to narrow the list of potentially active factors. Then the identified factors are studied in designs that focus on building a better understanding of interactions and quadratic effects. Sometimes there is a need to augment a design to resolve ambiguities relating to the factors responsible for effects. The steps outlined in this section relate to conducting and analyzing a single experiment. However, you may require a sequence of experiments to achieve your goals.

The example in [“The Coffee Strength Experiment”](#) on page 42 explicitly illustrates the steps in the DOE workflow process. It also shows how to use a data table script to analyze your experimental data. Many examples in the *Design of Experiments Guide* illustrate both the workflow that supports a good design and the analysis of the experimental data from the study.

The Coffee Strength Experiment

- [“Define the Study and Goals”](#)
- [“Create the Design”](#)
- [“Run the Experiment”](#)
- [“Analyze the Data”](#)

Define the Study and Goals

Your employer is a local mid-size coffee roaster. You need to address the strength of individually brewed twelve ounce cups of coffee. Your goal is to determine which factors have an effect on coffee strength and to find optimal settings for those factors.

Response

The response is coffee **Strength**. It is measured as total dissolved solids, using a refractometer. The coffee is brewed using a single cup coffee dripper and measured five minutes after the liquid is released from the grounds.

Previous studies indicate that a strength reading of 1.3 is most desirable, though the strength is still acceptable if it falls between 1.2 and 1.4.

Factors

Four factors are identified for the study: Grind, Temperature, Time, and Charge. Coffee is brewed at three stations in the work area. To account for variation due to brewing location, Station is included in the study as a blocking factor. The following describes the factors:

- Grind is the coarseness of the grind. Grind is set at two levels, Medium and Coarse.
- Temperature is the temperature in degrees Fahrenheit of the water measured immediately before pouring it over the grounds. Temperature is set at 195 and 205 degrees Fahrenheit.
- Time is the brewing time in minutes. Time is set at 3 or 4 minutes.
- Charge is the amount of coffee placed in the cone filter, measured in grams of coffee beans per ounce of water. Charge is set at 1.6 and 2.4.
- Station is the location where the coffee is brewed. The three stations are labeled as 1, 2, and 3.

Table 3.1 summarizes information about the factors and their settings. The factors and levels are also given in the *Coffee Factors.jmp* sample data table, located in the Design Experiment folder.

Table 3.1 Factors and Range of Settings for Coffee Experiment

Factor	Role	Range of Settings
Grind	Categorical	Medium, Coarse
Temperature	Continuous	195 - 205
Time	Continuous	3 - 4
Charge	Continuous	1.6 - 2.4

Table 3.1 Factors and Range of Settings for Coffee Experiment *(Continued)*

Factor	Role	Range of Settings
Station	Blocking	1, 2, 3

Note the following:

- Grind is categorical with two levels.
- Temperature, Time, and Charge are continuous.
- Station is a blocking factor with three levels.

All factors can be varied and reset for each run. There are no hard-to-change factors for this experiment.

The apparatus used in running the coffee experiment is shown in Figure 3.3. This is the setup at one of the three brewing stations. The two other stations have the same type of equipment.

Figure 3.3 Coffee Experiment Apparatus



Number of Runs

Based on the resources and time available, you determine that you can conduct 12 runs in all. Since there are three stations, you conduct 4 runs at each station.

Create the Design

Create the design following the steps in the design workflow process outlined in [“The DOE Workflow: Describe, Specify, Design”](#):

- [“Define Responses and Factors”](#)
- [“Specify the Model”](#)
- [“Generate the Design”](#)
- [“Evaluate the Design”](#)
- [“Make the Table”](#)

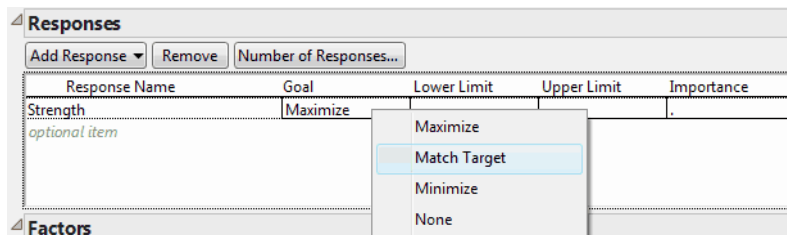
Define Responses and Factors

In the first outlines that appear, enter information about your response and factors.

Responses

1. Select **DOE > Custom Design**.
2. Double-click Y under Response Name and type **Strength**.
Note that the default Goal is **Maximize**. Your goal is to find factor settings that enable you to brew coffee with a target strength of 1.3, within limits of 1.2 and 1.4.
3. Click the default Goal of **Maximize** and change it to **Match Target**.

Figure 3.4 Selection of Match Target as the Goal



4. Click under **Lower Limit** and type 1.2.
5. Click under **Upper Limit** and type 1.4.
6. Leave the area under **Importance** blank.

Because there is only one response, that response is given Importance 1 by default.

The completed Responses outline appears in Figure 3.5.

Factors

Enter factors either manually or from a pre-existing table that contains the factors and settings. If you are designing a new experiment, you must first enter the factors manually. Once you have saved the factors to a data table using the Save Factors option, you can load them using the saved table.

For this example, you can choose either option. See [“Entering Factors Manually”](#) on page 46 or see [“Entering Factors Using Load Factors”](#) on page 47.

Entering Factors Manually

1. Click **Add Factor > Categorical > 2 Level**.

2. Type Grind over the default Name of X1.

Note that Role is set to Categorical, as requested. The Changes attribute is set to Easy by default, indicating that Grind settings can be reset for every run.

3. Click the default Values, L1 and L2, and change them to Coarse and Medium.

4. Type 3 next to **Add N Factors**. Then click **Add Factor > Continuous**.

5. Type the following names and values over the default entries:

- Temperature (195 and 205)
- Time (3 and 4)
- Charge (1.6 and 2.4)

6. Click **Add Factor > Blocking > 4 runs per block**.

Recall that your run budget allows for 12 runs. You want to balance these runs among the three stations.

7. Type Station over the default Name of X5.

Notice that Role is set to Blocking and that only one setting for Values appears. This is because JMP cannot determine the number of blocks until the desired number of runs is specified. Once you specify the Number of Runs in the Design Generation outline, JMP updates the number of levels for Station to what is required.

Figure 3.5 Completed Responses and Factors Outlines

The screenshot shows the Minitab Custom Design window with two main sections: Responses and Factors.

Responses Section:

- Buttons: Add Response, Remove, Number of Responses...
- Table:

Response Name	Goal	Lower Limit	Upper Limit	Importance
Strength	Match Target	1.2	1.4	.

optional item

Factors Section:

- Buttons: Add Factor, Remove, Add N Factors, 1
- Table:

Name	Role	Changes	Values
Grind	Categorical	Easy	Coarse Medium
Temperature	Continuous	Easy	195 205
Time	Continuous	Easy	3 4
Charge	Continuous	Easy	1.6 2.4
Station	Blocking	Easy	1

Specify Factors

Add a factor by clicking the Add Factor button. Double click on a factor name or level to edit it.

Continue

8. Click **Continue**.

The following outlines are added to the Custom Design window:

- Define Factor Constraints (not used in this example)
- Model
- Alias Terms
- Design Generation

Entering Factors Using Load Factors

To enter factors using a table containing factor information, proceed as follows:

1. Click the Custom Design red triangle and select **Load Factors**.
2. Select **Help > Sample Data Library** and open Design Experiment/Coffee Factors.jmp.

After loading the factors, the Custom Design window is updated. The following outlines are added to the Custom Design window:

- Define Factor Constraints (not used in this example)
- Model
- Alias Terms
- Design Generation

Define Factor Constraints

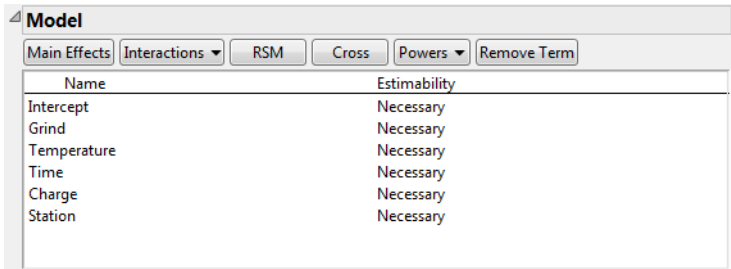
The Define Factor Constraints outline appears once you have entered your factors manually and clicked Continue, or once you have loaded the factors from the factor table. Adding factor constraints, if you have any, is part of the Responses and Factors step. Since there are no constraints on factor settings for this design, leave this outline unchanged.

Specify the Model

Model Outline

Figure 3.7 shows the Model outline. The Model outline is where you specify your assumed model, which contains the effects that you want to estimate. See “Specify” on page 41. The list that appears by default shows all main effects as **Necessary**, indicating that the design is capable of estimating all main effects. Because your main interest at this point is in the main effects of the factors, you do not add any effects to the Model outline.

Figure 3.6 Model Outline with Main Effects Only



Model	
Main Effects	Interactions
RSM	Cross
Powers	Remove Term
Name	Estimability
Intercept	Necessary
Grind	Necessary
Temperature	Necessary
Time	Necessary
Charge	Necessary
Station	Necessary

Steps to Duplicate Results (Optional)

Because the Custom Design algorithm begins with a random starting design, your design might differ from the one shown in Figure 3.8. To obtain a design with exactly the same runs, perform the following steps before generating your design:

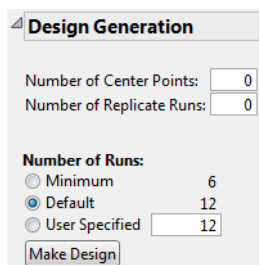
1. Click the Custom Design red triangle and select **Set Random Seed**.
2. Type 569534903.
3. Click **OK**.
4. Click the Custom Design red triangle and select **Number of Starts**.
5. Type 100.
6. Click **OK**.

Note: Setting the Random Seed and Number of Starts reproduces the exact design shown in this example. However, the rows in the design table might be in a different order. In constructing a design on your own, these steps are not necessary.

Generate the Design

In the Design Generation outline, you can enter additional details about the structure and size of your design. The Default design is shown as having 12 runs. Recall that your design budget allows for 12 runs ([“Number of Runs”](#) on page 44).

Figure 3.7 Design Generation Outline



Design Generation

Number of Center Points:

Number of Replicate Runs:

Number of Runs:

☐ Minimum 6

☒ Default 12

☐ User Specified

Make Design

1. Click **Make Design**.

The Design and Design Evaluation outlines are added to the Custom Design window. The Output Options panel also appears.

The Design outline shows the design (Figure 3.8). If you did not follow the steps in [“Steps to Duplicate Results \(Optional\)”](#) on page 48, your design might be different from the one in Figure 3.8. This is because the algorithm begins with a random starting design.

Figure 3.8 Design for Coffee Experiment

Design					
Run	Grind	Temperature	Time	Charge	Station
1	Coarse	195	3	1.6	1
2	Medium	195	4	2.4	2
3	Medium	195	4	1.6	3
4	Medium	205	4	2.4	1
5	Coarse	205	3	2.4	2
6	Medium	205	3	2.4	3
7	Medium	205	3	1.6	1
8	Coarse	205	4	1.6	2
9	Coarse	205	4	1.6	3
10	Coarse	195	4	2.4	1
11	Medium	195	3	1.6	2
12	Coarse	195	3	2.4	3

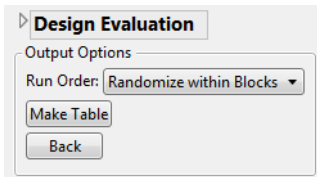
Evaluate the Design

The Design Evaluation outline provides various ways to evaluate your design. This is an important topic, but for simplicity, it is not covered in the context of this example. See the [“Evaluate Designs”](#) chapter on page 433.

Make the Table

Specify the order of runs in your data table using the Output Options panel. The default selection, **Randomize within Blocks**, is appropriate. This selection arranges the runs in a random order for each Station.

Figure 3.9 Output Options



- 1. Click **Make Table**.

The data table shown in Figure 3.10 opens. Keep in mind that, if you did not follow the steps in [“Steps to Duplicate Results \(Optional\)”](#) on page 48, your design table might be different. Your design table represents another optimal design.

Figure 3.10 Custom Design Table

	Grind	Temperature	Time	Charge	Station	Strength
1	Coarse	195	4	2.4	1	•
2	Medium	205	3	1.6	1	•
3	Medium	205	4	2.4	1	•
4	Coarse	195	3	1.6	1	•
5	Medium	195	4	2.4	2	•
6	Medium	195	3	1.6	2	•
7	Coarse	205	4	1.6	2	•
8	Coarse	205	3	2.4	2	•
9	Medium	195	4	1.6	3	•
10	Medium	205	3	2.4	3	•
11	Coarse	205	4	1.6	3	•
12	Coarse	195	3	2.4	3	•

Note the asterisks in the Columns panel to the right of the factors and response. These indicate column properties that have been saved to the columns in the data table. These column properties are used in the analysis of the data. See [“Factors”](#) on page 88 and [“Factor Column Properties”](#) on page 94.

Run the Experiment

At this point, you perform the experiment. At each Station, four runs are conducted in the order shown in the design table. Equipment and material are reset between runs. For example, if two consecutive runs require water at 195 degrees, separate 12-ounce batches of water are heated to 195 degrees after the heating container cools. The Strength measurements are recorded.

Your design and the experimental results for Strength are given in the Coffee Data.jmp sample data table (Figure 3.11), located in the Design Experiment folder.

Figure 3.11 Coffee Design with Strength Results

	Grind	Temperature	Time	Charge	Station	Strength
1	Mediu...	205	4	2.4	1	1.78
2	Coarse	195	3	1.6	1	1.25
3	Mediu...	205	3	1.6	1	1.10
4	Coarse	195	4	2.4	1	1.63
5	Coarse	205	4	1.6	2	1.26
6	Mediu...	195	4	2.4	2	1.63
7	Mediu...	195	3	1.6	2	1.22
8	Coarse	205	3	2.4	2	1.51
9	Coarse	205	4	1.6	3	1.07
10	Coarse	195	3	2.4	3	1.26
11	Mediu...	195	4	1.6	3	1.13
12	Mediu...	205	3	2.4	3	1.25

Analyze the Data

The Custom Design platform facilitates the task of data analysis by saving a Model script to the design table that it creates (Figure 3.10). Run this script after you conduct your experiment and enter your data. The script opens a Fit Model window containing the effects that you specified in the Model outline of the Custom Design window.

Fit the Model

1. Select **Help > Sample Data Library** and open Design Experiment/Coffee Data.jmp.

In the Table panel, notice the Model script created by Custom Design.

2. Click the green triangle next to the Model script.

The Model Specification window shows the effects that you specified in the Model outline.

Figure 3.12 Model Specification Window

Model Specification

Select Columns
6 Columns
Grind
Temperature
Time
Charge
Station
Strength

Pick Role Variables

Y: Strength (optional)

Weight: optional numeric

Freq: optional numeric

Validation: optional

By: optional

Personality: Standard Least Squares

Emphasis: Effect Screening

Help Run

Recall ☐ Keep dialog open

Remove

Construct Model Effects

Add Cross Nest Macros

Grind
Temperature
Time
Charge
Station

Degree: 2

Attributes ☒

Transform ☒

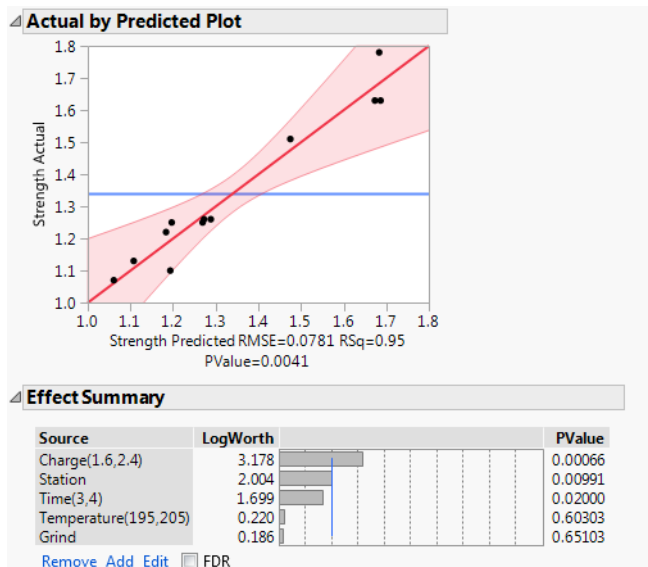
☐ No Intercept

3. Select the **Keep dialog open** option.
4. Click **Run**.

Analyze the Model

The Effect Summary and Actual by Predicted Plot reports give high-level information about the model.

Figure 3.13 Effect Summary and Actual by Predicted Plot for Full Model



Note the following:

- The Actual by Predicted Plot shows no evidence of lack of fit.
- The model is significant, as indicated by the Actual by Predicted Plot. The notation **P = 0.0041**, shown below the plot, gives the significance level of the overall model test.
- The Effect Summary report shows that Charge, Station, and Time are significant at the 0.05 level.
- The Effect Summary report also shows that Temperature and Grind are not significant.

Reduce the Model

Because Temperature and Grind appear not to be active, they contribute random noise to the model. Refit the model without these effects to obtain more precise estimates of the model parameters associated with the active effects.

1. In the Model Specification window, select Temperature and Grind in the Construct Model Effects list.
2. Click **Remove**.
3. Confirm that the model **Emphasis** is set to **Effect Screening**.

The Effect Screening emphasis presents reports (such as the Prediction Profiler) that are useful for analyzing experimental designs.

4. Click **Run**.

Figure 3.14 Partial Output for the Reduced Model

Parameter Estimates					
Term	Estimate	Std Error	t Ratio	Prob> t	
Intercept	1.3408333	0.020047	66.88	<.0001*	
Time(3,4)	0.0758333	0.020047	3.78	0.0069*	
Charge(1,6,2,4)	0.1691667	0.020047	8.44	<.0001*	
Station[1]	0.0991667	0.028351	3.50	0.0100*	
Station[2]	0.0641667	0.028351	2.26	0.0580	

Effect Tests					
Source	Nparm	DF	Sum of Squares	F Ratio	Prob > F
Time(3,4)	1	1	0.06900833	14.3093	0.0069*
Charge(1,6,2,4)	1	1	0.34340833	71.2078	<.0001*
Station	2	2	0.16251667	16.8494	0.0021*

Scaled Estimates					
Nominal factors expanded to all levels					
Term	Scaled Estimate		Std Error	t Ratio	Prob> t
Intercept	1.3408333		0.020047	66.88	<.0001*
Time(3,4)	0.0758333		0.020047	3.78	0.0069*
Charge(1,6,2,4)	0.1691667		0.020047	8.44	<.0001*
Station[1]	0.0991667		0.028351	3.50	0.0100*
Station[2]	0.0641667		0.028351	2.26	0.0580
Station[3]	-0.163333		0.028351	-5.76	0.0007*

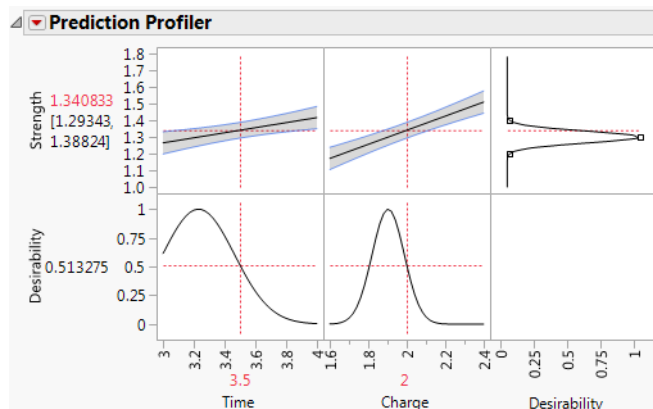
Note the following:

- The Effect Tests report shows that all three effects remain significant.
- The Scaled Estimates report further indicates that the Station[1] and Station[3] means differ significantly from the average response of Strength.
- Note that the Estimates that appear in the Parameter Estimates report are identical to their counterparts in the Scaled Estimates report. This is because the effects are coded. See [“Coding”](#) on page 742 in the “Column Properties” appendix.
- The estimate of the Station[3] effect only appears in the Scaled Estimates report, where nominal factors are expanded to show estimates for all their levels.
- The Parameter Estimates report gives estimates for the model coefficients where the model is specified in terms of the coded effects.

Explore the Model

The Prediction Profiler appears at the bottom of the report.

Figure 3.15 Prediction Profiler



Recall that, in designing your experiment, you set a response Goal of Match Target with limits of 1.2 and 1.4. JMP uses this information to construct a desirability function to reflect your specifications. See “Factors” on page 88.

Note the following in Figure 3.15:

- The first two plots in the top row of the graph show how Strength varies for one of the factors, given the setting of the other factor. For example, when Charge is 2, the line in the plot for Time shows how predicted Strength changes with Time.
- The values to the left of the top row of plots give the Predicted Strength (in red) and a confidence interval for the mean Strength for the selected factor settings.
- The right-most plot in the top row shows the desirability function for Strength. The desirability function indicates that the target of 1.3 is most desirable. Desirability decreases as you move away from that target. Desirability is close to 0 at the limits of 1.2 and 1.4.
- The plots in the bottom row show the desirability trace for each factor at the setting of the other factor.
- The value to the left of the bottom row of plots gives the Desirability of the response value for the selected factor settings.

Explore various factor settings by dragging the red dashed vertical lines in the columns for Time and Charge. Since there are no interactions in the model, the profiler indicates that increasing Charge increases Strength. Also, Strength seems to be more sensitive to changes in Charge than to changes in Time.

Since Station is a blocking factor, it does not appear in the Prediction Profiler. However, you might like to see how predicted Strength varies by Station. To include Station in the Prediction Profiler, follow these steps:

1. Click the Prediction Profiler red triangle and select **Reset Factor Grid**.

A Factor Settings window appears with columns for Time, Charge, and Station. Under Station, notice that the box corresponding to **Show** is not selected. This indicates that Station is not shown in the Prediction Profiler.

2. Select the box under Station in the row corresponding to **Show**.
3. Deselect the box under Station in the row corresponding to **Lock Factor Setting**.

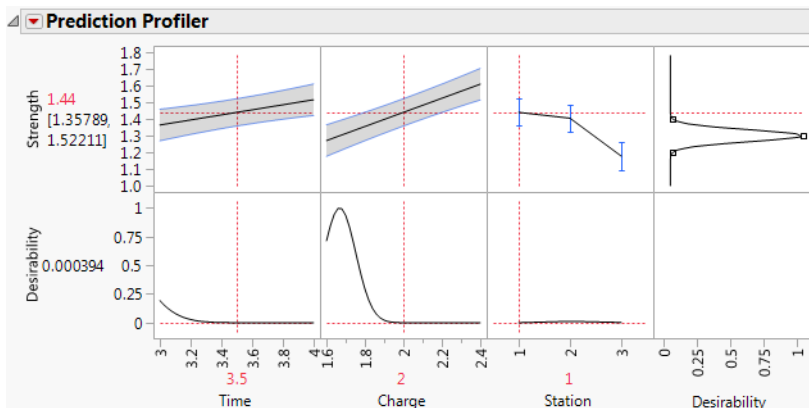
Figure 3.16 Factor Settings Window

Factor	Time	Charge	Station
Current Value:	3.5	2	
Minimum Setting:	3	1.6	
Maximum Setting:	4	2.4	
Number of Plotted Points:	41	41	
Show	☒	☒	☒
Lock Factor Setting:	☐	☐	☐

OK Cancel

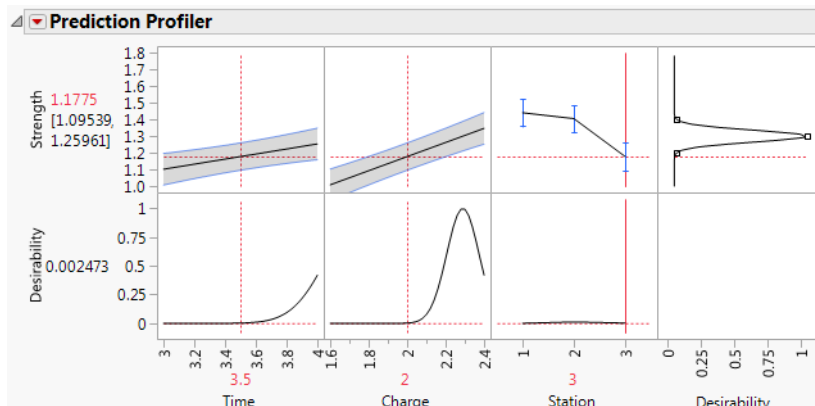
4. Click **OK**.
- Plots for Station appear in the Prediction Profiler.
5. Click in either plot above Station to insert a dashed red vertical line.
6. Move the dashed red vertical line to Station 1.

Figure 3.17 Prediction Profiler Showing Results for Station 1



7. Move the dashed red vertical line to Station 3.

Figure 3.18 Prediction Profiler Showing Results for Station 3



The predicted Strength in the center of the design region for Station 1 is 1.44. For Station 3, the predicted Strength is about 1.18. The magnitude of the difference indicates that you need to address Station variability. Better control of Station variation should lead to more consistent Strength. Once Station consistency is achieved, you can determine common optimal settings for Time and Charge.

The process that you used to construct the design for the coffee experiment followed the steps in the DOE workflow. The next section describes the DOE workflow in more detail.

The DOE Workflow: Describe, Specify, Design

The DOE platforms are structured as a series of steps that present the workflow that is intrinsic to designing experiments. Once you complete each step, you click **Continue** to move to the next step. The elements described in this section are common to nine of the design of experiments platforms. These are the platforms that are addressed in this section:

- Custom Design
- Definitive Screening Design
- Screening Design
- Response Surface Design
- Full Factorial Design
- Mixture Design
- **JMP PRO** Covering Array
- Space Filling Design
- Taguchi Arrays

Three special-purpose platforms differ substantially: Choice Designs, Accelerated Life Test Design, and Nonlinear Design. These three platforms are not addressed in this section.

This section describes the steps in the DOE workflow. It also discusses their implementation in the various design platforms.

Define Responses and Factors

In the Describe step of the experimental design framework:

- You identify the responses and factors of interest.
- You determine your goals for the experiment. Do you want to maximize the response, or hit a target? What is that target? Or do you simply want to identify which factors have an effect on the response?
- You identify factor settings that describe your experimental range or design space.

When they open, most of the JMP DOE platforms display outlines where you can list your responses and your factors. The Responses outline is common across platforms. There you insert your responses and additional information, such as the response goal, lower limit, upper limit, and importance.

The Factors outline varies across platforms. This is to accommodate the types of factors and specific design situations that each platform addresses. In certain platforms, once responses and factors are entered, a Define Factor Constraints outline appears after you click Continue. In this outline, you can constrain the values of the factors that are available for the design.

Figure 3.19 shows the Responses and Factors outline using the Custom Design platform for constructing the design in the Box Corrosion Split-Plot.jmp sample data table, located in the Design Experiment folder. Also shown is the Define Factor Constraints outline, which appears once you click Continue. The Define Constraints outline enables you to specify restrictions that your factor settings must satisfy.

Figure 3.19 Responses and Factors for Box Corrosion Split-Plot Experiment

Custom Design

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Corrosion Resistance	Maximize	.	.	.

optional item

Factors

Add Factor Remove Add N Factors 1

Name	Role	Changes	Values
Furnace Temp	Categorical	Hard	360 370 380
Coating	Categorical	Easy	C1 C2 C3 C4

Define Factor Constraints

☒ None
☐ Specify Linear Constraints
☐ Use Disallowed Combinations Filter
☐ Use Disallowed Combinations Script

Specify the Model

Once you have completed filling in the Responses and Factors outlines, click the **Continue** button. This brings you to the next phase of design construction, where you either explicitly or implicitly choose an assumed model.

The Custom Design platform enables you to explicitly specify the model that you want to fit. The design that is generated is optimal for this model. The other design platforms do not allow you to explicitly specify your model. For example, in the screening platform, one option enables you to choose from a list of full factorial, fractional factorial, and Plackett-Burman designs. The aliasing relationships in these designs implicitly define the models that you can fit.

In Custom Design, when you click Continue after filling in the Responses and Factors, you see the Model outline. An example, for the design used in the Box Corrosion Split-Plot.jmp sample data table, is shown in Figure 3.20. The assumed model requires that the Furnace Temp and Coating main effects, and their interaction, be estimable. The design that is generated guarantees estimability of these effects.

In most other platforms, clicking Continue gives you a collection of designs to choose from. In Full Factorial, Continue takes you directly to Output Options, since the design is determined once the Factors outline is completed.

Figure 3.20 Model Outline for Box Corrosion Split-Plot Experiment

Model	
Main Effects	Interactions ▼
RSM	Cross
Powers ▼	Remove Term
Name	Estimability
Intercept	Necessary
Furnace Temp	Necessary
Coating	Necessary
Furnace Temp*Coating	Necessary

Generate the Design

Most of the DOE platforms give you some control over the size of the final design. In Custom Design, you can specify the number of runs and, when appropriate, the number of center points and replicate runs. In other platforms, you have various degrees of flexibility. Often you can specify the number of center points, replicate runs, or replicates of the design.

Once you have specified your options in terms of the number of runs, click Make Design. The DOE window is updated to show your design in a Design outline.

The Design outline for a 24-run custom design for the Box Corrosion Split-Plot.jmp experiment is shown in Figure 3.21. Because **Changes** for Furnace Temp was specified as **Hard**, a Whole Plots factor is constructed to represent the random blocks of settings for Furnace Temp.

Figure 3.21 Design Outline for Box Corrosion Split-Plot Experiment

Design			
Run	Whole Plots	Furnace Temp	Coating
1	1	370	C2
2	1	370	C3
3	1	370	C1
4	1	370	C4
5	2	370	C3
6	2	370	C4
7	2	370	C2
8	2	370	C1
9	3	360	C3
10	3	360	C2
11	3	360	C4
12	3	360	C1
13	4	380	C4
14	4	380	C3
15	4	380	C2
16	4	380	C1
17	5	380	C4
18	5	380	C2
19	5	380	C1
20	5	380	C3
21	6	360	C3
22	6	360	C1
23	6	360	C4
24	6	360	C2

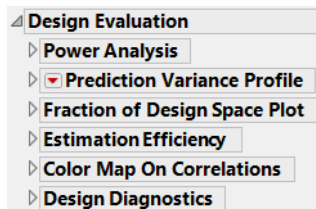
Note: For a custom design, once you have created your design, the controls in the Model and Alias Outlines are disabled. Use the Back button to change model or alias terms.

Evaluate the Design

When you click Make Design, in most platforms, a Design Evaluation outline appears. Here you can explore the design that you created in terms of the following: its power to detect effects, its prediction variance, its estimation efficiency, its aliasing relationships, the correlations between effects, and other design efficiency measures. The Design Evaluation outline for a Custom Design is shown in Figure 3.22. Design Evaluation is covered in the [“Evaluate Designs”](#) chapter on page 433.

For some platforms, other types of design diagnostics are appropriate. For example, Space Filling Design provides a Design Diagnostics outline with metrics specific to space-filling designs. Covering Array provides a Metrics outline with measures that are specific to coverage.

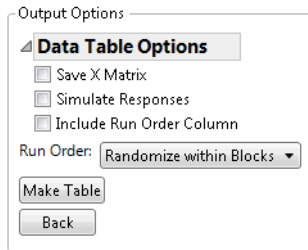
Figure 3.22 Design Evaluation Outline in Custom Design



Make the Table

Most platforms provide an Output Options node or panel. Depending on the platform and the design, you can use the Output Options panel to specify additional design structure. For example, you can specify the number of runs, center points, replicates, or the order in which you want the design runs to appear in the generated data table.

The Output Options panel shown in Figure 3.23 is for the experiment in the Wine.jmp sample data table, located in the Design Experiment folder. In this example, you can choose various Run Order options and construct the design data table. Or, you can choose to go Back and restructure your design.

Figure 3.23 Output Options Panel for Wine Experiment

Output Options

Data Table Options

☐ Save X Matrix

☐ Simulate Responses

☐ Include Run Order Column

Run Order: Randomize within Blocks ▼

Make Table

Back

Principles and Guidelines for Experimental Design

Certain principles underlie the design of experiments and the analysis of experimental data. The principles of effect hierarchy, effect heredity, and effect sparsity relate primarily to model selection. These principles help you reduce the set of possible models in searching for a best model. See Hamada and Wu (1992), Wu and Hamada (2009), and Goos and Jones (2011).

- “Effect Hierarchy”
- “Effect Heredity”
- “Effect Sparsity”
- “Center Points, Replicate Runs, and Testing”

Effect Hierarchy

In regression modeling, the principle of *effect hierarchy* maintains that main (first-order) effects tend to account for the largest amounts of variation in the response. Second-order effects, that is, interaction effects and quadratic terms, are next in terms of accounting for variation. Then come higher-order terms, in hierarchical order.

Here are the implications for modeling: main effects are more likely to be important than second-order effects; second-order effects are more likely to be important than third-order effects; and so on, for higher-order terms.

Effect Heredity

The principle of *effect heredity* relates to the inclusion in the model of lower-order components of higher-order effects. The motivation for this principle is observational evidence that factors with small main effects tend not to have significant interaction effects.

Strong effect heredity requires that all lower-order components of a model effect be included in the model. Suppose that a three-way interaction (ABC) is in the model. Then all of its component main effects and two-way interactions (A, B, C, AB, AC, BC) must also be in the model.

Weak effect heredity requires that only a sequence of lower-order components of a model effect be included. If a three-way interaction is in the model, then the model must contain one of the factors involved and one two-way interaction involving that factor. Suppose that the three-way interaction ABC is in the model. Then if B and BC are also in the model, the model satisfies weak effect heredity.

For continuous factors, effect heredity ensures that the model is invariant to changes in the location and scale of the factors.

Effect Sparsity

The principle of *effect sparsity* asserts that most of the variation in the response is explained by a relatively small number of effects. Screening designs, where many effects are studied, rely heavily on effect sparsity. Experience shows that the number of runs used in a screening design should be at least twice the number of effects that are likely to be significant.

Center Points, Replicate Runs, and Testing

Several DOE platforms enable you to add center points (for continuous factors), replicate runs, or full replicates of the design, to your design. Here is some background relative to adding design points.

Adding Center Points

Center points for continuous factors enable you to test for lack of fit due to nonlinear effects. Testing for lack of fit helps you determine whether the error variance estimate has been inflated due to a missing model term. This can be a wise investment of runs.

You can replicate runs solely at center points or you can replicate other design runs. JMP uses replicate runs to construct a model-independent error estimate (pure-error estimate). This pure-error estimate enables you to test for lack of fit.

Be aware that center points do not help you obtain more precise estimates of model effects. They enable you to test for evidence of curvature, but do not identify the responsible nonlinear effects.

To identify the source of curvature, you must set continuous factors at a minimum of three levels. Definitive screening designs are three-level designs with the ability to detect and identify any factors causing strong nonlinear effects on the response. See [Chapter 7](#), “Definitive Screening Designs”.

Adding Replicate Runs

If your run budget allows, you can either replicate runs or distribute new runs optimally within the design space. Adding replicate runs adds precision for some estimates and improves the power of the lack of fit test. However, for a given run budget, adding replicate runs generally lowers the ability of the design to estimate model effects. You are not able to estimate as many terms as you could by distributing the runs optimally within the design space.

Testing for Lack of Fit

Designed experiments are typically constructed to require as few runs as possible, consistent with the goals of the experiment. With too few runs, only extremely large effects can be detected. For example, for a given effect, the t test statistic is the ratio of the change in response means to their standard error. If there is only one error degree of freedom (df), then the critical value of the test exceeds 12. So, for such a nearly saturated design to detect an effect, it has to be very large.

A similar observation applies to the lack-of-fit test. The power of this test to detect lack-of-fit depends on the numbers of degrees of freedom in the numerator and denominator. If you have only 1 df of each kind, you need an F value that exceeds 150 to declare significance at the 0.05 level. If you have 2 df of each kind, then the F value must exceed 19. In order for the test to be significant in this second case, the lack-of-fit mean square must be 19 times larger than the pure error mean square. It is also true that the lack-of-fit test is sensitive to outliers.

For more information about the Lack of Fit test, see the Standard Least Squares Report and Options chapter in *Fitting Linear Models*.

Determining the Number of Runs

In industrial applications, each run is often very costly, so there is incentive to minimize the number of runs. To estimate the fixed effects of interest, you need only as many runs as there are terms in the model. To determine whether the effects are active, you need a reasonable estimate of the error variance. Unless you already have a good estimate of this variance, consider adding at least 4 runs to the number required to estimate the model terms.

Chapter 4

Custom Designs

Construct Designs That Meet Your Needs

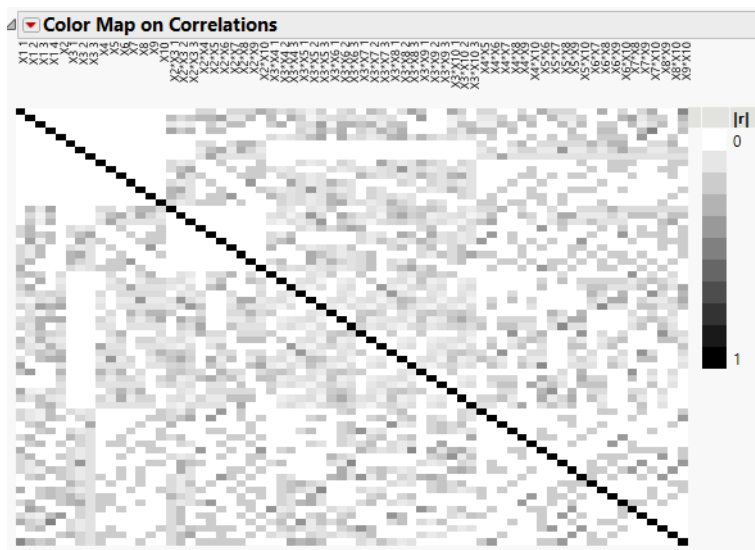
Use the Custom Design platform to construct optimal designs that are custom built for your specific experimental setting. Generally, a custom design is more cost-effective than a design obtained using alternative methods. You can perform the following tasks:

- Enter factors of many different types.
- Specify constraints on the design space.
- Indicate which effects are necessary to estimate and which are desirable to estimate, if possible, given the number of runs.
- Specify a number of experimental runs that matches your budget.

The Custom Design platform constructs a wide variety of designs, including these special design types: Screening, Response Surface, Mixture, Random Block, Split Plot, Split-Split-Plot, and Two-Way Split Plot.

This chapter contains a detailed example of how to use the Custom Design platform, followed by information about the platform. See also the [“Examples of Custom Designs”](#) chapter on page 133.

Figure 4.1 Color Map for Absolute Correlations



Contents

Overview of Custom Design	67
Example of a Custom Design	68
Create the Design	70
Analyze the Data	80
Custom Design Window	86
Responses	86
Factors	88
Define Factor Constraints	95
Model	98
Alias Terms	100
Design Generation	101
Design	103
Design Evaluation	103
Output Options	104
Custom Design Options	107
Description of Options	107
Simulate Responses	110
Save X Matrix	112
Number of Starts	115
Design Search Time	115
Set Delta for Power	116
Technical Details for Custom Designs	117
Designs with Randomization Restrictions	117
Covariates with Hard-to-Change Levels	126
Numbers of Whole Plots and Subplots	126
Optimality Criteria	127
D-Efficiency	132
Coordinate-Exchange Algorithm	132

Overview of Custom Design

Use the Custom Design platform to construct an optimal design custom built for your specific experimental needs.

You can include a wide range of factor types, including the following:

- Continuous
- Discrete numeric (with any number of levels)
- Categorical (with any number of levels)
- Blocking (with an specified number of runs per block)
- Covariate
- Mixture
- Constant
- Uncontrolled

Specify the Region of Operability

You can restrict your experimental region to reflect your operating conditions using linear factor constraints or disallowed combinations. In particular, restrictions can be specified for categorical, continuous, and discrete numeric factors. See [“Define Factor Constraints”](#) on page 95.

Specify Factors with Hard-to-Change Levels

For continuous, discrete numeric, categorical, and mixture factors, you can indicate two levels of difficult-to-change factors. These difficulty levels are represented by whole plots or whole plots and split plots. You can also specify hard-to-change covariates.

Specify the Effects of Primary Interest

You can explicitly specify your *assumed* model. Your assumed model is an initial model that ideally contains all the effects that you want to estimate. Your model can contain any combination of main effects, interactions, response surface effects, and polynomial effects (up to the fifth power). You can specify the effects for which estimability is necessary and those for which estimability is desired. Custom Design uses a Bayesian optimality approach to estimate effects whose estimability is desired, subject to the number of runs. See [“Model”](#) on page 98.

Specify the Number of Runs

The Custom Design platform enables you to specify the number of runs that matches the budget for your experimental situation. The platform indicates the minimum number of runs that can be used to estimate the required effects and provides a default number of runs. These values can serve as a guide for determining a feasible number of runs. See [“Design Generation”](#) on page 101.

Construct the Appropriate Design Type

Custom Design can construct a wide variety of design types. These include classical designs and random block designs. For examples of different design types, see the [“Examples of Custom Designs”](#) chapter on page 133.

Construct an Optimal Design

Given your specific requirements, the Custom Design platform constructs a design that is optimal. The algorithm supports several optimality criteria:

- D optimality
- I optimality
- Bayesian D and I optimality (using If Possible effects)
- A optimality
- Alias optimality

See [“Optimality Criteria”](#) on page 127.

Designs are constructed using the coordinate-exchange algorithm (Meyer and Nachtsheim, 1995). See [“Coordinate-Exchange Algorithm”](#) on page 132.

Example of a Custom Design

The following example describes a wine tasting experiment. Your employer grows two varieties of Pinot Noir grapes that can be processed in different ways. Your goal is to determine which factors affect the taste of Pinot Noir wine. Before the grapes are processed, you set up your experimental design. Once processed, the wine samples are aged for 12 months, then filtered and bottled. At this point, the wine samples are rated for quality by expert wine tasters.

Response

Most of your vineyard's product is sold to five large wine distributors. You arrange for a wine-tasting expert from each distributor to evaluate the wine samples for quality. To maximize the number of factors that you can study, you decide that each expert must rate eight different samples. This means that your design needs to have 40 wine samples, or runs.

The ratings follow a 0 – 20 scale, where 0 is the worst and 20 is the best. Rating, the variable consisting of the experts' ratings, is the response of interest. You want to identify the wine-related factors that maximize the response.

Blocking Factor

A blocking factor is used to account for variation that is not necessarily of direct interest. A blocking factor is particularly effective when observations taken at one factor level are expected to be more similar than observations at different levels. In your experiment, ratings by one expert are likely to have similar characteristics and to differ from ratings by a different expert. Yet, you are interested in which properties of the wine lead to high ratings by all experts.

Because each rater tastes eight wines, Rater is a blocking factor with eight runs per block. For this experiment, only these five raters are of concern. You are not interested in generalizing to a larger population of raters.

Process Factors

You have identified nine process factors for the study. These include the grape variety, the field on which the grapes were grown, and seven other factors related to processing. You can experiment with any combination of these factors. Also, the factors can be varied at will as part of the experiment. Relative to the experiment, these factors are all "Easy" to change. For information about specifying factor changes, see ["Changes and Random Blocks"](#) on page 92.

The factors and their levels appear in Table 4.1. Note that all of these factors are categorical. The factors and their levels are also given in the factor table Wine Factors.jmp in the Design Experiment folder of Sample Data.

To experiment with all possible combinations of these factors would require a staggering $4 \times 2^8 = 1024$ runs. However, in this example, you are able to construct a compelling design in only 40 runs.

Table 4.1 Process Factors and Levels for Wine Tasting Experiment

Factor	Levels
Variety	Bernard, Dijon

Table 4.1 Process Factors and Levels for Wine Tasting Experiment *(Continued)*

Factor	Levels
Field	1, 2, 3, 4
De-Stem	No, Yes
Yeast	Cultured, Wild
Temperature	High, Low
Press	Hard, Soft
Barrel Age	New, 2 Years
Barrel Seasoning	Air, Kiln
Filtering	No, Yes

Create the Design

Note: In order to introduce and describe the Custom Design outlines, this example works through the outlines in succession.

To create the custom design, follow the steps in these sections:

- [“Responses”](#) on page 71
- [“Factors”](#) on page 71
- [“Model”](#) on page 73
- [“Alias Terms”](#) on page 74
- [“Duplicate Results \(Optional\)”](#) on page 74
- [“Design Generation”](#) on page 75
- [“Design”](#) on page 75
- [“Design Evaluation”](#) on page 76
- [“Output Options”](#) on page 79

For information about the complete DOE workflow, see [“The DOE Workflow: Describe, Specify, Design”](#) on page 57 in the “Starting Out with DOE” chapter.

Responses

Add your response, the response Goal, and, if appropriate, the Lower Limit, Upper Limit, and Importance. Here, the response is Rating.

1. Select **DOE > Custom Design**.
2. Double-click Y under Response Name and type Rating.
Note that the default Goal is Maximize. Because you want to maximize the taste rating, do not change the goal.
3. Click under Lower Limit and type 0.
The least desirable rating is 0.
4. Click under Upper Limit and type 20.
The most desirable rating is 20.
5. Leave the area under Importance blank.
Because there is only one response, that response is given Importance 1 by default.

Figure 4.2 on page 72 shows the completed Responses outline.

Factors

Enter factors either manually or automatically using a pre-existing table that contains the factors and settings.

- If you are designing a new experiment, you must first enter the factors manually. See [“Entering Factors Manually”](#) on page 71.
- Once you have saved the factors using the Save Factors option, you can load them automatically using the saved table. See [“Entering Factors Using Load Factors”](#) on page 73.

Both methods add these four outlines to the Custom Design window: Define Factor Constraints, Model, Alias Terms, and Design Generation.

Entering Factors Manually

1. First, add the blocking factor, Rater. Click **Add Factor > Blocking > 8 runs per block**.
2. Type Rater over the default Name of X1.
Note that Role is set to Blocking. Note also that only one setting for Values appears. This is because the number of blocks cannot be determined until the desired number of runs is specified. Once you specify the Number of Runs in the Design Generation outline, the number of levels for Rater updates to what is required.
3. Click **Add Factor > Categorical > 2 Level**.
4. Type Variety over the default Name of X2.

Note that Role is set to Categorical, as requested, and that Changes is set to Easy by default.

5. Click L1 and L2 and change them to Bernard and Dijon.
6. Click **Add Factor > Categorical > 4 Level**.
7. Type Field over the default Name of X3.
8. Click L1, L2, L3, and L4, and change them to 1, 2, 3, and 4.
9. Click **Add Factor > Categorical > 2 Level**.
10. Type De-Stem over the default Name of X4.
11. Click L1 and L2 and change them to No and Yes.

Add the rest of the factors as follows:

12. Type 6 next to **Add N Factors**, and then click **Add Factor > Categorical > 2 Level**.
13. Type the following names and values over the default ones:
 - Yeast (Cultured and Wild)
 - Temperature (High and Low)
 - Press (Hard and Soft)
 - Barrel Age (New and Two Years)
 - Barrel Seasoning (Air and Kiln)
 - Filtering (No and Yes)

Figure 4.2 Completed Responses and Factors Outlines

Responses

Add Response
Remove
Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Rating	Maximize	0	20	.

optional item

Factors

Add Factor
Remove
Add N Factors
1

Name	Role	Changes	Values
Rater	Blocking	Easy	1
Variety	Categorical	Easy	Bernard Dijon
Field	Categorical	Easy	1 2 3 4
De-Stem	Categorical	Easy	No Yes
Yeast	Categorical	Easy	Cultured Wild
Temperature	Categorical	Easy	High Low
Press	Categorical	Easy	Hard Soft
Barrel Age	Categorical	Easy	New 2 Years
Barrel Seasoning	Categorical	Easy	Air Kiln
Filtering	Categorical	Easy	No Yes

14. Click **Continue**.

The following outlines are added to the Custom Design window:

- Define Factor Constraints (not used in this example)
- Model
- Alias Terms
- Design Generation

Entering Factors Using Load Factors

To enter factors using a table containing factor information, proceed as follows:

1. Select **Help > Sample Data Library** and open Design Experiment/Wine Factors.jmp.
2. Click the Custom Design red triangle and select **Load Factors**.

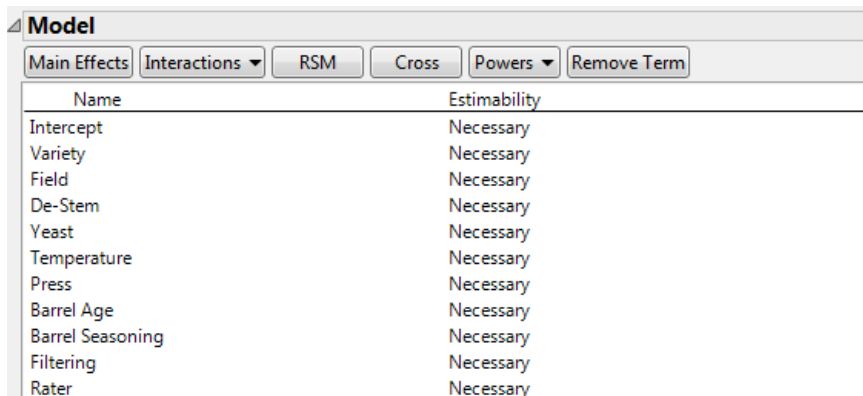
After loading the factors, the Custom Design window automatically updates. The following outlines are added to the Custom Design window:

- Define Factor Constraints (not used in this example)
- Model
- Alias Terms
- Design Generation

Model

The Model outline shows all main effects as Necessary, indicating that the design needs to be capable of estimating all main effects. For this example, your assumed model reflects your interest in main effects only. However, if you wanted to estimate other effects, you could add them to the Model outline. For more information about models, see [“Model”](#) on page 98.

Figure 4.3 Model Outline

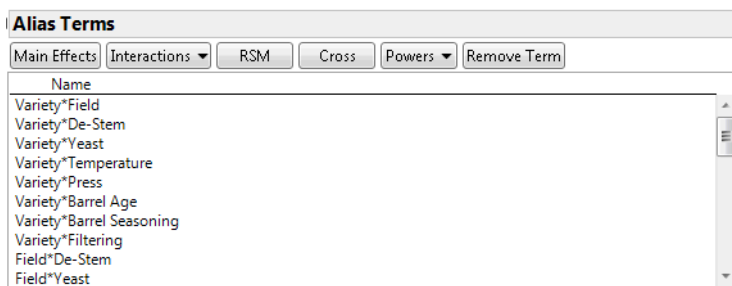


Name	Estimability
Intercept	Necessary
Variety	Necessary
Field	Necessary
De-Stem	Necessary
Yeast	Necessary
Temperature	Necessary
Press	Necessary
Barrel Age	Necessary
Barrel Seasoning	Necessary
Filtering	Necessary
Rater	Necessary

Alias Terms

The Alias Terms outline specifies the effects to be shown in the Alias Matrix, which appears later. The Alias Matrix shows the aliasing relationships between the Model terms and the effects listed in the Alias Terms outline. Open the Alias Terms outline node to verify that all two-factor interactions are listed. For more information about the alias matrix, see [“Alias Matrix”](#) on page 78.

Figure 4.4 Partial View of the Alias Terms Outline



Duplicate Results (Optional)

In the next step, you generate your design. Because the Custom Design algorithm begins with a random starting design, your design might differ from the one shown in Figure 4.5. If you want to obtain a design with exactly the same runs and run order, perform the following steps:

1. Click the Custom Design red triangle and select **Set Random Seed**.
2. Type 100526291 (the random seed).
3. Click **OK**.
4. Click the Custom Design red triangle and select **Number of Starts**.
5. Type 2.
6. Click **OK**.

Note: Setting the Random Seed and Number of Starts reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

Proceed to the Design Generation section.

Design Generation

In the Design Generation outline, you can enter additional details about the structure and size of your design. In this example, the Default design shows 16 runs. But you have five raters, each of whom can sample eight wines. This means that you want a design with 40 runs. Change the number of runs as follows:

1. Under **Number of Runs**, type **40** in the **User Specified** box.
Because you do not want to replicate runs, leave the **Number of Replicate Runs** set to **0**.
2. Click **Make Design**.

The Design and Design Evaluation outlines are added to the Custom Design window. The Output Options panel also appears.

Design

The Design outline shows the runs in the design that you have constructed. Later, you are able to randomize the order under Output Options. For now, verify that this design is appropriate for your experiment. For example, check that each of five Raters evaluates eight wines, that all necessary factors are shown, and that none of the settings represent infeasible combinations.

Figure 4.5 Design for Wine Experiment

Design											
Run	Rater	Variety	Field	De-Stem	Yeast	Temperature	Press	Barrel Age	Barrel Seasoning	Filtering	
1	1	Dijon	3	No	Wild	High	Soft	New	Kiln	Yes	
2	2	Bernard	4	Yes	Cultured	Low	Hard	New	Kiln	No	
3	3	Bernard	2	Yes	Wild	High	Hard	New	Air	No	
4	4	Bernard	1	Yes	Wild	Low	Hard	2 Years	Air	Yes	
5	5	Dijon	1	Yes	Wild	Low	Soft	New	Kiln	No	
6	1	Dijon	2	Yes	Cultured	Low	Soft	2 Years	Air	Yes	
7	2	Dijon	1	No	Cultured	High	Hard	New	Air	No	
8	3	Dijon	4	No	Cultured	High	Soft	New	Air	Yes	
9	4	Dijon	4	Yes	Wild	High	Hard	New	Kiln	Yes	
10	5	Bernard	2	No	Wild	Low	Hard	New	Air	Yes	
11	1	Bernard	3	Yes	Wild	High	Hard	New	Air	No	
12	2	Dijon	1	Yes	Wild	High	Soft	2 Years	Air	Yes	
13	3	Bernard	4	Yes	Wild	Low	Soft	2 Years	Kiln	No	
14	4	Bernard	3	No	Cultured	Low	Soft	New	Air	No	
15	5	Dijon	2	No	Cultured	High	Soft	2 Years	Kiln	No	
16	1	Dijon	1	Yes	Cultured	Low	Hard	2 Years	Kiln	Yes	
17	2	Bernard	2	No	Cultured	Low	Hard	2 Years	Air	No	
18	3	Dijon	2	Yes	Cultured	High	Hard	New	Kiln	No	
19	4	Bernard	1	No	Cultured	Low	Hard	New	Kiln	Yes	
20	5	Bernard	3	Yes	Cultured	High	Hard	2 Years	Air	Yes	
21	1	Bernard	4	No	Wild	High	Hard	2 Years	Kiln	No	
22	2	Bernard	4	Yes	Cultured	High	Soft	2 Years	Kiln	Yes	
23	3	Dijon	1	No	Wild	Low	Hard	2 Years	Kiln	No	
24	4	Dijon	2	Yes	Wild	High	Soft	2 Years	Air	No	
25	5	Bernard	1	No	Cultured	High	Soft	New	Kiln	No	
26	1	Bernard	2	No	Wild	Low	Soft	New	Kiln	Yes	
27	2	Dijon	3	No	Wild	High	Hard	New	Kiln	Yes	
28	3	Bernard	1	No	Wild	High	Soft	2 Years	Air	Yes	
29	4	Bernard	3	Yes	Cultured	High	Soft	2 Years	Kiln	No	
30	5	Dijon	3	Yes	Wild	Low	Hard	2 Years	Kiln	No	
31	1	Dijon	4	No	Cultured	Low	Hard	2 Years	Air	No	
32	2	Bernard	2	Yes	Wild	Low	Soft	New	Kiln	Yes	
33	3	Bernard	3	No	Cultured	Low	Soft	2 Years	Kiln	Yes	
34	4	Dijon	2	No	Cultured	High	Hard	2 Years	Kiln	Yes	
35	5	Dijon	4	Yes	Cultured	Low	Soft	New	Air	Yes	
36	1	Bernard	1	Yes	Cultured	High	Soft	New	Air	No	
37	2	Dijon	3	No	Wild	Low	Soft	2 Years	Air	No	
38	3	Dijon	3	Yes	Cultured	Low	Hard	New	Air	Yes	
39	4	Dijon	4	No	Wild	Low	Soft	New	Air	No	
40	5	Bernard	4	No	Wild	High	Hard	2 Years	Air	Yes	

Design Evaluation

The Design Evaluation outline provides different ways to evaluate your design.

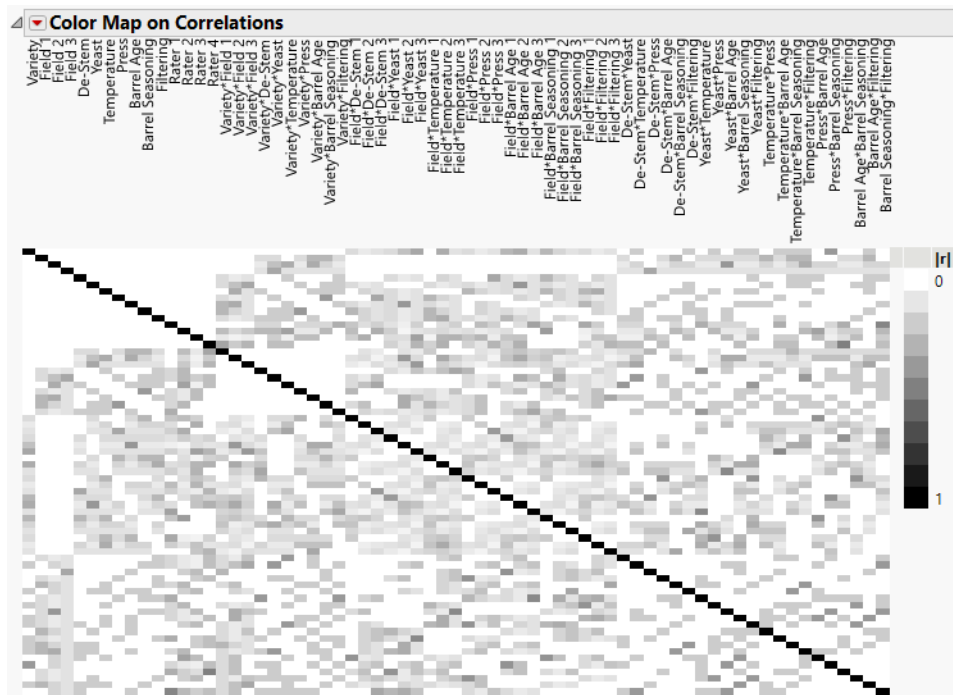
Note: For more information about the Design Evaluation outline, see the [“Evaluate Designs”](#) chapter on page 433.

For this example, open the Design Evaluation outline, and examine the Color Map on Correlations, the Alias Matrix, and Design Diagnostics.

Color Map on Correlations

The Color Map on Correlations shows the absolute value of the correlation between any two effects that appear in either the Model or the Alias Terms outline.

Figure 4.6 Color Map on Correlations



The main effects are represented by the 15 terms in the upper left corner of the map. The white corresponding to the correlations of main effects with other main effects indicate correlations of 0. This means that all main effects are orthogonal and can be estimated independently.

The only black in Figure 4.6 is on the main diagonal. The color indicates absolute correlations of one, reflecting that each term is perfectly correlated with itself. It follows that no main effect is completely confounded with any two-way interaction. In fact, the absolute values of the correlations of main effects with two-way interactions are fairly low. This means that estimates of main effects might be only slightly biased by the presence of active two-way interactions.

Tip: Position your mouse pointer over cells in the color map to see the absolute correlations between effects. Right-click below the legend to save the correlations to a data table.

Alias Matrix

In the Alias Matrix, model effects are listed in the column on the left. For a given model effect, a column entry indicates the degree to which the column effect (if active) biases the estimate of the model effect.

Figure 4.7 Partial View of Alias Matrix

Alias Matrix									
Effect	Variety*Field 1	Variety*Field 2	Variety*Field 3	Variety*De-Stem	Variety*Yeast	Variety*Temperature	Variety*Press	Variety*Barrel Age	
Intercept	0	0	0	0	0	0	0	0	0
Variety	0	0	0	0	0	0	0	0	0
Field 1	0	0	0	0.141	0.141	0	-0.14	0.283	
Field 2	0	0	0	0.082	-0.24	-0.33	0.082	0.327	
Field 3	0	0	0	-0.12	0.346	0.115	-0.12	-0.12	
De-Stem	0.141	0.082	-0.12	0	0	-0.2	0	0	
Yeast	0.141	-0.24	0.346	0	0	0	-0.2	0	
Temperature	0	-0.33	0.115	-0.2	0	0	0	-0.2	
Press	-0.14	0.082	-0.12	0	-0.2	0	0	0	
Barrel Age	0.283	0.327	-0.12	0	0	-0.2	0	0	
Barrel Seasoning	0.283	0.163	0.115	0	-0.2	0.2	0.4	0	
Filtering	-0.14	-0.08	0.115	-0.2	0.2	0.2	0	0.2	
Rater 1	0	0	0	-0.16	-0.16	0	-0.16	0.158	
Rater 2	-0.26	0.149	-0.21	-0.27	0.274	-0.37	-0.09	-0.09	
Rater 3	0.091	-0.05	0.075	0	-0.19	-0.06	-0.26	-0.26	
Rater 4	0.354	-0.2	0.289	0	0.25	-0.25	0	0	

For example, consider the model effect Barrel Seasoning. If Variety*Press is active, then the expected value of the estimate for the Barrel Seasoning effect differs from an unbiased estimate of that effect. The amount by which it differs is equal to 0.4 times the effect of Variety*Press. Therefore, what appears to be a significant Barrel Seasoning estimated effect could in reality be a significant Variety*Press effect.

Design Diagnostics

The Design Diagnostics outline provides information about the efficiency of the design. Efficiency measures compare your design to a theoretically optimal design, which might not exist. The efficiency values are ratios, expressed as percents, of the efficiency of your design to the efficiency of this optimal design. For more information about the efficiency measures, see [“Estimation Efficiency”](#) on page 460 in the “Evaluate Designs” chapter.

Figure 4.8 Design Diagnostics Outline

Design Diagnostics	
D Optimal Design	
D Efficiency	100
G Efficiency	100
A Efficiency	100
Average Variance of Prediction	0.4
Design Creation Time (seconds)	0.05

Notice that the D-, G-, and A-efficiency values are all 100%. Because your design is orthogonal for main effects, the design is optimal for the main effects model relative to all three efficiency criteria.

The first line in the Design Diagnostics outline indicates that your design was constructed to optimize the D-efficiency criterion. See the Optimality Criterion description in [“Custom Design Options”](#) on page 107. In this case, your design has D Efficiency of 100%.

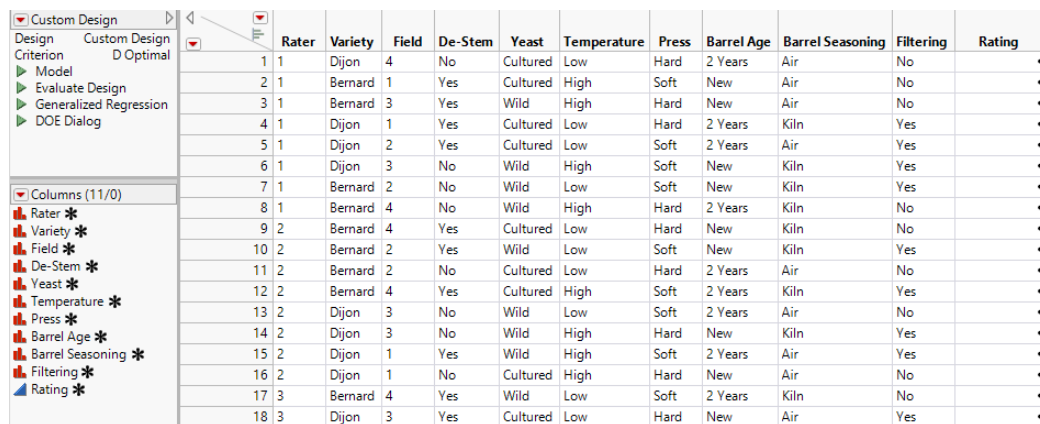
Output Options

Specify the order of runs in your data table using the Output Options panel. The default selection, **Randomize within Blocks**, is appropriate for this example. Simply click **Make Table**.

A Custom Design table is created and opens, similar to the one in Figure 4.9.

Note: Your table might look different because the algorithm that creates it uses a random starting design. To obtain the precise table shown in Figure 4.9, follow the steps in [“Duplicate Results \(Optional\)”](#) on page 74.

Figure 4.9 Custom Design Table



	Rater	Variety	Field	De-Stem	Yeast	Temperature	Press	Barrel Age	Barrel Seasoning	Filtering	Rating
1	1	Dijon	4	No	Cultured	Low	Hard	2 Years	Air	No	•
2	1	Bernard	1	Yes	Cultured	High	Soft	New	Air	No	•
3	1	Bernard	3	Yes	Wild	High	Hard	New	Air	No	•
4	1	Dijon	1	Yes	Cultured	Low	Hard	2 Years	Kiln	Yes	•
5	1	Dijon	2	Yes	Cultured	Low	Soft	2 Years	Air	Yes	•
6	1	Dijon	3	No	Wild	High	Soft	New	Kiln	Yes	•
7	1	Bernard	2	No	Wild	Low	Soft	New	Kiln	Yes	•
8	1	Bernard	4	No	Wild	High	Hard	2 Years	Kiln	No	•
9	2	Bernard	4	Yes	Cultured	Low	Hard	New	Kiln	No	•
10	2	Bernard	2	Yes	Wild	Low	Soft	New	Kiln	Yes	•
11	2	Bernard	2	No	Cultured	Low	Hard	2 Years	Air	No	•
12	2	Bernard	4	Yes	Cultured	High	Soft	2 Years	Kiln	Yes	•
13	2	Dijon	3	No	Wild	Low	Soft	2 Years	Air	No	•
14	2	Dijon	3	No	Wild	High	Hard	New	Kiln	Yes	•
15	2	Dijon	1	Yes	Wild	High	Soft	2 Years	Air	Yes	•
16	2	Dijon	1	No	Cultured	High	Hard	New	Air	No	•
17	3	Bernard	4	Yes	Wild	Low	Soft	2 Years	Kiln	No	•
18	3	Dijon	3	Yes	Cultured	Low	Hard	New	Air	Yes	•

Note the following:

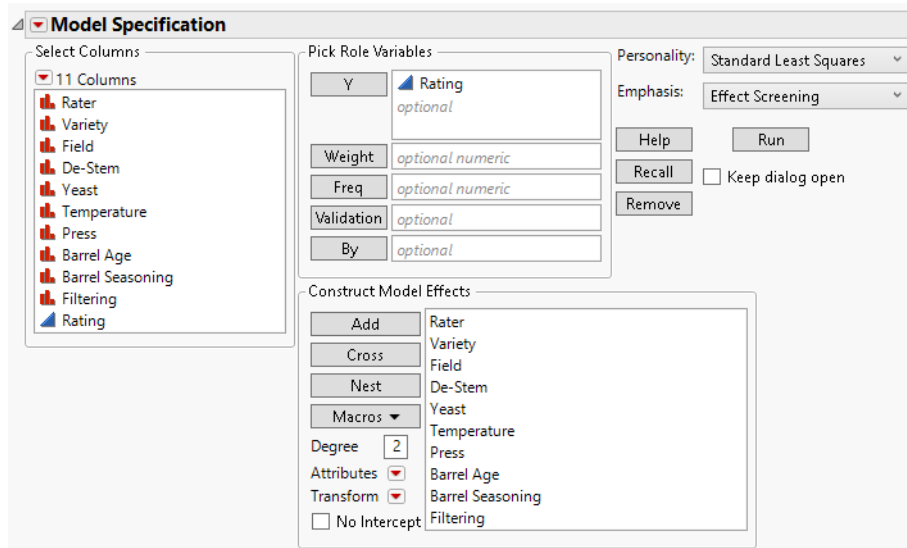
- In the Table panel, the Model, Evaluate Design, and DOE Dialog scripts are added during the design creation process. The Model script opens a Fit Model window containing the effects that you specified as Necessary in the Custom Design dialog. The DOE Dialog script re-creates the window used to generate the design table.
- In the Columns panel, the asterisks to the right of the factors and response indicate column properties that have been saved to the columns in the data table. These column properties are used in the analysis of the data. For more information about column properties, see [“Factors”](#) on page 88 and [“Factor Column Properties”](#) on page 94.

Analyze the Data

Now you are ready to run your experiment, gather the Rating data, and insert the results in the Rating column of your Custom Design table.

1. Select **Help > Sample Data Library** and open Design Experiment/Wine Data.jmp.
The Wine Data.jmp table is exactly the same as the Custom Design table shown in Figure 4.9, except that it contains your recorded experimental results.
2. In the Table panel, click the green triangle next to the **Model** script.

Figure 4.10 Fit Model Dialog for Wine Experiment



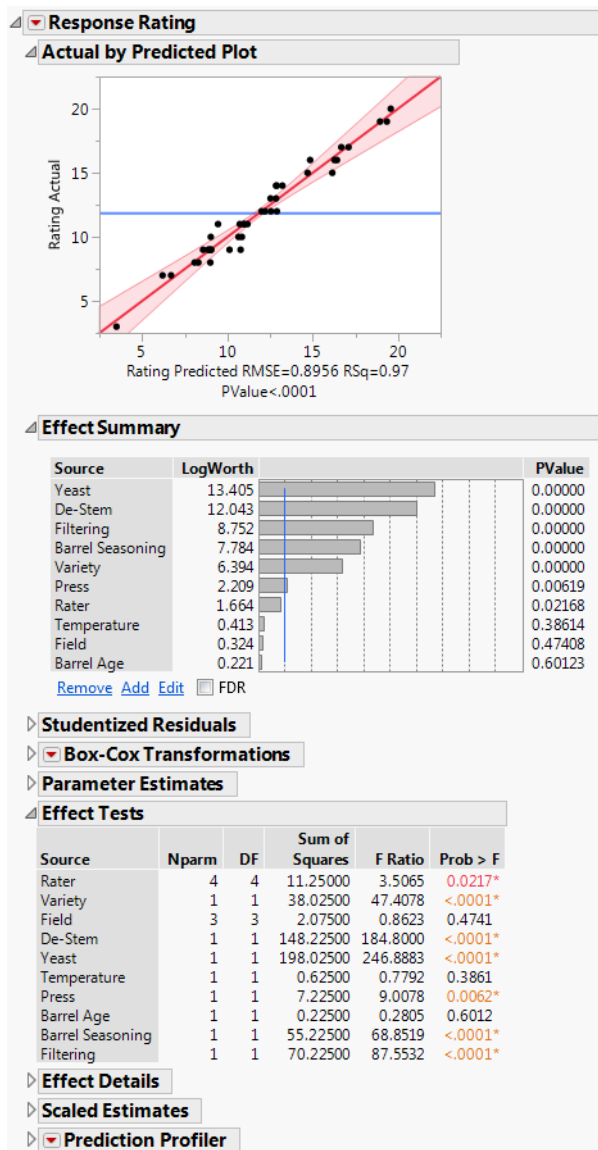
Notice that Rater, the blocking factor, is added as a fixed effect, rather than as a random block effect. This is appropriate because the five raters were specifically chosen and are not a random sample from a larger population.

3. Click **Run**.

Interpret the Full Model Results

The results are shown below.

Figure 4.11 Partial Model Fit Results



Note the following:

- The Actual by Predicted Plot shows no obvious evidence of lack of fit.
- The model is significant, as indicated by the Actual by Predicted Plot and by the p -value beneath it.
- The Effect Tests report indicates that seven of the model terms are significant at the 0.05 level. Field, Temperature, and Barrel Age are not significant.

- The Effect Summary report lists these effects in decreasing order of significance. Larger LogWorth values correspond to smaller PValues and greater significance.

Reduce the Model

Reduce the model by removing the effects that you identified as inactive:

1. In the Effect Summary report, press the Control key and hold it as you select Temperature, Field, and Barrel Age.
2. Click **Remove**.

The report updates to show the model fit with these three effects removed.

Interpret the Reduced Model Results

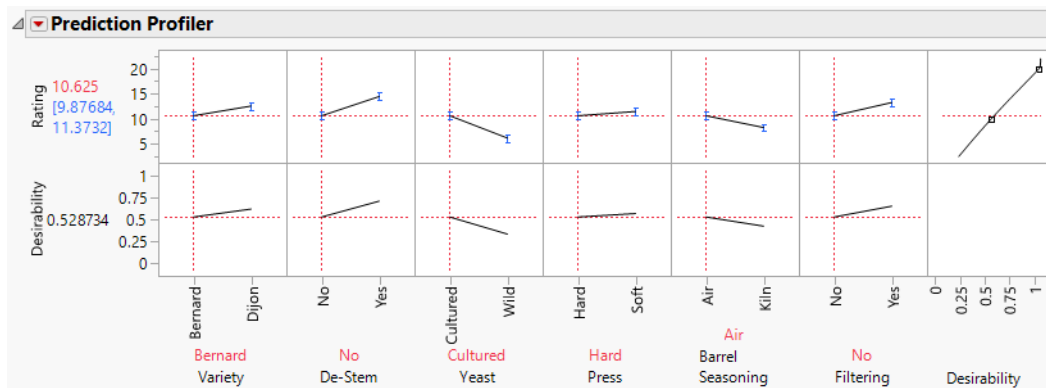
The Actual by Predicted Plot for the reduced model shows no lack of fit issues. The Effect Summary and the Effect Test report show that the remaining seven terms are significant at the 0.05 level.

Figure 4.12 shows the Prediction Profiler. Recall that you specified a response goal of Maximize, with lower and upper limits of 0 and 20. Setting these limits caused a Response Limits column property to be saved to the Rating column in the Custom Design table. The Prediction Profiler uses the Response Limits information to construct a Desirability function, which appears in the right-most plot in the top row in Figure 4.12. The bottom row displays Desirability traces.

The first six plots in the top row show traces of the predicted model. For each factor, the line in the plot shows how Rating varies when all other factors are set at the values defined by the red dashed vertical lines. By default, the profiler appears with categorical factors set at their low settings. By varying the settings for the factors, you can see how the predicted Rating for wines changes. Notice that a confidence interval is given for the mean predicted Rating.

Observe that Rater is not included among the factors shown in the profiler. This is because Rater is a block variable. You included Rater to explain variation, but Rater is not of direct interest in terms of optimizing process factor settings. The predicted Rating for a wine with the given settings is the average of the predicted ratings for that wine by all raters.

Figure 4.12 Profiler for Reduced Model



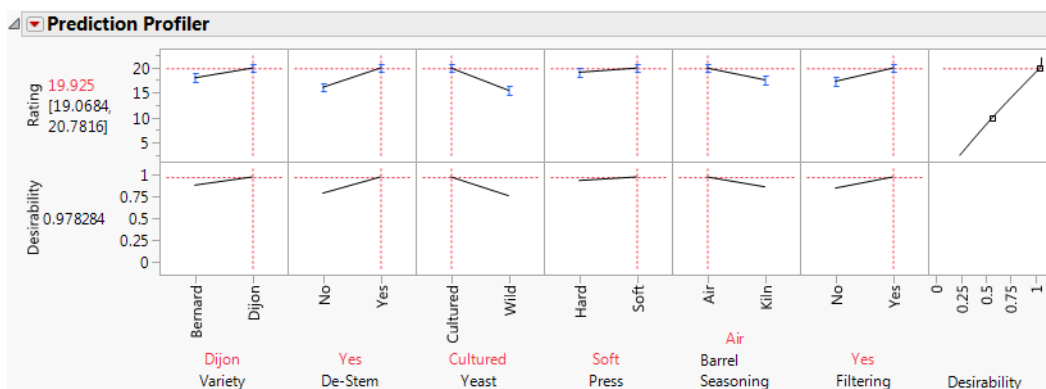
Optimize Factor Settings

You would like to identify settings that maximize Rating across raters.

1. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

The red dashed vertical lines in the Prediction Profiler update to show optimal settings for each factor. The optimal settings result in a predicted rating of 19.925. In general, there can be different sets of factor settings that result in the same optimal value.

Figure 4.13 Prediction Profiler with Factor Settings Optimized



2. To see predicted ratings for all runs, save the Prediction Formula. Click the Response Rating red triangle and select **Save Columns > Prediction Formula**.

A column called Pred Formula Rating is added to the data table. Note that one of the runs, row 33, was given the maximum rating of 20 by Rater 5. The predicted rating for that run by Rater 5 is 19.550. But the row 33 trial was run at the optimal settings. The predicted

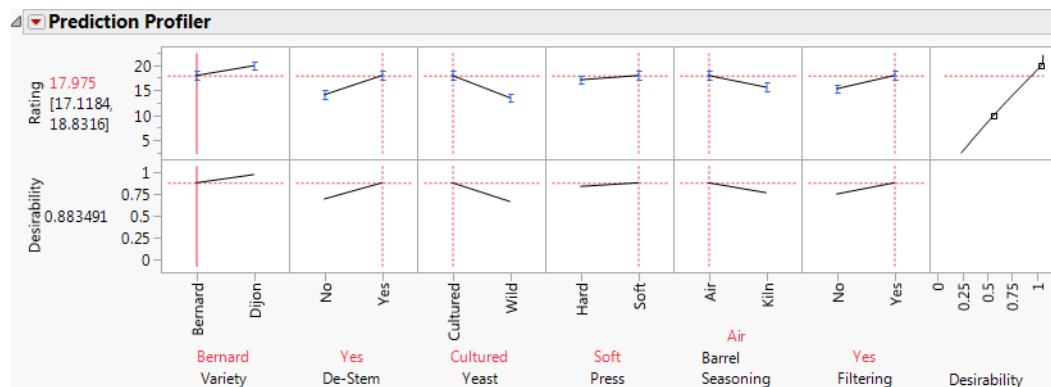
value of 19.925 given for these settings in the Prediction Profiler is obtained by averaging the predicted ratings for that run over all five raters.

Lock a Factor Level

When you maximized desirability, you learned that the optimal rating is achieved with the Dijon variety of grapes (Figure 4.13). Your manager points out that it would be cost-prohibitive to replant the fields that are growing Bernard grapes with young Dijon vines. Therefore, you need to find optimal process settings and the predicted rating for Bernard grapes.

1. In the Variety plot of the Prediction Profiler, drag the red dashed vertical line to Bernard.
2. Press Control and click in one of the Variety plots.
The Factor Settings window appears.
3. Select **Lock Factor Setting** and click **OK**.
4. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

Figure 4.14 Prediction Profiler with Optimal Settings for Bernard Variety



The optimal settings are unchanged because the model contains no interaction terms. The predicted rating at these settings is 17.975.

Profiler with Rater

If you want to see the Profiler traces for the levels of Rater, perform the following steps:

1. Click the Prediction Profiler red triangle and select **Reset Factor Grid**.

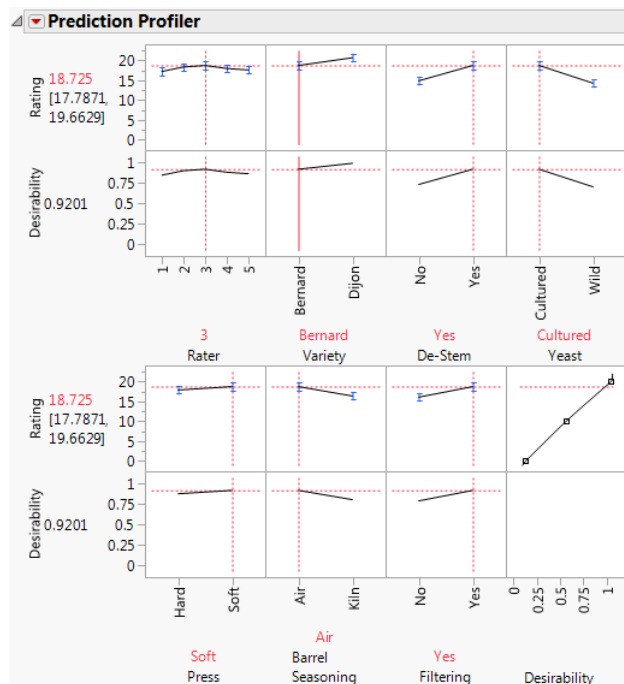
A Factor Settings window appears with columns for all of the factors, including **Rater**. The box under Rater and next to **Show** is not checked. This indicates that Rater is not shown in the Prediction Profiler.

2. Check the box under Rater in the row corresponding to **Show**.
3. Deselect the box under Rater in the row corresponding to **Lock Factor Setting**.
4. Click **OK**.

The Profiler updates to show a plot for Rater.

5. Click in either plot above Rater.

Figure 4.15 Profiler for Reduced Model Showing Rater



A dashed vertical red line appears. Drag this line to see the traces for each of the raters. Keep in mind that Variety is still locked at Bernard. To unlock Variety, press Control and click in one of the Variety plots. In the Factor Settings window that appears, deselect **Lock Factor Setting**.

Summary

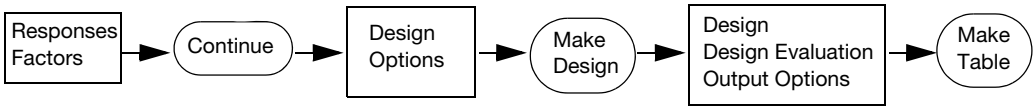
In your wine tasting experiment, using only 40 runs, you have identified six (out of nine) factors that have an effect on ratings for Pinot Noir grapes. You found that you could achieve a predicted rating of 19.925 (out of a possible 20) at the optimal settings for those factors. You also identified optimal settings for both varieties of grapes.

In this section, you constructed a design using the outlines in the Custom Design window. The next section explains each outline and the design steps in more detail.

Custom Design Window

The Custom Design window updates as you work through the design steps. The outlines that appear, separated by buttons that update the window, follow the flow in Figure 4.16.

Figure 4.16 Custom Design Flow



This section describes the outlines in the Custom Design window.

Responses

Use the Responses outline to specify one or more responses.

Tip: When you have completed the Responses outline, consider selecting **Save Responses** from the red triangle menu. This option saves the response names, goals, limits, and importance values in a data table that you can later reload in DOE platforms.

Figure 4.17 Responses Outline

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.

optional item

- Add Response** Enters a single response with a goal type of Maximize, Match Target, Minimize, or None. If you select Match Target, enter limits for your target value. If you select Maximize or Minimize, entering limits is not required but can be useful if you intend to use desirability functions.
- Remove** Removes the selected responses.
- Number of Responses** Enters additional responses so that the number that you enter is the total number of responses. If you have entered a response other than the default Y, the Goal for each of the additional responses is the Goal associated with the last response entered. Otherwise, the Goal defaults to Match Target. Click the Goal type in the table to change it.

The Responses outline contains the following columns:

Response Name The name of the response. When added, a response is given a default name of Y, Y2, and so on. To change this name, double-click it and enter the desired name.

Goal, Lower Limit, Upper Limit The Goal tells JMP whether you want to maximize your response, minimize your response, match a target, or that you have no response goal. JMP assigns a Response Limits column property, based on these specifications, to each response column in the design table. It uses this information to define a desirability function for each response. The Profiler and Contour Profiler use these desirability functions to find optimal factor settings. For further details, see the Profiler chapter in *Profilers* and [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

- A Goal of Maximize indicates that the best value is the largest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Minimize indicates that the best value is the smallest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Match Target indicates that the best value is a specific target value. The default target value is assumed to be midway between the Lower Limit and Upper Limit.
- A Goal of None indicates that there is no goal in terms of optimization. No desirability function is constructed.

Note: If your target response is not midway between the Lower Limit and the Upper Limit, you can change the target after you generate your design table. In the data table, open the Column Info window for the response column (**Cols > Column Info**) and enter the desired target value.

Importance When you have several responses, the Importance values that you specify are used to compute an overall desirability function. These values are treated as weights for the responses. If there is only one response, then specifying the Importance is unnecessary because it is set to 1 by default.

Editing the Responses Outline

In the Responses outline, note the following:

- Double-click a response to edit the response name.
- Click the goal to change it.
- Click on a limit or importance weight to change it.
- For multiple responses, you might want to enter values for the importance weights.

Response Limits Column Property

The Goal, Lower Limit, Upper Limit, and Importance that you specify when you enter a response are used in finding optimal factor settings. For each response, the information is saved in the generated design data table as a Response Limits column property. JMP uses this information to define the desirability function. The desirability function is used in the Prediction Profiler to find optimal factor settings. For further details about the Response Limits column property and examples of its use, see [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

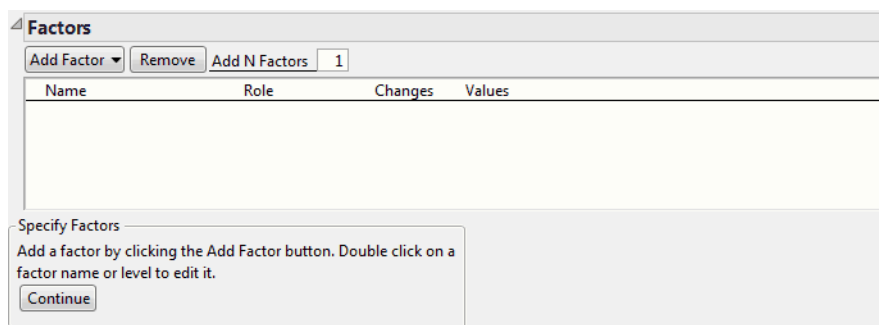
If you do not specify a Lower Limit and Upper Limit, JMP uses the range of the observed data for the response to define the limits for the desirability function. Specifying the Lower Limit and Upper Limit gives you control over the specification of the desirability function. For more details about the construction of the desirability function, see the Profiler chapter in *Profilers*.

Factors

Add factors in the Factors outline.

Tip: When you have completed the Factors outline, consider selecting **Save Factors** from the red triangle menu. This saves the factor names, roles, changes, and values in a data table that you can later reload.

Figure 4.18 Factors Outline



Add Factor Select the factor type. See [“Factor Types”](#) on page 89.

Remove Removes the selected factors.

Note: If you attempt to remove all factors after clicking the Continue or Back button, one continuous factor remains. You can delete it after you add new factors.

Add N Factors Adds multiple factors. Enter the number of factors to add, click **Add Factor**, and then select the factor type. Repeat **Add N Factors** to add multiple factors of different types.

Factors Outline

The Factors outline contains the following columns:

Name The name of the factor. When added, a factor is given a default name of X1, X2, and so on. To change this name, double-click it and enter the desired name.

Role Specifies the Design Role of the factor. The Design Role column property for the factor is saved to the data table. This property ensures that the factor type is modeled appropriately.

Changes Indicates whether the factor levels are Easy, Hard, or Very Hard to change. Click the default value of Easy to change it. When you specify factors as Hard or Very Hard to change, your design reflects these restrictions on randomization. A factor cannot be designated as Very Hard unless the Factors list contains a factor designated as Hard. The Factor Changes column property is saved to the data table. See [“Changes and Random Blocks”](#) on page 92.

Values The experimental settings for the factors. To insert Values, click the default values and enter the desired values.

Editing the Factors Outline

In the Factors outline, note the following:

- To edit a factor name, double-click the factor name.
- Categorical factors have a down arrow to the left of the factor name. Click the arrow to add a level.
- To remove a factor level, click the value, click **Delete**, and click outside the text box.
- To modify the entry under Changes, click the value in the Changes column and select the appropriate entry.
- To edit a value, click the value in the Values column.

Factor Types

To choose a factor type, click **Add Factor** in Custom Design.

Note: A Design Role column property containing each factor's role is added to that factor's column in the design table that is generated. The Design Role column property ensures that the factor is modeled correctly.

Continuous Numeric data types only. A continuous factor is a factor that you can conceptually set to any value between the lower and upper limits you supply, given the limitations of your process and measurement system.

Discrete Numeric Numeric data types only. A discrete numeric factor can assume only a discrete number of values. These values have an implied order.

The default values for a discrete numeric factor with k levels, where $k > 2$, are the integers 1, 2, ..., k . The default values for a discrete numeric factor with $k = 2$ levels are -1 and 1. Replace the default values with the settings that you plan to use in your experiment.

Note: Not all levels of a discrete numeric factor appear in the design. The levels that appear are determined by your specifications in the Model outline. If you need all levels to appear in your design, consider using the Screening Design platform.

In the assumed model, the effects for a discrete numeric factor with k levels include polynomial terms in that effect through order $k-1$. For k greater than 6, powers up to the 5th level are included. The Estimability for polynomial effects (powers of two or higher) is set to If Possible. This allows the algorithm to use the multiple levels as permitted by the run size. If the polynomial terms are not included, then a main effects only design is created. For more information about how discrete numeric factors are treated in the assumed model, see [“Model”](#) on page 98.

Fit Model treats a discrete numeric factor as a continuous predictor. The Model script that is saved to the design table does not contain any polynomial terms of order greater than two.

Categorical Either numeric or character data types. The data type in the resulting data table is categorical. The value order of the levels is the order of the values, as entered from left to right. This ordering is saved in the Value Order column property after the design data table is created.

Blocking Either numeric or character data types. A blocking factor is a special type of categorical factor that can enter the model only as a main effect. When you define a blocking factor, you specify the number of runs per block. The RunsPerBlock column property is saved to the design table. The Default run size always assumes that there are at least two blocks. If you specify a run size that is not an integer multiple of the number of runs per block, JMP tries to balance the design to the extent possible. In balancing the design, JMP ensures that there are at least two runs per block.

Covariate Either numeric or character data types. The values of a covariate factor are measurements on experimental units that are known in advance of an experiment. Covariate values are selected to ensure the optimality of the resulting design relative to the

optimality criterion. See [“Changes and Random Blocks”](#) on page 92 and [“Covariates with Hard-to-Change Levels”](#) on page 126.

JMP obtains the covariate factors and their values from a data table that contains the measured covariates for the available experimental units. Make this data table your current data table. When you select Covariate, a list of columns in the current data table opens, and you select the columns containing covariates from this list. Your design must have no more runs than the number of rows in the covariate table. If the design has fewer runs than the number of covariate rows then the design table includes a Covariate Row Index column. This column indicates the row from the covariate table that corresponds to each experimental run.

In some situations, you might want to select a small set of design points from a larger set of candidate settings. For example, you might have multiple measurement columns (factors) for a large batch of units. You want to treat the measurements for each unit as a candidate run. From these candidate runs, you want to select a small but optimal collection for which you measure a response. In this case, make the data table of all candidate runs the active table, select Add Factor > Covariate, and enter all of your measurement columns as covariates. Specify your desired run size. The Custom Design platform identifies an optimal collection of design settings.

Note: You cannot specify a Number of Runs or Number of Whole Plots that exceeds the number of rows in the covariate’s data table.

Mixture Continuous factors that represent ingredients in a mixture. The values for a mixture factor must sum to a constant. By default, the values for all mixture factors sum to one. To set the sum of the mixture components to some other positive value, select **Advanced Options > Mixture Sum** from the red triangle menu. The Mixture column property is saved to the data table.

Constant Either numeric or character data types. A constant factor is a factor whose values are fixed during an experiment. Constant factors are not included in the Model outline or in the Model script that is saved to the data table.

Uncontrolled Either numeric or character data types. An uncontrolled factor is one whose values cannot be controlled during production, but it is a factor that you want to include in the model. It is assumed that you can record the factor’s value for each experimental run.

An empty column with a Continuous Modeling Type is created in the design table. You can change the column’s Data Type and Modeling Type in the Column Info window if required. Enter your data in this column. Uncontrolled factors are included in the Model outline and the Model script that is saved to the data table.

Changes and Random Blocks

Specifying the relative difficulty of changing a factor from run to run is useful in industrial experimentation. It is often convenient to make several runs while keeping factors that are hard-to-change fixed at some setting. A Changes value of Hard results in a split-plot design. A Changes value of Very Hard results in a split-split-plot design or a two-way split-plot design.

You can set Changes for Continuous, Discrete Numeric, Categorical, and Mixture factors to Hard and Very Hard. To set a factor to Very Hard, the list must contain another factor that is set to Hard.

You can set Changes for a Covariate factor to Hard. In this case, all other covariates are also set to Hard and the remaining factors are set to Easy. The algorithm requires a combination of row exchange and coordinate exchange. For this reason, even moderately sized designs might take some time to generate.

For designs with Hard or Very Hard to change factors, Custom Design strives to find a design that is optimal, given your specified optimality criterion. See [“Optimality Criteria”](#) on page 127. For more information about the methodology used to generate split-plot designs, see Jones and Goos (2007). For more information about designs with hard-to-change covariates, see Jones and Goos (2015).

Figure 4.19 shows a split-split-plot scenario, using the factors from the Cheese Factors.jmp sample data table (located in the Design Experiment folder).

Figure 4.19 Factors and Design Generation Outline for a Split-Split-Plot Design

Custom Design

Responses

Factors

Add Factor ▼ Remove Add N Factors

Name	Role	Changes	Values
storage 1	Continuous	Very Hard	-1 1
storage 2	Categorical	Very Hard	L1 L2
curds 1	Continuous	Hard	-1 1
curds 2	Continuous	Hard	-1 1
curds 3	Categorical	Hard	L1 L2
curds 4	Continuous	Hard	-1 1
curds 5	Categorical	Hard	L1 L2 L3 L4
cheese 1	Continuous	Easy	-1 1
cheese 2	Continuous	Easy	-1 1

Define Factor Constraints

Model

Alias Terms

Design Generation

☐ Hard to change factors can vary independently of Very Hard to change factors.

Number of Whole Plots

Number of Subplots

Number of Runs:

☐ Minimum 14

☐ Default 30

☒ User Specified

Make Design

If you assign Changes as Hard for one or more factors, but no factors are assigned Changes that are Very Hard, a categorical factor called Whole Plots is added to the design. This situation results in a split-plot design:

- Each level of Whole Plots corresponds to a block of constant settings of the hard-to-change factors.
- The Model script in the design table applies the Random Effect attribute to the factor Whole Plots.
- The factor Whole Plots is assigned the Design Role column property with a value of Random Block.

When you designate Changes as both Hard and Very Hard, categorical factors called Subplots and Whole Plots are added to the design. This situation results in a split-split-plot design:

- Each level of Subplots corresponds to a block of constant settings of the hard-to-change factors.
- Each level of Whole Plots corresponds to a block of constant settings of the very-hard-to-change factors.
- The Model script in the design table applies the Random Effect attribute to the Whole Plots and Subplots effects.

- The levels of the hard-to-change factor are assumed to be nested within the levels of the very-hard-to-change factor by default.
- In the design table, both of the factors Whole Plots and Subplots are assigned the Design Role column property with a value of Random Block.

To construct a two-way split-plot design, select the **Hard to change factors can vary independently of Very Hard to change factors** option under Design Generation. The option crosses the levels of the hard-to-change factor with the levels of the very-hard-to-change factor. See [“Two-Way Split-Plot Designs”](#) on page 123.

Use the Number of Whole Plots and Number of Subplots text boxes to specify values for the numbers of whole plots or subplots. These boxes are initialized to suggested numbers of whole plots and subplots. For information about how these values are obtained, see [“Numbers of Whole Plots and Subplots”](#) on page 126.

For more information and scenarios that illustrate random block split-plot, split-split-plot, and two-way split-plot designs, see [“Designs with Randomization Restrictions”](#) on page 117. For more information about designs with hard-to-change covariates, see [“Covariates with Hard-to-Change Levels”](#) on page 126.

Factor Column Properties

For each factor, various column properties are saved to the data table. You can find more information about these column properties and related examples in [Appendix A, “Column Properties”](#).

Design Role Each factor is given the Design Role column property. The Role that you specify in defining the factor determines the value of its Design Role column property. When you add a random block under Design Generation, that factor is assigned the Random Block value. The Design Role property reflects how the factor is intended to be used in modeling the experimental data. Design Role values are used in the Augment Design platform. See [“Design Role”](#) on page 740 in the “Column Properties” appendix.

Factor Changes Each factor is assigned the Factor Changes column property. The value that you specify under Changes determines the value of its Factor Changes column property. The Factor Changes property reflects how the factor is used in modeling the experimental data. Factor Changes values are used in the Augment Design and Evaluate Design platforms. See [“Factor Changes”](#) on page 755 in the “Column Properties” appendix.

Coding If the Role is Continuous, Discrete Numeric, a continuous Covariate, or Uncontrolled, the Coding column property for the factor is saved. This property transforms the factor values so that the low and high values correspond to -1 and $+1$, respectively. See [“Coding”](#) on page 742 in the “Column Properties” appendix.

Value Order If the Role is Categorical or Blocking, the Value Order column property for the factor is saved. This property determines the order in which levels of the factor appear. See [“Value Order”](#) on page 759 in the “Column Properties” appendix.

Mixture If the Role is Mixture, the Mixture column property for the factor is saved. This property indicates the limits for the factor and the mixture sum. It also enables you to choose the coding for the mixture factors. See [“Mixture”](#) on page 750 in the “Column Properties” appendix.

RunsPerBlock For a blocking factor, indicates the maximum allowable number of runs in each block. When a Blocking factor is specified in the Factors outline, the RunsPerBlock column property is saved for that factor. See [“RunsPerBlock”](#) on page 765 in the “Column Properties” appendix.

Define Factor Constraints

Note: If you are working in Covering Arrays, see the [“Covering Arrays”](#) chapter on page 579.

Use Define Factor Constraints to restrict the design space. Unless you have loaded a constraint or included one as part of a script, the **None** option is selected. To specify constraints, select one of the other options:

Specify Linear Constraints Specifies inequality constraints on linear combinations of factors. Available only for factors with a Role of Continuous or Mixture. See [“Specify Linear Constraints”](#).

Note: When you save a script for a design that involves a linear constraint, the script expresses the linear constraint as a *less than or equal to* inequality ($\leq$).

Use Disallowed Combinations Filter Defines sets of constraints based on restricting values of individual factors. You can define both AND and OR constraints. See [“Use Disallowed Combinations Filter”](#).

Use Disallowed Combinations Script Defines disallowed combinations and other constraints as Boolean JSL expressions in a script editor box. See [“Use Disallowed Combinations Script”](#).

Specify Linear Constraints

In cases where it is impossible to vary continuous factors independently over the design space, you can specify linear inequality constraints. Linear inequalities describe factor level settings that are allowed.

Click **Add** to enter one or more linear inequality constraints.

Add Adds a template for a linear expression involving all the continuous factors in your design. Enter coefficient values for the factors and select the direction of the inequality to reflect your linear constraint. Specify the constraining value in the box to the right of the inequality. To add more constraints, click **Add** again.

Note: The Add option is disabled if you have already constrained the design region by specifying a Sphere Radius.

Remove Last Constraint Removes the last constraint.

Check Constraints Checks the constraints for consistency. This option removes redundant constraints and conducts feasibility checks. A JMP alert appears if there is a problem. If constraints are equivalent to bounds on the factors, a JMP alert indicates that the bounds in the Factors outline have been updated.

Use Disallowed Combinations Filter

This option uses an adaptation of the Data Filter to facilitate specifying disallowed combinations. For detailed information, see the JMP Reports chapter in *Using JMP*.

Select factors from the Add Filter Factors list and click **Add**. Then specify the disallowed combinations by using the slider (for continuous factors) or by selecting levels (for categorical factors).

The red triangle options for the Add Filter Factors menu are those found in the Select Columns panel of many platform launch windows. See the Get Started chapter in *Using JMP*.

When you click Add, the Disallowed Combinations control panel shows the selected factors and provides options for further control. Factors are represented as follows, based on their modeling types:

Continuous Factors For a continuous factor, a double-arrow slider that spans the range of factor settings appears. An expression that describes the range using an inequality appears above the slider. You can specify disallowed settings by dragging the slider arrows or by clicking on the inequality bounds in the expression and entering your desired constraints. In the slider, a solid blue highlight represents the disallowed values.

Categorical Factor For a categorical factor, the possible levels are displayed either as labeled blocks or, when the number of levels is large, as list entries. Select a level to disallow it. To select multiple levels, hold the Control key. The block or list entries are highlighted to indicate the levels that have been disallowed. When you add a categorical factor to the Disallowed Combinations panel, the number of levels of the categorical factor is given in parentheses following the factor name.

Disallowed Combinations Options

The control panel has the following controls:

Clear Clears all disallowed factor level settings that you have specified. This does not clear the selected factors.

Start Over Removes all selected factors and returns you to the initial list of factors.

AND Opens the Add Filter Factors list. Selected factors become an AND group. Any combination of factor levels specified within an AND group is disallowed.

To add a factor to an AND group later on, click the group's outline to see a highlighted rectangle. Select AND and add the factor.

To remove a single factor, select **Delete** from its red triangle menu.

OR Opens the Add Filter Factors list. Selected factors become a separate AND group. For AND groups separated by OR, a combination is disallowed if it is specified in at least one AND group.

Red Triangle Options for Factors

A factor can appear in several OR groups. An occurrence of the factor in a specific OR group is referred to as an *instance* of the factor.

Delete Removes the selected instance of the factor from the Disallowed Combinations panel.

Clear Selection Clears any selection for that instance of the factor.

Invert Selection Deselects the selected values and selects the values not previously selected for that instance of the factor.

Display Options Available only for categorical factors. Changes the appearance of the display. Options include showing each level as a block, list, or single category, or adding a check box next to each value.

Find Available only for categorical factors. Provides a text box beneath the factor name where you can enter a search string for levels of the factor. Press the Enter key or click outside the text box to perform the search. Once **Find** is selected, Find options appear in the red triangle menu, such as Does not contain, Match Case, Starts with or Ends with, or Clear Find.

Use Disallowed Combinations Script

Use this option to disallow particular combinations of factor levels using a JSL script. This option can be used with continuous factors or mixed continuous and categorical factors.

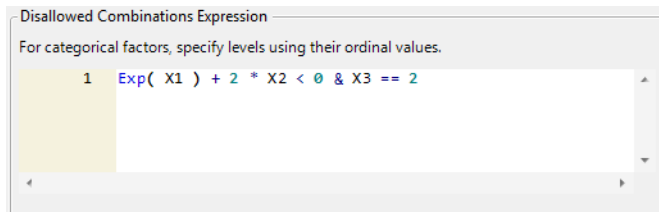
This option opens a script window where you insert a script that identifies the combinations that you want to disallow. The script must evaluate as a Boolean expression. When the expression evaluates as true, the specified combination is disallowed.

When forming the expression for a categorical factor, use the ordinal value of the level or the name of the level. If a factor's levels are high, medium, and low, specified in that order in the Factors outline, their associated ordinal values are 1, 2, and 3. For example, suppose that you have two continuous factors, X1 and X2, and a categorical factor X3 with three levels: L1, L2, and L3, in order. You want to disallow levels where the following is true:

$$e^{X_1} + 2X_2 < 0 \text{ and } X_3 = L2$$

Enter the expression `(Exp(X1) + 2*X2 < 0) & (X3 == 2)` into the script window.

Figure 4.20 Expression in Script Editor



(In the figure, unnecessary parentheses were removed by parsing.) Notice that functions can be entered as part of the Boolean expression. The expression `(Exp(X1) + 2*X2 < 0) & (X3 == "L2")` is also valid.

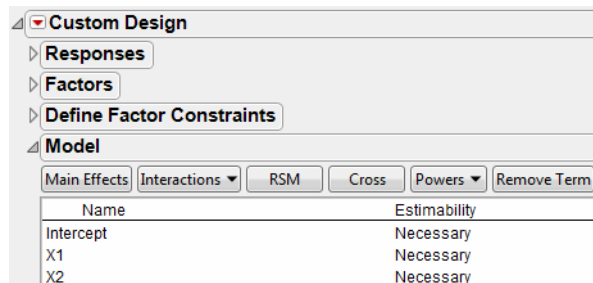
Model

Specify your assumed model (which contains all the effects that you want to estimate) in the Model outline. For each effect that you specify, you can designate that effect's Estimability. The Estimability value indicates whether it is Necessary to estimate that effect, or if you are content to estimate that effect If Possible.

The initial Model outline includes the main effects for all factors. If you have entered a discrete numeric factor with three or more levels, polynomial terms are also included in the initial model. The Estimability of second-and higher-order terms is set to If Possible. If you want to ensure that these terms are estimable, change their Estimability to Necessary.

Note: You can ensure that the estimability of discrete numeric polynomial terms is always set to Necessary. Select **File > Preferences > Platforms > DOE**. Check Discrete Numeric Powers Set to Necessary.

Figure 4.21 Model Outline



When you construct your design table, JMP saves a Model script to the data table. Except for discrete numeric factors, the Model script contains the effects shown in the Model outline. For a discrete numeric factor, the Model script contains only its main effect and quadratic term.

The Model outline contains the following buttons and fields:

Main Effects Adds main effects for all factors in the model, and polynomial terms for discrete numeric factors.

Interactions Adds interaction effects. If no factors are selected in the Factors outline, select 2nd, 3rd, 4th, or 5th to add all appropriate interactions up to that order. Add interactions up to a given order for specific factors by selecting the factor names in the Factors outline, selecting Interactions, and then specifying the appropriate order. Interactions between non-mixture and mixture factors, and interactions with blocking and constant factors, are not added.

RSM Adds interaction and quadratic terms up to the second order (response surface model terms) for all continuous factors. Categorical factors are not included in RSM terms. Main effects for non-mixture factors that interact with all the mixture factors are removed.

Cross Adds specific interaction terms. Select factor names in the Factors outline and effect names in the Model outline. Click Cross to add the crossed terms to the Model outline.

Powers Adds polynomial terms. If no factor names are selected in the Factors outline, adds polynomial terms for all continuous factors. If factor names are selected in the Factors outline, adds polynomial terms for only those factors. Select 2nd, 3rd, 4th, or 5th to add polynomial terms of that order.

Scheffé Cubic (Available for mixture factors.) Adds Scheffé cubic terms for all mixture factors. These terms are used to specify a mixture model with third-degree polynomial terms.

Remove Term Removes selected effects.

Name The name of the effect.

Estimability A designation of your need to estimate the effect.

- If Estimability is set to Necessary, the algorithm ensures that the effect is estimable.
- If Estimability is set to If Possible, the algorithm attempts to make that effect estimable, as permitted by the number of runs that you specify.

Except for polynomial terms for discrete numeric factors, all effects are specified as Necessary by default. Click an effect's Estimability value to change it.

Bayesian D-Optimality and Estimation of If Possible Effects

The Bayesian D-Optimal design approach obtains precise estimation of all Necessary terms while providing omnibus detectability (and some estimability) for If Possible terms. For more detail, see [“Response Surface Experiments”](#) on page 156 in the “Examples of Custom Designs” chapter and [“Bayesian D-Optimality”](#) on page 128.

Alias Terms

It is possible that effects *not* included in your assumed model are active. In the Alias Terms outline, add potentially active effects that are not in your assumed model but might bias the estimates of model terms. Once you generate your design, the Alias Matrix outline appears under Design Evaluation. The Alias Matrix entries represent the degree of bias imparted to model parameters by the effects that you specified in the Alias Terms outline. See the [“The Alias Matrix”](#) on page 771 in the “Technical Details” appendix.

By default, the Alias Terms outline includes all two-way interaction effects that are not in your Model outline (with the exception of terms involving blocking factors). Add terms using the buttons. For a description of how to use these buttons to add effects to the Alias Terms table, see [“Model”](#) on page 98.

Note: For interaction terms, only interactions of the selected order are added. This differs from the Model outline where interactions of the selected order and all lower orders are added to the model.

For example, suppose that you specify a design with three continuous factors. Your assumed model, specified in the Model outline, contains only those three main effects. You can afford only six runs. You want to see how estimates of the main effects might be biased by active two-way interactions and the three-way interaction.

The Alias Terms table includes all two-way interactions by default. You can add the three-way interaction by selecting **Interactions > 3rd**.

Figure 4.22 Alias Terms Outline

The 'Alias Terms' window has tabs for 'Main Effects', 'Interactions', 'RSM', 'Cross', 'Powers', and 'Remove Term'. The 'Interactions' tab is selected, showing a list of alias terms under the 'Name' column:

Name
X1*X2
X1*X3
X2*X3
X1*X2*X3

Once you specify six runs in the Design Generation outline and click **Make Design**, the Design Evaluation outline appears. Open the Design Evaluation outline and the Alias Matrix outline.

Figure 4.23 Alias Matrix

Effect	X1*X2	X1*X3	X2*X3	X1*X2*X3
Intercept	0	0	0.5	-0.5
X1	0	0	-0.5	0.5
X2	0.5	-0.5	0	0
X3	-0.5	0.5	0	0

The Alias Matrix indicates that each main effect is partially aliased with two of the interactions. See [“Alias Matrix”](#) on page 462 in the “Evaluate Designs” chapter and [“The Alias Matrix”](#) on page 771 in the “Technical Details” appendix.

Design Generation

The Design Generation outline gives you choices relating to the size and structure of the design. Typically, the input area has two parts:

- Design structure options
- Number of runs options

Figure 4.24 Design Generation Outline

The 'Design Generation' window contains the following options:

- ☐ Group runs into random blocks of size:
- Number of Center Points:
- Number of Replicate Runs:
- Number of Runs:**
 - ☐ Minimum 3
 - ☐ Default 8
 - ☒ User Specified
- Make Design**

Design Structure Options

Group runs into random blocks of size (Not available if a blocking factor is specified.) To construct a random block design, enter the number of runs that you want in each block.

When you specify the sample size, a factor called Random Block is created. Its levels define blocks of a size that is consistent with the block size that you entered, given the specified number of runs. If the number of runs is an integer multiple of the block size, the block sizes equal your specified value.

Number of Whole Plots Appears when you specify a hard or very-hard-to-change factor. The factor Whole Plots corresponds to the very-hard-to-change factors (split-split-plot design), if there are any, otherwise to the hard-to-change factors (split-plot design). JMP suggests a value for the number of whole plots that maximizes the information about the coefficients in the model. Or, you can enter a value for the number of whole plots. See [“Numbers of Whole Plots and Subplots”](#) on page 126.

Number of Subplots Appears when you specify a very-hard-to-change factor. The factor Subplots corresponds to the hard-to-change factors in the split-split-plot design. JMP suggests values for the number of whole plots and subplots that maximize the information about the coefficients in the model. Or, you can enter a value for the number of subplots. See [“Numbers of Whole Plots and Subplots”](#) on page 126.

Hard to change factors can vary independently of Very Hard to change factors Select this option to create a strip-plot (also known as two-way split-plot or split block) design. This option creates a design where the hard-to-change factors are randomized within the levels of the very-hard-to-change factors. They are *not* nested within the very-hard-to-change factors.

Number of Center Points Appears only if the design contains factors with a Continuous or Mixture factor type. Specify how many additional runs you want to add as center points to the design. A center point is a run whose setting for each continuous factor is midway between the high and low settings. See [“Center Points, Replicate Runs, and Testing”](#) on page 63 in the “Starting Out with DOE” chapter.

If a design contains both continuous and other types of factors, center points might not be balanced relative to the levels of the other factors. Custom Design chooses the center points to maximize the D-, I-, or alias efficiency of the design.

Number of Replicate Runs Specify the number of replicate trials that you want to add to the design. This does not replicate the entire design, but chooses the optimal design points to replicate. See [“Center Points, Replicate Runs, and Testing”](#) on page 63 in the “Starting Out with DOE” chapter.

Number of Runs Options

Minimum A lower bound on the number of runs necessary to avoid failures in design generation. When you select Minimum, the resulting design is saturated. There are no degrees of freedom for error.

Note: If you select the Minimum number of runs, there will be no error term for testing. You will not be able to test parameter estimates. This choice is appropriate only when the cost of additional runs is prohibitive.

Default Suggests the number of runs. This value is based on heuristics for creating a balanced design with at least four runs more than the Minimum number of runs.

User Specified Specify the number of runs that you want. Enter that value into the Number of Runs text box. This option enables you to balance the cost of additional runs against the potential gain in information.

Number of Runs This is the only option that appears when a covariate factor with Changes set to Easy is specified. The number of runs shown is the number of rows in the data table associated with your covariate or covariates. You can specify a smaller number of runs. In that case, the covariate runs that are selected optimize the design criterion.

Make Design

Once you have completed the Design Generation outline, click **Make Design**. Custom Design generates the design, presents it in the Design outline, and provides evaluation information in the Design Evaluation outline. Notice the controls in the Model and Alias Outlines are disabled. Use the Back button to change model or alias terms. The Output Options panel also appears, enabling you to create the design table.

Note: Sometimes several designs can optimize the optimality criterion. When this is the case, the design algorithm might generate different designs when you click the **Back** and **Make Design** buttons repeatedly.

Design

The Design outline shows the runs for a design that is optimal, given the conditions that you have specified. The runs might not appear to be randomized. You can select Run Order options in the Output Options panel before generating your design table.

Design Evaluation

The Design Evaluation outline provides a number of ways to evaluate the properties of the generated design. Open the Design Evaluation outline to see the following options:

Power Analysis Enables you to explore your ability to detect effects of given sizes.

Prediction Variance Profile Shows the prediction variance over the range of factor settings.

Fraction of Design Space Plot Shows how much of the model prediction variance lies below (or above) a given value.

Prediction Variance Surface Shows a surface plot of the prediction variance for any two continuous factors.

Estimation Efficiency For each parameter, gives the fractional increase in the length of a confidence interval compared to that of an idealized (orthogonal) design, which might not exist. Also gives the relative standard error of the parameters.

Alias Matrix Gives coefficients that indicate the degree by which the model parameters are biased by effects that are potentially active, but not in the model. You specify the terms representing potentially active effects in the Alias Terms table. See [“The Alias Matrix”](#) on page 771 in the “Technical Details” appendix.

Color Map on Correlations Shows the absolute correlation between effects on a plot using an intensity scale.

Design Diagnostics Indicates the optimality criterion used to construct the design. Also gives efficiency measures for your design. See Optimality Criterion in [“Custom Design Options”](#) on page 107 and [“Optimality Criteria”](#) on page 127.

Note: The Design Diagnostics outline does not provide the following statistics when the model includes factors with Changes set to Hard or Very Hard or with Estimability set to If Possible: D Efficiency, G Efficiency, A Efficiency.

For more information about the Design Evaluation outline, see [“Design Evaluation”](#) on page 449 in the “Evaluate Designs” chapter.

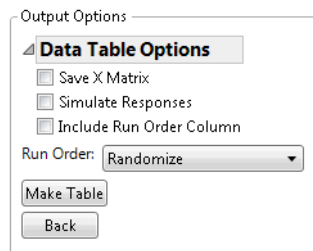
Output Options

Use the Output Options panel to perform the following tasks:

- specify how you want the custom design data table to appear
- construct the design table
- return to a previous point in the Custom Design window

In most cases, the Output Options panel appears as shown in Figure 4.25.

Figure 4.25 Output Options Panel



The Output Options panel contains these options:

- [“Data Table Options”](#) on page 105
- [“Run Order”](#) on page 105
- [“Make Table”](#) on page 106
- [“Back”](#) on page 106

Data Table Options

The **Data Table Options** includes the **Save X Matrix** option ([“Save X Matrix”](#) on page 112), the **Simulate Response** option ([“Simulate Responses”](#) on page 110) as well as an **Include Run Order Column** option.

Include Run Order Column Adds a column to the design data table that corresponds to the design order in the Design outline.

Tip: Use the Include Run Order Column when you save your data with a sorted run order. This provides a column of the randomized run order from the design outline.

Run Order

The **Run Order** options determine the order of the runs in the design table. Choices include the following:

Keep the Same Rows in the design table are in the same order as in the Design outline.

Sort Left to Right Columns in the design table are sorted from left to right.

Randomize Rows in the design table are in random order.

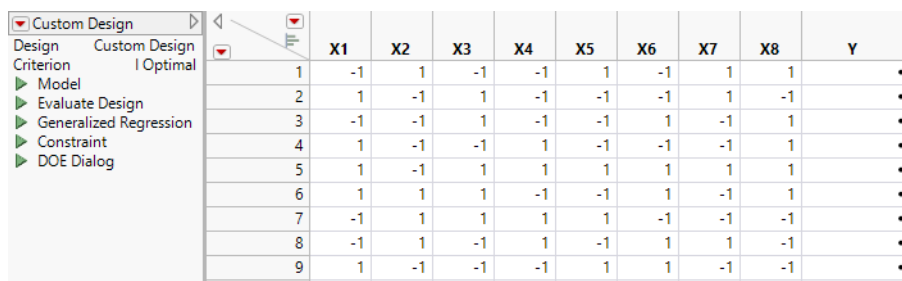
Sort Right to Left Columns in the design table are sorted from right to left.

Randomize within Blocks Rows in the design table are in random order within the blocks.

Make Table

Click **Make Table** to construct the custom design data table. In the Custom Design table, the Table panel (in the upper left) can contain scripts, as appropriate given your design. The Model, Evaluate Design, and DOE Dialog scripts are always provided. To run a script, click the green triangle next to the script name. If your design includes covariates, the design table includes a covariate row index column unless all covariate rows are used in the design.

Figure 4.26 Custom Design Table Showing Scripts



The screenshot shows the 'Custom Design' window. On the left, the 'Table' panel lists several scripts: 'Design', 'Custom Design', 'Criterion', 'I Optimal', 'Model', 'Evaluate Design', 'Generalized Regression', 'Constraint', and 'DOE Dialog'. Each script has a green triangle icon next to it. To the right of the scripts is a table with 10 columns: 'X1', 'X2', 'X3', 'X4', 'X5', 'X6', 'X7', 'X8', and 'Y'. The table contains 9 rows of data, each representing a design point. The 'Y' column contains a dot (•) for each row.

	X1	X2	X3	X4	X5	X6	X7	X8	Y
1	-1	1	-1	-1	1	-1	1	1	•
2	1	-1	1	-1	-1	-1	1	-1	•
3	-1	-1	1	-1	-1	1	-1	1	•
4	1	-1	-1	1	-1	-1	-1	1	•
5	1	-1	1	1	1	1	1	1	•
6	1	1	1	-1	-1	1	-1	1	•
7	-1	1	1	1	1	-1	-1	-1	•
8	-1	1	-1	1	-1	1	1	-1	•
9	1	-1	-1	-1	1	1	-1	-1	•

Possible scripts include the following:

Model Runs the **Analyze > Fit Model** platform. The model described by the script is determined by your choices in the Model outline and by the type of design.

Evaluate Design Runs the **DOE > Design Diagnostics > Evaluate Design** platform. The model described by the script is determined by your choices in the Model outline and by the type of design.

Constraint Shows model constraints that you entered in the Define Factor Constraints outline using the **Specify Linear Constraints** option.

Disallowed Combinations Shows model constraints that you entered in the Define Factor Constraints outline using the **Use Disallowed Combinations Filter** or the **Use Disallowed Combinations Script** options.

DOE Dialog Re-creates the Custom Design window that you used to generate the design table. The script also contains the random seed used to generate your design.

Back

The Back button takes you back to where you were before clicking Make Design. You can make changes to the previous outlines and regenerate the design.

Note: If you attempt to remove all factors after clicking the Back button, one continuous factor remains. You can delete the continuous factor after new factors are added.

Custom Design Options

- “Description of Options”
- “Simulate Responses”
- “Save X Matrix”
- “Number of Starts”
- “Design Search Time”
- “Set Delta for Power”

Description of Options

The Custom Design red triangle menu contains the following options:

Save Responses Saves the information in the Responses panel to a new data table. You can then quickly load the responses and their associated information into most DOE windows. This option is helpful if you anticipate reusing the responses.

Load Responses Loads responses that you saved using the Save Responses option.

Save Factors Saves the information in the Factors panel to a new data table. Each factor’s column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Note: It is possible to create a factors table by entering data into an empty table, but remember to assign each column an appropriate Design Role. Do this by right-clicking on the column name in the data grid and selecting **Column Properties > Design Role**. In the Design Role area, select the appropriate role.

Load Factors Loads factors that you saved using the Save Factors option.

Save Constraints (Unavailable for some platforms) Saves factor constraints that you defined in the Define Factor Constraints or Linear Constraints outline into a data table, with a column for each constraint. You can then quickly load the constraints into most DOE windows.

In the constraint table, the first rows contain the coefficients for each factor. The last row contains the inequality bound. Each constraint’s column contains a column property called **ConstraintState** that identifies the constraint as a “less than” or a “greater than” constraint. See “[ConstraintState](#)” on page 766 in the “Column Properties” appendix.

Load Constraints (Unavailable for some platforms) Loads factor constraints that you saved using the Save Constraints option.

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction
- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.
- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click Apply, a formula for the simulated response values is saved in a new column called <Y> Simulated, where Y is the name of the response. Clicking Apply again updates the formula and values in <Y> Simulated.

For additional details, see [“Simulate Responses”](#) on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Save X Matrix Saves the Moments Matrix and Model Matrix to table scripts in the design data table. These scripts contain the moments and design matrices. See [“Save X Matrix”](#) on page 112.

Caution: For a design with nominal factors, the Model Matrix saved by the Save X Matrix option is *not* the coding matrix used in fitting the linear model. You can obtain the coding matrix used for fitting the model by selecting the option Save Columns > Save Coding Table in the Fit Model report that you obtain when you run the Model script.

Optimality Criterion Changes the design optimality criterion. The default criterion, **Recommended**, specifies D-optimality for all design types, unless you added quadratic effects using the RSM button in the Model outline. For more information about the D-, I-, and alias-optimal designs, see [“Optimality Criteria”](#) on page 127.

Note: You can set a preference to always use a given optimality criterion. Select File > Preferences > Platforms > DOE. Check Optimality Criterion and select your preferred criterion.

Number of Starts Enables you to specify the number of random starts used in constructing the design. See [“Number of Starts”](#) on page 115.

Design Search Time Maximum number of seconds spent searching for a design. The default search time is based on the complexity of the design. See [“Design Search Time”](#) on page 115 and [“Number of Starts”](#) on page 115.

If the iterations of the algorithm require more than a few seconds, a Computing Design progress window appears. If you click **Cancel** in the progress window, the calculation stops and gives the best design found at that point. The progress window also displays D-efficiency for D-optimal designs that do not include factors with Changes set to Hard or Very Hard or with Estimability set to If Possible.

Note: You can set a preference for Design Search Time. Select File > Preferences > Platforms > DOE. Check Design Search Time and enter the maximum number of seconds. In certain situations where more time is required, JMP extends the search time.

Sphere Radius Constrains the continuous factors in a design to a hypersphere. Specify the radius and click **OK**. Design points are chosen so that their distance from 0 equals the Sphere Radius. Select this option before you click Make Design.

Note: Sphere Radius constraints cannot be combined with constraints added using the Specify Linear Constraints option. Also, the option is not available when hard-to-change factors are included (split-plot designs).

Advanced Options > Mixture Sum Set the sum of the mixture factors to any positive value. Use this option to keep a component of a mixture constant throughout an experiment.

Advanced Options > Split Plot Variance Ratio Specify the ratio of the variance of the random whole plot and the subplot variance (if present) to the error variance. Before

setting this value, you must define a hard-to-change factor for your split-plot design, or hard and very-hard-to-change factors for your split-split-plot design. Then you can enter one or two positive numbers for the variance ratios, depending on whether you have specified a split-plot or a split-split-plot design.

Advanced Options > Prior Parameter Variance (Available only when the Model outline is available.) Specify the weights that are used for factors whose Estimability is set to If Possible. The option updates to show the default weights when you click Make Design. Enter a positive number for each of the terms for which you want to specify a weight. The value that you enter is the square root of the reciprocal of the prior variance. A larger value represents a smaller variance and therefore more prior information that the effect is not active.

Bayesian D- or I-optimality is used in constructing designs with If Possible factors. The default values used in the algorithm are 0 for Necessary terms, 4 for interactions involving If Possible terms, and 1 for If Possible terms. See [“The Alias Matrix”](#) on page 771 in the “Technical Details” appendix and [“Optimality Criteria”](#) on page 127.

Advanced Options > Parameter Weights (Use for A-Optimal designs.) Specify weights for the model parameters. This allows you to place more weight on the variance of the main effects over say 2nd order effects. For more information on parameter weights see Morgan and Stallings (2017).

Advanced Options > D Efficiency Weight Specify the relative importance of D-efficiency to alias optimality in constructing the design. Select this option to balance reducing the variance of the coefficients with obtaining a desirable alias structure. Values should be between 0 and 1. Larger values give more weight to D-Efficiency. The default value is 0.5. This option has an effect only when you select Make Alias Optimal Design as your Optimality Criterion.

For the definition of D-efficiency, see [“Optimality Criteria”](#) on page 127. For more information about alias optimality, see [“Alias Optimality”](#) on page 131.

Advanced Options > Set Delta for Power Specify the difference in the mean response that you want to detect for model effects. See [“Set Delta for Power”](#) on page 116.

Save Script to Script Window Creates the script for the design that you specified in the Custom Design window and places it in an open script window.

Simulate Responses

When you click Make Table to create your design table, the Simulate Responses option does the following for each response:

- It adds random response values to the response column in your design table.

- It adds a new a column containing a simulation model formula to the design table. The formula and values are based on the model that is specified in the design window.

A Model window opens where you can specify parameter values and select a response distribution for simulation. When you click Apply in the Model window, each column containing a simulation model formula is updated.

Control Window

Figure 4.27 shows the Model window for a design with one continuous factor (X1) and one three-level categorical factor (X2). Notice that X2 is represented by two model terms.

Figure 4.27 Simulate Responses Control Window

Effects	Y
Intercept	1
X1	1
X2 1	1
X2 2	-1

Distribution

☒ Normal Error σ:

☐ Binomial

☐ Poisson

Apply

The initial window shows values for the coefficients of either 1 or -1, and a Normal distribution with error standard deviation equal to 1. If you have set Anticipated Coefficients as part of Power Analysis under Design Evaluation in the DOE window, then the default values in the Simulate Responses outline are the values that you specified as Anticipated Coefficients and Anticipated RMSE (Error Std) in the Power Analysis outline. If it is not possible to fit the model specified in the data table's Model script, the intercept and coefficients have default values of 0.

Simulate Responses

To specify a model for simulated values, do the following:

1. For each term in the list of Effects, enter coefficients for the linear model used to simulate the response values. These define a linear function, $L(\mathbf{x}, \beta) = \mathbf{x}'\beta$. See the Simulate Responses outline in Figure 4.27:
 - The vector $\mathbf{x}$ consists of the terms that define the effects listed under Effects.
 - The vector β is the vector of model coefficients that you specify under Y.
2. Under Distribution, select a response distribution.

3. Click **Apply**. A <Y> Simulated column containing simulated values and their formula is added to the design table, where Y is the name of the response column.

Distribution

Choose from one of the following distributions in the Simulate Responses window:

Normal Simulates values from a normal distribution. Enter a value for Error σ , the standard deviation of the normal error distribution. If you have designated factors to have Changes of Hard in the Factors outline, you can enter a value for Whole Plots σ , the whole plot error. If you have designated factors to have Changes of Hard and Very Hard, you can enter values for both the subplot and whole plot errors. When you click Apply, random values and a formula containing a random response vector based on the model are entered in the column <Y> Simulated.

Binomial Simulates values from a binomial distribution. Enter a value for N, the number of trials. Random integer values are generated according to a binomial distribution based on N trials with probability of success $1/(1 + \exp(-L(\mathbf{x}, \beta)))$. When you click Apply, random values and their formula are entered in the column <Y> Simulated. A column called N Trials that contains the value N is also added to the data table.

Poisson Simulates random integer values according to a Poisson distribution with parameter $\exp(L(\mathbf{x}, \beta))$. When you click Apply, random values and their formula are entered in the column <Y> Simulated.

Note: You can set a preference to simulate responses every time you click Make Table. To do so, select **File > Preferences > Platforms > DOE**. Select **Simulate Responses**.

Save X Matrix

This option saves scripts called Moments Matrix and Model Matrix that contain the moments matrix and the model matrix to the design data table. The moments matrix and the model matrix are used to calculate the Average Variance of Prediction, which appears in the Design Diagnostics section of the Design Evaluation outline. See Goos and Jones (2011). If the design is a split-plot design, a V Inverse script is also saved. The V Inverse script contains the inverse of the covariance matrix of the responses.

Caution: For a design with nominal factors, the Model Matrix saved by the Save X Matrix option is *not* the coding matrix used in fitting the linear model. You can obtain the coding matrix used for fitting the model by selecting the option Save Columns > Save Coding Table in the Fit Model report that you obtain when you run the Model script.

Note: You can set a preference to always save the matrix script. Select **File > Preferences > Platforms > DOE**. Check **Save X Matrix**.

Model Matrix

The *model matrix* describes the design for the experiment. The model matrix has a row for each run and a column for each term of the model specified in the Model outline. For each run, the corresponding row of the model matrix contains the coded values of the model terms:

- Continuous terms are coded to span the range from -1 to 1.
- Nominal terms are coded by applying the Gram-Schmidt orthogonalization procedure to the coding vectors used in fitting linear models.

For more information about the Gram-Schmidt orthogonalization procedure, see Horn and Johnson (2012).

Note: Coding for power analysis matches that of fitting linear models. For more information about the coding used for nominal terms in fitting linear models, see the Standard Least Squares Report and Options chapter in *Fitting Linear Models*.

Moments Matrix

The *moments matrix* is dependent upon the model effects but is independent of the design. It is defined as follows:

$$\mathbf{M} = \int_R f(\mathbf{x})f(\mathbf{x})'d\mathbf{x}$$

where $f(x)$ denotes the model effects corresponding to factor combinations of the vector of factors, x , and R denotes the design space. For more information about moments and design matrices, see Goos and Jones (2011) and Myers et al. (2009). Note that the moments matrix is called a matrix of region moments in Myers et al. (2009).

Scripts

From the Custom Design red triangle menu, select **Save X Matrix**. After the design and the table are created, in the Custom Design table, the Moments Matrix and Model Matrix scripts, and if the design is a split plot, the V Inverse script, are saved as table scripts.

- Right-click and Select **Edit** from either the Moments Matrix, Model Matrix, or V Inverse script. The script shows the corresponding matrix. You can copy this matrix into scripts that you write.
- When you run the Moments Matrix script, the log shows the number of rows in the moments matrix, called Moments.

- When you run the script Model Matrix, the log displays the number of rows in the model matrix, called X.
- When you run the script V Inverse, the log displays the number of rows in the inverse covariance matrix, called V Inverse.

Example

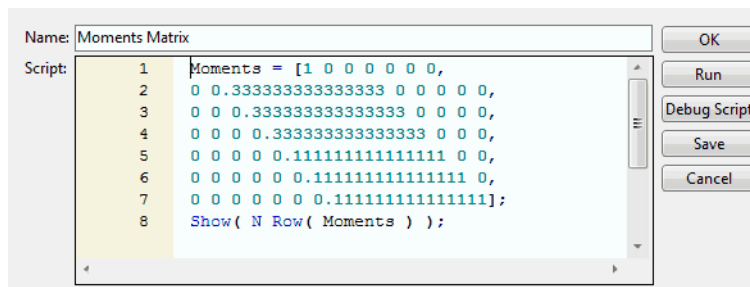
Follow these steps to illustrate these features:

Tip: To see the log, select **View > Log (Window > Log** on macOS).

1. Select **DOE > Custom Design**.
2. Add 3 continuous factors and click **Continue**.
3. Click **Interactions > 2nd**.
4. Click the Custom Design red triangle and select **Save X Matrix**.
5. Using the Default Number of Runs (12), click **Make Design** and then **Make Table**.
6. In the Table panel, right-click the Moments Matrix script and select **Edit**.

The script appears in a script window. The script shows the moments matrix, which is called **Moments**.

Figure 4.28 Moments Matrix Script



7. If it is not already open, select **View > Log (Window > Log** on the macOS).
8. In the Table panel, click the green triangle next to the Moments Matrix script.
The number of rows appear in the log as `N Row(Moments)=7`.
9. In the Table panel, right-click the Model Matrix script and select **Edit**.
The script appears in a script window. The script shows the model matrix, which is called **X**.
10. Click **Run**.
The number of rows appears in the log as `N Row(X)=12`.

11. To view the Model Matrix as a data table, add these lines to the script:

```
dt = New Table( "Model Matrix" );  
dt << Set Matrix( X );
```

12. Click **Run**.

Number of Starts

The number of starts is the number of times that the coordinate-exchange algorithm initiates with a new design. See [“Coordinate-Exchange Algorithm”](#) on page 132. You can specify your own value using the **Number of Starts** option. Increasing the number of random starts tends to improve the optimality of the resulting design.

Unless you specify a value for Number of Starts and click OK, the number of starts is controlled by Design Search Time. To see how many starts were used to construct a design, click Make Design. Then select Number of Starts. The value in the text box is the number of starts used to construct the specific design.

In certain special cases, the globally optimal design is known from theory. If the coordinate-exchange algorithm detects that it has found an optimal design, it stops searching and returns that design.

Tip: To reproduce a specific design, you need to specify the Number of Starts and the Random Seed originally used to produce the design. Obtain these values from the red triangle options after you click Make Design.

In examples of custom designs in the documentation, the random seed and number of starts are often provided so that you can reconstruct the exact design being discussed.

Design Search Time

Design Search Time is the amount of time allocated to finding an optimal design. Custom Design’s coordinate-exchange algorithm consists of finding near-optimal designs based on random starting designs. See [“Coordinate-Exchange Algorithm”](#) on page 132. The Design Search Time determines how many designs are constructed based on random starting designs.

You can specify your own value using the **Design Search Time** option. Increasing the search time tends to improve the optimality of the resulting design.

Keep in mind that designs produced by rerunning the algorithm can differ. Even with the same random seed, the numbers of starting designs used to construct the final design might differ because of variations in computing speed and other factors.

Note: The number of starting designs is given by the value in the Number of Starts text box. However, this value is not updated until after you construct your design by clicking Make Design.

In certain special cases, the globally D-optimal design is known from theory. These cases include:

- Two-level fractional factorial designs or nonregular orthogonal arrays. These are globally D-optimal for all main effect and two-factor interaction models.
- Latin-square designs. These are D-optimal for main effect models assuming the right sample size and numbers of levels of the factors.
- Plackett-Burman designs. These are D-optimal for main effect models.

If the coordinate-exchange algorithm detects that it has found an optimal design, it stops searching and returns that design.

Set Delta for Power

This option specifies the difference in the mean response that you want to be able to detect for model effects. Power calculations appear in the Power Analysis outline within the Design Evaluation outline. Power is calculated for each model parameter based on detecting the specified difference of delta. For categorical effects, the power calculation is based on detecting a maximum change of delta between any two levels.

For example, suppose that you want to detect a change of 3 units in the mean response. All of your factors are continuous. Because your factors are expressed in coded units (the coded levels are -1 and 1), a change of 3 units in the response corresponds to parameter coefficient values of 1.5. When you specify 3 in the Set Delta for Power text box, the Anticipated Coefficients in the Power Analysis outline are set to 1.5. For each parameter, the probability of detecting the change of 3 units in the response appears in the Power column to the right of the parameter.

By default, delta is set to 2. The default coefficient for each continuous effect is set to 1. An n -level categorical factor is represented by $n-1$ indicator variables. The default coefficients for the $n-1$ terms (which represent the categorical factor) are alternating values of 1 and -1. The default coefficients for an interaction effect with more than one degree of freedom are also alternating values of 1 and -1.

Note: The order in which parameters appear in the Power Analysis report might not be identical to their order in the Parameter Estimates report obtained using Standard Least Squares. This difference can occur only when the model contains an interaction with more than one degree of freedom.

Given a specified value of delta, each coefficient in the Anticipated Coefficients list is set at $\text{delta}/2$ multiplied by the default coefficient. For a continuous factor, this assignment ensures that a difference of delta is detected with the calculated power. For a categorical factor, this assignment of coefficients ensures that a maximum difference of delta between any two levels is detected with the calculated power.

Technical Details for Custom Designs

- [“Designs with Randomization Restrictions”](#)
- [“Covariates with Hard-to-Change Levels”](#)
- [“Numbers of Whole Plots and Subplots”](#)
- [“Optimality Criteria”](#)
- [“D-Efficiency”](#)
- [“Coordinate-Exchange Algorithm”](#)

Designs with Randomization Restrictions

This section describes how the Custom Design platform handles various types of designs where random assignment of experimental units to factor level settings is restricted. Random block designs and various types of split-plot designs are included.

Random Block Designs

A random block design groups the runs of an experiment into blocks that are considered to be randomly chosen from a larger population. Runs within a block of runs are usually more homogeneous than runs in different blocks. In these instances, you are often better able to discern other effects if you account for the variation explained by the blocking variables.

Scenario for a Random Block Design

Goos (2002) presents an example involving a pastry dough mixing experiment. The purpose of the experiment is to understand how certain properties of the dough depend on three factors: feed flow rate, initial moisture content, and rotational screw speed. Since it was possible to conduct only four runs a day, the experiment required several days to run. It is likely that random day-to-day differences in environmental variables have some effect on all of the runs that are performed on a given day. To account for the day-to-day variation, the runs were grouped into blocks of size four so that this variation would not compromise the information about the three factors.

The blocking factor, Day, consists of each day's runs. The days on which the trials were conducted are representative of a large population of days with different environmental conditions. It follows that Day is a random blocking factor.

Setup for a Random Block Design

To create a random block design, use the Custom Design platform to enter responses and factors and define your model as usual. In the Design Generation outline, select the **Group runs into random blocks of size** option and enter the number of runs you want in each block. See [“Design Structure Options”](#) on page 101.

Note: To define a fixed blocking factor, enter a blocking factor in the Factors outline. To define a random blocking factor, do not enter a blocking factor in the Factors outline. Instead, select the **Group runs into random blocks of size** option under Design Generation.

Split-Plot Designs

Split-plot designs are used in situations where the settings of certain factors are held constant for groups of runs. In industry, these are usually factors that are difficult or expensive to change from run to run. Factors whose settings need to be held constant for groups of runs are classified as *hard-to-change* in JMP.

Because certain factors are hard-to-change, it is not practical to randomly allocate them to experimental units. Instead, they are allocated to groups of units. This imposes a restriction on randomization that must be considered in generating a design and in analyzing the results.

Scenario for a Split-Plot Design

Box et al. (2005) presents an experiment to study the corrosion resistance of steel bars. The bars are placed in a furnace for curing. Afterward, a coating is applied to increase resistance to corrosion. The two factors of interest are:

- Furnace Temp in degrees centigrade, with levels 360, 370, and 380
- Coating, with levels C1, C2, C3, and C4 depicting four different types of coating

Furnace Temp is a hard-to-change factor, due to the time it takes to reset the temperature in the furnace. For this reason, four bars are processed for each setting of furnace temperature. At a later stage, the four coatings are randomly assigned to the four bars.

The experimental units are the bars. Furnace Temp is a hard-to-change factor whose levels define whole plots. Within each whole plot, the Coating factor is randomly assigned to the experimental units to which the whole plot factor was applied.

Figure 4.29 Factors and Design Outlines for Split-Plot Design

Factors

Add Factor Remove Add N Factors 1

Name	Role	Changes	Values
Furnace Temp	Categorical	Hard	360 370 380
Coating	Categorical	Easy	C1 C2 C3 C4

Define Factor Constraints

Model

Alias Terms

Design

Run	Whole Plots	Furnace Temp	Coating
1	1	370	C1
2	1	370	C3
3	1	370	C2
4	2	370	C4
5	2	370	C2
6	2	370	C1
7	3	380	C3
8	3	380	C1
9	3	380	C4
10	4	360	C4
11	4	360	C2
12	4	360	C3
13	5	380	C1
14	5	380	C4
15	5	380	C2

The Factors outline for the corrosion experiment has Changes set to Hard for Furnace Temp and Easy for Coating. The 15-run design consists of five whole plots, within which the settings of Temperature are held constant.

Setup for a Split-Plot Design

In general, several factors can be applied to a processing step where settings are hard-to-change. In the furnace example, you might consider a furnace location factor, as well as temperature. In the Factors outline, under the Changes column, you would specify a Changes value of Hard for such factors.

When a custom design involves only easy-to-change and hard-to-change factors, the runs of the hard-to-change factors are grouped using a new factor called Whole Plots. The values of Whole Plots designate blocks of runs with identical settings for the hard-to-change factors. The Model script that is saved to the design table treats Whole Plots as a random effect. See [“Changes”](#) on page 89 and [“Design Structure Options”](#) on page 101.

For an example of creating a split-plot design and analyzing the experimental data, see [“Split-Plot Experiment”](#) on page 195 in the [“Examples of Custom Designs”](#) chapter.

Split-Split-Plot Designs

A split-split-plot design is used when there are two levels of factors that are hard-to-change. In industry, such designs often occur when batches of material or experimental units from one processing stage pass to a second processing stage. Factors are applied to batches of material at the first stage. Then those batches are divided for second-stage processing, where additional factors are studied. The first stage factors are considered *very-hard-to-change*, and the second-stage factors are considered *hard-to-change*. Additional factors can be applied to experimental units after the second processing stage. These factors are considered *easy-to-change*.

In a split-split-plot design, the batches are considered to be random blocks. Since the batches are divided for second-stage processing, the second-stage factors are nested within the first-stage factors.

Scenario for a Split-Split-Plot Design

Schoen (1999) presents an example of a split-split-plot design that relates to cheese quality. The factors are given in the Cheese Factors.jmp data table found in the Design Experiment folder. The experiment consists of three stages of processing:

- Milk is received from farmers and stored in a large tank.
- Milk from this tank is distributed to smaller tanks used for curd processing.
- The curds from each tank are transported to presses for processing individual cheeses.

The experiment consists of testing:

- Two factors that are applied when the milk is in the large storage tank.
- Five factors that are applied to the smaller curd processing tanks.
- Three factors that are applied to the individual cheeses from a curds processing tank.

Notice that the levels of factors applied to the curd processing tanks (subplots) are nested within the levels of factors applied to the milk storage tank (whole plots).

The Factors outline for the cheese experiment have Changes set as follows:

- Very Hard for the two storage tank factors.
- Hard for the five curd processing tank factors.
- Easy for the three factors that can be randomly assigned to cheeses.

Figure 4.30 Factors and Design Generation Outline for Split-Split-Plot Design

Custom Design

Responses

Factors

Add Factor
Remove
Add N Factors
1

Name	Role	Changes	Values
storage 1	Continuous	Very Hard	-1 1
storage 2	Categorical	Very Hard	L1 L2
curds 1	Continuous	Hard	-1 1
curds 2	Continuous	Hard	-1 1
curds 3	Categorical	Hard	L1 L2
curds 4	Continuous	Hard	-1 1
curds 5	Categorical	Hard	L1 L2 L3 L4
cheese 1	Continuous	Easy	-1 1
cheese 2	Continuous	Easy	-1 1

Define Factor Constraints

Model

Alias Terms

Design Generation

☐ Hard to change factors can vary independently of Very Hard to change factors.

Number of Whole Plots
5

Number of Subplots
10

Number of Runs:

☐ Minimum 14

☐ Default 30

☒ User Specified 22

Make Design

The default number of whole plots is 5 and the default number of subplots is 10. The number of runs is set to 22. If only 10 subplots are used the design does not have enough whole plots to estimate the subplot variance. Change the number of subplots to 11 and click **Make Design** to see a 22-run design.

Figure 4.31 Split-Split-Plot Design for Cheese Scenario

Design													
Run	Whole Plots	Subplots	storage 1	storage 2	curds 1	curds 2	curds 3	curds 4	curds 5	cheese 1	cheese 2	cheese 3	
1	1	1	-1	L2	1	1	L1	-1	L2	1	-1	L3	
2	1	1	-1	L2	1	1	L1	-1	L2	-1	1	L1	
3	1	2	-1	L2	-1	-1	L2	-1	L3	1	1	L1	
4	1	2	-1	L2	-1	-1	L2	-1	L3	-1	-1	L3	
5	1	3	-1	L2	1	-1	L2	1	L1	1	1	L3	
6	2	4	1	L1	-1	1	L1	-1	L1	1	-1	L2	
7	2	4	1	L1	-1	1	L1	-1	L1	1	1	L3	
8	2	4	1	L1	-1	1	L1	-1	L1	-1	1	L1	
9	2	5	1	L1	1	1	L1	1	L3	-1	-1	L1	
10	2	5	1	L1	1	1	L1	1	L3	1	1	L2	
11	3	6	1	L2	1	-1	L2	-1	L2	1	-1	L1	
12	3	6	1	L2	1	-1	L2	-1	L2	-1	1	L2	
13	3	7	1	L2	-1	1	L1	1	L4	1	1	L2	
14	3	7	1	L2	-1	1	L1	1	L4	-1	-1	L3	
15	4	8	1	L1	-1	1	L2	1	L2	1	1	L3	
16	4	8	1	L1	-1	1	L2	1	L2	-1	-1	L1	
17	4	9	1	L1	1	-1	L1	-1	L4	1	-1	L1	
18	4	9	1	L1	1	-1	L1	-1	L4	-1	1	L2	
19	5	10	-1	L1	1	1	L2	-1	L4	-1	-1	L2	
20	5	10	-1	L1	1	1	L2	-1	L4	1	1	L1	
21	5	11	-1	L1	-1	-1	L1	1	L2	-1	1	L3	
22	5	11	-1	L1	-1	-1	L1	1	L2	1	-1	L2	

The five whole plots correspond to the storage factors; **storage 1** and **storage 2**. The settings of the storage factors are constant within a whole plot. If consecutive whole plots have the same setting for a whole plot factor, the factor should be reset between the plots. For example, you should reset the level for **storage 1** between runs 10 and 11 and between runs 14 and 15, and you should reset the level for **storage 2** between runs 18 and 19. Resetting the factor between whole plots, even when the specified settings are the same, is required in order to capture whole plot variation.

The 11 subplots correspond to the curds factors. Within a subplot, the settings of the curds factors are constant. Each level of Subplots appears only within one level of Whole Plots, indicating that the levels of Subplots are nested within the levels of Whole Plots.

Levels of the cheese factors vary randomly from run to run.

Setup for a Split-Split-Plot Design

In a split-split-plot design, the Factors outline contains factors with Changes set to Very Hard and Hard. The design can also contain factors with Changes set to Easy. Two factors are created:

- A factor called **Whole Plots** represents the blocks of constant levels of the factors with Changes set to Very Hard.
- A factor called **Subplots** represents the blocks of constant levels of the factors with Changes set to Hard.
- The factor **Subplots** reflects the nesting of the levels of the factors with Changes set to Hard within the levels of the factors with Changes set to Very Hard.

- The levels of factors with Changes set to Easy are randomly assigned to units within subplots.
- The factors Whole Plots and Subplots are treated as random effects in the Model script that is saved to the design table.

See the Changes description under “[Factors Outline](#)” on page 89 and “[Design Structure Options](#)” on page 101.

Two-Way Split-Plot Designs

A two-way split-plot (also known as strip plot or split block) design consists of two split-plot components. In industry, these designs arise when batches of material or experimental units from one processing stage pass to a second processing stage. But, after the first processing stage, it is possible to divide the batches into sub-batches. The second-stage processing factors are applied randomly to these sub-batches. For a specific second-stage experimental setting, all of the sub-batches assigned to that setting can be processed simultaneously. Additional factors can be applied to experimental units after the second processing stage.

In contrast to a split-split-plot design, the second-stage factors are *not nested* within the first-stage factors. After the first stage, the batches are subdivided and formed into new batches. Therefore, both the first- and second-stage factors are applied to whole batches.

Although factors at both stages might be equally hard-to-change, to distinguish these factors, JMP denotes the first stage factors as *very-hard-to-change*, and the second-stage factors as *hard-to-change*. Additional factors applied to experimental units after the second processing stage are considered *easy-to-change*.

Scenario for a Two-Way Split-Plot Design

Vivacqua and Bisgaard (2004) describe an experiment to improve the open circuit voltage in battery cells. Two stages of processing are of interest:

- First stage: A continuous assembly process
- Second stage: A curing process with a 5-day cycle time

The engineers want to study six two-level factors:

- Four factors, X1, X2, X3, and X4, that are applied to the assembly process
- Two factors, X5 and X6, that are applied to the curing process

A full factorial design with all factors at two levels would require $2^6 = 64$ runs, and would require a prohibitive $64 \times 5 = 320$ days. Also, it is not practical to vary assembly conditions for individual batteries. However, assembly conditions can be changed for large batches, such as batches of 2000 batteries.

Both the first- and second-stage factors are hard-to-change. In a sense, there are two split-plot designs. However, the batches of 2,000 batteries from the first-stage experiment can be divided into four sub-batches of 500 batteries each. These sub-batches can be randomly assigned to the four settings of the two second-stage factors. All of the batches assigned to a given set of curing conditions can be processed simultaneously. In other words, the first- and second-stage factors are *crossed*.

To distinguish between the first- and second-stage factors, you designate the Changes for the first-stage factors as Very Hard, and the Changes for the second-stage factors as Hard (Figure 4.32). Also, under Design Generation, note the following option: **Hard to change factors can vary independently of Very Hard to change factors**. If this is not checked, the design is treated as a split-split-plot design, with nesting of factors at the two levels. Check this option to create a two-way split-plot design.

Figure 4.32 Factors and Design Generation Outline for Two-Way Split Plot Design

The screenshot shows the Minitab Custom Design dialog box. The 'Factors' section lists six factors (X1-X6) with their roles, change levels, and values. The 'Design Generation' section has the checkbox 'Hard to change factors can vary independently of Very Hard to change factors' checked. The 'Number of Whole Plots' is 7 and 'Number of Subplots' is 14. Under 'Number of Runs', 'Default' (28) is selected. A 'Make Design' button is at the bottom.

Name	Role	Changes	Values
X1	Continuous	Very Hard	-1 1
X2	Continuous	Very Hard	-1 1
X3	Continuous	Very Hard	-1 1
X4	Continuous	Very Hard	-1 1
X5	Continuous	Hard	-1 1
X6	Continuous	Hard	-1 1

Design Generation

☒ Hard to change factors can vary independently of Very Hard to change factors.

Number of Whole Plots: 7

Number of Subplots: 14

Number of Runs:

☐ Minimum 15

☒ Default 28

☐ User Specified 28

Make Design

The default number of whole plots is 7; the default number of subplots is 14. Click **Make Design** to see the 28-run design.

Figure 4.33 Two-Way Split-Plot Design for Battery Cells

Design									
Run	Whole Plots	Subplots	X1	X2	X3	X4	X5	X6	
1	1	1	1	1	1	-1	1	-1	
2	1	2	1	1	1	-1	-1	1	
3	1	3	1	1	1	-1	-1	-1	
4	1	4	1	1	1	-1	1	1	
5	2	5	-1	-1	-1	-1	-1	1	
6	2	6	-1	-1	-1	-1	1	-1	
7	2	7	-1	-1	-1	-1	-1	-1	
8	2	8	-1	-1	-1	-1	1	1	
9	3	9	1	1	-1	-1	-1	1	
10	3	10	1	1	-1	-1	1	-1	
11	3	11	1	1	-1	-1	-1	-1	
12	3	12	1	1	-1	-1	1	1	
13	4	13	1	-1	-1	1	-1	-1	
14	4	14	1	-1	-1	1	1	1	
15	4	1	1	-1	-1	1	1	-1	
16	4	2	1	-1	-1	1	-1	1	
17	5	3	-1	1	-1	1	-1	-1	
18	5	4	-1	1	-1	1	1	1	
19	5	5	-1	1	-1	1	-1	1	
20	5	6	-1	1	-1	1	1	-1	
21	6	7	1	-1	1	1	-1	-1	
22	6	8	1	-1	1	1	1	1	
23	6	9	1	-1	1	1	-1	1	
24	6	10	1	-1	1	1	1	-1	
25	7	11	-1	-1	1	-1	-1	-1	
26	7	12	-1	-1	1	-1	1	1	
27	7	13	-1	-1	1	-1	-1	-1	
28	7	14	-1	-1	1	-1	1	1	

The seven whole plots correspond to the first-stage factors, X1, X2, X3, and X4. The settings of these factors are constant within a whole plot. The 14 subplots correspond to the second-stage factors, X5 and X6. For example, the sub-batches for runs 1 and 15 (from different whole plots) are subject to the same subplot treatment, where X5 is set at 1 and X6 at -1.

Setup for a Two-Way Split-Plot Design

A two-way split-plot design requires factors with Changes set to Very Hard and to Hard. As described in [“Setup for a Split-Split-Plot Design”](#) on page 122, factors called Whole Plots and Subplots are created. However, in a two-way split-plot design, Subplots does not nest the levels of factors with Changes set to Hard within the levels of factors with Changes set to Very Hard. Both Whole Plots and Subplots are treated as random effects in the Model script that is saved to the design table.

You need to ensure that the factor Subplots is *not nested* within the factor Whole Plots. Select the option **Hard to change factors can vary independently of Very Hard to change factor** in the Design Generation outline (Figure 4.32). See [“Changes”](#) on page 89 and [“Design Structure Options”](#) on page 101.

For an example of creating a split-plot design and analyzing the experimental data, see [“Two-Way Split-Plot Experiment”](#) on page 200 in the “Examples of Custom Designs” chapter.

Covariates with Hard-to-Change Levels

Suppose that you have measurements on batches of material that are available for use in testing experimental factors. Or suppose that you have measurements on individuals who might be selected to participate in testing experimental factors. The measurements on batches or individuals are known in advance of the experiment and are considered to be covariates.

The batches or individuals correspond to whole plots. You might want to use only some of these whole plots in your experiment. Because information about the whole plots in the form of covariates is available, the design should choose the whole plots in an optimal fashion.

In the Factors outline, the Custom Design platform enables you to designate covariates as hard-to-change. The model, as given by the terms that you include in the Model outline, can include interactions and powers constructed using covariates and experimental factors.

Note: When you set Changes for a Covariate factor to Hard, all other covariates are also set to Hard. The remaining factors must be set to Easy. Because the algorithm requires a combination of row exchange and coordinate exchange, even moderately sized designs might take some time to generate.

Scenario for an Experiment with a Hard-to-Change Covariate

An experiment involving batches of polypropylene plates is discussed in Goos and Jones (2011, Chapter 9) and Jones and Goos (2015). Large batches of polypropylene plates are produced according to various formulations determined by several variables. Some plates are used immediately, and the remainder are stored for future experimental purposes. The compositions of these stored batches are known.

A customer has certain requirements regarding the plate formulation. Future experiments involve customizing the gas plasma treatment to the types of formulations required by the customer. The composition variables are considered hard-to-change covariates. Gas plasma treatment factors can be applied to sub-batches of plates with a given formulation.

The optimal design identifies the batches (defined by the covariates) to use, determines the number of plates from each batch to use, and provides settings for the gas plasma levels. Note that the optimal number of batches and plates from a given batch depend on the covariates.

An example is provided in the [“Examples of Custom Designs”](#) chapter on page 133.

Numbers of Whole Plots and Subplots

JMP suggests default values for the Number of Whole Plots and Number of Subplots. These values are based on heuristics for creating a balanced design that allows for estimation of the effects specified in the Model outline.

If you enter missing values for Number of Whole Plots or Number of Subplots, JMP chooses values that maximize the D-efficiency of the design. The algorithm uses the values specified in the Split Plot Variance Ratio option. See [“Advanced Options > Split Plot Variance Ratio”](#) on page 109. The D-efficiency is given by the determinant of $\mathbf{X}'\mathbf{V}^{-1}\mathbf{X}$, where $\mathbf{V}^{-1}$ is the inverse of the variance matrix of the responses. See Goos (2002).

If you enter values for the Number of Whole Plots and Number of Subplots, Custom Design attempts to maximize the optimality of the resulting design. For more information about split-plot designs, see Jones and Goos (2007). For more information about designs with hard-to-change covariates, see Jones and Goos (2015).

Optimality Criteria

This section provides information about the following designs:

- [“D-Optimality”](#) on page 127
- [“Bayesian D-Optimality”](#) on page 128
- [“I-Optimality”](#) on page 129
- [“Bayesian I-Optimality”](#) on page 130
- [“A-Optimality”](#) on page 130
- [“Alias Optimality”](#) on page 131

D-Optimality

By default, the Custom Design platform optimizes the D-optimality criterion except when a full quadratic model is created using the RSM button. In that case, an I-optimal design is constructed.

The D-optimality criterion minimizes the determinant of the covariance matrix of the model coefficient estimates. It follows that D-optimality focuses on precise estimates of the effects. This criterion is desirable in the following cases:

- screening designs
- experiments that focus on estimating effects or testing for significance
- designs where identifying the active factors is the experimental goal

The D-optimality criterion is dependent on the assumed model. This is a limitation because often the form of the true model is not known in advance. The runs of a D-optimal design optimize the precision of the coefficients of the assumed model. In the extreme, a D-optimal design might be saturated, with the same number of runs as parameters and no degrees of freedom for lack of fit.

Specifically, a D-optimal design maximizes D , where D is defined as follows:

$$D = \det[\mathbf{X}'\mathbf{X}]$$

and where $\mathbf{X}$ is the model matrix as defined in “[Simulate Responses](#)” on page 110.

D-optimal split-plot designs maximize D , where D is defined as follows:

$$D = \det[\mathbf{X}'\mathbf{V}^{-1}\mathbf{X}]$$

and $\mathbf{V}^{-1}$ is the block diagonal covariance matrix of the responses (Goos [2002](#)).

Since a D-optimal design focuses on minimizing the standard errors of coefficients, it might not allow for checking that the model is correct. For example, a D-optimal design does not include center points for a first-order model. When there are potentially active terms that are not included in the assumed model, a better approach is to specify If Possible terms and to use a Bayesian D-optimal design.

Bayesian D-Optimality

Bayesian D-optimality is a modification of the D-optimality criterion. The Bayesian D-optimality criterion is useful when there are potentially active interactions or non-linear effects. See DuMouchel and Jones ([1994](#)) and Jones et al ([2008](#)).

Bayesian D-optimality estimates a specified set of model parameters precisely. These are the effects whose Estimability you designate as Necessary in the Model outline. But at the same time, Bayesian D-optimality has the ability to estimate other, typically higher-order effects, as allowed by the run size. These are the effects whose Estimability you designate as If Possible in the Model outline. To the extent possible given the run size restriction, a Bayesian D-optimal design allows for detecting inadequacy in a model that contains only the Necessary effects.

The Bayesian D-optimality criterion is most effective when the number of runs is larger than the number of Necessary terms, but smaller than the sum of the Necessary and If Possible terms. When this is the case, the number of runs is smaller than the number of parameters that you would like to estimate. Using prior information in the Bayesian setting allows for precise estimation of all of the Necessary terms while providing the ability to detect and estimate some If Possible terms.

To allow for a meaningful prior distribution to apply to the parameters of the model, responses and factors are scaled to have certain properties (DuMouchel and Jones, [1994](#), Section 2.2).

Consider the following notation:

- $\mathbf{X}$ is the model matrix as defined in “[Simulate Responses](#)” on page 110
- $\mathbf{K}$ is a diagonal matrix with values as follows:

- $k = 0$ for Necessary terms
- $k = 1$ for If Possible main effects, powers, and interactions involving a categorical factor with more than two levels
- $k = 4$ for all other If Possible terms

The prior distribution imposed on the vector of If Possible parameters is multivariate normal, with mean vector $\mathbf{0}$ and diagonal covariance matrix with diagonal entries $1/k^2$. Therefore, a value k^2 is the reciprocal of the prior variance of the corresponding parameter.

The values for k are empirically determined. If Possible main effects, powers, and interactions with more than one degree of freedom have a prior variance of 1. Other If Possible terms have a prior variance of $1/16$. In the notation of DuMouchel and Jones (1994) $k = 1/\tau$.

To control the weights for If Possible terms, select **Advanced Options > Prior Parameter Variance** from the red triangle menu. See “[Advanced Options > Prior Parameter Variance](#)” on page 110.

The posterior distribution for the parameters has the covariance matrix $(\mathbf{X}'\mathbf{X} + \mathbf{K}^2)^{-1}$. The Bayesian D-optimal design is obtained by maximizing the determinant of the inverse of the posterior covariance matrix:

$$\det(\mathbf{X}'\mathbf{X} + \mathbf{K}^2)$$

I-Optimality

I-optimal designs minimize the average variance of prediction over the design space. The I-optimality criterion is more appropriate than D-optimality if your primary experimental goal is not to estimate coefficients, but rather to do the following:

- predict a response
- determine optimum operating conditions
- determine regions in the design space where the response falls within an acceptable range

In these cases, precise prediction of the response takes precedence over precise estimation of the parameters.

The prediction variance relative to the unknown error variance at a point x_0 in the design space can be calculated as follows:

$$\text{var}(\hat{\mathbf{Y}}|\mathbf{x}_0) = f(\mathbf{x}_0)'(\mathbf{X}'\mathbf{X})^{-1}f(\mathbf{x}_0)$$

where $\mathbf{X}$ is the model matrix as defined in “[Simulate Responses](#)” on page 110.

I-optimal designs minimize the integral I of the prediction variance over the entire design space, where I is given as follows:

$$I = \int_R f(\mathbf{x})' (\mathbf{X}'\mathbf{X})^{-1} f(\mathbf{x}) d\mathbf{x} = \text{trace}[(\mathbf{X}'\mathbf{X})^{-1} \mathbf{M}]$$

Here $\mathbf{M}$ is the moments matrix:

$$\mathbf{M} = \int_R f(\mathbf{x}) f(\mathbf{x})' d\mathbf{x}$$

See “[Simulate Responses](#)” on page 110. For further information about simulating responses, see Goos and Jones (2011b).

The moments matrix does not depend on the design and can be computed in advance. The row vector $f(\mathbf{x})'$ consists of a 1 followed by the effects corresponding to the assumed model. For example, for a full quadratic model in two continuous factors, $f(\mathbf{x})'$ is defined as follows:

$$f(\mathbf{x})' = (1, x_1, x_2, x_1x_2, x_1^2, x_2^2)$$

A-Optimality

The A-optimality criterion minimizes the trace of the of the covariance matrix of the model coefficient estimates. The trace is the sum of the main diagonal elements of a matrix. An A-optimal design minimizes the sum of the variances of the regression coefficients.

Bayesian I-Optimality

The Bayesian I-optimal design minimizes the average prediction variance over the design region for Necessary and If Possible terms.

Consider the following notation:

- $\mathbf{X}$ is the model matrix, defined in “[Simulate Responses](#)” on page 110
- $\mathbf{K}$ is a diagonal matrix with values as follows:
 - $k = 0$ for Necessary terms
 - $k = 1$ for If Possible main effects, powers, and interactions involving a categorical factor with more than two levels
 - $k = 4$ for all other If Possible terms

The prior distribution imposed on the vector of If Possible parameters is multivariate normal, with mean vector $\mathbf{0}$ and diagonal covariance matrix with diagonal entries $1/k^2$. (See “[Bayesian D-Optimality](#)” on page 128 for more information about the values k .)

The posterior variance of the predicted value at a point $\mathbf{x}_0$ is as follows:

$$\text{var}(\hat{\mathbf{Y}}|\mathbf{x}_0) = f(\mathbf{x}_0)'(\mathbf{X}'\mathbf{X} + \mathbf{K}^2)^{-1}f(\mathbf{x}_0)$$

The Bayesian I-optimal design minimizes the average prediction variance over the design region, as follows:

$$I_B = \text{trace}[(\mathbf{X}'\mathbf{X} + \mathbf{K}^2)^{-1}\mathbf{M}]$$

where $\mathbf{M}$ is the moments matrix. See [“Simulate Responses”](#) on page 110.

Alias Optimality

Alias optimality seeks to minimize the aliasing between effects that are in the assumed model and effects that are not in the model but are potentially active. Effects that are not in the model but that are of potential interest are called *alias effects*. For more information about alias-optimal designs, see Jones and Nachtsheim (2011b).

Specifically, let $\mathbf{X}_1$ be the model matrix corresponding to the terms in the assumed model, as defined in [“Simulate Responses”](#) on page 110. The design defines the model that corresponds to the alias effects. Denote the matrix of model terms for the alias effects by $\mathbf{X}_2$.

The *alias matrix* is the matrix $\mathbf{A}$, defined as follows:

$$\mathbf{A} = (\mathbf{X}_1'\mathbf{X}_1)^{-1}\mathbf{X}_1'\mathbf{X}_2$$

The entries in the alias matrix represent the degree of bias associated with the estimates of model terms. See [“The Alias Matrix”](#) on page 771 in the “Technical Details” appendix for the derivation of the alias matrix.

The sum of squares of the entries in $\mathbf{A}$ provides a summary measure of bias. This sum of squares can be represented in terms of a trace as follows:

$$\text{trace}(\mathbf{A}'\mathbf{A})$$

Designs that reduce the trace criterion generally have lower D-efficiency than the D-optimal design. Consequently, alias optimality seeks to minimize the trace of $\mathbf{A}'\mathbf{A}$ subject to a lower bound on D-efficiency. For the definition of D-efficiency, see [“Optimality Criteria”](#) on page 127. The lower bound on D-efficiency is given by the D-efficiency weight, which you can specify under Advanced Options. See [“Advanced Options > D Efficiency Weight”](#) on page 110.

D-Efficiency

Let $\mathbf{X}$ denote the design, or model, matrix for a given assumed model with p parameters. For the definition of the model matrix, see “[Simulate Responses](#)” on page 110. Let $\mathbf{X}_D$ denote the model matrix for a D-optimal design for the assumed model. Then the D-efficiency of the design given by $\mathbf{X}$ is as follows:

$$\text{D-Efficiency} = \left[\frac{\det(\mathbf{X}'\mathbf{X})}{\det(\mathbf{X}_D'\mathbf{X}_D)} \right]^{1/p}$$

Coordinate-Exchange Algorithm

Custom Design constructs a design that seeks to optimize one of several optimality criteria. (See “[Optimality Criteria](#)” on page 127.) To optimize the criterion, Custom Design uses the *coordinate-exchange algorithm* (Meyer and Nachtsheim, 1995). The algorithm begins by randomly selecting values within the specified design region for each factor and each run to construct a starting design.

Suppose your study requires continuous factors, no factor constraints, and a main-effects model. An iteration consists of testing each value of the model matrix, as follows:

- The current value of each factor is replaced by its two most extreme values.
- The optimality criterion is computed for both of these replacements.
- If one of the values increases the optimality criterion, this value replaces the old value.

The process continues until no replacement occurs for an entire iteration.

Appropriate adjustments are made to the algorithm to account for polynomial terms, nominal factors, and factor constraints.

The design obtained using this process is optimal in a large class of neighboring designs. But it is only *locally* optimal. To improve the likelihood of finding a globally optimal design, the coordinate-exchange algorithm is repeated a large number of times. Goos and Jones (2011, p. 36) recommend using at least 1,000 random starts for all but the most trivial design situations. The number of starting designs is controlled by the Number of Starts option. See “[Number of Starts](#)” on page 115. Custom Design provides the design that maximizes the optimality criterion among all the constructed designs.

Examples of Custom Designs

Perform Experiments That Meet Your Needs

Use the Custom Design platform as your primary tool for constructing a wide range of experimental designs. You can construct a variety of design types and fine tune them to your specific experimental needs and resource budget.

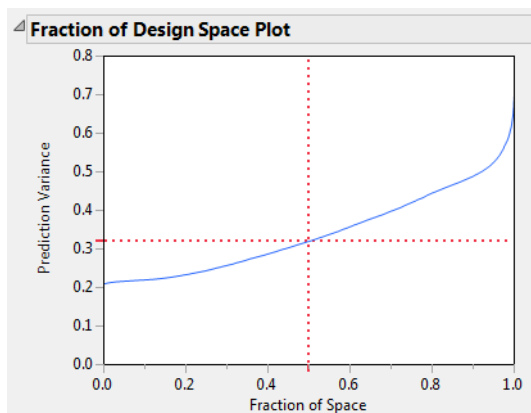
Custom Design provides more options and control than the Screening, Response Surface, Full Factorial, and Mixture Design platforms. The designs that you construct are created specifically to meet your goals. This eliminates the struggle to find a classical design that only comes close to meeting your goals.

The flexible special-purpose designs that you can construct using Custom Design include:

- Screening designs, including supersaturated screening designs
- Response surface designs, including those with categorical factors
- Mixture designs, including those with process factors, and mixture of mixture designs
- Designs that include covariates or that are robust to linear time trends
- Fixed and random block designs
- Split-plot, split-split-plot, and two-way split-plot (strip-plot) designs

In this chapter you construct most of these design types within the Custom Design platform. In many cases, you also analyze the experimental results. For help with using the Custom Design platform, see the “[Custom Designs](#)” chapter on page 65.

Figure 5.1 Fraction of Design Space Plot



Contents

Screening Experiments	135
Design That Estimates Main Effects Only	135
Design That Estimates All Two-Factor Interactions	138
Design That Avoids Aliasing of Main Effects and Two-Factor Interactions	141
Supersaturated Screening Designs	144
Design for Fixed Blocks	152
Response Surface Experiments	156
Response Surface Design	156
Response Surface Design with Flexible Blocking	163
Comparison of a D-Optimal and an I-Optimal Response Surface Design	166
Response Surface Design With Constraints and Categorical Factor	170
Mixture Experiments	176
Mixture Design with Nonmixture Factors	176
Mixture of Mixtures Design	181
Experiments with Covariates	185
Design with Fixed Covariates	185
Design with Hard-to-Change Covariates	188
Design with a Linear Time Trend	192
Experiments with Randomization Restrictions	195
Split-Plot Experiment	195
Two-Way Split-Plot Experiment	200
Experiments for Robust Process and Product Design	207

Screening Experiments

In the early stages of studying a process, you identify a list of factors that potentially affect your response or responses. You are interested in identifying the *active* factors, that is, the factors that actually do affect your response or responses. A *screening design* helps you determine which factors are likely to be active. Once the active factors are identified, you can construct more sophisticated designs, such as response surface designs, to model interactions and curvature.

Screening designs constructed using the Custom Design platform are often equivalent to the classical designs provided in the Screening Design platform. However, Custom Design can construct designs for cases where classical screening designs are not available.

The Custom Design platform constructs screening designs using either the D-optimality or Bayesian D-optimality criterion. The D-optimality criterion minimizes the determinant of the covariance matrix of the model coefficient estimates. It follows that D-optimality focuses on precise estimates of the effects. See [“Optimality Criteria”](#) on page 127 in the “Custom Designs” chapter.

- [“Design That Estimates Main Effects Only”](#)
- [“Design That Estimates All Two-Factor Interactions”](#)
- [“Design That Avoids Aliasing of Main Effects and Two-Factor Interactions”](#)
- [“Supersaturated Screening Designs”](#)
- [“Design for Fixed Blocks”](#)

Design That Estimates Main Effects Only

Note: For more information about main effects only screening designs, see the [“Screening Designs”](#) chapter on page 289.

In this example, you are interested in studying the main effects of six factors. You construct a screening design where all of the main effects are orthogonal. However, the main effects are aliased with two-factor interactions.

1. Select **DOE > Custom Design**.
2. In the Factors outline, type 6 next to **Add N Factors**.
3. Click **Add Factor > Continuous**.
4. Click **Continue**.

The Model outline appears. It includes only main effects with Estimability designated as Necessary. This means that all main effects are estimable in the design that is generated.

Figure 5.2 Custom Design Window Showing Model Outline

Custom Design

Responses

Factors

Add Factor

Remove

Add N Factors

1

Name	Role	Changes	Values
X1	Continuous	Easy	-1
X2	Continuous	Easy	-1
X3	Continuous	Easy	-1
X4	Continuous	Easy	-1
X5	Continuous	Easy	-1

Define Factor Constraints

Model

Main Effects

Interactions

RSM

Cross

Powers

Remove Term

Name	Estimability
Intercept	Necessary
X1	Necessary
X2	Necessary
X3	Necessary
X4	Necessary
X5	Necessary
X6	Necessary

Alias Terms

Design Generation

☐ Group runs into random blocks of size: 2

Number of Center Points: 0

Number of Replicate Runs: 0

Number of Runs:

☐ Minimum 7
 ☐ Default 12
 ☒ User Specified 12

Make Design

Keep the default of 12 runs.

Note: Setting the Random Seed in step 5 and Number of Starts in step 6 reproduces the exact results shown in this example. When constructing a design on your own, these steps are not necessary.

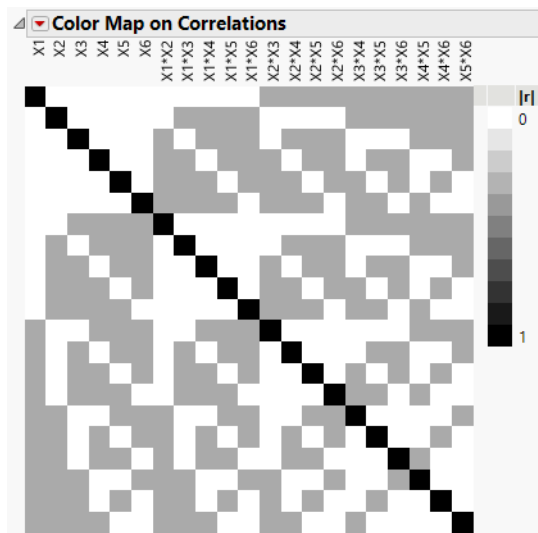
- (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
- (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 1, and click **OK**.
- Click **Make Design**.

Figure 5.3 Design for Main Effects Only

Design						
Run	X1	X2	X3	X4	X5	X6
1	-1	-1	-1	-1	-1	-1
2	1	-1	1	-1	-1	-1
3	1	1	-1	1	-1	-1
4	-1	1	1	-1	1	-1
5	1	-1	1	1	-1	1
6	1	1	-1	1	1	-1
7	1	1	1	-1	1	1
8	-1	1	1	1	-1	1
9	-1	-1	1	1	1	-1
10	-1	-1	-1	1	1	1
11	1	-1	-1	-1	1	1
12	-1	1	-1	-1	-1	1

8. Open the **Design Evaluation > Color Map on Correlations** outline.

Figure 5.4 Color Map on Correlations



Notice the following:

- The main effects are represented by the six terms in the upper left corner of the map.
- The white coloring corresponding to the correlations of the six main effects with other main effects indicates correlations of 0. This means that all main effects are orthogonal and can be estimated independently of each other.
- The light gray color in the squares corresponding to some two-way interactions indicates that the corresponding effects are correlated. This means that these effects cannot be estimated independently of other effects. Place your cursor over these squares to see the exact correlation.

- Notice that no effects are completely confounded with each other. The black squares, indicating absolute correlations of 1, are only on the main diagonal.

9. Open the **Design Evaluation > Alias Matrix** outline.

Figure 5.5 Alias Matrix

Alias Matrix																
Effect	X1*X2	X1*X3	X1*X4	X1*X5	X1*X6	X2*X3	X2*X4	X2*X5	X2*X6	X3*X4	X3*X5	X3*X6	X4*X5	X4*X6	X5*X6	
Intercept	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
X1	0	0	0	0	0	-0.33	0.333	0.333	-0.33	-0.33	-0.33	0.333	-0.33	-0.33	0.333	0
X2	0	-0.33	0.333	0.333	-0.33	0	0	0	0	-0.33	0.333	0.333	-0.33	-0.33	-0.33	0
X3	-0.33	0	-0.33	-0.33	0.333	0	-0.33	0.333	0.333	0	0	0	-0.33	0.333	-0.33	0
X4	0.333	-0.33	0	-0.33	-0.33	-0.33	0	-0.33	-0.33	0	-0.33	0.333	0	0	-0.33	0
X5	0.333	-0.33	-0.33	0	0.333	0.333	-0.33	0	-0.33	-0.33	0	-0.33	0	-0.33	0	0
X6	-0.33	0.333	-0.33	0.333	0	0.333	-0.33	-0.33	0	0.333	-0.33	0	-0.33	0	0	0

The Alias Matrix shows how the coefficients of the main effect terms in the model are biased by potentially active two-factor interaction effects. The column labels identify interactions. For example, in the X1 row, the column X2*X3 has a value of -0.33 and the column X2*X4 has a value of 0.333. This means that the expected value of the main effect of X1 is the sum of the main effect of X1 plus -0.33 times the effect of X2*X3, plus 0.333 times the effect of X2*X4, and so on, for the rest of the X1 row. In order for the estimate of the main effect of X1 to be meaningful, you must assume that these interactions are negligible in size compared to the effect of X1.

Tip: The Alias Matrix is a generalization of the confounding pattern in fractional factorial designs.

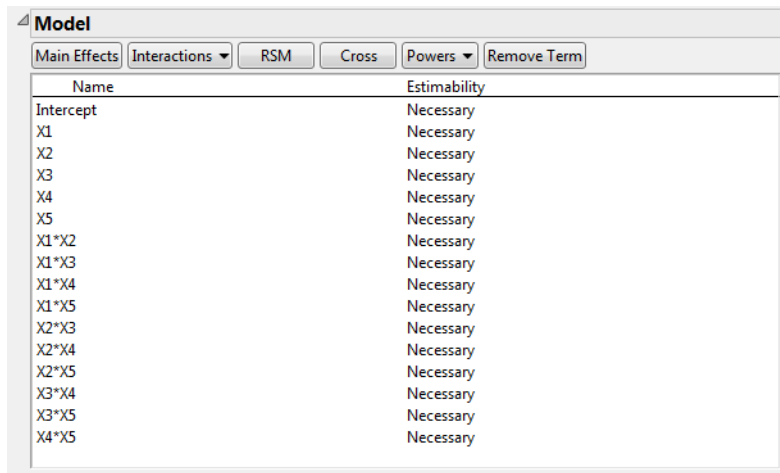
Design That Estimates All Two-Factor Interactions

The Alias Matrix in Figure 5.5 shows partial aliasing of effects. In other cases, main effects might be fully aliased, or *confounded*, with two-factor interactions. In both of these cases, strong two-factor interactions can confuse the results of main effects only experiments. To avoid this risk, create a design that resolves all two-factor interactions.

In this example, you create a resolution V screening design. Two-factor interactions are orthogonal, but they are confounded with three-factor interactions.

1. Select **DOE > Custom Design**.
2. Type 5 next to **Add N Factors**.
3. Click **Add Factor > Continuous**.
4. Click **Continue**.
5. In the Model outline, select **Interactions > 2nd**.

Figure 5.6 Model Outline Showing Interactions



The screenshot shows the 'Model' window in Minitab. At the top, there are tabs for 'Main Effects', 'Interactions', 'RSM', 'Cross', 'Powers', and 'Remove Term'. The 'Interactions' tab is selected. Below the tabs is a table with two columns: 'Name' and 'Estimability'. The table lists the following terms and their estimability:

Name	Estimability
Intercept	Necessary
X1	Necessary
X2	Necessary
X3	Necessary
X4	Necessary
X5	Necessary
X1*X2	Necessary
X1*X3	Necessary
X1*X4	Necessary
X1*X5	Necessary
X2*X3	Necessary
X2*X4	Necessary
X2*X5	Necessary
X3*X4	Necessary
X3*X5	Necessary
X4*X5	Necessary

- Click **Minimum** to accept 16 for the number of runs.

Note: Setting the Random Seed in step 7 and Number of Starts in step 8 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

- (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
- (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 1, and click **OK**.
- Click **Make Design**.

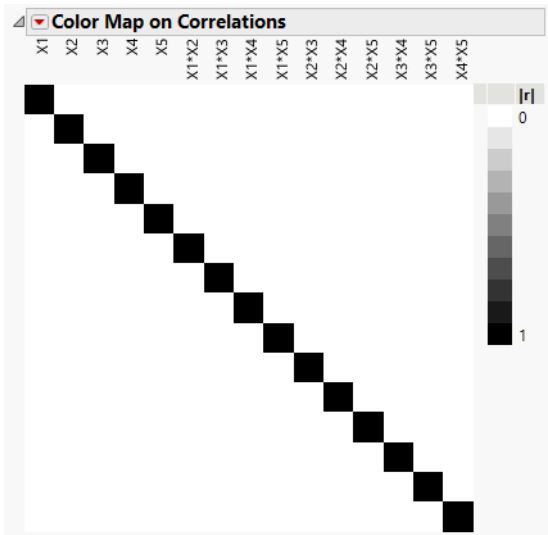
Figure 5.7 shows the runs of the design. All main effects and two-factor interactions are estimable because their Estimability was designated as Necessary (by default) in the Model outline.

Figure 5.7 Design to Estimate All Two-Factor Interactions

Design					
Run	X1	X2	X3	X4	X5
1	-1	-1	1	-1	-1
2	-1	-1	-1	1	-1
3	1	1	1	1	1
4	-1	-1	1	1	1
5	1	-1	1	-1	1
6	1	1	-1	1	-1
7	1	-1	-1	-1	-1
8	1	-1	1	1	-1
9	-1	1	1	-1	1
10	-1	-1	-1	-1	1
11	-1	1	-1	1	1
12	1	1	-1	-1	1
13	1	1	1	-1	-1
14	-1	1	1	1	-1
15	1	-1	-1	1	1
16	-1	1	-1	-1	-1

10. Open the **Design Evaluation > Color Map on Correlations** outline.

Figure 5.8 Color Map on Correlations



The Color Map indicates that the five main effects and the ten two-way interactions are all mutually orthogonal.

Design That Avoids Aliasing of Main Effects and Two-Factor Interactions

Suppose that your primary interest is in estimating the main effects of six continuous factors. However, you want to do this in a way that minimizes aliasing of main effects with potentially active two-factor interactions.

Your budget allows for only 16 runs. With six factors, there are 15 possible two-factor interactions. The minimum number of runs required to fit the constant, the six main effects, and the 15 two-factor interactions is 22.

In this example, you find a compromise between an 8-run main effects only design (see [“Design That Estimates Main Effects Only”](#) on page 135) and a 22-run design capable of fitting all the two-factor interactions. You use Alias Optimality as the optimality criterion to achieve your goal.

1. Select **DOE > Custom Design**.
2. Type 6 next to **Add N Factors**.
3. Click **Add Factor > Continuous**.
4. Click **Continue**.

The model includes the main effect terms by default. The default estimability of these terms is Necessary. In the Alias Terms outline, notice that second-order interactions are added. By default, all two-way interactions not included in the assumed model are added to the Alias Terms list.

5. Click the Custom Design red triangle and select **Optimality Criterion > Make Alias Optimal Design**.

The Make Alias Optimal Design selection tells JMP to generate a design that balances reduction in aliasing with D-efficiency. See [“Alias Optimality”](#) on page 131 in the “Custom Designs” chapter.

6. Click **User Specified** and change the number of runs to 16.

Figure 5.9 Factors, Model, Alias Terms, and Number of Runs

Custom Design

Responses

Factors

Add Factor Remove Add N Factors 1

Name	Role	Changes	Values
X1	Continuous	Easy	-1 1
X2	Continuous	Easy	-1 1
X3	Continuous	Easy	-1 1
X4	Continuous	Easy	-1 1
X5	Continuous	Easy	-1 1

Define Factor Constraints

Model

Main Effects Interactions RSM Cross Powers Remove Term

Name	Estimability
Intercept	Necessary
X1	Necessary
X2	Necessary
X3	Necessary
X4	Necessary
X5	Necessary
X6	Necessary

Alias Terms

Main Effects Interactions RSM Cross Powers Remove Term

Name
X1*X2
X1*X3
X1*X4
X1*X5
X1*X6
X2*X3
X2*X4
X2*X5

Design Generation

☐ Group runs into random blocks of size: 2

Number of Center Points: 0

Number of Replicate Runs: 0

Number of Runs:

☐ Minimum 7

☐ Default 12

☒ User Specified 16

Make Design

Note: Setting the Random Seed in step 7 and Number of Starts in step 8 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

- (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
- (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 10, and click **OK**.
- Click **Make Design**.

10. Open the **Design Evaluation > Alias Matrix** outline.

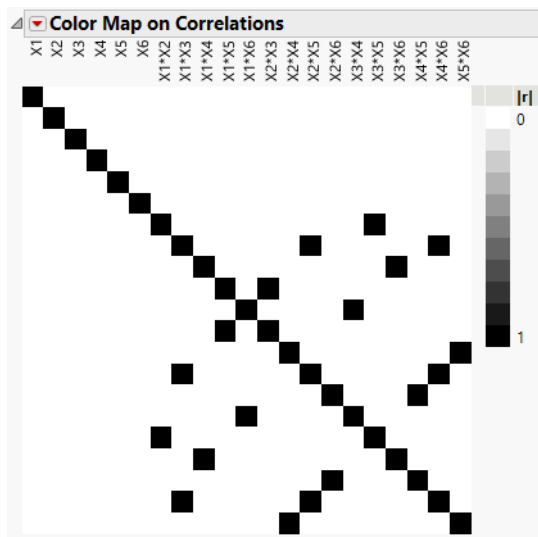
Figure 5.10 Alias Matrix

Alias Matrix															
Effect	X1*X2	X1*X3	X1*X4	X1*X5	X1*X6	X2*X3	X2*X4	X2*X5	X2*X6	X3*X4	X3*X5	X3*X6	X4*X5	X4*X6	X5*X6
Intercept	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
X1	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
X2	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
X3	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
X4	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
X5	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0
X6	0	0	0	0	0	0	0	0	0	0	0	0	0	0	0

All rows contain only zeros, which means that the Intercept and main effect terms are not biased by any two-factor interactions.

11. Open the **Design Evaluation > Color Map on Correlations** outline.

Figure 5.11 Color Map on Correlations



The Color Map on Correlations shows that main effects can be estimated independently of two-way interactions. However, some two-way interactions are fully aliased with other two-way interactions. Place your cursor over the off-diagonal red squares to see which two-way interactions are confounded.

It turns out that this particular design is a resolution IV orthogonal screening design. Main effects are not aliased with each other or with two-way interactions. But two-way interactions are fully aliased with other two-way interactions.

Supersaturated Screening Designs

It is common for brainstorming sessions to identify dozens of potentially active factors. Rather than reduce the list without the benefit of data, you can use a supersaturated design.

In a *saturated* design, the number of runs equals the number of model terms. In a *supersaturated* design, the number of model terms exceeds the number of runs (Lin, 1993). A supersaturated design can examine dozens of factors using fewer than half as many runs as factors. This makes it an attractive choice for factor screening when there are many factors and experimental runs are expensive.

Alternatively, use a group orthogonal supersaturated design for improved active effect identification over traditional supersaturated designs. See the [“Group Orthogonal Supersaturated Designs”](#) chapter on page 713.

Limitations of Supersaturated Designs

There are drawbacks to supersaturated designs:

- If the number of *active* factors is more than half the number of runs in the experiment, then it is likely that these factors will be impossible to identify. A general rule is that the number of runs should be at least four times larger than the number of active factors. In other words, if you expect that there might be as many as five active factors, you should plan on at least 20 runs.
- Analysis of supersaturated designs cannot yet be reduced to an automatic procedure. However, using forward stepwise regression is reasonable. In addition, the Screening platform (**DOE > Classical > Two Level Screening > Fit Two Level Screening**) offers a streamlined analysis.

Generate a Supersaturated Design

In this example, you want to construct a supersaturated design to study 12 factors in 8 runs. To create a supersaturated design, you set the Estimability of all model terms (except the intercept) to If Possible.

Note: This example is for illustration only. You should have at least 14 runs in any supersaturated design. If there are as many as four active factors, it is very difficult to interpret the results of an 8-run design. See [“Limitations of Supersaturated Designs”](#) on page 144.

1. Select **DOE > Custom Design**.
2. Type 12 next to **Add N Factors**.
3. Click **Add Factor > Continuous**.
4. Click **Continue**.

5. In the Model outline, select all terms except the Intercept.
6. Click Necessary next to any effect and change it to If Possible.
Setting the effects to If Possible ensures that JMP uses the Bayesian D-optimality criterion to obtain the design.

Figure 5.12 Factors, Model, and Number of Runs

Factors

Add Factor ▼ Remove Add N Factors 1

Name	Role	Changes	Values
X1	Continuous	Easy	-1 1
X2	Continuous	Easy	-1 1
X3	Continuous	Easy	-1 1
X4	Continuous	Easy	-1 1
X5	Continuous	Easy	-1 1
X6	Continuous	Easy	-1 1
X7	Continuous	Easy	-1 1
X8	Continuous	Easy	-1 1
X9	Continuous	Easy	-1 1
X10	Continuous	Easy	-1 1
X11	Continuous	Easy	-1 1
X12	Continuous	Easy	-1 1

Define Factor Constraints

Model

Main Effects Interactions ▼ RSM Cross Powers ▼ Remove Term

Name	Estimability
Intercept	Necessary
X1	If Possible
X2	If Possible
X3	If Possible
X4	If Possible
X5	If Possible
X6	If Possible
X7	If Possible
X8	If Possible
X9	If Possible
X10	If Possible
X11	If Possible
X12	If Possible

Alias Terms

Design Generation

☐ Group runs into random blocks of size: 2

Number of Center Points: 0

Number of Replicate Runs: 0

Number of Runs:

☐ Minimum 1

☒ Default 8

☐ User Specified 8

Make Design

7. In the Alias Terms outline, select all effects and click **Remove Term**.

This ensures that only the main effects appear in the Color Map on Correlations. This plot is constructed once the design is created.

8. Click the Custom Design red triangle and select **Simulate Responses**.

This option generates random responses that appear in your design table. You will use these responses to see how to analyze experimental data.

Keep the Number of Runs set to the Default of 8.

Note: Setting the Random Seed in step 9 and Number of Starts in step 10 reproduces the design shown in this example. In constructing a design on your own, these steps are not necessary.

9. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
10. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 5, and click **OK**.
11. Click **Make Design**.
12. Click **Make Table**.

Do not close your Custom Design window. You return to it later in this example.

Figure 5.13 Design Table with Simulated Responses

	X1	X2	X3	X4	X5	X6	X7	X8	X9	X10	X11	X12	Y	Y Simulated
1	1	1	-1	1	-1	-1	-1	1	1	-1	1	1	3.5435384836	3.54353848
2	1	-1	1	-1	-1	1	-1	1	-1	-1	-1	1	-0.260441268	-0.2604413
3	-1	1	1	1	1	1	-1	-1	1	-1	1	-1	2.4203423585	2.42034236
4	1	1	1	-1	1	-1	1	-1	-1	1	1	1	6.577962661	6.57796266
5	-1	-1	-1	-1	1	-1	-1	1	1	1	-1	-1	-3.722945159	-3.7229452
6	-1	-1	1	1	-1	-1	1	1	-1	1	1	-1	1.6889943423	1.68899434
7	-1	1	-1	-1	-1	1	1	-1	-1	-1	-1	-1	-4.493514962	-4.493515
8	1	-1	-1	1	1	1	1	-1	1	1	-1	1	5.4294393499	5.42943935

The response columns, Y and Y Simulated, are initiated with the same simulated values. The values are random values from a $N(0, s)$ distribution, where s is the RMSE from the power analysis dialog with a default of 1. The Y Simulated values are updated with randomly generated values using the model defined by the parameter values in the Simulate Responses window. The Y column is intended for your true responses after the experiment has been run.

Figure 5.14 Simulate Responses Window

The screenshot shows the 'Simulate Responses' window. It contains a table with 'Effects' and 'Y' columns. The 'Effects' column lists Intercept, X1, X2, X3, X4, X5, X6, X7, X8, X9, X10, X11, and X12. The 'Y' column contains the value 1 for each effect. Below the table is a 'Reset Coefficients' button. Below that is the 'Distribution' section, which has three radio buttons: 'Normal' (selected), 'Binomial', and 'Poisson'. To the right of the 'Normal' radio button is a text box labeled 'Error σ ' with the value 1. At the bottom is an 'Apply' button.

Effects	Y
Intercept	1
X1	1
X2	1
X3	1
X4	1
X5	1
X6	1
X7	1
X8	1
X9	1
X10	1
X11	1
X12	1

Reset Coefficients

Distribution

☒ Normal Error σ : 1

☐ Binomial

☐ Poisson

Apply

The Simulate Responses window shows default coefficients of 1 for all terms, a Normal distribution selection and Error σ of 1. The values in the Y and Y Simulated column currently reflect only random variation.

13. Change the values of the coefficients in the Simulate Responses window as shown in Figure 5.15.

Figure 5.15 Parameter Values for Simulated Responses

Simulate Responses

Effects	Y
Intercept	100
X1	10
X2	0
X3	0
X4	0
X5	0
X6	0
X7	0
X8	0
X9	0
X10	0
X11	10
X12	0

Reset Coefficients

Distribution

☒ Normal

Error σ : 1

☐ Binomial

☐ Poisson

Apply

14. Click **Apply**.

The response values in the Y Simulated column change.

Note: The response values are randomly generated. Your values will not match those in Figure 5.16 exactly.

Figure 5.16 Y Simulated Response Column with X1 and X11 Active

	X1	X2	X3	X4	X5	X6	X7	X8	X9	X10	X11	X12	Y	Y Simulated
1	1	1	-1	1	-1	-1	-1	1	1	-1	1	1	2.6497553994	113.318701
2	1	-1	1	-1	-1	1	-1	1	-1	-1	-1	1	-0.399454412	112.516952
3	-1	1	1	1	1	1	-1	-1	1	-1	1	-1	5.5131126481	64.5883626
4	1	1	1	-1	1	-1	1	-1	-1	1	1	1	4.4095312768	113.211358
5	-1	-1	-1	-1	1	-1	-1	1	1	1	-1	-1	-2.738425301	66.4591064
6	-1	-1	1	1	1	-1	-1	1	1	-1	1	-1	0.3707954895	67.5814445
7	-1	1	-1	-1	-1	1	1	-1	-1	-1	-1	-1	-4.42155356	66.0901851
8	1	-1	-1	1	1	1	1	-1	1	1	-1	1	2.2250808685	113.770165

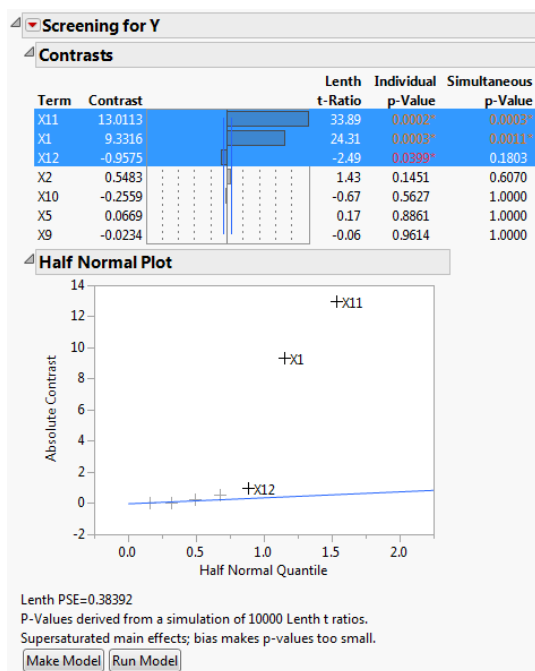
In your simulation, you specified X1 and X11 as active factors with large effects relative to the error variation. For this reason, your analysis of the data should identify these two factors as active.

Analyze a Supersaturated Design Using the Screening Platform

The Screening platform provides a way to identify active factors. Use the screening platform to analyze the Y Simulated values in the design table (Figure 5.16). The Screening platform is located under the **DOE > Classical** menu.

1. Select **DOE > Classical > Two Level Screening > Fit Two Level Screening**.
2. Select Y Simulated and click **Y**.
3. Select X1 through X12 ,click **X** and click **OK**.

Figure 5.17 Screening Report for Supersaturated Design



The factors X1 and X11 have large contrast and Lenth t-Ratio values. Also, their Simultaneous p-Values are small. In the Half Normal Plot, both X1 and X11 fall far from the line. The Contrasts and the Half Normal Plot reports indicate that X1 and X11 are active. Although X12 has an Individual p-Value less than 0.05, its effect is much smaller than that of X1 and X11.

Because the design is supersaturated, *p*-values might be smaller than they would be in a model where all effects are estimable. This is because effect estimates are biased by other potentially active main effects. In Figure 5.17, a note directly above the **Make Model** button warns you of this possibility.

You might also want to check whether the effects that appear active could be highly correlated with other effects. When this occurs, one effect can mask the true significance of

another effect. The Color Map in Figure 5.19 displays absolute correlations between effects.

4. Click **Make Model**.

The constructed model contains only the effects X1, X11, and X12.

5. Click **Run** in the Model Specification window.

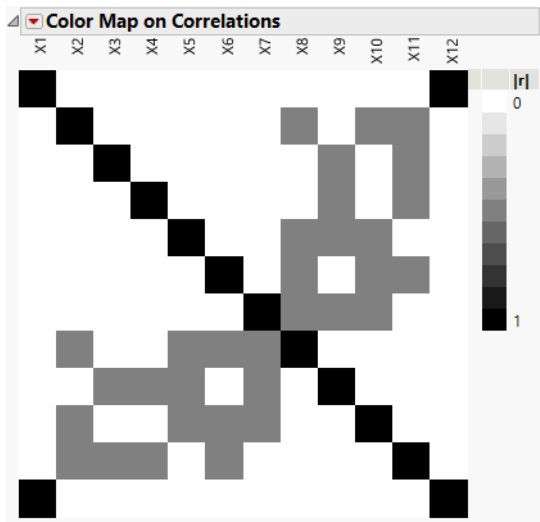
Figure 5.18 Parameter Estimates for Model

Parameter Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Intercept	100.12645	0.319502	313.38	<.0001*
X11	11.347787	0.360447	31.48	<.0001*
X1	9.8209584	0.319502	30.74	<.0001*
X12	-1.1329	0.360447	-3.14	0.0347*

Note that the parameter estimates for X11 and X1 are close to the theoretical values that you used to simulate the model. See Figure 5.15, where you specified a model with X1 = 10 and X11 = 10. The significance of the factor X12 is an example of a false positive.

6. In your Custom Design window, open the **Design Evaluation > Color Map on Correlations** outline.

Figure 5.19 Color Map on Correlations Outline



With your cursor, place your mouse pointer over cells to see the absolute correlations. Notice that X1 has correlations as high as 0.5 with other main effects (X4, X5, X7). (Figure 5.19 uses JMP default colors.)

Analyze a Supersaturated Design Using Stepwise Regression

Stepwise regression is another way to identify active factors. The design table in Figure 5.16 contains three scripts. The Model script analyzes your data using stepwise regression in the Fit Model platform.

1. In the Table panel of the design table, click the green triangle next to the **Model** script.
2. Change the **Personality** from **Standard Least Squares** to **Stepwise**.
3. Click **Run**.
4. In the Stepwise Fit for Y report, change the **Stopping Rule** to **Minimum AICc**.

For designed experiments, BIC is typically a more lenient stopping rule than AICc as it tends to allow inactive effects into the model.

5. Click **Go**.

Figure 5.20 Stepwise Regression for Supersaturated Design

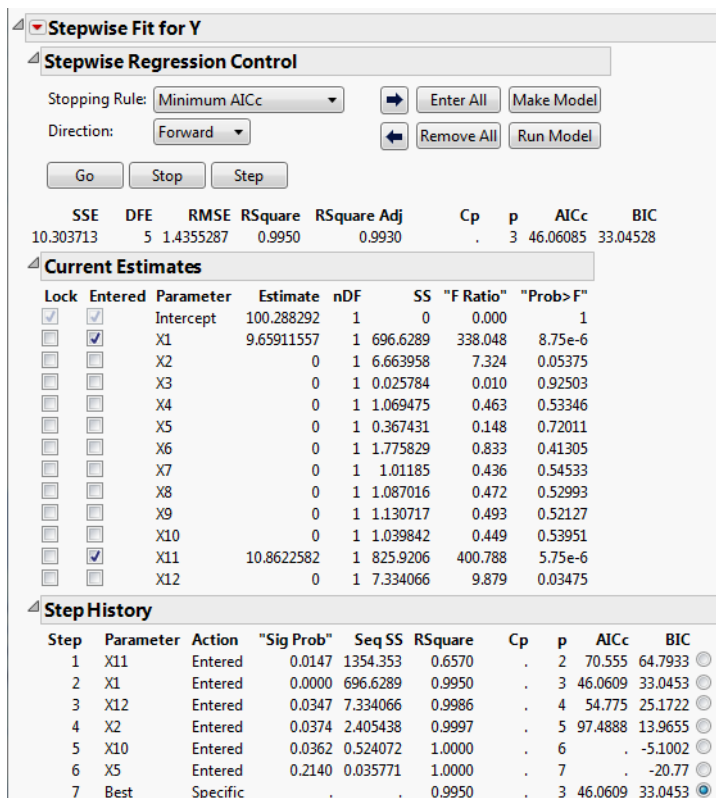


Figure 5.20 shows that the selected model consists of the two active factors, X1 and X11. The step history appears in the bottom part of the report. Keep in mind that correlations

between X1 and X11 and other factors could mask the effects of other active factors (Figure 5.19).

Note: This example defines two large main effects and sets the rest to zero. Real-world situations can be less likely to have such clearly differentiated effects.

Design for Fixed Blocks

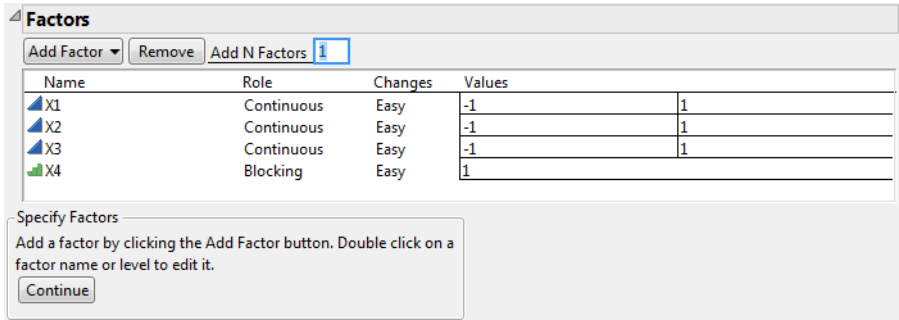
Traditional screening designs require block sizes to be a power of two. However, the Custom Design platform can create designs with fixed blocks of any size.

Suppose that you want to study three factors. You can run only three trials per day and you expect substantial day-to-day variation. Consequently, you need to block your design over multiple days. Also, in this study, you are interested in estimating all two-factor interactions. In this example, you construct a design with three runs per block.

- 1. Select **DOE > Custom Design**.
- 2. In the Factors outline, type 3 next to **Add N Factors**.
- 3. Click **Add Factor > Continuous**.
- 4. Click **Add Factor > Blocking > 3 runs per block**.

The blocking factor X4 shows only one level under Values. This is because the run size is unknown at this point.

Figure 5.21 Factors Outline Showing One Block for X4



- 5. Click **Continue**.

Figure 5.22 Factors Outline Showing Three Blocks for X4

The screenshot shows the Minitab Custom Design interface. The 'Factors' section is expanded, showing a table of factors and their values. Factor X4 is a blocking factor with three levels, indicated by the three columns under its 'Values' header.

Name	Role	Changes	Values
X1	Continuous	Easy	-1 1
X2	Continuous	Easy	-1 1
X3	Continuous	Easy	-1 1
X4	Blocking	Easy	1 2 3

Below the table, the 'Design Generation' section is visible, showing options for the number of center points, replicate runs, and the number of runs (set to 9).

The Factors outline now shows an appropriate number of blocks, calculated as the Default run size divided by the number of runs per block. For this example, the default sample size of 9 requires three blocks. The Factors outline now shows that X4 has three values, indicating the three blocks.

Note: If you specify a different number of runs, the Factors outline updates to show the appropriate number of values for the blocking factor.

6. Select the three continuous factors, X1, X2, and X3, in the Factors outline.
7. In the Model outline, click **Interactions > 2nd**.

Figure 5.23 Factors Outline Showing Six Blocks for X4

Factors

Add Factor ▼ Remove Add N Factors 1

Name	Role	Changes	Values
X1	Continuous	Easy	-1 1
X2	Continuous	Easy	-1 1
X3	Continuous	Easy	-1 1
X4	Blocking	Easy	1 2 3 4 5 6

Define Factor Constraints

Model

Main Effects Interactions ▼ RSM Cross Powers ▼ Remove Term

Name	Estimability
Intercept	Necessary
X1	Necessary
X2	Necessary
X3	Necessary
X4	Necessary
X1*X2	Necessary
X1*X3	Necessary
X2*X3	Necessary

Alias Terms

Design Generation

Number of Center Points: 0
Number of Replicate Runs: 0

Number of Runs:

☐ Minimum 12
☒ Default 18
☐ User Specified 18

Make Design

The Number of Runs panel now shows that 18 is the Default run size. The Factors outline now shows six values for X4, indicating six blocks.

Note: Setting the Random Seed in step 8 and Number of Starts in step 9 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

8. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
9. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 5, and click **OK**.
10. Click **Make Design**.

Figure 5.24 Fixed Block Design

Design				
Run	X1	X2	X3	X4
1	1	-1	-1	1
2	-1	-1	-1	2
3	1	1	-1	3
4	-1	-1	-1	4
5	1	1	1	5
6	-1	-1	1	6
7	-1	1	-1	1
8	1	1	-1	2
9	-1	1	1	3
10	1	-1	1	4
11	-1	-1	1	5
12	-1	1	-1	6
13	-1	-1	1	1
14	-1	1	1	2
15	1	-1	1	3
16	-1	1	1	4
17	1	-1	-1	5
18	1	1	1	6

In the design, look at the blocking factor, X4. The six blocks are represented. When you conduct your experiment, each day you will run three trials, where X4 = 1 on the first day, X4 = 2 on the second day, and so on. So you want the design table to randomize the trials within blocks. In the Output Options panel, the Randomize within Blocks option is already selected for Run Order.

11. Click **Make Table**.

Figure 5.25 Design Table for Fixed Block Design

	X1	X2	X3	X4	Y
1	1	-1	-1	1	•
2	-1	1	-1	1	•
3	-1	-1	1	1	•
4	-1	1	1	2	•
5	1	1	-1	2	•
6	-1	-1	-1	2	•
7	1	-1	1	3	•
8	-1	1	1	3	•
9	1	1	-1	3	•
10	-1	1	1	4	•
11	1	-1	1	4	•
12	-1	-1	-1	4	•
13	1	-1	-1	5	•
14	-1	-1	1	5	•
15	1	1	1	5	•
16	-1	-1	1	6	•
17	-1	1	-1	6	•
18	1	1	1	6	•

The rows in the design table are grouped by each day's runs. This design enables you to estimate the block effect, all main effects, and two-factor interactions.

Response Surface Experiments

Response surface experiments typically involve a small number (generally 2 to 8) of continuous factors that have been identified as active. The main goal of a response surface experiment is to develop a predictive model of the relationship between the factors and the response. Often, you use the predictive model to find better operating settings for your process. For this reason, your assumed model for a response surface experiment is usually quadratic.

Because a screening design is focused on identifying active effects, a measure of its quality is the size of the relative variance of the coefficients. You want these relative variances to be small. D-optimality addresses these relative variances.

In response surface experiments, the prediction variance over the range of the factors is more important than the variance of the coefficients. The prediction variance over the design space is addressed by I-optimality. An I-optimal design tends to place fewer runs at the extremes of the design space than does a D-optimal design. For more information about D- and I-optimality, see [“Optimality Criteria”](#) on page 127 in the “Custom Designs” chapter.

By default, Custom Design uses the Recommended option for the Optimality Criterion. Custom Design uses the I-optimality criterion as the Recommended criterion whenever you add quadratic effects using the RSM button in the Model outline. Otherwise, Custom Design uses the D-optimality criterion as the Recommended criterion. See [“Optimality Criteria”](#) on page 127 in the “Custom Designs” chapter.

- [“Response Surface Design”](#)
- [“Response Surface Design with Flexible Blocking”](#)
- [“Comparison of a D-Optimal and an I-Optimal Response Surface Design”](#)
- [“Response Surface Design With Constraints and Categorical Factor”](#)

Response Surface Design

The following example contains these sections:

- [“Construct a Response Surface Design”](#) on page 157
- [“Analyze the Experimental Results”](#) on page 160

Construct a Response Surface Design

Construct a response surface design for three continuous factors that you have identified as active. You want to find process settings to maintain your response(Y) within specifications. The lower and upper specification limits for Y are 54 and 56, respectively, with a target of 55.

1. Select **DOE > Custom Design**.
2. In the Responses outline, click **Maximize** and select **Match Target**.
3. Type 54 as the **Lower Limit** and 56 as the **Upper Limit**.
4. Leave **Importance** blank.

Because there is only one response, the Importance value is set to 1 by default.

5. Type 3 next to **Add N Factors**.
6. Click **Add Factor > Continuous**.

This adds three continuous factors: X1, X2, and X3.

7. Click **Continue**.
8. In the Model outline, click the **RSM** button.

This adds quadratic and interaction terms to the model. It also sets the value of the Recommended optimality criterion to I-optimality. You can verify this in the Design Diagnostics outline once you click Make Design.

Leave the Default Number of Runs set to 16.

Note: Setting the Random Seed in step 9 and Number of Starts in step 10 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

9. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 929281409, and click **OK**.
10. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 40, and click **OK**.
11. Click **Make Design**.

The Design outline shows the design.

Figure 5.26 RSM Design

Design			
Run	X1	X2	X3
1	0	-1	0
2	1	-1	1
3	0	0	0
4	-1	1	1
5	-1	0	0
6	1	1	1
7	-1	1	-1
8	-1	-1	-1
9	0	0	1
10	1	0	0
11	0	1	0
12	0	0	0
13	1	1	-1
14	0	0	-1
15	-1	-1	1
16	1	-1	-1

In order to estimate quadratic effects, a response surface design uses three levels for each factor. Note that the design in Figure 5.26 is a face-centered Central Composite Design with two center points.

12. Open the **Design Evaluation > Design Diagnostics** outline.

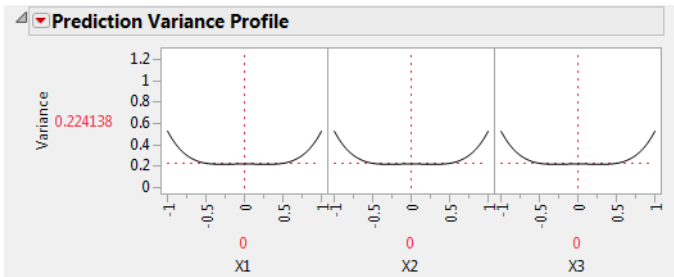
Figure 5.27 Design Diagnostics Outline

Design Diagnostics	
I Optimal Design	
D Efficiency	42.99905
G Efficiency	88.62743
A Efficiency	30.68134
Average Variance of Prediction	0.340517
Design Creation Time (seconds)	0.033333

The first line in the Design Diagnostics outline identifies the optimality criterion being used. This design is I-optimal.

13. Open the **Design Evaluation > Prediction Variance Profile** outline.

Figure 5.28 Prediction Variance Profile

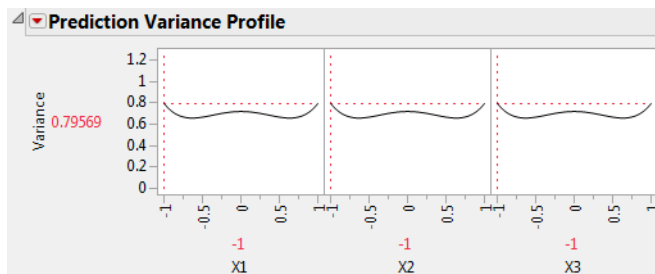


The vertical axis shows the relative prediction variance of the expected value of the response. The *relative prediction variance* is the prediction variance divided by the error variance. When the relative prediction variance is one, its absolute variance equals the error variance of the regression model.

The profiler shows values of the relative prediction variance over the design space. You can move the sliders to explore the prediction variance's behavior. The prediction variance is smallest in the center of the design space. It is fairly constant, with values only slightly larger than 0.2, for factor settings between -0.5 and 0.5. The prediction variance increases as the settings approach the design space boundaries.

14. Click the Prediction Variance Profile red triangle and select **Maximize Variance**.

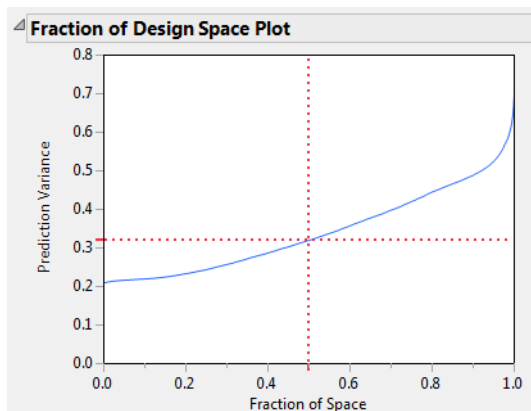
Figure 5.29 Prediction Variance Profile with Relative Variance Maximized



The profiler shows that the maximum value of the relative prediction variance is 0.79569.

15. Open the **Design Evaluation > Fraction of Design Space Plot** outline.

Figure 5.30 Fraction of Design Space Plot



The blue curve in the plot shows the relative prediction variance as a function of the fraction of design space. The red dashed cross hairs indicate that, for 50% of the design space, the prediction variance is about 0.32 or less. Use the cross hair tool to draw other inferences. For example, when the Fraction of Space is 0.95, the Prediction Variance is about 0.52. This means that for 95% of the design space, the relative prediction variance is below 0.52.

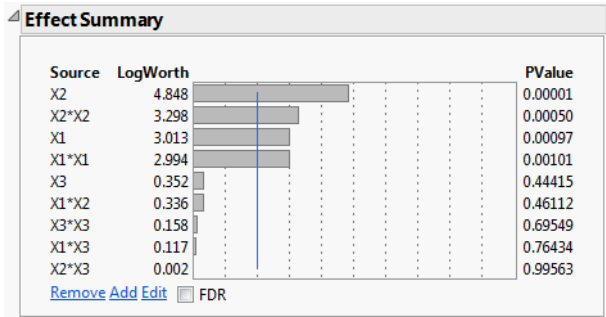
Analyze the Experimental Results

The Custom RSM.jmp sample data table contains the results of the experiment. The Model script opens a Fit Model window showing all of the effects specified in the DOE window’s Model outline. This script was saved to the data table by the Custom Design platform.

- 1. Select **Help > Sample Data Library** and open Design Experiment/Custom RSM.jmp.
- 2. In the Table panel, click the green triangle next to the **Model** script.
- 3. Click **Run**.

The Effect Summary report shows the LogWorth and PValue for each effect in the model. The vertical blue line in the plot is set at the value 2. A LogWorth that exceeds 2 is significant at the 0.01 level.

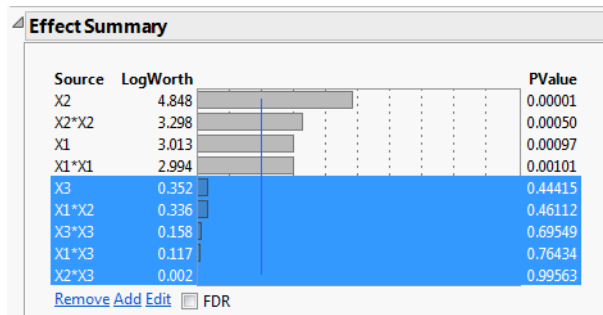
Figure 5.31 Effect Summary Report



The report shows that X1, X2, X1*X1, and X2*X2 are significant at the 0.01 level. None of the other effects are significant at even the 0.10 level. Reduce the model by removing these insignificant effects.

- 4. In the Effect Summary report, select X3, X1*X2, X3*X3, X1*X3, and X2*X3.

Figure 5.32 Effect Summary Report with Insignificant Effects Selected



5. Click **Remove**.

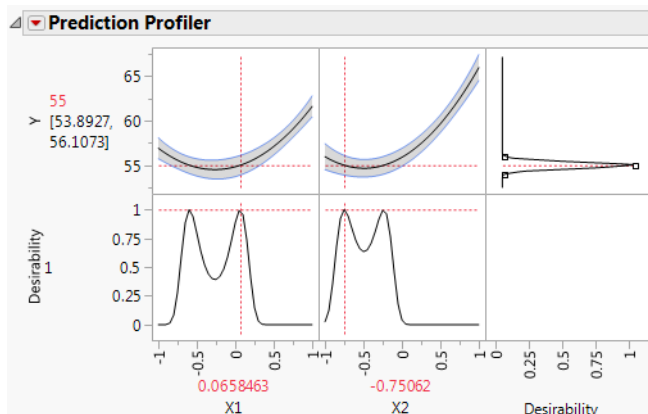
The Fit Least Squares report is updated to show a model containing only the significant effects: X1, X2, X1*X1, and X2*X2.

Use the Prediction Profiler (at the bottom of the Fit Least Squares window) to explore how the predicted response (Y) changes as you vary the factors X1 and X2. Note the quadratic behavior of Y across the values of X1 and X2.

Remember that you entered response limits for Y in the Responses outline of the Custom Design window. As a result, the Response Limit column property is attached to the Y column in the design table. The Desirability function for Y (in the top plot at right) is based on the information contained in the Response Limit column property. JMP uses this function to calculate Desirability as a function of the settings of X1 and X2. The traces of the Desirability function appear in the bottom row of plots.

6. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

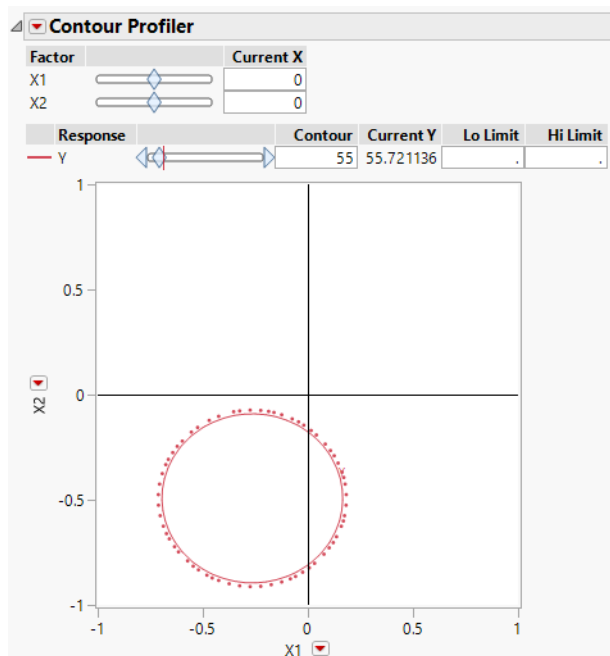
Figure 5.33 Prediction Profiler with Desirability Maximized



The predicted response achieves the target value of 55 at the process settings shown in red above X1 and X2. Figure 5.33 shows that a value of X1 near -0.65 also achieves a predicted value of 55 when $X2 = -0.75062$. In fact, your Prediction Profiler might show different settings as those that maximize desirability. This is because the predicted response is 55 for many settings of X1 and X2.

7. Click the Response Y red triangle and select **Factor Profiling > Contour Profiler**.
8. In the Contour Profiler report, type 55 as the value for **Contour**.

Figure 5.34 Contour Profiler



The settings of X1 and X2 that correspond to the red contour have predicted response values of 55. You might want to select from among these process settings based on cost efficiency.

Response Surface Design with Flexible Blocking

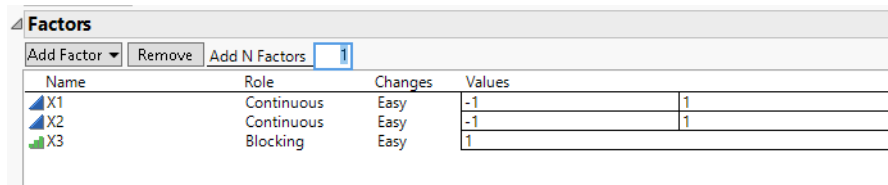
When optimizing a process, you might need to include qualitative factors in your experiment as well as continuous factors. You might need to block by qualitative factors such as batch or day, or include qualitative factors such as machine or delivery mechanism. But the Response Surface Design platform supports only continuous factors. To obtain a response surface design with a qualitative factor, you can replicate the design over each level of the factor. However, this is inefficient. The Custom Design platform constructs an optimal design with fewer runs.

In this example, you construct a response surface design that accommodates two continuous factors and a blocking factor with four runs per block. You can include categorical or discrete numeric factors in a similar fashion.

1. Select **DOE > Custom Design**.
2. Type 2 next to **Add N Factors**.
3. Click **Add Factor > Continuous**.
4. Click **Add Factor > Blocking > 4 runs per block**.

Notice that only one level appears under Values. This is because the number of blocks cannot be determined until the number of runs is determined.

Figure 5.35 Factors Outline with Two Continuous Factors and a Blocking Factor



Name	Role	Changes	Values
X1	Continuous	Easy	-1 1
X2	Continuous	Easy	-1 1
X3	Blocking	Easy	1

5. Click **Continue**.

The Default number of runs is 12. The Factors outline updates to show three levels for the Blocking factor, X3. Because you required X3 to have four runs per block, the 12 runs allow three blocks.

6. Click **RSM**.

An informational JMP Alert window reminds you that the blocking factor cannot appear in interaction or quadratic terms. JMP adds only the appropriate RSM terms to the list.

7. Click **OK** to dismiss the message.

Quadratic and interactions terms for X1 and X2 are added to the model. Because you added RSM terms, the Recommended optimality criterion changes from D-Optimal to I-Optimal. You can see this later in the Design Diagnostics outline.

Figure 5.36 Model Outline with Response Surface Effects

Model	
Main Effects	Interactions ▼ RSM Cross Powers ▼ Remove Term
Name	Estimability
Intercept	Necessary
X1	Necessary
X2	Necessary
X3	Necessary
X1*X1	Necessary
X1*X2	Necessary
X2*X2	Necessary

Note: Setting the Random Seed in step 8 and Number of Starts in step 9 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

8. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
9. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 5, and click **OK**.
10. Click **Make Design**.
11. Open the **Design Evaluation > Design Diagnostics** outline.

Figure 5.37 Design Diagnostics Outline

Design Diagnostics	
I Optimal Design	
D Efficiency	50.85382
G Efficiency	56.88003
A Efficiency	37.51264
Average Variance of Prediction	0.49973
Design Creation Time (seconds)	0

The first line in the Design Diagnostics outline identifies the optimality criterion being used. This design is I-optimal.

12. Click **Make Table**.

Figure 5.38 Design Table with Blocking Factor

	X1	X2	X3	Y
1	0	-1	1	•
2	1	1	1	•
3	-1	0	1	•
4	0	0	1	•
5	0	0	2	•
6	1	0	2	•
7	-1	1	2	•
8	0	-1	2	•
9	0	0	3	•
10	1	-1	3	•
11	-1	-1	3	•
12	0	1	3	•

Because the default Run Order was Randomize within Blocks, the levels of the blocking factor (X3) are sorted.

13. In the Table panel of the design table, click the green triangle next to the **Model** script.

Figure 5.39 Fit Model Window

Model Specification

Select Columns: 4 Columns (X1, X2, X3, Y)

Pick Role Variables:

- Y: optional
- Weight: optional numeric
- Freq: optional numeric
- Validation: optional
- By: optional

Personality: Standard Least Squares

Emphasis: Effect Screening

Buttons: Help, Run, Recall, Remove

Keep dialog open: ☐

Construct Model Effects:

- Add: X1 & RS, X2 & RS, X3, X1*X1, X1*X2, X2*X2
- Cross: X3
- Nest: X1*X1, X1*X2, X2*X2
- Macros: X3
- Degree: 2
- Attributes: ☒
- Transform: ☒
- No Intercept: ☐

Notice the following:

- The blocking factor (X3) is entered as an effect.
- No interactions involving X3 are included.
- The other five effects define a response surface model for X1 and X2.

Comparison of a D-Optimal and an I-Optimal Response Surface Design

In this example, you explore the differences between I-optimality and D-optimality in the context of a two-factor response surface design.

I-Optimal Design

1. Select **DOE > Custom Design**.
2. Type 2 next to **Add N Factors**.
3. Click **Add Factor > Continuous**.
4. Click **Continue**.
5. Click **RSM**.

Quadratic and interactions terms for X1 and X2 are added to the model. Because you added RSM terms, the Recommended optimality criterion changes from D-Optimal to I-Optimal. You can see this later in the Design Diagnostics outline.

Note: Setting the Random Seed in step 6 and Number of Starts in step 7 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

6. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 383570403, and click **OK**.
7. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 8, and click **OK**.
8. Click **Make Design**.

Figure 5.40 I-Optimal Design

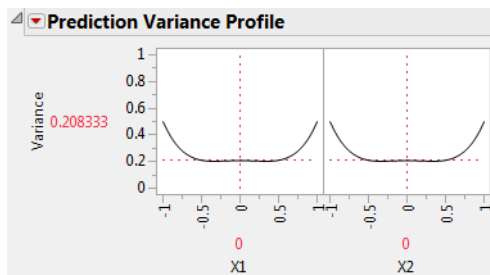
Design		
Run	X1	X2
1	0	0
2	1	0
3	-1	-1
4	0	0
5	0	1
6	-1	0
7	0	0
8	-1	1
9	1	-1
10	0	0
11	0	-1
12	1	1

In this I-optimal design, runs 1, 4, 7, and 10 are at the center point ($X1 = 0$ and $X2 = 0$). I-optimal designs tend to place more runs in the center (and consequently fewer runs at

the extremes) of the design space than do D-optimal designs. You can compare this design to the D-optimal design shown in Figure 5.42.

9. Open the **Design Evaluation > Prediction Variance Profile** outline.

Figure 5.41 Prediction Variance Profile for I-Optimal Model



The relative prediction variance of the expected response is smallest in the center of the design space.

10. Open the **Fraction of Design Space Plot** outline.

The Fraction of Design Space Plot appears on the left in Figure 5.44. When the Fraction of Space is 0.95, the vertical coordinate of the blue curve is about 0.5. This means that for about 95% of the design space, the relative prediction variance is below 50% of the error variance.

This Custom Design window contains your I-optimal design. Keep this window open. In the next section, you generate a D-optimal design, and compare the two.

D-Optimal Design

To compare Prediction Variance Profile and Fraction of Design Space plots for the I- and D-optimal designs:

1. In the Custom Design window containing your I-optimal design, click the Custom Design red triangle and select **Save Script to Script Window**.
A window appears, showing a script that reproduces your work.
2. In this new script window, select **Edit > Run Script**.
A duplicate Custom Design window appears, but with the Design Evaluation outlines closed.
3. In this new Custom Design window, click **Back**.
4. Click the Custom Design red triangle and select **Optimality Criterion > Make D-Optimal Design**.
5. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 383570403, and click **OK**.

6. Click **Make Design**.

You current Custom Design window contains your D-optimal design.

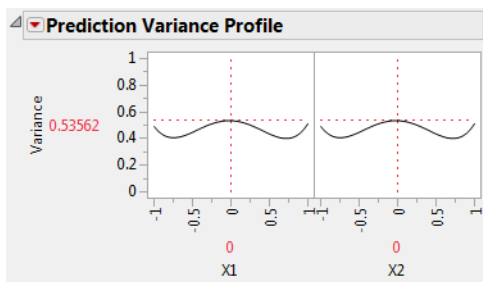
Figure 5.42 D-Optimal Design

Design		
Run	X1	X2
1	0	0
2	1	0
3	-1	-1
4	1	-1
5	0	1
6	-1	0
7	-1	-1
8	-1	1
9	1	-1
10	-1	1
11	0	-1
12	1	1

In this D-optimal design, run 1 is the only run at the center point. D-optimal designs tend to place more runs at the extremes of the design space than do I-optimal designs. Recall that the I-optimal design had four center runs (Figure 5.40).

7. Open the **Design Evaluation > Prediction Variance Profile** outline.

Figure 5.43 Prediction Variance Profile for D-Optimal Model



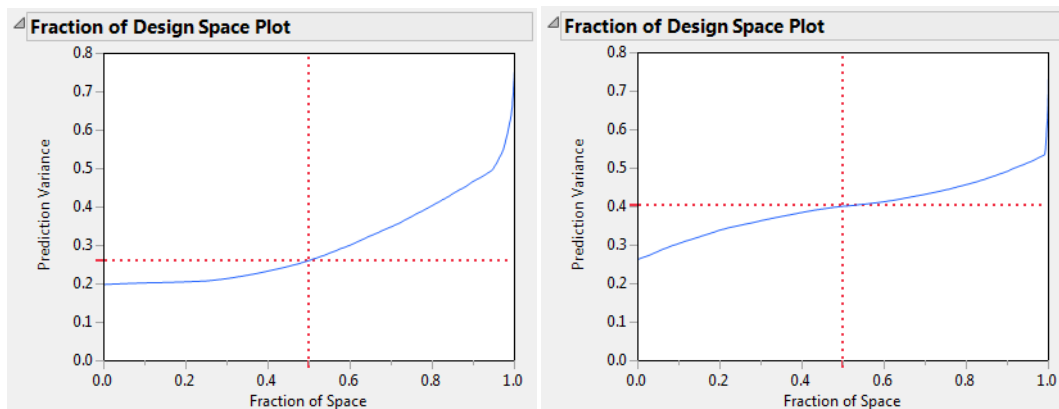
At the center of the design region, the relative prediction variance is 0.53562, as compared to 0.208333 for the I-optimal design (Figure 5.41). This means that the relative standard error is 0.732 for the D-optimal design and 0.456 for the I-optimal design. All else being equal, at the center of the design region, confidence intervals for the expected response based on the D-optimal design are about 60% wider than those based on the I-optimal design.

The Design outline shows that the D-optimal design has nine design points, one for every combination of X1 and X2 set to -1, 0, 1. The D-optimality criterion attempts to keep the relative prediction variance low at each of these design points. Explore the variance at the extremes of the design region by moving the sliders for X1 and X2 to -1 and 1. Note that the variance at these extreme points is usually smaller than the variance for the I-optimal design at these points.

8. Open the **Design Evaluation > Fraction of Design Space Plot** outline.

The Fraction of Design Space Plot appears on the right in Figure 5.44.

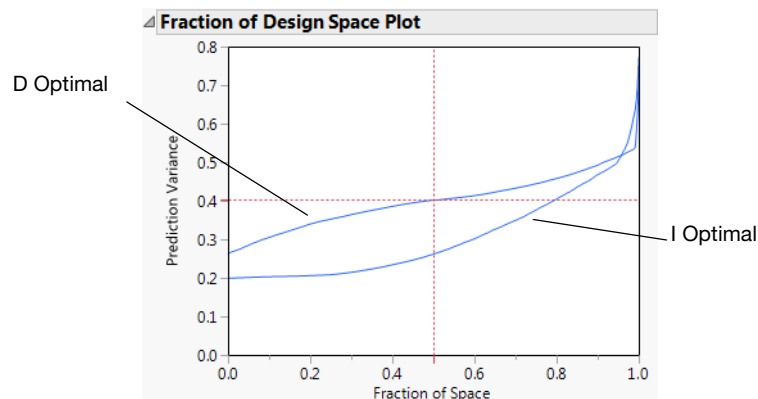
Figure 5.44 Fraction of Design Space Plots (I-Optimal on left, D-Optimal on right)



The red cross-hairs in each plot indicate the maximum prediction variance for 50% of the design space. For 50% of the design space, the prediction variance for the I-optimal design falls below about 0.27. For the D-optimal design, the prediction variance is about 0.4.

9. Right-click in the Fraction of Design Space Plot for your I-optimal design. Select **Edit > Copy Frame Contents**.
10. Right-click in the Fraction of Design Space Plot for your D-optimal design. Select **Edit > Paste Frame Contents**.

Figure 5.45 Fraction of Design Space Plots Superimposed



The variance curve for the I-optimal design is below the curve for the D-optimal design over at least 90% of the design space. This reflects the fact that I-optimality attempts to

minimize prediction variance over all of the design space. In contrast, D-optimality focuses on reducing prediction variance at the design points.

Response Surface Design With Constraints and Categorical Factor

In this example, you create a design to optimize the yield of a chemical reaction. Your experimental factors include a categorical factor at three levels, where constraints involve two of the levels. In this example, you will use the Disallowed Combinations Filter to enter the constraints.

Your response is Yield. Your factors are the following:

- Time: The range of interest is 500 to 560 seconds.
- Temperature: The range of interest is 350 to 750 degrees Kelvin.
- Catalyst: Three catalysts A, B, and C, must be tested.

Your design must allow for constraints on two of the levels of Catalyst:

- When catalyst B is used, the temperature must be above 400.
- When catalyst C is used, the temperature must be below 650.

Define the Response and Factors

1. Select **DOE > Custom Design**.
2. In the Response outline, double-click Y and change it to Yield.
Because your goal is to maximize Yield, leave the **Goal** set to Maximize.
3. In the Factors outline, type 2 next to **Add N Factors**.
4. Click **Add Factor > Continuous**.
5. Click **Add Factor > Categorical > 3 Level**.
6. Rename the factors Time, Temperature, and Catalyst.
7. Change the **Values** for Time to 500 and 560.
8. Change the **Values** for Temperature to 350 and 750.
9. Change the **Values** for Catalyst to A, B, and C.

Figure 5.46 Factor Settings

Factors				
Add Factor ▼ Remove Add N Factors ⓘ				
Name	Role	Changes	Values	
Time	Continuous	Easy	500	560
Temperature	Continuous	Easy	350	750
▼ Catalyst	Categorical	Easy	A	B C

10. Click **Continue**.

Define the Constraints

1. Select **Use Disallowed Combinations Filter** in the Define Factor Constraints outline.
2. Select Temperature and Catalyst from the **Add Filter Factors** list.
3. Click **Add**.
4. Click 750 in the equation that appears above the Temperature slider and change it to 400. Press Enter (or click elsewhere).
This disallows factor settings with Temperature values below 400.
5. Click the B block under Catalyst.
Together with the constraint on Temperature, this disallows factor settings for which Catalyst is B and Temperature is below 400.
6. Click **OR**.
This allows you to define your second constraint.
7. Select Temperature and Catalyst from the **Add Filter Factors** list. (Your earlier selection may have been retained.)
8. Click **Add**.
9. In the panel that appears beneath the word OR, click 350 in the equation that appears above the Temperature slider and change it to 650. Press Enter (or click elsewhere).
This disallows factor settings with Temperature values above 650.
10. Click the C block under Catalyst.
Together with the constraint on Temperature, this disallows factor settings for which Catalyst is C and Temperature is above 650.

Figure 5.47 Constraints Defined

Define Factor Constraints

☐ None
☐ Specify Linear Constraints
☒ Use Disallowed Combinations Filter
☐ Use Disallowed Combinations Script

Disallowed Combinations

Clear Help

Temperature x

350 400

Catalyst (3) x

A B C

OR

Temperature x

650 750

Catalyst (3) x

A B C

AND OR

Add Response Surface Terms and Make Design

1. In the Model outline, select **RSM**.

A JMP Alert informs you that only quadratic terms for continuous factors are being added to the model.

2. Click **OK** to dismiss the alert.

JMP adds only the appropriate response surface terms to the model.

Note: Setting the Random Seed in step 3 and Number of Starts in step 4 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

3. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 654321, and click **OK**.
4. (Optional) Click the Custom Design red triangle, select **Number of Starts** and set it to 1000. Click **OK**.
5. Click **Make Design**.

Figure 5.48 Design Satisfying Constraints

Design			
Run	Time	Temperature	Catalyst
1	500	750	B
2	500	410	B
3	560	750	A
4	521	750	A
5	542	350	A
6	500	648.7299	C
7	500	350	A
8	533	630	C
9	506	350	C
10	560	530	A
11	560	410	B
12	530	570	B
13	530	550	A
14	530	570	B
15	557	645.1294	C
16	500	570	A
17	560	750	B
18	560	350	C

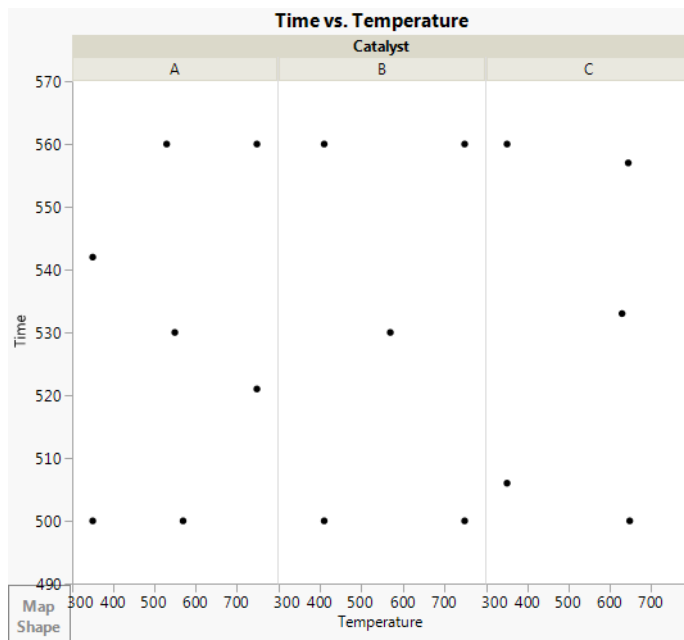
Because you added RSM terms to the Model outline, this is an I-optimal design. It satisfies your constraints:

- When catalyst B is used, the temperature must be above 400.
- When catalyst C is used, the temperature must be below 650.

View the Design

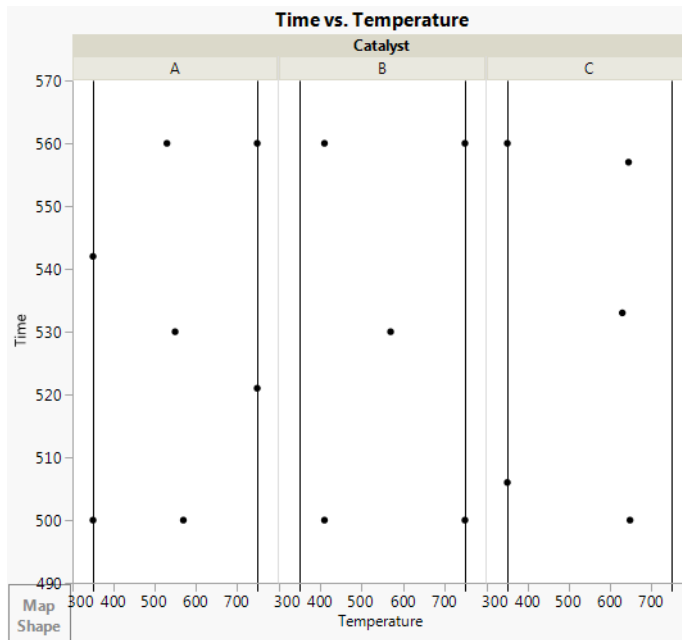
Recall that the design region consists of settings of Time between 500 and 560 and Temperature between 350 and 750. To obtain a geometric view of the design points within the design region, do the following:

1. Click **Make Table**.
2. Select **Graph > Graph Builder**.
3. Select Time and Temperature and drag them to the center of the template.
4. Deselect the **Smoother** icon. This is the second icon from the left above the template.
5. Select Catalyst and drag it to the **Group X** zone at the top of the template.

Figure 5.49 Design Points for Three Levels of Catalyst


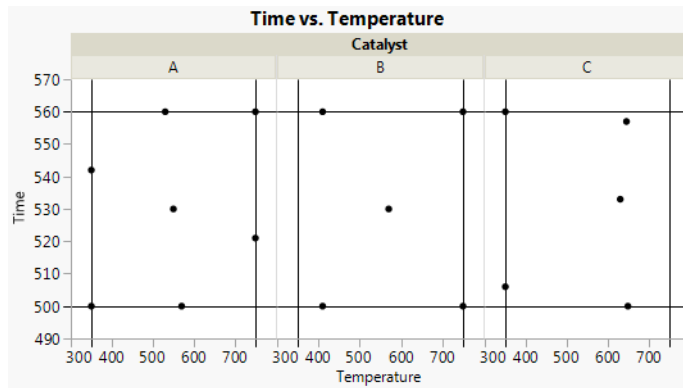
6. Double-click the Temperature axis.
7. Under Reference Lines, do the following:
 - Next to **Value**, enter 350.
 - Click **Add**.
 - Next to **Value**, enter 750.
 - Click **Add**.
8. Click **OK**.

Figure 5.50 Design Bounds on Temperature for Three Levels of Catalyst



9. Double-click the Time axis.
10. Under Reference Lines, do the following:
 - Next to **Value**, enter 500.
 - Click **Add**.
 - Next to **Value**, enter 560.
 - Click **Add**.
11. Click **OK**.
12. Click **Done**.

Note: The plot in Figure 5.51 is vertically resized.

Figure 5.51 Design Regions for Three Levels of Catalyst

Six of the settings for Catalyst A fall on the edges of the design region. All Temperature setting for Catalyst B are above the 400 degree constraint and all Temperature settings for Catalyst C are below the 650 degree constraint.

Mixture Experiments

Both the Custom Design and Mixture Design platforms construct designs for situations where all of your factors are ingredients in a mixture. However, mixture experiments can involve non-mixture process variables, or *process factors*. In addition to mixture only designs, the Custom Design platform can construct designs to accommodate both mixture and process factors. The Custom Design platform also allows the mixture components to sum to any positive number. See the Mixture description in “[Factor Types](#)” on page 89 in the “Custom Designs” chapter.

- “[Mixture Design with Nonmixture Factors](#)”
- “[Mixture of Mixtures Design](#)”

Mixture Design with Nonmixture Factors

In the following example from Atkinson and Donev (1992), you create a design for an experiment involving both mixture factors and process factors. The design is an 18-run design that is balanced with respect to the levels of a categorical factor. The design enables you to fit a full response surface. You use Design Evaluation plots and results to examine the relative prediction variance of the design.

The response and factors involved in the design are as follows:

- The response is **Damping**, which measures the electromagnetic damping of an acrylonitrile powder.
- The three mixture ingredients are:
 - **CuSO₄** (copper sulphate), ranging from 0.2 to 0.8
 - **Na₂S₂O₃** (sodium thiosulphate), ranging from 0.2 to 0.8
 - **Glyoxal** (glyoxal), ranging from 0 to 0.6
- The nonmixture environmental factor of interest is **Wavelength** (the wavelength of an electromagnetic wave) at three levels denoted L1, L2, and L3.

Wavelength is a continuous variable. However, the researchers were interested only in predictions at three specific wavelengths. For this reason, you treat Wavelength as a categorical factor with three levels.

This example contains the following steps:

- [“Create the Design”](#) on page 177
- [“Evaluate the Design”](#) on page 179

Create the Design

1. Select **DOE > Custom Design**.
2. Double-click Y under Response Name and type **Damping**.
3. Click **Maximize** under Goal and change it to **None**.

The goal is set to None because the authors of the study do not mention how much damping is desirable. Select **Help > Sample Data Library** and open **Design Experiment/Donev Mixture Factors.jmp**.

4. Click the Custom Design red triangle and select **Load Factors**.

This loads the three mixture ingredients and the categorical environmental factor. Note that the bounds on the values of the three mixture factors are also loaded.

Figure 5.52 Responses Outline and Factors Outline

Responses

Add Response ▼ Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Damping <i>optional item</i>	None	NA	NA	NA

Factors

Add Factor ▼ Remove Add N Factors 1

Name	Role	Changes	Values
▲ CuSO4	Mixture	Easy	0.2 0.8
▲ Na2S2O3	Mixture	Easy	0.2 0.8
▲ Glyoxal	Mixture	Easy	0 0.6
▼ Wavelength	Categorical	Easy	L1 L2 L3

- In the Model outline, click **Interactions > 2nd**.

An informational JMP Alert window reminds you that JMP removes the main effect terms for non-mixture factors that interact with all the mixture factors. This means that the main effect of Wavelength is removed, but all two-way interactions of mixture factors with Wavelength are added.

- Click **OK** to dismiss the message.

The effects in the Model outline define a response surface model in the mixture ingredients along with the additive effect of the wavelength. See Scheffé (1958).

Figure 5.53 Model and Design Generation Outlines

Model

Main Effects Interactions ▼ Cross Powers ▼ Scheffe Cubic Remove Term

Name	Estimability
CuSO4	Necessary
Na2S2O3	Necessary
Glyoxal	Necessary
CuSO4*Na2S2O3	Necessary
CuSO4*Glyoxal	Necessary
CuSO4*Wavelength	Necessary
Na2S2O3*Glyoxal	Necessary
Na2S2O3*Wavelength	Necessary
Glyoxal*Wavelength	Necessary

Alias Terms

Design Generation

☐ Group runs into random blocks of size: 2

Number of Center Points: 0

Number of Replicate Runs: 0

Number of Runs:

☐ Minimum 12

☒ Default 18

☐ User Specified 18

Make Design

7. Leave the default number of runs at 18.

The choice of 18 runs allows six runs for each of the three levels of the wavelength factor.

Note: Setting the Random Seed in step 8 and Number of Starts in step 9 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

8. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
9. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 5, and click **OK**.
10. Click **Make Design**.

Figure 5.54 Design Outline Showing 18-Run Design

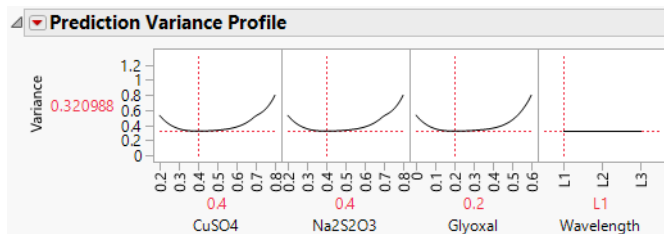
Design				
Run	CuSO4	Na2S2O3	Glyoxal	Wavelength
1	0.8	0.2	0	L2
2	0.5	0.2	0.3	L1
3	0.8	0.2	0	L3
4	0.2	0.8	0	L3
5	0.2	0.8	0	L1
6	0.2	0.2	0.6	L1
7	0.2	0.2	0.6	L3
8	0.2	0.5	0.3	L1
9	0.2	0.2	0.6	L2
10	0.5	0.2	0.3	L2
11	0.5	0.5	0	L3
12	0.8	0.2	0	L1
13	0.2	0.8	0	L2
14	0.5	0.2	0.3	L3
15	0.2	0.5	0.3	L2
16	0.5	0.5	0	L1
17	0.2	0.5	0.3	L3
18	0.5	0.5	0	L2

You can check that there are six runs for each level of Wavelength.

Evaluate the Design

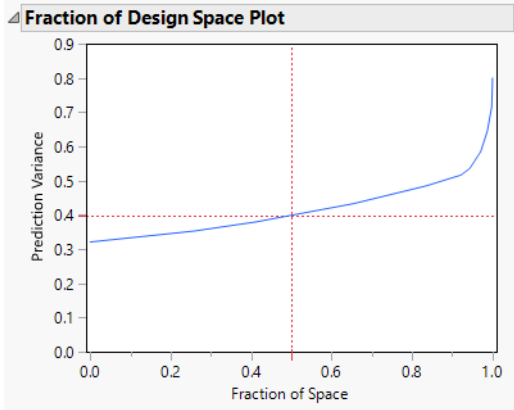
1. Open the **Design Evaluation > Prediction Variance Profile** outline.

Figure 5.55 Prediction Variance Profile for 18-Run Design



- Move the slider for Wavelength to verify that the relative prediction variance profiles for the mixture factors do not change across the levels of Wavelength. Move the slider for any one of the mixture factors. The sliders for the other two mixture factors adjust to make the mixture ingredients sum to one. Notice that the smallest relative prediction variances occur near the center settings for the mixture factors.
- Click the Prediction Variance Profile red triangle and select **Maximize Variance**.
Notice that the maximum relative prediction variance over the design space is 0.8 times the error variance.
 - Open the **Fraction of Design Space Plot** outline.

Figure 5.56 Fraction of Design Space Plot for 18-Run Design



- Over the entire design space, the relative prediction variance is below 0.8. The minimum relative prediction variance is about 0.32. As seen in Figure 5.55, the minimum occurs near the center settings for the mixture factors.
- Open the **Design Diagnostics** outline.

Figure 5.57 Design Diagnostics Outline for 18-Run Design

Design Diagnostics	
D Optimal Design	
D Efficiency	3.627407
G Efficiency	83.33333
A Efficiency	0.328715
Average Variance of Prediction	0.412864
Design Creation Time (seconds)	0.1

The design is optimal relative to the D-optimality criterion, even though its D-efficiency is very low (3.6%). Because mixture designs are far from orthogonal due to the mixture constraint, they typically have very low D-efficiencies. The Average (relative) Variance of Prediction is 0.412864. This is consistent with the Fraction of Design Space plot in Figure 5.56.

Mixture of Mixtures Design

In this example, construct a design for a mixture of mixtures situation.

Consider the ingredients that go into a cake. Dry ingredients include flour, sugar, and cocoa. Wet ingredients include milk, melted butter, and eggs. The wet and dry components of the cake are two mixtures that are first mixed separately and then blended together. Table 5.1 lists the factors and the ranges over which you vary them as part of your experiment.

Table 5.1 Dry and Wet Components and Experimental Ranges

Mixture	Ingredient	Lower and Upper Levels
Dry	Cocoa	0.1 - 0.2
	Sugar	0 - 0.15
	Flour	0.2 - 0.3
Wet	Butter	0.1 - 0.2
	Milk	0.25 - 0.35
	Eggs	0.05 - 0.20

The dry components (the mixture of cocoa, sugar, and flour) comprise 45% of the combined mixture. The wet components (butter, milk, and eggs) comprise 55%.

Note: With the constraints on the dry and wet ingredients, if you include all 6 factors in the mixture model, a model singularity exists.

The goal of your experiment is to optimize a Taste rating. Taste is rated on a scale of 1 to 10, with 10 representing the best taste.

You construct a 10-run design to fit a main effects model. Because of the constraint on the proportions of dry and wet ingredients, you need to include only five factors in the Model outline to avoid singularity. The choice of which factor not to include is arbitrary.

This example contains the following steps:

- [“Create the Design”](#) on page 181
- [“Analyze the Experimental Results”](#) on page 184

Create the Design

1. Select **DOE > Custom Design**.

2. Double-click Y under Response Name and type Taste.
Note that the default goal is Maximize. Because you want to maximize the Taste rating, do not change the goal.
3. Click under Lower Limit and type 0.
The least desirable rating is 0.
4. Click under Upper Limit and type 10.
The most desirable rating is 10.
5. Leave the area under Importance blank.
Because there is only one response, that response is given Importance 1 by default.
6. Select **Help > Sample Data Library** and open Design Experiment/Cake Factors.jmp.
7. Click the Custom Design red triangle and select **Load Factors**.

Figure 5.58 Completed Responses and Factors Outlines

Responses

Add Response ▼ Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Taste <i>optional item</i>	Maximize	0	10	.

Factors

Add Factor ▼ Remove Add N Factors 1

Name	Role	Changes	Values
▲ Cocoa	Mixture	Easy	0.1 0.2
▲ Sugar	Mixture	Easy	0 0.15
▲ Flour	Mixture	Easy	0.2 0.3
▲ Butter	Mixture	Easy	0.1 0.2
▲ Milk	Mixture	Easy	0.25 0.35
▲ Eggs	Mixture	Easy	0.05 0.2

Note that the factors are all mixture factors. The Values that define the range of settings for the experiment vary from factor to factor.

8. In the Define Factor Constraints outline, select **Specify Linear Constraints**.
9. In the Linear Constraints panel, click **Add** twice.
10. Enter the constraints shown in Figure 5.59.

Figure 5.59 Define Factor Constraints

Define Factor Constraints

☐ None
☒ Specify Linear Constraints
☐ Use Disallowed Combinations Filter
☐ Use Disallowed Combinations Script

Linear Constraints

1	Cocoa +	1	Sugar +	1	Flour +	0	Butter +	0	Milk +	0	Eggs	≤	0.45
-1	Cocoa +	-1	Sugar +	-1	Flour +	0	Butter +	0	Milk +	0	Eggs	≤	-0.45

Note that the second constraint is equivalent to specifying that the sum of Cocoa, Sugar, and Flour is greater than or equal to 0.45. The two constraints together imply that Cocoa, Sugar, and Flour comprise exactly 45% of the mixture, ensuring that the wet factors constitute the remaining 55%.

11. In the Model outline, select Eggs and click **Remove Term**.

Because of the equality constraint, a model containing all six effects would be singular.

12. Type 10 next to **User Specified**.

Your experiment requires baking 10 cakes.

Note: Setting the Random Seed in step 13 and Number of Starts in step 14 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

13. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
14. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 40, and click **OK**.
15. Click **Make Design**.

A JMP Alert informs you that your factor constraints include an equality constraint. This was what you intended, because the sum of the dry ingredient proportions is constrained to 45%.

16. Click **OK** to dismiss the JMP Alert.
17. Click **Make Table**.

Figure 5.60 Mixture of Mixtures Design

	Cocoa	Sugar	Flour	Butter	Milk	Eggs	Taste
1	0.1	0.15	0.2	0.2	0.3	0.05	•
2	0.1	0.15	0.2	0.1	0.35	0.1	•
3	0.1	0.15	0.2	0.1	0.25	0.2	•
4	0.1	0.05	0.3	0.2	0.25	0.1	•
5	0.15	0	0.3	0.1	0.35	0.1	•
6	0.1	0.05	0.3	0.15	0.35	0.05	•
7	0.2	0.05	0.2	0.1	0.25	0.2	•
8	0.2	0	0.25	0.2	0.25	0.1	•
9	0.1	0.05	0.3	0.1	0.25	0.2	•
10	0.2	0.05	0.2	0.15	0.35	0.05	•

The settings for the dry ingredients, Cocoa, Sugar, and Flour, sum to 45% of the mixture and the settings for the wet ingredients, Butter, Milk, and Eggs, sum to 55% of the mixture. The settings also conform to the upper and lower limits given in the Factors outline.

Analyze the Experimental Results

The Cake Data.jmp sample data table shows the results of the experiment. The design table contains a Model script that opens a Fit Model window showing the five main effects specified in the DOE window's Model outline. Notice that the main effect of Eggs is not included in the Model outline for this design. The level of Eggs can be determined by the settings of the other 5 factors and the constraint on the dry ingredients. This script was saved to the data table when it was created by Custom Design.

1. Open the Cake Data.jmp sample data table, located in the Design Experiment folder.
2. In the Table panel of the design table, click the green triangle next to the **Model** script.

The main effect due to Eggs is not included. To avoid model singularities, the Eggs effect was excluded from the Model outline in the Custom Design window. The five effects are designated as Response Surface and Mixture effects.

3. Click **Run**.

A JMP Alert appears, notifying you that the Profiler cannot be shown because of the mixture of mixtures constraint on the subsets of factors.

4. Click **OK** to dismiss the JMP Alert.

The Parameter Estimates report indicates that Sugar, Flour, and Butter are significant at the 0.05 level.

Figure 5.61 Parameter Estimates Report

Parameter Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
(Cocoa-0.1)/0.35	5.1056115	2.360215	2.16	0.0829
Sugar/0.35	16.873989	1.675197	10.07	0.0002*
(Flour-0.2)/0.35	15.750757	2.207353	7.14	0.0008*
(Butter-0.1)/0.35	5.3740398	2.654689	2.02	0.0988
(Milk-0.25)/0.35	1.5973046	2.435648	0.66	0.5409

Experiments with Covariates

Sometimes measurements on the experimental units that are intended for an experiment are available. These measurements might affect the experimental response. It is useful to include these variables, called *covariates*, as design factors. Although you cannot directly control these values, you can ensure that the levels of the other design factors are chosen to yield the most precise estimates of all the effects.

The Custom Design platform constructs a design that selects covariate values in an optimal fashion. Covariate values are selected from an existing data table that provides covariate information about the potential experimental units. You can specify a number of runs that is smaller than the number of experimental units listed in your data table. You can also specify covariates that are hard-to-change. When you make your design and the design has fewer runs than the number of rows in the covariate table, the design table includes a Covariate Row Index column. This column indicates the row from the covariate table that corresponds to each experimental run.

Note: The number of rows in the covariate data table where covariate factors have nonmissing values must be greater than or equal to the specified Number of Runs.

- “Design with Fixed Covariates”
- “Design with Hard-to-Change Covariates”
- “Design with a Linear Time Trend”

Design with Fixed Covariates

In this example, you are interested in modeling the Shrinkage of parts produced by an injection molding process. The Thermoplastic.jmp sample data table in the Design Experiment folder lists 25 batches of raw (thermoplastic) material for potential use in your study. For each batch, material was removed to obtain measurements of Specific Gravity and Tensile Strength. A third covariate, Supplier, is also available.

You want to study the effects of three controllable factors, Temperature (mold temperature), Speed (screw speed), and Time (hold time), on Shrinkage. But you also want to study the effects of the covariates: Specific Gravity, Tensile Strength, and Supplier. Your resources allow for 12 runs.

Create the Design

1. Select **Help > Sample Data Library** and open Design Experiment/Thermoplastic.jmp.
2. Select **DOE > Custom Design**.
3. Double-click Y under Response Name and type Shrinkage.
4. Click **Maximize** under Goal and change it to **Minimize**.
5. Click **Add Factor** and select **Covariate**.
6. Select Specific Gravity, Tensile Strength, and Supplier from the list and click **OK**.

These are covariates and cannot be controlled.

7. Type 3 next to **Add N Factors**.
8. Click **Add Factor > Continuous**.
9. Rename the three continuous factors Temperature, Speed, and Time.

These are factors that can be controlled.

Figure 5.62 Responses and Factors Outlines

The screenshot shows the Minitab Custom Design dialog box with two main sections: Responses and Factors.

Responses Section:

- Buttons: Add Response, Remove, Number of Responses...
- Table:

Response Name	Goal	Lower Limit	Upper Limit	Importance
Shrinkage	Minimize	.	.	.

Factors Section:

- Buttons: Add Factor, Remove, Add N Factors (set to 1)
- Table:

Name	Role	Changes	Values
Specific Gravity	Covariate	Easy	1.004
Tensile Strength	Covariate	Easy	6.42
Supplier	Covariate	Easy	A B C
Temperature	Continuous	Easy	-1 1
Speed	Continuous	Easy	-1 1
Time	Continuous	Easy	-1 1

10. Click **Continue**.

The Number of Runs shows the number of rows with covariate values available. You have 25 batches with measured covariates.

11. Type 12 next to **Number of Runs**.

Note: Setting the Random Seed in step 12 and Number of Starts in step 13 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

12. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 84951, and click **OK**.
13. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 40, and click **OK**.
14. Click **Make Design**.

Figure 5.63 Twelve-Run Optimal Design

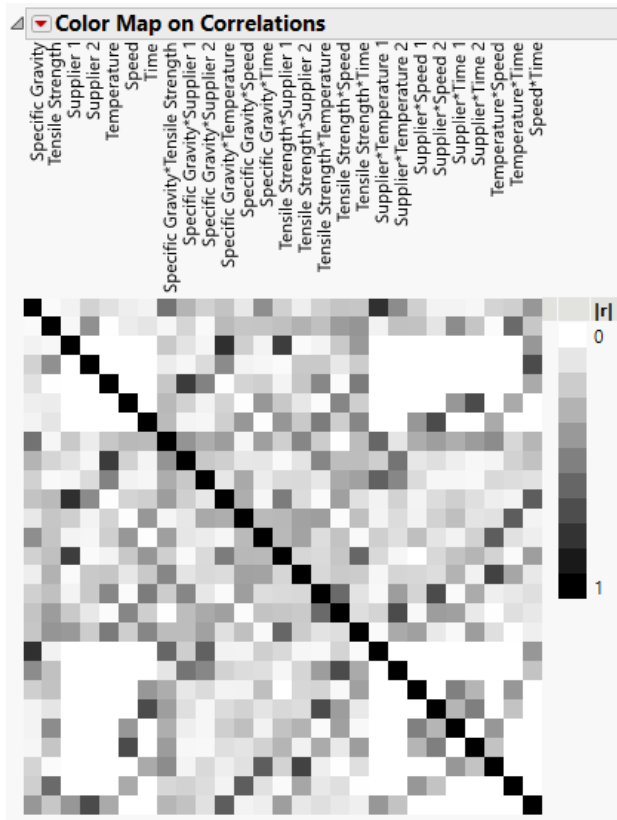
Design						
Run	Specific Gravity	Tensile Strength	Supplier	Temperature	Speed	Time
1	1.107	8.07	B	1	1	1
2	1.144	6.82	B	1	-1	-1
3	1.004	6.52	C	1	1	-1
4	1.015	7.46	A	-1	1	-1
5	1.129	6.97	A	1	-1	-1
6	1.139	6.94	C	-1	1	-1
7	1.004	7.12	C	1	-1	1
8	1.067	6.7	B	-1	1	1
9	1.047	6.42	A	-1	-1	1
10	1.03	7.54	B	-1	-1	-1
11	1.094	6.48	A	1	1	1
12	1.113	6.89	C	-1	-1	1

This design is D-optimal, given the potential covariate values. It selects the best sets of covariate values and the best settings for the three controllable factors.

Evaluate the Design

1. Open the **Design Evaluation > Color Map On Correlations** outline.

Figure 5.64 Color Map on Correlations



The seven terms corresponding to main effects appear in the upper left corner of the color map. Notice that these seven terms are close to orthogonal. The largest absolute correlation is between Tensile Strength and Supplier 2. This absolute correlation of about 0.43 is a consequence of the available covariate values. (Figure 5.64 uses JMP default colors.)

Design with Hard-to-Change Covariates

In this example, you construct a design for developing a running shoe for serious runners that has good wear (Wear) properties. Your experimental factors include the following:

- sole thickness (Thickness)
- amount of gel cushioning (Gel)
- outsole material (Outsole)
- midsole material (Midsole)

To obtain generalizable conclusions, you need to test your shoes on a broad base of serious runners. To accommodate your experimental budget, each runner must test several experimental combinations.

Your company has collected data on 100 suitable runners willing to participate in your study. The concomitant variables (*covariates*) measured on these runners are average daily miles run (Miles), weight (Weight), and the foot's strike point (Strike Point).

Create your design as follows:

1. Select **Help > Sample Data Library** and open Design Experiment/Runners Covariates.jmp.
2. Select **DOE > Custom Design**.
3. Double-click Y under Response Name and type Wear.
4. Click **Maximize** under Goal and change it to **Minimize**.
5. Click **Add Factor** and select **Covariate**.
6. Select Miles, Weight, and Strike Point from the list and click **OK**.

These are the hard-to-change covariates associated with the runners.

7. For one of the factors Miles, Weight, and Strike Point, under Changes, click **Easy** and change it to **Hard**.

Note that Changes for all three covariates change to Hard.

To add the remaining factors manually, follow step 8 through step 16. Or, to load factors from a saved table, select **Load Factors** from the Custom Design red triangle. Open the Runners Factors.jmp sample data table, located in the Design Experiment folder. If you select **Load Factors**, skip step 8 through step 16.

8. Type 2 next to **Add N Factors**.
9. Click **Add Factor > Continuous**.
10. Rename the two factors Thickness and Gel.
11. Change the **Values** for Thickness to 5 and 20.
12. Change the **Values** for Gel to 1 and 10.
13. Type 2 next to **Add N Factors**.
14. Click **Add Factor > Categorical > 3 Level**.
15. Rename the two factors Outsole and Midsole.

Keep the default **Values** for these factors.

Figure 5.65 Responses and Factors Outlines

Responses

Add Response

Remove

Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Wear	Minimize	.	.	.
optional item				

Factors

Add Factor

Remove

Add N Factors

1

Name	Role	Changes	Values
Miles	Covariate	Hard	614.8
Weight	Covariate	Hard	138152
Strike Point	Covariate	Hard	ForefootHeelMidfoot
Thickness	Continuous	Easy	520
Gel	Continuous	Easy	110
Outsole	Categorical	Easy	L1L2L3
Midsole	Categorical	Easy	L1L2L3

16. Click **Continue**.

17. Select **Interactions > 2nd**.
The specified model fits all two-factor interactions, including covariate by experimental factor interactions.

18. Set the Number of Whole Plots, or runners, to 32 (if it is not already set to that number).

19. Type 64 next to User Specified under Number of Runs (if it is not already set to that number).

Note: Setting the Random Seed in step 20 and Number of Starts in step 21 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

20. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.

21. (Optional) Click the Custom Design red triangle, select **Number of Starts** and set it to 1 (if it is not already set to that number). Click **OK**.

22. Click **Make Design**.

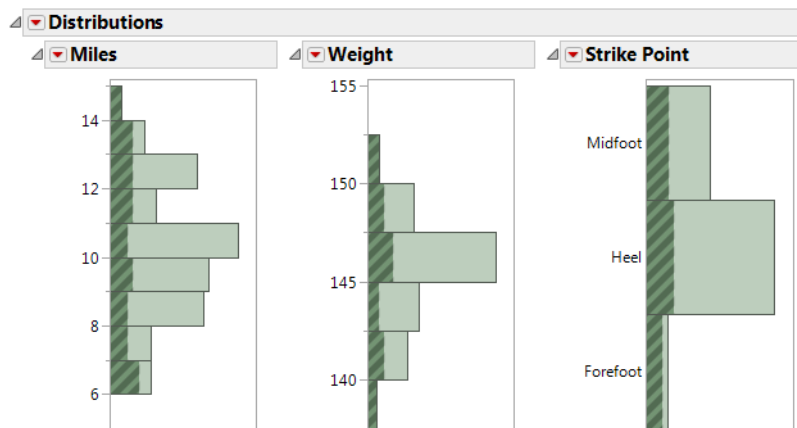
Figure 5.66 First 20 Runs of Design for Hard-to-Change Covariates

Design									
Run	Whole Plots	Miles	Weight	Strike Point	Thickness	Gel	Outsole	Midsole	
1	1	7.6	144	Midfoot	5	10	L3	L2	
2	1	7.6	144	Midfoot	20	1	L2	L1	
3	2	7.6	145	Midfoot	20	10	L3	L3	
4	2	7.6	145	Midfoot	5	1	L1	L1	
5	3	14.8	141	Midfoot	20	1	L1	L3	
6	3	14.8	141	Midfoot	5	10	L2	L1	
7	4	6.4	146	Heel	20	10	L1	L2	
8	4	6.4	146	Heel	5	1	L2	L3	
9	5	11.2	138	Heel	5	1	L2	L3	
10	5	11.2	138	Heel	20	10	L3	L2	
11	6	11.2	144	Forefoot	20	1	L3	L3	
12	6	11.2	144	Forefoot	5	10	L2	L1	
13	7	12.8	139	Heel	5	10	L3	L1	
14	7	12.8	139	Heel	20	1	L1	L2	
15	8	14	139	Midfoot	5	1	L1	L2	
16	8	14	139	Midfoot	20	10	L2	L3	
17	9	6	150	Heel	20	10	L2	L2	
18	9	6	150	Heel	5	1	L3	L3	
19	10	10.8	143	Forefoot	20	1	L3	L1	
20	10	10.8	143	Forefoot	5	10	L2	L2	

Of the 100 runners, 32 are selected based on their covariate values. The rows corresponding to the selected runners are selected in the RunnersCovariates.jmp sample data table. Settings of the experimental factors Thickness, Gel, Insole, and Outsole are determined so that the design is optimal for the model described in the Model Outline.

23. With the RunnersCovariates.jmp sample data table as the active table, select **Analyze > Distribution**.
24. Select all three columns as **Y, Columns**.
25. Check **Histograms Only**.
26. Click **OK**.

Figure 5.67 Histograms for 100 Runners with Selected Runner Data Shaded



The histograms illustrate how the design selected runners to balance runs across the covariates. The distribution of selected runners is similar to the overall distribution of all runners for each covariate.

Design with a Linear Time Trend

Often, experiments conducted in a time sequence experience a linear drift in the response. If you randomize the order of the runs, then the drift's effect does not generally bias the estimated factor effects. However, by accounting for the drift, you can reduce the variance of those effects.

Suppose that there is reason to suspect a strong linear trend in the response over time independent of the factor changes. Then you can construct a design that includes a linear covariate to account for the trend. The resulting design is optimal, given this trend covariate.

In this example, you design an experiment for 7 factors. You construct a 16-run design that is robust to linear trend.

1. Select **File > New > Data Table**.
2. Right-click Column1 and select **Column Info**.
3. Change the column name to Run Order.
4. From the list of **Initialize Data** options, select **Sequence Data**.
5. Type 16 next to **To**.
6. Click **OK**.

Consecutive integers from 1 to 16 have been entered in the data table.

7. Select **DOE > Custom Design**.
8. Click **Add Factor > Covariate**.
9. Select Run Order and click **OK**.
10. Type 7 next to **Add N Factors**.
11. Click **Add Factor > Continuous**.

Figure 5.68 Responses and Factors Outlines

The screenshot shows the Minitab 'Custom Design' dialog box. The 'Responses' section is expanded, showing a table with one response named 'Y' with a goal to 'Maximize'. The 'Factors' section is also expanded, showing a list of factors: 'Run Order' (Covariate), 'X2' through 'X8' (Continuous). Each factor has a 'Changes' column (Easy) and a 'Values' column (1, -1, 1, -1, 1, -1, 1, -1).

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.

Name	Role	Changes	Values
Run Order	Covariate	Easy	1
X2	Continuous	Easy	-1
X3	Continuous	Easy	-1
X4	Continuous	Easy	-1
X5	Continuous	Easy	-1
X6	Continuous	Easy	-1
X7	Continuous	Easy	-1
X8	Continuous	Easy	-1

12. Click **Continue**.
13. Open the **Alias Terms** outline.
14. Select all of the effects in the list and click **Remove Term**.

This omits the interaction effects from the correlation color map, leaving only the main effects.

Note: Setting the Random Seed in step 15 and Number of Starts in step 16 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

15. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 1084680980, and click **OK**.
16. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 30000, and click **OK**.
17. Click **Make Design**.

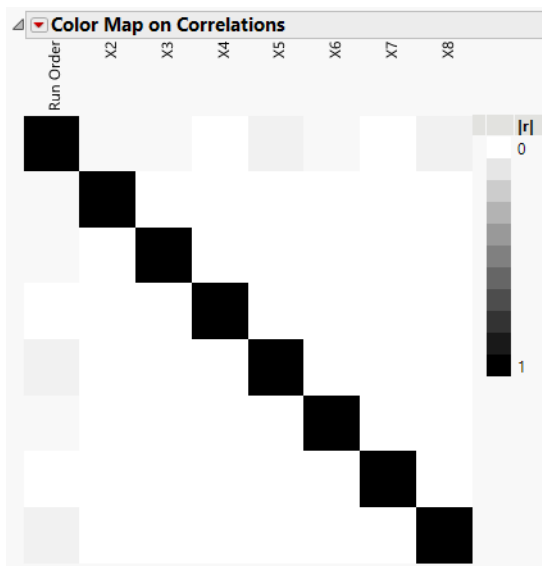
A progress bar displays elapsed time and D Efficiency.

Figure 5.69 Design Outline

Design								
Run	Run Order	X2	X3	X4	X5	X6	X7	X8
1	1	1	1	-1	-1	-1	-1	-1
2	2	1	-1	1	1	-1	-1	-1
3	3	1	-1	-1	1	1	1	1
4	4	-1	-1	1	-1	1	-1	1
5	5	-1	1	1	1	-1	1	-1
6	6	-1	1	-1	1	1	-1	1
7	7	-1	1	1	-1	-1	1	1
8	8	-1	1	-1	1	1	1	-1
9	9	1	1	1	-1	1	1	1
10	10	-1	-1	-1	-1	1	-1	-1
11	11	1	-1	1	-1	1	1	-1
12	12	1	-1	-1	1	-1	1	1
13	13	-1	-1	-1	-1	-1	1	-1
14	14	-1	-1	1	1	-1	-1	1
15	15	1	1	-1	-1	-1	-1	1
16	16	1	1	1	1	1	-1	-1

18. Open the **Design Evaluation > Color Map On Correlations** outline.

Figure 5.70 Color Map Showing Absolute Correlations with Run Order



The Color Map shows the following:

- The seven continuous factors, X2 through X8, are orthogonal to each other.
- Run Order, the linear time trend variable, has extremely low absolute correlation with X2 through X8.

19. Open the **Design Evaluation > Estimation Efficiency** outline.

Figure 5.71 Estimation Efficiency Outline

Estimation Efficiency		
Parameter	Fractional Increase in CI Length	Relative Std Error of Parameters
Intercept	0	0.25
Run Order	0.634	0.408
X2	4e-4	0.25
X3	4e-4	0.25
X4	0	0.25
X5	0.001	0.25
X6	4e-4	0.25
X7	0	0.25
X8	0.001	0.25

The small absolute correlations of Run Order with X2 through X8 result in very small increases in confidence interval lengths, relative to an ideal orthogonal design. The increases in the lengths of confidence intervals for X2 through X8 are all less than 0.1%.

In this example, the run order factor is nearly orthogonal to the factor effects. In some cases, your design might have more substantial correlations between the run order factor and other factors. Even in such a situation, including the run order as a factor accounts for any linear trend effect. Including the run order also allows for more precise estimation of the other factor effects.

Experiments with Randomization Restrictions

The Custom Design platform constructs split-plot, split-split-plot, and two-way split-plot (strip-plot) designs that are D-optimal or I-optimal. For more information about constructing these designs, see Goos (2002).

- [“Split-Plot Experiment”](#)
- [“Two-Way Split-Plot Experiment”](#)

Split-Plot Experiment

Split-plot designs originated in agriculture, but are commonplace in manufacturing and engineering studies. In a split-plot experiment, hard-to-change factors are reset only between one whole plot and the next whole plot. The whole plot is divided into subplots, and the levels of the easy-to-change factors are randomly assigned to each subplot.

The example in this section is adapted from Kowalski, Cornell, and Vining (2002). You are interested in the effects of five factors on the thickness of vinyl that is used to make automobile seat covers. The response and factors in the experiment are described below:

- The response is the thickness of the vinyl that is produced. You want to maximize thickness. A lower limit for thickness values is 10.

- The whole plot factors are the rate of extrusion (extrusion rate) and the temperature (temperature) of drying. These are process variables and are hard to change.
- The subplot factors are three plasticizers whose proportions (m1, m2, and m3) sum to one. These factors are mixture components.

Your experimental budget allows for running 7 settings of these whole plot factors. For each whole plot, you can conduct 4 runs of the subplot factors. This gives you a total of 28 runs.

Create the Design

1. Select **DOE > Custom Design**.
2. Double-click Y under Response Name and type thickness.
Keep the default goal set to Maximize.
3. Enter a Lower Limit of 10.

To add factors manually, follow step 4 through step 11. Or, to load factors from a saved table, select **Load Factors** from the Custom Design red triangle. Open the Vinyl Factors.jmp sample data table, located in the Design Experiment folder. If you select **Load Factors**, skip step 4 through step 11.

4. Type 2 next to **Add N Factors**.
5. Click **Add Factor > Continuous**.
6. Rename these factors extrusion rate and temperature.
Keep the default Values of -1 and 1 for these two factors.
7. Click **Easy** and select **Hard** for both extrusion rate and temperature.
This defines extrusion rate and temperature to be whole plot factors.
8. Type 3 next to **Add N Factors**.
9. Click **Add Factor > Mixture**.
10. Rename the three mixture factors m1, m2, and m3.
Keep the default Values of 0 and 1 for those three factors.

Figure 5.72 Responses and Factors Outlines

The screenshot shows the 'Custom Design' dialog box with two main sections: 'Responses' and 'Factors'.

Responses Section:

- Buttons: Add Response, Remove, Number of Responses...
- Table:

Response Name	Goal	Lower Limit	Upper Limit	Importance
thickness	Maximize	10	.	.

Factors Section:

- Buttons: Add Factor, Remove, Add N Factors (set to 1)
- Table:

Name	Role	Changes	Values
extrusion rate	Continuous	Hard	-1
temperature	Continuous	Hard	-1
m1	Mixture	Easy	0
m2	Mixture	Easy	0
m3	Mixture	Easy	0

11. Click **Continue**.
12. Click **Interactions > 2nd**.
13. Click **OK** to dismiss the informative message.

Note that 3 is the default value for the Number of Whole Plots.

14. Type 7 next to **Number of Whole Plots**.
15. Type 28 next to **User Specified**.

Note: Setting the Random Seed in step 16 and Number of Starts in step 17 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

16. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
17. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 5, and click **OK**.
18. Click **Make Design**.

Figure 5.73 Design Outline

Design						
Run	Whole Plots	extrusion rate	temperature	m1	m2	m3
1	1	-1	-1	0	0	1
2	1	-1	-1	0.6	0	0.4
3	1	-1	-1	0	1	3.55e-15
4	1	-1	-1	0.5	0.5	0
5	2	-1	1	0.4	0	0.6
6	2	-1	1	0	1	0
7	2	-1	1	1	0	0
8	2	-1	1	0	0.5	0.5
9	3	1	1	0	1	0
10	3	1	1	0	0	1
11	3	1	1	0.5	0	0.5
12	3	1	1	1	0	0
13	4	-1	1	0	0	1
14	4	-1	1	0.5	0.5	0
15	4	-1	1	1	0	0
16	4	-1	1	0	1	0
17	5	1	-1	0	0	1
18	5	1	-1	0	1	0
19	5	1	-1	0.5	0	0.5
20	5	1	-1	1	0	0
21	6	-1	-1	0	0.5	0.5
22	6	-1	-1	1	0	0
23	6	-1	-1	0	0	1
24	6	-1	-1	0	1	0
25	7	1	1	0	0	1
26	7	1	1	1	0	0
27	7	1	1	0	0.560358	0.439642
28	7	1	1	0.5	0.5	0

Note that the whole plot factors, extrusion rate and temperature, are reset seven times in accordance with the levels of the factor Whole Plots. Within each level of Whole Plots, the settings for the mixture ingredients, m1, m2, and m3, are assigned at random.

Analyze the Results

The Vinyl Data.jmp sample data table contains experimental results using a design created in a previous version of JMP.

1. Select **Help > Sample Data Library** and open Design Experiment/Vinyl Data.jmp.
This sample data table contains 28 runs and response values. The design settings in the table that you created using the Custom Design platform might differ from those used in the Vinyl Data.jmp design.
2. In the Table panel, click the green triangle next to the **Model** script.

Figure 5.74 Fit Model Window

Notice the following in the Fit Model window:

- The factor **Whole Plots** has the Attribute called **Random Effects (&Random)**. This specifies that the levels of **Whole Plots** are random realizations. They have an associated error term.
- The analysis method is **REML (Recommended)**. This method is specified precisely because the model contains a random effect. For more information about REML models, see the Standard Least Squares Report and Options chapter in *Fitting Linear Models*.

Tip: In the Fit Model window, JMP Pro users can change the Personality to Mixed Model.

3. Click **Run**.

Although factors at both stages might be equally hard-to-change, in order to distinguish these factors, JMP denotes the first stage factors as *very-hard-to-change* and the second-stage factors as *hard-to-change*. Additional factors applied to experimental units after the second processing stage are considered *easy-to-change*.

Scenario for a Two-Way Split-Plot Design

This example is based on an experiment to improve the open circuit voltage (OCV) in battery cells (Vivacqua and Bisgaard, 2004). You need to minimize the OCV in order to keep the cells from discharging on their own.

Battery cells move through two stages of processing:

- First stage: A continuous assembly process where batteries are processed in batches of 2000.
- Second stage: A curing process with a 5-day cycle time in a chamber that can accommodate 4000 batteries.

You want to study six two-level continuous factors:

- Four factors (A1, A2, A3, and A4) are applied to the assembly process. You can run 16 trials for the first-stage factors.
- Two factors (C5 and C6) are applied to the curing process. Because curing requires a 5-day cycle time, you can run only 6 cycles (30 days) for the second-stage factors. Using six curing cycles gives you partial replication of the curing settings, enabling you to test for curing effects.

Both the first- and second-stage factors are hard-to-change, suggesting two split-plots. However, the batches of 2,000 batteries from the first-stage experiment can be divided into sub-batches of 500 batteries each. Eight of these sub-batches can be randomly selected and processed simultaneously in the curing chamber.

The experiment has 48 experimental units. Note that the first- and second-stage factors are *crossed*.

Create the Design

To design a two-way split-plot experiment:

1. Select **DOE > Custom Design**.
2. Double-click Y under Response Name and type OCV.
3. Under Goal, click **Maximize** and select **Minimize**.
4. To add factors manually, follow step 5 through step 10. Or, to load factors from a saved table:
 - a. Click the Custom Design red triangle and select **Load Factors**.

- b. Open the Battery Factors.jmp sample data table, located in the Design Experiment folder.
 - c. Proceed to step 11.
5. Type 6 next to **Add N Factors**.
6. Click **Add Factor > Continuous**.
7. Rename the factors A1, A2, A3, A4, C1, and C2.
Keep the default Values of -1 and 1 for these factors.
8. For each of the factors A1, A2, A3, and A4, under Changes, click **Easy** and change it to **Very Hard**.
To distinguish between the first- and second-stage factors, you designate the Changes for the first-stage factors as Very Hard, and the Changes for the second-stage factors as Hard.
9. For each of the factors C1 and C2, under Changes, click **Easy** and change it to **Hard**.

Figure 5.76 Responses and Factors Outlines

The screenshot shows the JMP Design Experiment interface with two main sections: Responses and Factors.

Responses Section:

- Buttons: Add Response, Remove, Number of Responses...
- Table:

Response Name	Goal	Lower Limit	Upper Limit	Importance
OCV	Minimize	.	.	.

Factors Section:

- Buttons: Add Factor, Remove, Add N Factors, 1
- Table:

Name	Role	Changes	Values
A1	Continuous	Very Hard	-1, 1
A2	Continuous	Very Hard	-1, 1
A3	Continuous	Very Hard	-1, 1
A4	Continuous	Very Hard	-1, 1
C1	Continuous	Hard	-1, 1
C2	Continuous	Hard	-1, 1

10. Click **Continue**.
11. Select **Interactions > 2nd** in the Model outline.
12. In the Design Generation outline, select the option **Hard to change factors can vary independently of Very Hard to change factors**.
See Figure 5.77. Checking this option creates a two-way split-plot design. If this option is not checked, the design is treated as a split-split-plot design, with nesting of factors at the two levels.
13. Type 16 as the **Number of Whole Plots**.
This is the number of trials that you can run for the first-stage factors.
14. Type 6 as the **Number of Subplots**.

This is the number of trials that you can run for the second-stage factors.

15. Under Number of Runs, type 48 next to **User Specified**.

This is the total number of experimental units.

Figure 5.77 Design Generation Outline

Design Generation

☒ Hard to change factors can vary independently of Very Hard to change factors.

Number of Whole Plots: 16

Number of Subplots: 6

Number of Runs:

☐ Minimum 22

☐ Default 30

☒ User Specified 48

Make Design

Note: Setting the Random Seed in step 16 and Number of Starts in step 17 reproduces the exact results shown in this example. In constructing a design on your own, these steps are not necessary.

16. (Optional) Click the Custom Design red triangle, select **Set Random Seed**, type 1866762673, and click **OK**.
17. (Optional) Click the Custom Design red triangle, select **Number of Starts**, type 21, and click **OK**.
18. Click **Make Design**.
19. Click **Make Table**.

Figure 5.78 Partial View of Design Table

	Whole Plots	Subplots	A1	A2	A3	A4	C1	C2	OCV
1	1	1	-1	1	-1	1	-1	1	•
2	1	2	-1	1	-1	1	1	1	•
3	1	3	-1	1	-1	1	1	-1	•
4	2	4	1	1	-1	1	1	1	•
5	2	5	1	1	-1	1	-1	1	•
6	2	6	1	1	-1	1	-1	-1	•
7	3	1	-1	-1	-1	1	-1	1	•
8	3	2	-1	-1	-1	1	1	1	•
9	3	3	-1	-1	-1	1	1	-1	•
10	4	4	-1	-1	-1	-1	1	1	•
11	4	5	-1	-1	-1	-1	-1	1	•
12	4	6	-1	-1	-1	-1	-1	-1	•
13	5	1	1	1	-1	-1	-1	1	•
14	5	2	1	1	-1	-1	1	1	•
15	5	3	1	1	-1	-1	1	-1	•
16	6	4	1	-1	-1	-1	1	1	•

The design table shows 16 levels for Whole Plots. For each level of Whole Plots, the settings of the four assembly factors are constant. From each level of Whole Plots, three batches of 500 batteries (Subplots) are randomly assigned to settings of the curing factors. Two sets of curing conditions are replicated ($C1 = -1, C2 = 1$ and $C1 = 1, C2 = 1$). To see this, select columns C1 and C2, right-click in the header area, and select **Sort > Ascending**.

Analyze the Results

The Battery Data.jmp sample data table contains experimental results for the design that you generated.

1. Select **Help > Sample Data Library** and open Design Experiment/Battery Data.jmp.
2. In the Table panel, click the green triangle next to the **Model** script.

Notice the following in the Fit Model window:

- The factor Whole Plots has the Attribute called Random Effects (**&Random**). This specifies that the levels of Whole Plots are random realizations. They have an associated error term.
- The factor Subplots also has the Random Effects Attribute (**&Random**).
- The analysis Method is **REML (Recommended)**. This method is specified precisely because the model contains random effects. For more information about REML models, see the Standard Least Squares Report and Options chapter in *Fitting Linear Models*.

Tip: In the Fit Model window, JMP Pro users can change the Personality to Mixed Model.

3. Check the option to **Keep dialog open**.
4. Click **Run**.

Figure 5.79 Report for Full Model

Parameter Estimates						
Term	Estimate	Std Error	DFDen	t Ratio	Prob> t	
Intercept	38.183196	1.673404	2.075	22.82	0.0016*	
A1	-4.657571	1.093074	2.916	-4.26	0.0251*	
A2	0.147742	1.093074	2.916	0.14	0.9013	
A3	0.9486795	1.093074	2.916	0.87	0.4510	
A4	4.0716925	1.190175	3.214	3.42	0.0376*	
C1	-1.827487	1.606207	0.868	-1.14	0.4804	
C2	-15.73288	1.35446	0.719	-11.62	0.1065	
A1*A2	0.1289012	1.080526	2.784	0.12	0.9131	
A1*A3	-1.114015	1.080526	2.784	-1.03	0.3837	
A1*A4	-1.92398	1.176342	3.04	-1.64	0.1993	
A1*C1	-3.271232	0.480531	20.65	-6.81	<.0001*	
A1*C2	3.737258	0.478635	20.09	7.81	<.0001*	
A2*A3	0.3768179	1.080526	2.784	0.35	0.7520	
A2*A4	1.4693535	1.176342	3.04	1.25	0.2992	
A2*C1	1.9072313	0.480531	20.65	3.97	0.0007*	
A2*C2	-0.161179	0.478635	20.09	-0.34	0.7398	
A3*A4	0.0764369	1.176342	3.04	0.06	0.9522	
A3*C1	0.1761619	0.480531	20.65	0.37	0.7176	
A3*C2	-0.445242	0.478635	20.09	-0.93	0.3633	
A4*C1	0.5306856	0.511896	20.84	1.04	0.3118	
A4*C2	1.3506119	0.502894	20.12	2.69	0.0142*	
C1*C2	-0.425326	1.606207	0.868	-0.26	0.8401	

Random Effect Predictions						

REML Variance Component Estimates						
Random Effect	Var Ratio	Component	Std Error	95% Lower	95% Upper	Pct of Total
Whole Plots	1.6095523	15.451509	15.876576	-15.66601	46.569027	45.968
Subplots	0.8919389	8.562507	16.314188	-23.41271	40.537728	25.473
Residual		9.5998804	3.0299487	5.6239301	19.985723	28.559
Total		33.613897	16.824615	15.325965	123.59174	100.000

-2 LogLikelihood = 227.810084

Note: Total is the sum of the positive variance components.

Total including negative estimates = 33.613897

The Parameter Estimates report indicates that four two-way interactions, A1*C1, A1*C2, A2*C1, and A4*C2, and two main effects, A1 and A4, are significant at the 0.05 level.

- In the Table panel of Battery Data.jmp, click the green triangle next to the **Reduced Model 1** script.

The script opens a Fit Model window where insignificant interactions have been removed. The remaining effects are all main effects and the four two-way interactions A1*C1, A1*C2, A2*C1, and A4*C2. You are reducing the model in a conservative fashion.

- Click **Run**.

Figure 5.80 Report for Preliminary Reduced Model

Parameter Estimates					
Term	Estimate	Std Error	DFDen	t Ratio	Prob> t
Intercept	38.249531	1.373127	7.723	27.86	<.0001*
A1	-4.723906	1.175023	8.5	-4.02	0.0034*
A2	0.0497917	1.165418	8.228	0.04	0.9669
A3	0.7560417	1.165418	8.228	0.65	0.5342
A4	4.2328205	1.207654	9.16	3.50	0.0065*
C1	-2.195626	0.866042	1.083	-2.54	0.2234
C2	-15.79922	0.840903	1.008	-18.79	0.0331*
A1*C1	-3.283225	0.452505	25.17	-7.26	<.0001*
A1*C2	3.8035938	0.4498	24.42	8.46	<.0001*
A2*C1	1.8797279	0.452505	25.17	4.15	0.0003*
A4*C2	1.3877006	0.482556	24.95	2.88	0.0081*

Notice that the main effect C2 is now significant at the 0.05 level ($\text{Prob}>|t| = 0.0331$)

7. In the Fit Model window, remove A3.

The main effect A3 is the only main effect that is not significant and not involved in a two-way interaction.

8. Click Run.

Figure 5.81 Report for Reduced Model

Parameter Estimates					
Term	Estimate	Std Error	DFDen	t Ratio	Prob> t
Intercept	38.249531	1.362184	6.639	28.08	<.0001*
A1	-4.723906	1.137313	9.131	-4.15	0.0024*
A2	0.0497917	1.127389	8.818	0.04	0.9658
A4	4.2131272	1.173862	9.921	3.59	0.0050*
C1	-2.166553	0.899138	1.037	-2.41	0.2432
C2	-15.79922	0.874257	0.963	-18.07	0.0389*
A1*C1	-3.282201	0.452234	25.24	-7.26	<.0001*
A1*C2	3.8035938	0.449735	24.43	8.46	<.0001*
A2*C1	1.8843473	0.452234	25.24	4.17	0.0003*
A4*C2	1.3851686	0.4823	24.91	2.87	0.0082*

Random Effect Predictions					
REML Variance Component Estimates					
Random Effect	Var Ratio	Var Component	Std Error	95% Lower	95% Upper
Whole Plots	2.0231563	17.459497	9.7154555	-1.582446	36.50144
Subplots	0.3473621	2.9976761	5.8732929	-8.513766	14.509119
Residual		8.6298312	2.4692519	5.2816799	16.588508
Total		29.087005	9.1941655	17.028461	60.633509

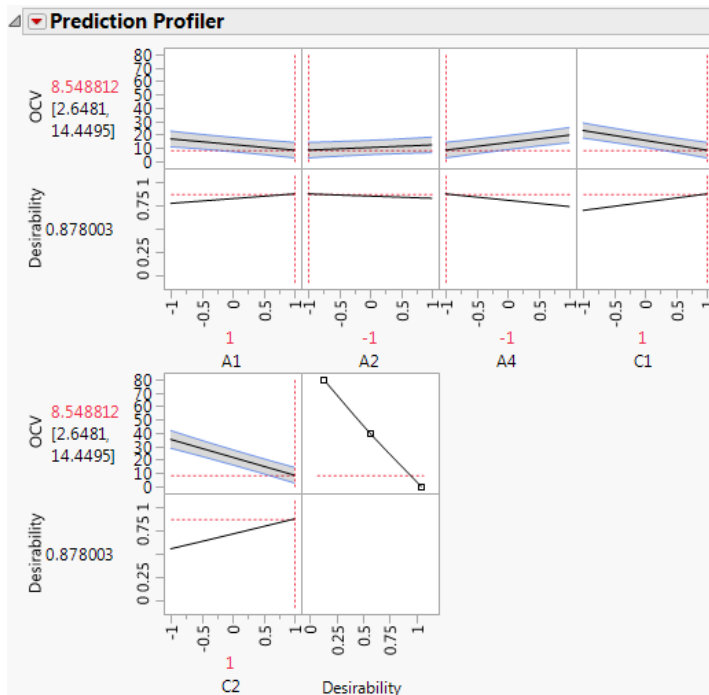
-2 LogLikelihood = 254.17861349
Note: Total is the sum of the positive variance components.
Total including negative estimates = 29.087005

The REML Variance Component Estimates report shows that the variance component associated with Whole Plots is about six times as large as the variance component for Subplots. This suggests that the assembly process is more variable than the curing process. Also, the within (Residual) error is larger than that for Subplots. Efforts to reduce variation should focus on the assembly process and on battery-to-battery differences.

9. Click the Response OCV red triangle and select **Factor Profiling > Profiler**.

10. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

Figure 5.82 Prediction Profiler with Settings That Minimize OCV



The profiler shows the five factors identified as active and settings that minimize OCV.

Experiments for Robust Process and Product Design

Often processes or products are impacted by noise factors. These are factors that are not easy or cost effective to control. Finding process or product settings of controllable factors that are least impacted by noise factors is desirable. Historically, Taguchi designs provided a method for experimenting in the presence of noise factors. Note that noise factors must be controlled during an experiment.

Alternatives to Taguchi designs include combined arrays and mixed resolution designs (Borror 2000). The mixed resolution designs are used to obtain designs that contain both control and noise factors. The goal of such a design is to find robust settings for control variables. The key is the ability to estimate specific effects. In particular, one is interested in estimating main effects of the control variables, main effects of the noise variables, control by control interactions, noise by control interactions, and quadratic terms for the control variables.

Use the custom designer to generate a design for studying robust process settings.

In this example, you are interested in finding optimal settings for four control settings in the presence of three noise factors.

1. Select **DOE > Custom Design**.
2. In the Factors outline, type 4 next to **Add N Factors**.
3. Click **Add Factor > Continuous**.
4. In the Factors outline, type 3 next to **Add N Factors**.
5. Click **Add Factor > Continuous**.
6. Double-click X5 and change it to Z1 to designate a noise factor. Repeat to change X6 and X7 to Z2 and Z3.

Figure 5.83 Factor Outline for Design with 4 Control and 3 Noise Factors

Factors				
Add Factor ▼	Remove	Add N Factors	1	
Name	Role	Changes	Values	
X1	Continuous	Easy	-1	1
X2	Continuous	Easy	-1	1
X3	Continuous	Easy	-1	1
X4	Continuous	Easy	-1	1
Z1	Continuous	Easy	-1	1
Z2	Continuous	Easy	-1	1
Z3	Continuous	Easy	-1	1

7. Click Continue.

Next you set up the model that for exploring optimal settings for the control factors that are least impacted by the noise factors.

8. The model contains all main effects by default.
9. Click **Interactions > 2nd** to add all two-way interactions.

This enters the control by control interactions, the control by noise interactions, and the noise by noise interactions. You do not want to estimate the interaction between noise factors.

10. Scroll to the bottom of the list, highlight Z1*Z2, Z1*Z3, and Z2*Z3 and click Remove Term. This removes the noise by noise interactions.
11. In the Factors outline, select X1, X2, X3, and X4.

12. In the Model outline, click Powers > 2nd to add the control variable quadratic terms.

Figure 5.84 Model Outline for Design with Noise Factors

Model	
Main Effects	Interactions ▼ RSM Cross Powers ▼ Remove Term
Name	Estimability
Intercept	Necessary
X1	Necessary
X2	Necessary
X3	Necessary
X4	Necessary
Z1	Necessary
Z2	Necessary
Z3	Necessary
X1*X2	Necessary
X1*X3	Necessary
X1*X4	Necessary
X1*Z1	Necessary
X1*Z2	Necessary
X1*Z3	Necessary
X2*X3	Necessary
X2*X4	Necessary
X2*Z1	Necessary
X2*Z2	Necessary
X2*Z3	Necessary
X3*X4	Necessary
X3*Z1	Necessary
X3*Z2	Necessary
X3*Z3	Necessary
X4*Z1	Necessary
X4*Z2	Necessary
X4*Z3	Necessary
X1*X1	Necessary
X2*X2	Necessary
X3*X3	Necessary
X4*X4	Necessary

13. Add two center points by entering 2 In the **Number of Center Points:** text box

14. Click **Make Design**.

The result is a 36 run design that enables you to evaluate and optimize your control factors in the presence of your three noise factors. The goal of your analysis is to determine which effects are important and then to select settings for the control factors that are least impacted by the noise factors.

Chapter 6

Augment Designs

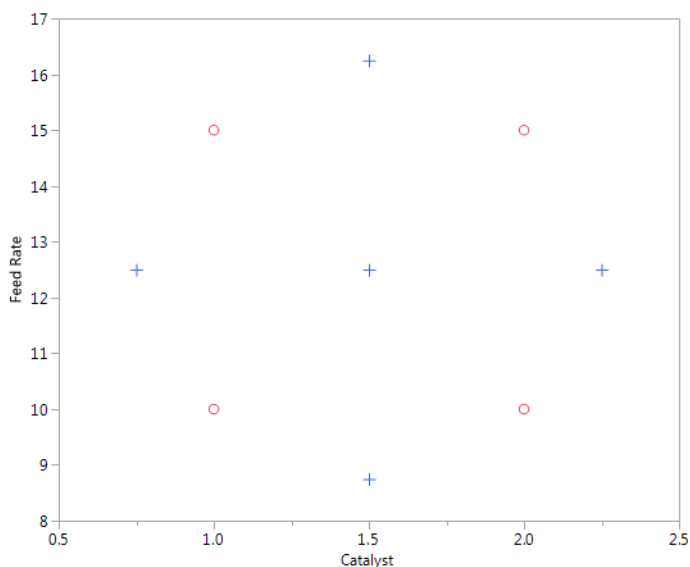
Use the Augment Design platform to add runs to an existing experimental design. Experimentation is often a sequential process and augmenting a designed experiment can help resolve ambiguities that result from a single design.

For an existing design table, the Augment Design platform constructs additional runs that optimize the overall design. You can add runs to accomplish the following objectives:

- Replicate the design a specified number of times.
- Add center points.
- Create a foldover design.
- Add axial points together with center points to transform a screening design to a response surface design.
- Add space filling points to a design.
- Add runs to the design in order to fit a specific model.

For more information about augmenting designs see Goos and Jones ([2011](#)).

Figure 6.1 Original Design Runs (+) with Augmented Runs (o)



Contents

Example Using the Augment Design Platform	213
Analyze the Augmented Design	215
Launch the Augment Design Platform	219
Augment Design Window	220
Factors	220
Define Factor Constraints	221
Augmentation Choices	224
Augment Design Platform Options	225
Additional Examples of Augmentation Choices	229
Replicate a Design	229
Add Center Points to a Design	231
Create a Foldover Design	231
Add Axial Points to a Design	232
Space Filling Augmentation	233
Statistical Details	235

Example Using the Augment Design Platform

This example demonstrates how to use the Augment Design platform to resolve ambiguities in the results of a screening design. In this study, a chemical engineer investigates the effects of five factors on the percent reaction of a chemical process. The original design was an 8-run design suitable for fitting main effects. This example augments the 8-run design such that all two-factor interactions can be estimated using a total of 16 runs.

1. Select **Help > Sample Data Library** and open Design Experiment/Reactor 8 Runs.jmp.
2. Select **DOE > Augment Design**.
3. Select Percent Reacted and click **Y, Response**.
4. Select Feed Rate, Catalyst, Stir Rate, Temperature, and Concentration, and click **X, Factor**.
5. Click **OK**.

Figure 6.2 Factors for the Reactor Example

Name	Role	Changes	Values
Feed Rate	Continuous	Easy	10 15
Catalyst	Continuous	Easy	1 2
Stir Rate	Continuous	Easy	100 120
Temperature	Continuous	Easy	140 180
Concentration	Continuous	Easy	3 6

☐ Group new runs into separate block

Define Factor Constraints

Augmentation Choices

Replicate Add Centerpoints Fold Over Add Axial Space Filling Augment

Note: You can select **Group new runs into separate block** to add a blocking factor to any design. However, recall that this example estimates all two-factor interactions in 16 runs, and that cannot be done when there is an additional blocking factor in the model.

6. Click **Augment**.

The model shown in Figure 6.3 is defined using the Model script in the data table. With 8 runs, you can estimate up to 7 effects. The model in the script contains all 5 main effects and two interaction terms. All terms are defined as necessary for the model. For more information about effect estimability, see “[Model](#)” on page 98 in the “Custom Designs” chapter.

Figure 6.3 Initial Augment Design Platform

Augment Design

Factors

Name	Role	Changes	Values
Feed Rate	Continuous	Easy	10 15
Catalyst	Continuous	Easy	1 2
Stir Rate	Continuous	Easy	100 120
Temperature	Continuous	Easy	140 180
Concentration	Continuous	Easy	3 6

☐ Group new runs into separate block

Define Factor Constraints

Model

Main Effects Interactions ▼ RSM Cross Powers ▼ Remove Term

Name	Estimability
Intercept	Necessary
Feed Rate	Necessary
Catalyst	Necessary
Stir Rate	Necessary
Temperature	Necessary
Concentration	Necessary
Feed Rate*Catalyst	Necessary
Feed Rate*Stir Rate	Necessary

Alias Terms

Factor Design

Run	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration
1	15	2	100	140	3
2	10	2	120	140	6
3	15	1	120	140	3
4	10	2	100	180	3
5	10	1	120	180	3
6	10	1	100	140	6
7	15	2	120	180	6
8	15	1	100	180	6

Design Generation

Enter Number of Runs (counting 8 included runs):

7. In the Model outline, select **Interactions > 2nd**.

This adds all the two-factor interactions to the model. The minimum number of runs given for the specified model is 16, as shown in the Design Generation text edit box.

Note: Setting the Random Seed in step 8 and Number of Starts in step 9 reproduces the exact results shown in Figure 6.4. When you are constructing a design on your own, these steps are optional.

8. (Optional) Click the Augment Design red triangle, select **Set Random Seed**, type 282322901, and click **OK**.
9. (Optional) Click the Augment Design red triangle, select **Number of Starts**, type 800, and click **OK**.
10. Click **Make Design**.
11. Click the **Design** disclosure icon to view the design.

Figure 6.4 16 Run Augmented Design

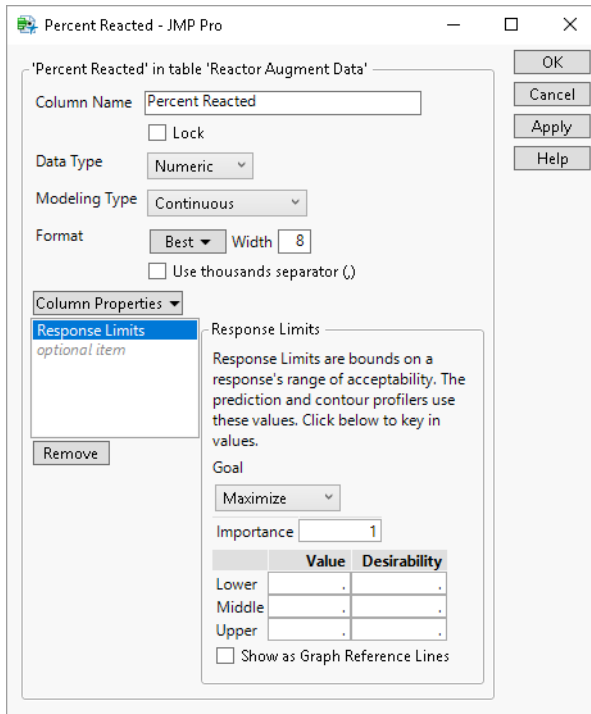
Design					
Run	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration
1	15	2	100	140	3
2	10	2	120	140	6
3	15	1	120	140	3
4	10	2	100	180	3
5	10	1	120	180	3
6	10	1	100	140	6
7	15	2	120	180	6
8	15	1	100	180	6
9	15	1	120	180	6
10	10	2	100	140	6
11	10	1	100	180	6
12	15	2	120	140	6
13	15	2	120	180	3
14	10	2	120	180	3
15	10	1	100	140	3
16	10	1	120	140	3

- Click **Make Table** to generate a design table containing the original design with results and the augmented runs.

Analyze the Augmented Design

Suppose you have conducted the additional experimental runs and have recorded results to analyze.

- Select **Help > Sample Data Library** and open Design Experiment/Reactor Augment Data.jmp. You want to maximize Percent Reacted. However, the column's Response Limits column property in this sample data table is set to **Minimize**.
- In the Columns panel of the data table, click the asterisk next to the Percent Reacted column name and select **Response Limits**, or from the Cols menu select **Column Info**.
- In the Column Info dialog that appears, change the response limit goal to **Maximize**.

Figure 6.5 Change the Response Limits Column Property for the Percent Reacted Column


- Click **OK**.

You are now ready to run the analysis.

- In the Table panel of the data table, click the green triangle next to the **Model** script. The **Model** script opens the Fit Model window with all main effects and two-factor interactions as effects.
- Change the fitting personality from **Standard Least Squares** to **Stepwise** and then click **Run**.
- In the Stepwise Regression Control panel, select **P-value Threshold** from the **Stopping Rule** menu and **Mixed** from the **Direction** menu. Enter 0.05 for the **Prob to Enter** and 0.1 for the **Prob to Leave**.
- In the Current Estimates panel, select the **Entered** check boxes for all the main effect terms (Feed Rate - Concentration) and then click **Go**.

Figure 6.6 Completed Stepwise Model

Stepwise Fit for Percent Reacted

Stepwise Regression Control

Stopping Rule: P-value Threshold ➔ Enter All Make Model

Prob to Enter: 0.05 ➔ Remove All Run Model

Prob to Leave: 0.1

Direction: Mixed

Rules: Combine

Go Stop Step

SSE	DFE	RMSE	RSquare	RSquare Adj	Cp	p	AICc	BIC
1257.1058	10	11.212073	0.6933	0.5399	.	6	143.2297	134.6378

Current Estimates

Lock	Entered	Parameter	Estimate	nDF	SS	"F Ratio"	"Prob>F"
☒	☒	Intercept	66.8125	1	0	0.000	1
☐	☒	Feed Rate(10,15)	4.4375	1	315.0625	2.506	0.14448
☐	☒	Catalyst(1,2)	10.1153846	1	1520.198	12.093	0.00595
☐	☒	Stir Rate(100,120)	-0.3653846	1	1.983516	0.016	0.90253
☐	☒	Temperature(140,180)	6.11538462	1	555.6264	4.420	0.06183
☐	☒	Concentration(3,6)	-1.7884615	1	41.58173	0.331	0.57791
☐	☐	Catalyst*Stir Rate	0	1	346.6875	3.427	0.09715
☐	☐	Catalyst*Concentration	0	1	46.44643	0.345	0.57124
☐	☐	Feed Rate*Catalyst	0	1	2.548951	0.018	0.89541
☐	☐	Feed Rate*Stir Rate	0	1	0.064103	0.000	0.98338
☐	☐	Feed Rate*Temperature	0	1	0.730769	0.005	0.9439
☐	☐	Feed Rate*Concentration	0	1	0.016484	0.000	0.99157
☐	☐	Catalyst*Temperature	0	1	667.5208	10.190	0.01097
☐	☐	Stir Rate*Temperature	0	1	22.5625	0.164	0.69453
☐	☐	Stir Rate*Concentration	0	1	2.571429	0.018	0.89495
☐	☐	Temperature*Concentration	0	1	631.1429	9.074	0.01466

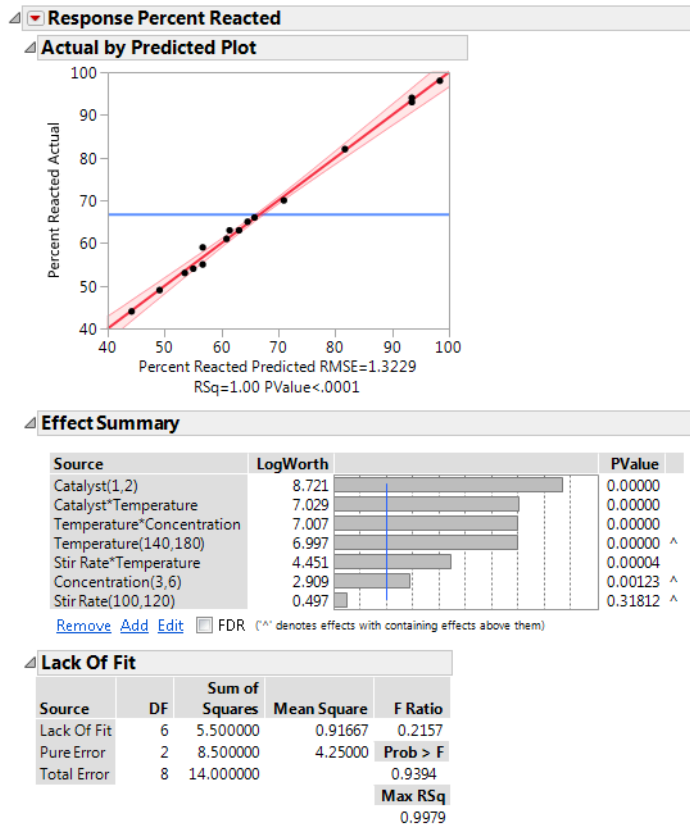
Stepwise regression, beginning with the main effects and using a p -value based search, results in a model with four main effects and three interaction terms. The main effect Feed Rate was removed from the model while the Catalyst*Temperature, Stir Rate*Temperature, and Temperature*Concentration interactions were added to the model.

- Click **Make Model** on the Stepwise Regression Control panel.

This launches the Fit Model platform for the model built using the stepwise procedure.

- Click **Run**.

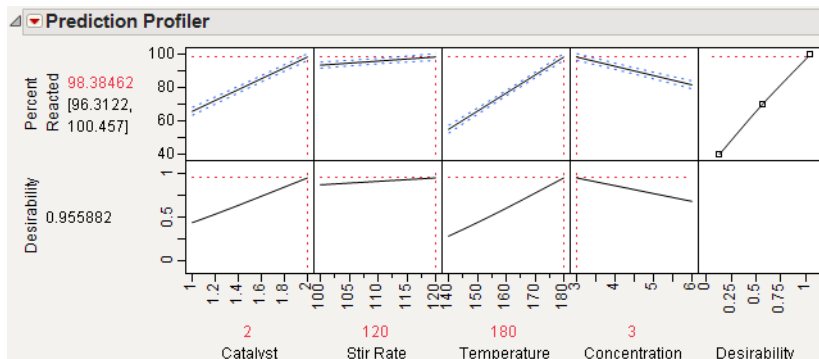
Figure 6.7 Prediction Model Analysis of Variance and Lack of Fit Tests



The Actual by Predicted Plot indicates that the overall model is significant (PValue < .0001). Both the Actual by Predicted Plot and the Lack of Fit tests show no evidence of model misspecification. The Effect Summary report shows that Catalyst is the most significant effect. However, note that the three two-factor interactions are also significant.

- Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

Figure 6.8 Maximum Percent Reacted



The prediction profile plot shows that the maximum percent reaction occurs at high Catalyst, Stir Rate, and Temperature and low Concentration. At these extreme settings, the estimate of Percent Reacted is 98.38.

Launch the Augment Design Platform

Launch the Augment Design platform by selecting **DOE > Augment Design** from the data table that contains the design that you want to augment.

Figure 6.9 Augment Design Launch Window

Add more runs to an existing data table. Replicate, add centerpoints, fold over, or add model terms.

Select Columns

▼ 7 Columns

- Pattern
- Feed Rate
- Catalyst
- Stir Rate
- Temperature
- Concentration
- Percent Reacted

Cast Selected Columns into Roles

Y, Response *required numeric continuous*
optional numeric continuous

X, Factor *required*
optional

Action

OK

Cancel

Remove

Recall

Help

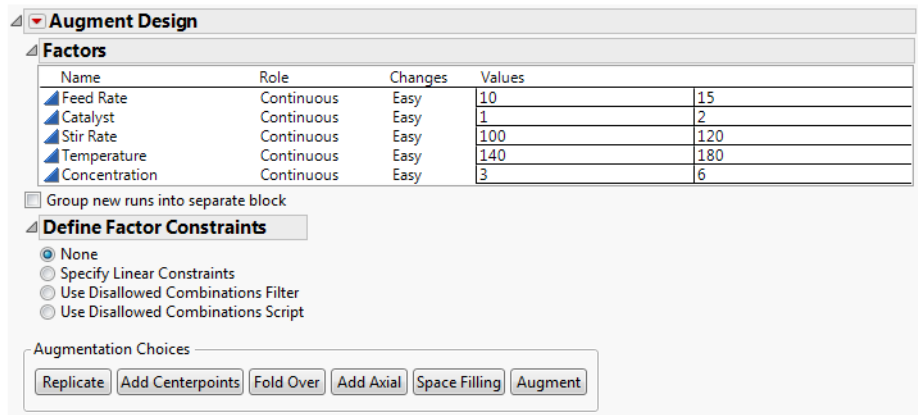
Y, Response Enter the numeric response column or columns. Entering a response is required.

X, Factor Enter the factor columns. Factors can be of any data type or modeling type.

Augment Design Window

The initial Augment Design window consists of the Factors and Define Factor Constraints outlines, and the Augmentation Choices panel.

Figure 6.10 Initial Augment Design Window Using Reactor 8 Runs.jmp



Factors

Name Lists all factors listed as **X, Factor** in the Augment Design launch window except for factors with the Random Block design role column property.

Role Specifies the Design Role of the factor specified in the column property. If the factor does not have a Design Role column property and is constant, then **Constant** appears in the Role column. Otherwise, the factor’s modeling type appears in the Role column.

Changes Indicates whether the factor levels are Easy, Hard, or Very Hard to change as specified by the Factor Changes column property in the data table. If the factor does not have a Factor Changes column property, then Changes is specified as Easy.

Note: If a factor has a Factor Changes column property that is set to Hard or Very Hard, then the corresponding whole plot factor must be included in the **X, Factor** list in the Augment Design launch window.

Values For continuous factors, shows the minimum and maximum values. For categorical factors, shows the levels.

Tip: Factors that have a role of Categorical or Constant appear in the Name column with a down arrow icon. Click the down arrow to add levels. If the factor is Constant and has a categorical modeling type, multiple levels can be added. If the factor is Constant and has a continuous modeling type, only one level can be added.

Group new runs into separate block Adds a blocking factor to the design with a block for the original design and a block for the augmented runs.

Define Factor Constraints

If you augment a design using the Space Filling or Augment options, you can define restrictions on the design space for the added runs.

Use Define Factor Constraints to restrict the design space. Unless you have loaded a constraint or included one as part of a script, the **None** option is selected. To specify constraints, select one of the other options:

Specify Linear Constraints Specifies inequality constraints on linear combinations of factors. Available only for factors with a Role of Continuous or Mixture. See [“Specify Linear Constraints”](#).

Note: When you save a script for a design that involves a linear constraint, the script expresses the linear constraint as a *less than or equal to* inequality ($\leq$).

Use Disallowed Combinations Filter Defines sets of constraints based on restricting values of individual factors. You can define both AND and OR constraints. See [“Use Disallowed Combinations Filter”](#).

Use Disallowed Combinations Script Defines disallowed combinations and other constraints as Boolean JSL expressions in a script editor box. See [“Use Disallowed Combinations Script”](#).

Specify Linear Constraints

In cases where it is impossible to vary continuous factors independently over the design space, you can specify linear inequality constraints. Linear inequalities describe factor level settings that are allowed.

Click **Add** to enter one or more linear inequality constraints.

Add Adds a template for a linear expression involving all the continuous factors in your design. Enter coefficient values for the factors and select the direction of the inequality to reflect your linear constraint. Specify the constraining value in the box to the right of the inequality. To add more constraints, click **Add** again.

Note: The Add option is disabled if you have already constrained the design region by specifying a Sphere Radius.

Remove Last Constraint Removes the last constraint.

Check Constraints Checks the constraints for consistency. This option removes redundant constraints and conducts feasibility checks. A JMP alert appears if there is a problem. If constraints are equivalent to bounds on the factors, a JMP alert indicates that the bounds in the Factors outline have been updated.

Use Disallowed Combinations Filter

This option uses an adaptation of the Data Filter to facilitate specifying disallowed combinations. For detailed information, see the JMP Reports chapter in *Using JMP*.

Select factors from the Add Filter Factors list and click **Add**. Then specify the disallowed combinations by using the slider (for continuous factors) or by selecting levels (for categorical factors).

The red triangle options for the Add Filter Factors menu are those found in the Select Columns panel of many platform launch windows. See the Get Started chapter in *Using JMP*.

When you click Add, the Disallowed Combinations control panel shows the selected factors and provides options for further control. Factors are represented as follows, based on their modeling types:

Continuous Factors For a continuous factor, a double-arrow slider that spans the range of factor settings appears. An expression that describes the range using an inequality appears above the slider. You can specify disallowed settings by dragging the slider arrows or by clicking on the inequality bounds in the expression and entering your desired constraints. In the slider, a solid blue highlight represents the disallowed values.

Categorical Factor For a categorical factor, the possible levels are displayed either as labeled blocks or, when the number of levels is large, as list entries. Select a level to disallow it. To select multiple levels, hold the Control key. The block or list entries are highlighted to indicate the levels that have been disallowed. When you add a categorical factor to the Disallowed Combinations panel, the number of levels of the categorical factor is given in parentheses following the factor name.

Disallowed Combinations Options

The control panel has the following controls:

Clear Clears all disallowed factor level settings that you have specified. This does not clear the selected factors.

Start Over Removes all selected factors and returns you to the initial list of factors.

AND Opens the Add Filter Factors list. Selected factors become an AND group. Any combination of factor levels specified within an AND group is disallowed.

To add a factor to an AND group later on, click the group's outline to see a highlighted rectangle. Select AND and add the factor.

To remove a single factor, select **Delete** from its red triangle menu.

OR Opens the Add Filter Factors list. Selected factors become a separate AND group. For AND groups separated by OR, a combination is disallowed if it is specified in at least one AND group.

Red Triangle Options for Factors

A factor can appear in several OR groups. An occurrence of the factor in a specific OR group is referred to as an *instance* of the factor.

Delete Removes the selected instance of the factor from the Disallowed Combinations panel.

Clear Selection Clears any selection for that instance of the factor.

Invert Selection Deselects the selected values and selects the values not previously selected for that instance of the factor.

Display Options Available only for categorical factors. Changes the appearance of the display. Options include showing each level as a block, list, or single category, or adding a check box next to each value.

Find Available only for categorical factors. Provides a text box beneath the factor name where you can enter a search string for levels of the factor. Press the Enter key or click outside the text box to perform the search. Once **Find** is selected, Find options appear in the red triangle menu, such as Does not contain, Match Case, Starts with or Ends with, or Clear Find.

Use Disallowed Combinations Script

Use this option to disallow particular combinations of factor levels using a JSL script. This option can be used with continuous factors or mixed continuous and categorical factors.

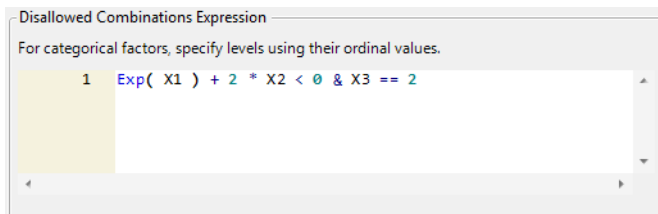
This option opens a script window where you insert a script that identifies the combinations that you want to disallow. The script must evaluate as a Boolean expression. When the expression evaluates as true, the specified combination is disallowed.

When forming the expression for a categorical factor, use the ordinal value of the level or the name of the level. If a factor's levels are high, medium, and low, specified in that order in the Factors outline, their associated ordinal values are 1, 2, and 3. For example, suppose that you have two continuous factors, X1 and X2, and a categorical factor X3 with three levels: L1, L2, and L3, in order. You want to disallow levels where the following is true:

$$e^{X_1} + 2X_2 < 0 \text{ and } X_3 = L2$$

Enter the expression `(Exp(X1) + 2*X2 < 0) & (X3 == 2)` into the script window.

Figure 6.11 Expression in Script Editor



(In the figure, unnecessary parentheses were removed by parsing.) Notice that functions can be entered as part of the Boolean expression. The expression `(Exp(X1) + 2*X2 < 0) & (X3 == "L2")` is also valid.

Augmentation Choices

The Augment Design platform has the following choices for augmentation:

Replicate Replicates the design a specified number of times. See [“Replicate a Design”](#) on page 229.

Add Centerpoints Adds center points. Specify how many runs you want to add as center points to the design. A center point is a run whose setting for each continuous factor is midway between the high and low settings. See [“Center Points, Replicate Runs, and Testing”](#) on page 63 in the “Starting Out with DOE” chapter.

If a design contains both continuous and other types of factors, center points might not be balanced relative to the levels of the other factors. Augment Design chooses the center points to maximize the D-, I-, or alias efficiency of the design.

See [“Add Center Points to a Design”](#) on page 231.

Fold Over Creates a foldover design. See [“Create a Foldover Design”](#) on page 231.

Add Axial Adds axial points and center points to a design. Add Axial can be used to generate a central composite design from a screening design. See [“Add Axial Points to a Design”](#) on page 232.

Space Filling Adds additional runs to any design consisting of continuous factors. Additional runs are constructed using the fast flexible filling methodology. See [“Space Filling Augmentation”](#) on page 233. Space Filling does not support grouping new runs in a separate block.

Augment Adds runs to the design using a model. This model typically has more terms than the original model. See [“Example Using the Augment Design Platform”](#) on page 213.

Augment Design Platform Options

This section describes the options available under the Augment Design red triangle menu.

Save Responses Saves the information in the Responses panel to a new data table. You can then quickly load the responses and their associated information into most DOE windows. This option is helpful if you anticipate reusing the responses.

Load Responses Loads responses that you saved using the Save Responses option.

Save Factors Saves the information in the Factors panel to a new data table. Each factor's column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Note: It is possible to create a factors table by entering data into an empty table, but remember to assign each column an appropriate Design Role. Do this by right-clicking on the column name in the data grid and selecting **Column Properties > Design Role**. In the Design Role area, select the appropriate role.

Load Factors Loads factors that you saved using the Save Factors option.

Save Constraints (Unavailable for some platforms) Saves factor constraints that you defined in the Define Factor Constraints or Linear Constraints outline into a data table, with a column for each constraint. You can then quickly load the constraints into most DOE windows.

In the constraint table, the first rows contain the coefficients for each factor. The last row contains the inequality bound. Each constraint's column contains a column property called **ConstraintState** that identifies the constraint as a “less than” or a “greater than” constraint. See [“ConstraintState”](#) on page 766 in the “Column Properties” appendix.

Load Constraints (Unavailable for some platforms) Loads factor constraints that you saved using the Save Constraints option.

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction
- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.
- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click Apply, a formula for the simulated response values is saved in a new column called <Y> Simulated, where Y is the name of the response. Clicking Apply again updates the formula and values in <Y> Simulated.

For additional details, see [“Simulate Responses”](#) on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Save X Matrix Saves scripts called Moments Matrix and Model Matrix to the design data table. These scripts contain the moments and design matrices. See [“Save X Matrix”](#) on page 112 in the “Custom Designs” chapter.

Caution: For a design with nominal factors, the matrix in the Model Matrix script saved by the Save X Matrix option is *not* the coding matrix used in fitting the linear model. You can obtain the coding matrix used for fitting the model by selecting the option Save Columns > Save Coding Table in the Fit Model report that you obtain when you run the Model script.

Optimality Criterion Changes the design optimality criterion. The default criterion, **Recommended**, specifies D-optimality for all design types, unless you added quadratic effects using the RSM button in the Model outline. For more information about the D-, I-, and alias-optimal designs, see [“Optimality Criteria”](#) on page 127 in the “Custom Designs” chapter.

Note: You can set a preference to always use a given optimality criterion. Select File > Preferences > Platforms > DOE. Select Optimality Criterion and choose your preferred criterion.

Number of Starts Enables you to specify the number of random starts used in constructing the design. See [“Number of Starts”](#) on page 115 in the “Custom Designs” chapter.

Design Search Time Maximum number of seconds spent searching for a design. The default search time is based on the complexity of the design. See [“Design Search Time”](#) on page 115 in the “Custom Designs” chapter and [“Number of Starts”](#) on page 115 in the “Custom Designs” chapter.

If the iterations of the algorithm require more than a few seconds, a Computing Design progress window appears. If you click **Cancel** in the progress window, the calculation stops and gives the best design found at that point. The progress window also displays D-efficiency for D-optimal designs that do not include factors with Changes set to Hard or Very Hard or with Estimability set to If Possible.

Note: You can set a preference for Design Search Time. Select File > Preferences > Platforms > DOE. Select Design Search Time and enter the maximum number of seconds. In certain situations where more time is required, JMP extends the search time.

Sphere Radius Constrains the continuous factors in a design to a hypersphere. Specify the radius and click **OK**. Design points are chosen so that their distance from 0 equals the Sphere Radius. Select this option before you click Make Design.

Note: Sphere Radius constraints cannot be combined with constraints added using the Specify Linear Constraints option. Also, the option is not available when hard-to-change factors are included (split-plot designs).

Advanced Options

Mixture Sum Sets the sum of the mixture factors to any positive value. Use this option to keep a component of a mixture constant throughout an experiment.

Split Plot Variance Ratio Specifies the ratio of the variance of the random whole plot and the subplot variance (if present) to the error variance. Before setting this value, you must define a hard-to-change factor for your split-plot design, or hard and very-hard-to-change factors for your split-split-plot design. Then you can enter one or two positive numbers for the variance ratios, depending on whether you have specified a split-plot or a split-split-plot design.

Prior Parameter Variance (Available only when the Model outline is available.) Specifies the weights that are used for factors whose Estimability is set to If Possible. The option is updated to show the default weights when you click Make Design. Enter a positive number for each of the terms for which you want to specify a weight. The value that you enter is the square root of the reciprocal of the prior variance. A larger value represents a smaller variance and therefore more prior information that the effect is not active.

Bayesian D- or I-optimality is used in constructing designs with If Possible factors. The default values used in the algorithm are 0 for Necessary terms, 4 for interactions involving If Possible terms, and 1 for If Possible terms. See [“The Alias Matrix”](#) on page 771 in the “Technical Details” appendix and [“Optimality Criteria”](#) on page 127 in the “Custom Designs” chapter.

D Efficiency Weight Specifies the relative importance of D-efficiency to alias optimality in constructing the design. Select this option to balance reducing the variance of the coefficients with obtaining a desirable alias structure. Values should be between 0 and 1. Larger values give more weight to D-Efficiency. The default value is 0.5. This option has an effect only when you select Make Alias Optimal Design as your Optimality Criterion.

For the definition of D-efficiency, see [“Optimality Criteria”](#) on page 127 in the “Custom Designs” chapter. For more information about alias optimality, see [“Alias Optimality”](#) on page 131 in the “Custom Designs” chapter.

Set Delta for Power Specifies the difference in the mean response that you want to detect for model effects. See [“Set Delta for Power”](#) on page 116 in the “Custom Designs” chapter.

Save Script to Script Window Creates the script for the design that you specified in the Custom Design window and places it in an open script window.

Additional Examples of Augmentation Choices

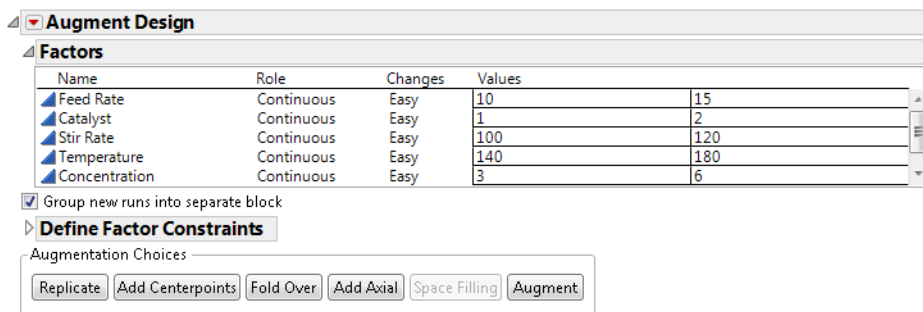
Replicate a Design

Replication provides a direct check on the assumption that the error variance is constant. It also reduces the variability of the regression coefficients in the presence of large process or measurement variability.

This example replicates an 8 run design.

1. Select **Help > Sample Data Library** and open Design Experiment/Reactor 8 Runs.jmp.
2. Select **DOE > Augment Design**.
3. Select Percent Reacted and click **Y, Response**.
4. Select Feed Rate, Catalyst, Stir Rate, Temperature, and Concentration and click **X, Factor**.
5. Click **OK**.

Figure 6.12 Choose an Augmentation Type



6. Select **Group new runs into separate block** to add a blocking variable to the design.
7. Click **Replicate**, enter 2, and then click **OK**.

Note: The replicate dialog box asks for the number of times each run is to be performed. Entering 2 specifies that you want each run to appear twice in the resulting design. This results in a single replicate of your original design.

Figure 6.13 Reactor Data Design Augmented with a Replicate

Factor Design						
Run	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration	Block
1	15	2	100	140	3	1
2	10	2	120	140	6	1
3	15	1	120	140	3	1
4	10	2	100	180	3	1
5	10	1	120	180	3	1
6	10	1	100	140	6	1
7	15	2	120	180	6	1
8	15	1	100	180	6	1
9	15	2	100	140	3	2
10	10	2	120	140	6	2
11	15	1	120	140	3	2
12	10	2	100	180	3	2
13	10	1	120	180	3	2
14	10	1	100	140	6	2
15	15	2	120	180	6	2
16	15	1	100	180	6	2

The original 8 runs are in Block 1 and the replicated runs are in Block 2.

Note: The full design can be evaluated from the Augment Design Platform. For more information about design evaluation, see [Chapter 15, “Evaluate Designs”](#).

8. Click **Make Table**.

Figure 6.14 The Replicated Design

	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration	Block	Percent Reacted
1	15	2	100	140	3	1	61
2	10	2	120	140	6	1	67
3	15	1	120	140	3	1	56
4	10	2	100	180	3	1	94
5	10	1	120	180	3	1	66
6	10	1	100	140	6	1	56
7	15	2	120	180	6	1	82
8	15	1	100	180	6	1	45
9	15	2	100	140	3	2	•
10	10	2	120	140	6	2	•
11	15	1	120	140	3	2	•
12	10	2	100	180	3	2	•
13	10	1	120	180	3	2	•
14	10	1	100	140	6	2	•
15	15	2	120	180	6	2	•
16	15	1	100	180	6	2	•

The table includes the original design and results in Block 1. The augmented runs are in Block 2. The percent reacted values can be added to the table after the augmented runs have been measured.

Add Center Points to a Design

Adding center points is useful to check for curvature and reduce the prediction error in the center of the factor region. Center points are usually replicated points that allow for an independent estimate of pure error, which can be used in a lack-of-fit test.

This example adds two center points to an 8-run design.

1. Select **Help > Sample Data Library** and open Design Experiment/Reactor 8 Runs.jmp.
2. Select **DOE > Augment Design**.
3. Select Percent Reacted and click **Y, Response**.
4. Select Feed Rate, Catalyst, Stir Rate, Temperature, and Concentration and click **X, Factor**.
5. Click **OK**.
6. Select **Group new runs into separate block** to add a blocking variable to the design.
7. Click **Add Centerpoints**, enter 2, and then click **OK**.
8. Click **Make Table**.

Figure 6.15 Design with Two Center Points Added

	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration	Block	Percent Reacted
1	15	2	100	140	3	1	61
2	10	2	120	140	6	1	67
3	15	1	120	140	3	1	56
4	10	2	100	180	3	1	94
5	10	1	120	180	3	1	66
6	10	1	100	140	6	1	56
7	15	2	120	180	6	1	82
8	15	1	100	180	6	1	45
9	12.5	1.5	110	160	4.5	2	•
10	12.5	1.5	110	160	4.5	2	•

After data has been collected for the two center runs you can then test for curvature in your model.

Create a Foldover Design

A foldover design removes the confounding of two-factor interactions and main effects. This is especially useful as a follow-up to saturated or near-saturated fractional factorial or Plackett-Burman designs.

This example creates a foldover design of an 8-run design.

1. Select **Help > Sample Data Library** and open Design Experiment/Reactor 8 Runs.jmp.
2. Select **DOE > Augment Design**.

3. Select **Percent Reacted** and click **Y, Response**.
4. Select **Feed Rate, Catalyst, Stir Rate, Temperature, and Concentration** and click **X, Factor**.
5. Click **OK**.
6. Select **Group new runs into separate block** to add a blocking variable to the design.
7. Click **Fold Over** and then click **OK**.

In the fold over dialog box, if you select no factors, then all design factors are included in the fold over. If you choose a subset of factors to fold over, the remaining factors are replicates of the original runs.

Figure 6.16 Foldover Design on All Factors

Factor Design						
Run	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration	Block
1	15	2	100	140	3	1
2	10	2	120	140	6	1
3	15	1	120	140	3	1
4	10	2	100	180	3	1
5	10	1	120	180	3	1
6	10	1	100	140	6	1
7	15	2	120	180	6	1
8	15	1	100	180	6	1
9	10	1	120	180	6	2
10	15	1	100	180	3	2
11	10	2	100	180	6	2
12	15	1	120	140	6	2
13	15	2	100	140	6	2
14	15	2	120	180	3	2
15	10	1	100	140	3	2
16	10	2	120	140	3	2

8. Click **Make Table**.

Each run in Block 2 can be obtained by changing the factor levels (from low to high or from high to low) of a run from Block 1.

Add Axial Points to a Design

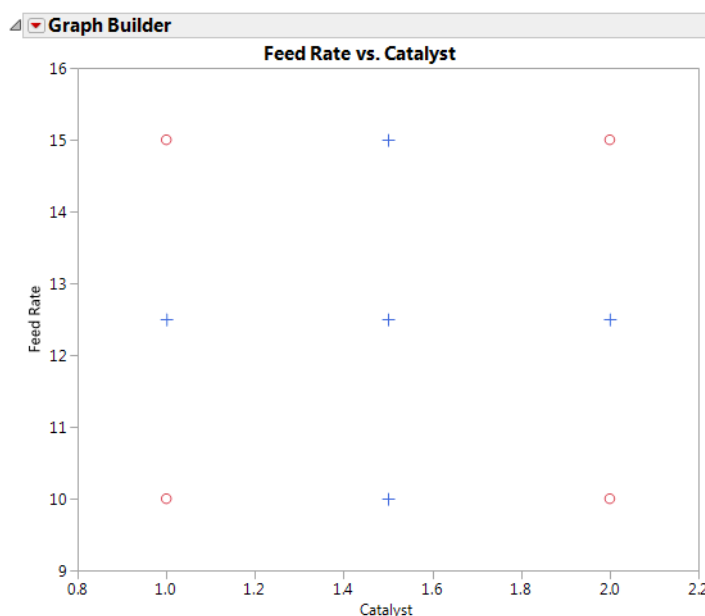
You can add axial and center points to augment a screening design in order to obtain a response surface design. This enables you to fit a response surface model to your data to optimize the settings of your factors with respect to your response.

This example augments an 8 run screening design with axial and center points.

1. Select **Help > Sample Data Library** and open **Design Experiment/Reactor 8 Runs.jmp**.
2. Select **DOE > Augment Design**.
3. Select **Percent Reacted** and click **Y, Response**.
4. Select **Feed Rate, Catalyst, Stir Rate, Temperature, and Concentration** and click **X, Factor**.
5. Click **OK**.

6. Select **Group new runs into separate block** to add a blocking variable to the design.
7. Click **Add Axial**, enter 1 for axial value, 2 for number of center points, and then click **OK**.
8. Click **Make Table**.
9. From the Augmented Design table, select **Rows > Color or Mark by Column**.
10. Select **Block**, from Markers select **Standard**, and then click **OK**.
11. Select **Graph > Graph Builder**.
12. Drag **Feed Rate** to Y and drag **Catalyst** to X. Then click the smoother icon to remove.

Figure 6.17 Graphical View of the Augmented Design for Two Factors



The “+” points in the graph of Feed Rate versus Catalyst are the added axial and center points. The axial points are on the face of the design space. Increasing or decreasing the axial value in the augmentation of the design would move the axial points beyond or within the design space.

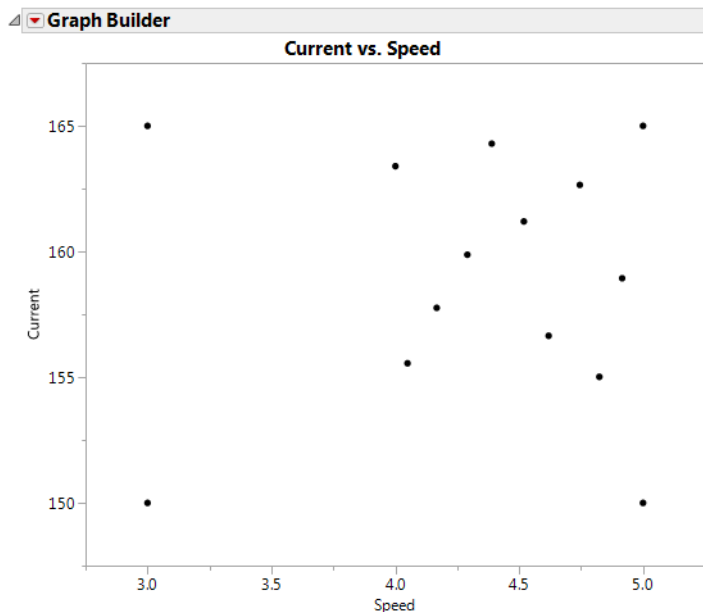
Space Filling Augmentation

The Space Filling augmentation choice adds points to a design consisting of continuous factors. For more information about the algorithm used, see [“Statistical Details”](#) on page 235. The Space Filling choice accommodates constraints on the design space. You can specify linear constraints or disallowed combinations.

This example augments a design to explore a constrained area of the design space where the response, Depth, was maximized, with added constraints on both factors.

1. Select **Help > Sample Data Library** and open Design Experiment/DOE Example 1.jmp.
2. Select **DOE > Augment Design**.
3. Select Depth and click **Y, Response**.
4. Select Speed and Current, click **X, Factor**, and then click **OK**.
5. Open the Define Factor Constraints outline.
6. Select **Specify Linear Constraints** and click **Add**.
7. Enter 1 for Speed, select greater than or equal to, and then enter 4 for the lower bound on Speed.
8. Click **Add**.
9. Enter 1 for Current, select greater than or equal to, and then enter 155 for the lower bound on Current.
10. Click **Space Filling**.
11. Enter 10 and click **OK**.
12. Click **Make Table**.
13. From the Augmented Design table, select Graph Builder.
14. Drag Speed to the X drop zone and drag Current to the Y drop zone.

Figure 6.18 Space Filling Augmented Runs



The augmented runs fill the space in the upper right corner of the design space. This is the constrained design space.

Statistical Details

This section contains statistical details about the algorithm used to select points for the space-filling augmentation choice.

The algorithm that is used to augment designs begins by generating a large number of random points within the specified design region. These points are then clustered using a Fast Ward algorithm into a number of clusters that equals the Number of Additional Runs that you specify.

The final design points are obtained by optimizing the MaxPro (*maximum projection*) criterion over the existing and additional runs. For p factors and n equal to the number of existing and additional runs, the MaxPro criterion strives to find points in the clusters that minimize the following criterion:

$$C_{MaxPro} = \sum_i^{n-1} \sum_{j=i+1}^n \left[1 / \prod_{k=1}^p (x_{ik} - x_{jk})^2 \right]$$

The MaxPro criterion maximizes the product of the distances between design points in a way that involves all factors. This supports the goal of providing good space-filling properties on projections of factors. See Joseph et al. (2015).

Chapter 7

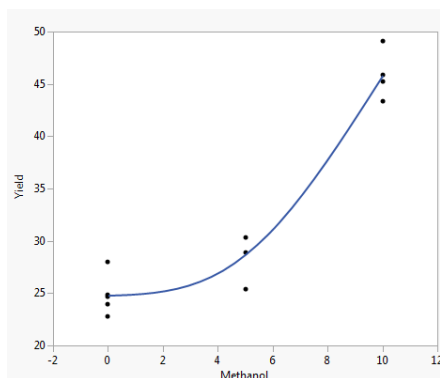
Definitive Screening Designs

Definitive screening designs (DSD) are screening designs. They are appropriate for early stage experimentation work, typically with four or more factors. DSD can be used for combinations of continuous or two-level categorical factors. They work best when most of the factors are continuous. Each continuous factor has three levels allowing one to investigate quadratic model terms for continuous factors.

Here are areas where definitive screening designs are superior to standard screening designs:

- They help identify the causes of nonlinear effects by fielding each continuous factor at three levels. In standard screening designs, continuous factors have only two levels. You can add center points to screening designs, but these points establish only if curvature exists. They do not allow you to identify the factors responsible for quadratic effects.
- They avoid confounding between any effects up through the second order. For continuous factors, definitive screening designs have main effects that are orthogonal to each other and orthogonal to two-factor interactions and quadratic effects. Two-factor interactions are not completely confounded with each other. Confounding occurs in many standard screening designs with a similar number of runs.
- They avoid the need for costly additional experimentation to resolve ambiguity from the initial results of standard screening designs.

Figure 7.1 Plot of Response against Factor Values Showing Curvature



Contents

- Overview of Definitive Screening Design 239
- Examples of Definitive Screening Designs 241
 - Definitive Screening Design 242
 - Comparison with a Fractional Factorial Design 244
 - Definitive Screening Design with Blocking 246
 - Comparison of a Definitive Screening Design with a Plackett-Burman Design 250
- Definitive Screening Design Window 253
 - Responses 254
 - Factors 256
 - Design Options 258
 - Design 260
 - Design Evaluation 261
 - Output Options 261
- Definitive Screening Design Options 263
 - Simulate Responses 265
- Technical Details 268
 - Structure of Definitive Screening Designs 268
 - Analysis of Experimental Data 270

Overview of Definitive Screening Design

Investigators use screening designs when they want to identify the factors that have the most substantial effects on a response. A screening design enables you to study a large number of factors in a fairly small experiment.

Many standard screening designs focus on estimating main effects. Definitive screening designs offer advantages over standard screening designs. They avoid confounding of effects and can identify factors having a nonlinear effect on the response. For more information about the advantages and construction of definitive screening designs, see Jones and Nachtsheim (2011a).

For designs containing only continuous factors, compare these properties of definitive screening designs versus standard screening designs:

Note: When quadratic effects are mentioned, the standard screening designs are assumed to have center points.

- Main effects are orthogonal to two-factor interactions.
 - Definitive Screening Designs: Always
 - Standard Screening Designs: Only for Resolution IV or higher
- No two-factor interaction is completely confounded with any other two-factor interaction.
 - Definitive Screening Designs: Always
 - Standard Screening Designs: Only for Resolution V or higher
- All quadratic effects are estimable in models containing only main and quadratic effects.
 - Definitive Screening Designs: Always
 - Standard Screening Designs: Never

These properties are described more fully in the remainder of this section.

Standard Screening Designs

Standard screening designs, such as fractional factorial or Plackett-Burman designs, attempt to study many factors with a relatively small allocation of resources. However, standard screening designs have several undesirable features:

- They can alias some main effects with two-factor interactions. In Plackett-Burman designs, for example, main effects are correlated with several two-factor interactions. If one or more two-factor interaction effects are substantial, then the experimenter must perform additional runs to resolve the ambiguities.

- They can also confound some two-factor interactions with each other. Consequently, if a two-factor interaction effect is substantial, then the experimenter must perform additional runs to resolve the remaining ambiguities.
- Continuous factors are usually set at only two levels (low and high). However, engineers and scientists often prefer designs where continuous factors are set at three levels (low, middle, and high). This is because two levels are not sufficient to detect nonlinearity, which is common in physical systems. You can use a traditional screening design with added center points to detect nonlinearity, but such a design does not identify the responsible factors.

Definitive Screening Designs

Using definitive screening designs, you can do the following:

- Avoid model ambiguity, enabling you to identify important factors more quickly and efficiently.
- Identify the cause of nonlinear effects while avoiding confounding any terms up to second order. So not only can you *detect* nonlinearity, as you might with center points in a traditional screening design, but you can *identify* the responsible factors.

Definitive screening designs offer the following advantages:

- Definitive screening designs require only a small number of runs. For six or more factors, the minimum number of required runs is usually only a few more than twice the number of factors. For more detail on the number of runs, see [“Conference Matrices and the Number of Runs”](#) on page 269.
- Main effects are orthogonal to two-factor interactions. This means that estimates of main effects are not biased by the presence of active two-factor interactions, whether these interactions are included in the model or not. Note that resolution III screening designs confound some main and interaction effects. Also, Plackett-Burman designs produce biased main effect estimates if there are active two-factor interactions.
- No two-factor interaction is completely confounded with any other two-factor interaction. However, a two-factor interaction might be correlated with other two-factor interactions. Note that resolution IV screening designs completely confound some two-factor interaction effects.
- All quadratic effects are estimable in models comprised only of main effects and quadratic terms. This enables you to identify the factors that account for nonlinearity. Note that traditional screening designs with added center points do *not* allow estimation of all quadratic effects in models consisting of main and quadratic effects.
- Quadratic effects are orthogonal to main effects and not completely confounded with two-factor interactions. A quadratic effect might be correlated with interaction effects.
- For 6 through at least 30 factors, it is possible to estimate the parameters of any full quadratic model involving three or fewer factors with high precision.

- For 18 factors or more, they can fit full quadratic models in any 4 factors. For 24 factors or more, they can fit full quadratic models in any 5 factors.

Definitive Screening Design Platform

The Definitive Screening Design platform enables you to construct definitive screening designs for continuous factors and for two-level categorical factors. It also enables you to construct blocked designs. You can add extra non-center runs that enhance the ability of the design to reliably detect effects when many effects are active.

To view the absolute values of the correlations among effects, use the Color Map on Correlations provided as part of the Design Evaluation outline in the Definitive Screening Design window. You can compare the aliasing structure of definitive screening designs to that of other designs by comparing their color maps on correlations. See [“Color Map on Correlations”](#) on page 463 in the “Evaluate Designs” chapter.

For more information about the structure of definitive screening designs, see [“Structure of Definitive Screening Designs”](#) on page 268. For information about definitive screening designs with blocks, see [“Blocking in Definitive Screening Designs”](#) on page 259. For suggestions on how to analyze data obtained using definitive screening designs, see [“Analysis of Experimental Data”](#) on page 270.

Fit Definitive Screening Platform

After you run a Definitive Screening Design (DSD), analyze your results using the Fit Definitive Screening platform. Standard model selection methods applied to DSDs can fail to identify active effects. To identify active main effects and second-order effects, the Fit Definitive Screening platform uses an algorithm called *Effective Model Selection for DSDs*. This algorithm leverages the special structure of DSDs. See [Chapter 8, “The Fit Definitive Screening Platform”](#).

If you create your DSD in JMP, the design table contains a script called Fit Definitive Screening that automatically runs an analysis using the Effective Model Selection for DSDs methodology.

Examples of Definitive Screening Designs

- [“Definitive Screening Design”](#)
- [“Comparison with a Fractional Factorial Design”](#)
- [“Definitive Screening Design with Blocking”](#)
- [“Comparison of a Definitive Screening Design with a Plackett-Burman Design”](#)

Definitive Screening Design

Suppose that you need to determine which of six factors have an effect on the yield of an extraction process.

Create the Design

The factors and their settings are given in the data table Extraction Factors.jmp. You create a definitive screening design to investigate.

1. Select **DOE > Definitive Screening > Definitive Screening Design**.
2. Double-click Y under Response Name and type Yield.
3. Select **Help > Sample Data Library** and open Design Experiment/Extraction Factors.jmp.
4. Click the red triangle next to Definitive Screening Design and select **Load Factors**

The factor names and ranges are added to the Factors outline.

Figure 7.2 Responses and Factors for Extraction Design

The screenshot shows the JMP software interface for creating a Definitive Screening Design. The 'Responses' panel is at the top, and the 'Factors' panel is below it.

Responses Panel:

- Buttons: Add Response, Remove, Number of Responses...
- Table:

Response Name	Goal	Lower Limit	Upper Limit	Importance
Yield	Maximize	.	.	.

Factors Panel:

- Buttons: Continuous, Categorical, Remove, Add N Factors (set to 1)
- Table:

Name	Role	Values
Methanol	Continuous	0 10
Ethanol	Continuous	0 10
Propanol	Continuous	0 10
Butanol	Continuous	0 10
pH	Continuous	6 9
Time	Continuous	1 2

Specify Factors:

Add a Continuous or Categorical factor by clicking its button. Double click on a factor name or level to edit it.

Continue

5. Click **Continue**.

The Design Options outline opens. Here you can specify a blocking structure. There is no need to block in this example, so you accept the default selection of **No Blocks Required**.

You can also choose to add Extra Runs, which greatly enhance your ability to detect second-order effects. A minimum of four Extra Runs is highly recommended and is the default.

6. Click **Make Design**.

The Definitive Screening Design window updates to show a Design outline and a Design Evaluation outline.

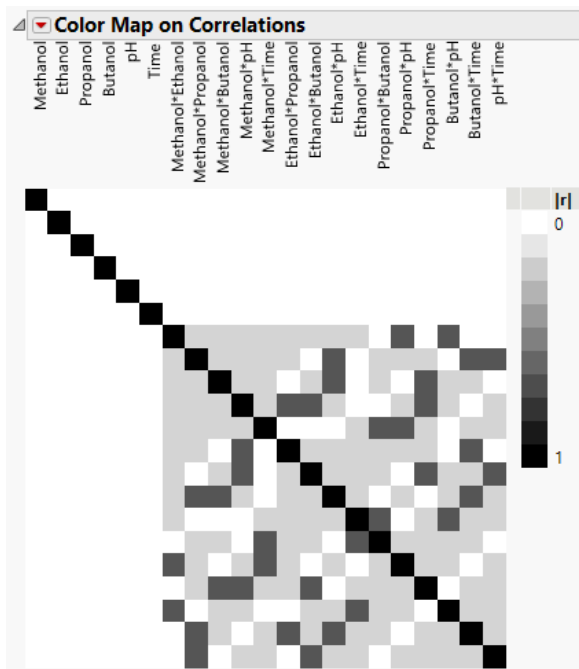
Figure 7.3 Design Outline

Design							
Run	Methanol	Ethanol	Propanol	Butanol	pH	Time	
1	5	10	10	10	9	2	
2	5	0	0	0	6	1	
3	10	5	10	10	6	2	
4	0	5	0	0	9	1	
5	10	0	5	10	9	1	
6	0	10	5	0	6	2	
7	10	0	0	5	9	2	
8	0	10	10	5	6	1	
9	10	10	0	0	7.5	2	
10	0	0	10	10	7.5	1	
11	10	0	10	0	6	1.5	
12	0	10	0	10	9	1.5	
13	10	10	0	10	6	1	
14	0	0	10	0	9	2	
15	10	10	10	0	9	1	
16	0	0	0	10	6	2	
17	5	5	5	5	7.5	1.5	

- Open the **Design Evaluation > Color Map on Correlations** outline.

The Color Map on Correlations assigns a color intensity scale to the absolute values of correlations among all main effects and two-factor interactions.

Figure 7.4 Color Map on Correlations for Extraction Design



Note the following:

- The solid deep blue area shows that there is no correlation between main effects or between main effects and two-factor interactions.
- The lighter blue and gray areas indicate that the absolute correlations between two-factor interactions are small.
- The solid red squares indicate absolute correlations of 1. These all appear on the diagonal, reflecting the expected correlation of an effect with itself.

In the Output Options panel, note that the Run Order is set to **Randomize**.

8. Click **Make Table** to obtain the data table shown in Figure 7.5.

Note: The runs in your design might appear in a different order than the order shown in Figure 7.5.

Figure 7.5 Definitive Screening Design for Extraction Process

	Methanol	Ethanol	Propanol	Butanol	pH	Time	Yield
1	5	0	0	0	6	1	•
2	0	0	10	0	9	2	•
3	5	10	10	10	9	2	•
4	0	5	0	0	9	1	•
5	10	0	10	0	6	1.5	•
6	10	10	10	0	9	1	•
7	10	5	10	10	6	2	•
8	0	10	0	10	9	1.5	•
9	0	10	5	0	6	2	•
10	5	5	5	5	7.5	1.5	•
11	10	10	0	10	6	1	•
12	10	10	0	0	7.5	2	•
13	10	0	5	10	9	1	•
14	0	10	10	5	6	1	•
15	10	0	0	5	9	2	•
16	0	0	0	10	6	2	•
17	0	0	10	10	7.5	1	•

Comparison with a Fractional Factorial Design

Suppose that you had chosen a traditional screening design instead of the definitive screening design in “[Definitive Screening Design](#)” on page 242. This example compares the two designs in terms of confounding.

1. Select **DOE > Classical > Two Level Screening > Screening Design**.
2. Double-click Y under Response Name and type Yield.
3. Select **Help > Sample Data Library** and open Design Experiment/Extraction Factors.jmp.
4. Click the red triangle next to and select **Load Factors**.

The factor names and ranges are added to the Factors outline.

5. Click **Continue**.
6. Select **Choose from a list of fractional factorial designs**.
7. Click **Continue**.

Potential designs appear in the Design List.

Figure 7.6 Screening Design List for Six Continuous Factors

Design List
Choose a design by clicking on its row in the list.

Number Of Runs	Block Size	Design Type	Resolution - what is estimable
8		Fractional Factorial	3 - Main Effects Only
8	4	Fractional Factorial	3 - Main Effects Only
12		Plackett-Burman	3 - Main Effects Only
16		Fractional Factorial	4 - Some 2-factor interactions
16	8	Fractional Factorial	4 - Some 2-factor interactions
16	4	Fractional Factorial	4 - Some 2-factor interactions
16	2	Fractional Factorial	4 - Some 2-factor interactions
32		Fractional Factorial	5+ - All 2-factor interactions
32	16	Fractional Factorial	5+ - All 2-factor interactions
32	8	Fractional Factorial	4 - Some 2-factor interactions
32	4	Fractional Factorial	4 - Some 2-factor interactions
32	2	Fractional Factorial	4 - Some 2-factor interactions
64		Full Factorial	>6 - Full Resolution
64	32	Full Factorial	5+ - All 2-factor interactions

8. Select the sixteen-run fractional factorial design with no blocks, shown highlighted in Figure 7.6.
9. Click **Continue**.
10. Open the **Display and Modify Design > Aliasing of Effects** outline.

Figure 7.7 Aliasing of Effects for Fractional Factorial Design

Aliasing of Effects

Effects	Aliases
Methanol*Ethanol	= pH*Time
Methanol*Propanol	= Butanol*Time
Methanol*Butanol	= Propanol*Time
Methanol*pH	= Ethanol*Time
Methanol*Time	= Ethanol*pH = Propanol*Butanol
Ethanol*Propanol	= Butanol*pH
Ethanol*Butanol	= Propanol*pH

The Aliasing of Effects outline for the 16-run fractional factorial design shows that every two-factor interaction is confounded with at least one other two-factor interaction. In this fractional factorial design, the Ethanol*Time interaction is confounded with Methanol*pH. To determine which interaction is active, you need to run additional trials. If the factors had been entered in a different order, the Ethanol*Time interaction might have been aliased with two other two-factor interactions.

In the section [“Definitive Screening Design”](#) on page 242, you constructed a 17-run definitive screening design. The Color Map on Correlations for this DSD (Figure 7.4) shows that no two-factor interactions are confounded with any other two-factor interactions. For the fractional factorial design, there are seven instances of confounded two-factor interactions. If you suspect that there are active two-factor effects, the DSD is the better choice.

You can conduct a more thorough comparison of the two designs using the Compare Designs platform (DOE > Design Diagnostics > Compare Designs). See [Chapter 16, “Compare Designs”](#).

Definitive Screening Design with Blocking

Suppose that, due to raw material constraints, the extraction experiment requires that you run it using material from two separate lots. You can generate a definitive screening design with a blocking variable to account for the potential lot variation.

Create the Design

The extraction factors and their settings are given in the data table *Extraction Factors.jmp*. Generate a definitive screening design with a block as follows.

1. Select **DOE > Definitive Screening > Definitive Screening Design**.
2. Double-click Y under Response Name and type Yield.
3. Select **Help > Sample Data Library** and open *Design Experiment/Extraction Factors.jmp*.
4. Click the red triangle next to Definitive Screening Design and select **Load Factors**.

The factor names and ranges are added to the Factors outline.

5. Click **Continue**.

The Design Options outline opens. Here you can specify a blocking structure.

6. Select **Add Blocks with Center Runs to Estimate Quadratic Effects**.

Leave **Number of Blocks** set at 2.

You are recreating the design for the *Extraction2 Data.jmp* sample data table, which was created without Extra Runs. Although four Extra Runs are strongly recommended, you will not add Extra Runs in this example.

7. Next to **Number of Extra Runs**, select 0.
8. Click **Make Design**.

The Definitive Screening Design window updates to show a Design outline and a Design Evaluation outline.

Check that Block has been added to the Factors outline and to the Design.

9. In the Factors outline, Double-click Block and type Lot.

In the Output Options panel, note that the Run Order is set to Randomize within Blocks.

10. Click **Make Table**.

Note: The runs in your design might appear in a different order than the order shown in Figure 7.8.

Figure 7.8 Definitive Screening Design with Block for Extraction Process

	Lot	Methanol	Ethanol	Propanol	Butanol	pH	Time	Yield
1	1	5	5	5	5	7.5	1.5	•
2	1	5	10	10	10	9	2	•
3	1	5	0	0	0	6	1	•
4	1	10	0	5	0	9	2	•
5	1	0	0	0	10	7.5	2	•
6	1	0	10	5	10	6	1	•
7	1	10	10	10	0	7.5	1	•
8	2	10	5	0	10	9	1	•
9	2	0	10	0	0	9	1.5	•
10	2	0	0	10	5	9	1	•
11	2	10	0	10	10	6	1.5	•
12	2	0	5	10	0	6	2	•
13	2	10	10	0	5	6	2	•
14	2	5	5	5	5	7.5	1.5	•

Notice that run 1 is a center point run in Lot 1 and run 14 is a center point run in Lot 2.

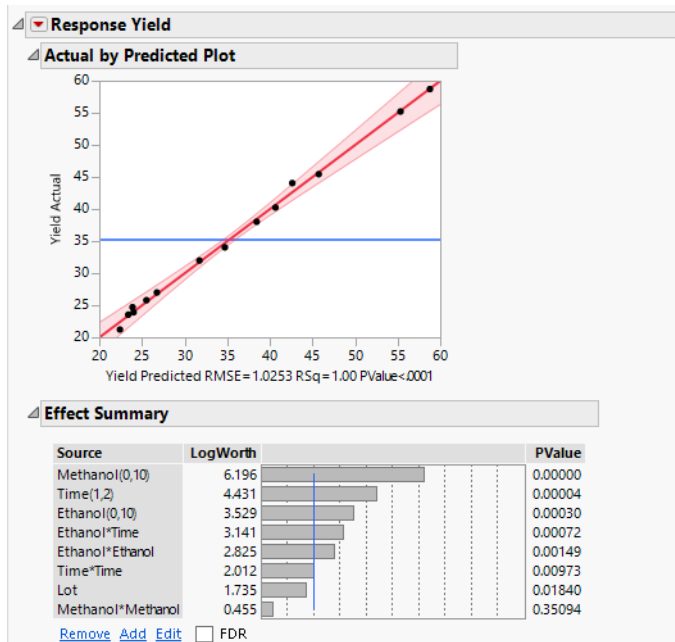
Analyze the Experimental Data

At this point, you conduct your experiment and record your data in the Yield column of the design table (Figure 7.8). The Extraction2 Data.jmp sample data table contains your experimental results. The runs in the Extraction2 Data.jmp sample data table are in a different order than those in Figure 7.8.

To explore all second-order effects, one option is to use All Possible Models regression. Another option is to use forward stepwise regression. However, these standard methods often fail to identify active effects. For this reason, you use the Fit Definitive Screening platform.

1. Select **Help > Sample Data Library** and open Design Experiment/Extraction2 Data.jmp.
2. In the Table panel of the design table, click the green triangle next to the **Fit Definitive Screening** script.

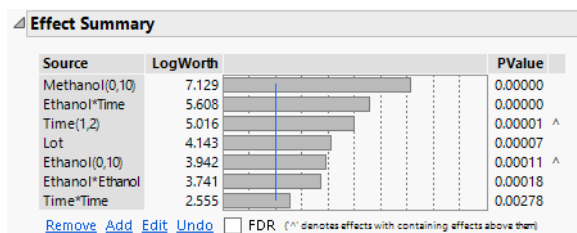
Figure 7.10 Standard Least Squares Report



The Actual by Predicted Plot shows no evidence of lack of fit. The Effect Summary report shows that Methanol*Methanol is not significant. You decide to remove this effect from the model.

4. Select Methanol*Methanol in the Effect Summary report and click **Remove**.

Figure 7.11 Final Set of Active Effects



The remaining effects are all significant. You conclude that these are the active effects.

Comparison of a Definitive Screening Design with a Plackett-Burman Design

Plackett-Burman designs are an alternative to fractional factorials for screening. However, Plackett-Burman designs have complex aliasing of the main effects by two-factor interactions.

This example shows how to compare a definitive screening with a Plackett-Burman design using the Evaluate Design platform. For an extensive example using the Compare Designs platform, see “[Designs of Same Run Size](#)” on page 471 in the “Compare Designs” chapter.

The Definitive Screening Design

- 1. Select **DOE > Definitive Screening > Definitive Screening Design**.
- 2. Type **4** in the **Add N Factors** box and click **Continuous**.
- 3. Type **2** in the **Add N Factors** box and click **Categorical**.

Your window should appear as shown in Figure 7.12.

Figure 7.12 Definitive Screening Dialog with 4 Continuous and 2 Categorical Factors

Definitive Screening Design

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.
optional item				

Factors

Continuous Categorical Remove Add N Factors 2

Name	Role	Values
X1	Continuous	-1 1
X2	Continuous	-1 1
X3	Continuous	-1 1
X4	Continuous	-1 1
X5	Categorical	L1 L2
X6	Categorical	L1 L2

Specify Factors

Add a Continuous or Categorical factor by clicking its button. Double click on a factor name or level to edit it.

Continue

- 4. Click **Continue**.

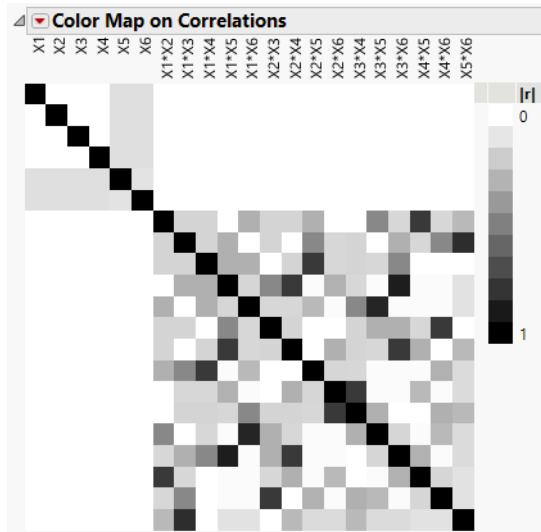
This example does not require a block. Under the Design Options Outline, check that the No Blocks Required option is selected.

In order to compare designs of approximately equal sizes. do not add Extra Runs.

- 5. Next to **Number of Extra Runs**, select 0.

6. Click **Make Design**.
The design that is generated has 14 runs.
7. Open the **Design Evaluation > Color Map On Correlations** outline.

Figure 7.13 Color Map for Definitive Screening Design



Notice that the categorical main effects have small correlations with each other and with the continuous factors' main effects. These correlations lead to a small reduction in the precision of the estimates.

8. Do not close your Definitive Design Screening window until you compare the color map with that of the Plackett-Burman design, below.

The Plackett-Burman Design

Now create a Plackett-Burman design using the same factor structure.

1. Select **DOE > Classical > Two Level Screening > Screening Design**.
2. Type 4 in the **Add N Factors** box and click **Continuous**.
3. Type 2 in the **Add N Factors** box and click **Categorical > 2 Level**.
4. Click **Continue**.
5. Select **Choose from a list of fractional factorial designs** and click **Continue**.

Potential designs appear in the Design List.

Figure 7.14 Plackett-Burman Design

Screening Design

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.

optional item

Factors

Continuous Discrete Numeric Categorical Remove Add N Factors 2

Name	Role	Values
X1	Continuous	-1 1
X2	Continuous	-1 1
X3	Continuous	-1 1
X4	Continuous	-1 1
X5	Categorical	L1 L2
X6	Categorical	L1 L2

Design List

Choose a design by clicking on its row in the list.

Number Of Runs	Block Size	Design Type	Resolution - what is estimable
8		Fractional Factorial	3 - Main Effects Only
8	4	Fractional Factorial	3 - Main Effects Only
12		Plackett-Burman	3 - Main Effects Only
16		Fractional Factorial	4 - Some 2-factor interactions
16	8	Fractional Factorial	4 - Some 2-factor interactions
16	4	Fractional Factorial	4 - Some 2-factor interactions
16	2	Fractional Factorial	4 - Some 2-factor interactions
32		Fractional Factorial	5+ - All 2-factor interactions
32	16	Fractional Factorial	5+ - All 2-factor interactions
32	8	Fractional Factorial	4 - Some 2-factor interactions
32	4	Fractional Factorial	4 - Some 2-factor interactions
32	2	Fractional Factorial	4 - Some 2-factor interactions
64		Full Factorial	>6 - Full Resolution
64	32	Full Factorial	5+ - All 2-factor interactions

6. Select the 12 run Plackett-Burman design (Figure 7.14).
7. Click **Continue**.
8. Open the **Design Evaluation > Color Map On Correlations** outline.
Compare the color map for the 12-run Plackett-Burman design to the color map for the 14-run definitive screening design.

Figure 7.15 Plackett-Burman Correlations (left) and Definitive Screening Correlations (right)

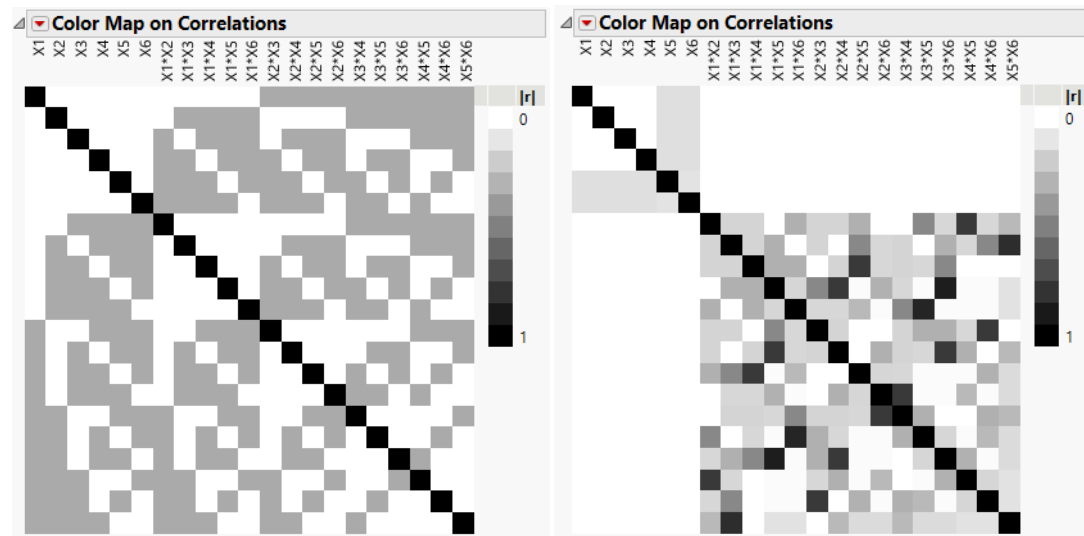


Figure 7.15 shows both color maps, but shows only the portion of the Plackett-Burman color map that involves main effects and two-way interactions. (To construct the color map for the Plackett-Burman design without the three-way interactions, construct the design. Then obtain the color map using Evaluate Design.)

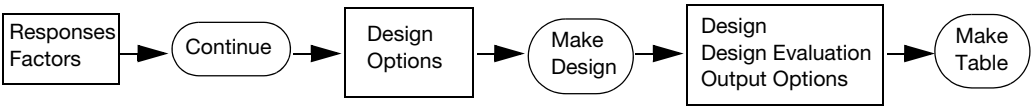
In the color map for the Plackett-Burman design on the left, you see that most two-factor interactions are correlated with main effects. This means that any non-negligible two-factor interaction will bias several main effects. This can lead to a failure to identify an active main effect or the false conclusion that an inactive main effect is active.

Contrast this with the color map for the definitive screening design on the right. With only two additional runs, the definitive screening design trades off a small increase in the variance of the main effects for complete independence of main effects and two-factor interactions.

Definitive Screening Design Window

The definitive screening design window updates as you work through the design steps. See [“The DOE Workflow: Describe, Specify, Design”](#) on page 57 in the “Starting Out with DOE” chapter. The outlines, separated by buttons that update the outlines, follow the flow in Figure 7.16.

Figure 7.16 Definitive Screening Design Flow



Responses

Use the Responses outline to specify one or more responses.

Tip: When you have completed the Responses outline, consider selecting **Save Responses** from the red triangle menu. This option saves the response names, goals, limits, and importance values in a data table that you can later reload in DOE platforms.

Figure 7.17 Responses Outline

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.

optional item

Add Response Enters a single response with a goal type of Maximize, Match Target, Minimize, or None. If you select Match Target, enter limits for your target value. If you select Maximize or Minimize, entering limits is not required but can be useful if you intend to use desirability functions.

Remove Removes the selected responses.

Number of Responses Enters additional responses so that the number that you enter is the total number of responses. If you have entered a response other than the default Y, the Goal for each of the additional responses is the Goal associated with the last response entered. Otherwise, the Goal defaults to Match Target. Click the Goal type in the table to change it.

The Responses outline contains the following columns:

Response Name The name of the response. When added, a response is given a default name of Y, Y2, and so on. To change this name, double-click it and enter the desired name.

Goal, Lower Limit, Upper Limit The Goal tells JMP whether you want to maximize your response, minimize your response, match a target, or that you have no response goal. JMP assigns a Response Limits column property, based on these specifications, to each response column in the design table. It uses this information to define a desirability

function for each response. The Profiler and Contour Profiler use these desirability functions to find optimal factor settings. For further details, see the Profiler chapter in *Profilers* and “[Response Limits](#)” on page 732 in the “Column Properties” appendix.

- A Goal of Maximize indicates that the best value is the largest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Minimize indicates that the best value is the smallest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Match Target indicates that the best value is a specific target value. The default target value is assumed to be midway between the Lower Limit and Upper Limit.
- A Goal of None indicates that there is no goal in terms of optimization. No desirability function is constructed.

Note: If your target response is not midway between the Lower Limit and the Upper Limit, you can change the target after you generate your design table. In the data table, open the Column Info window for the response column (**Cols > Column Info**) and enter the desired target value.

Importance When you have several responses, the Importance values that you specify are used to compute an overall desirability function. These values are treated as weights for the responses. If there is only one response, then specifying the Importance is unnecessary because it is set to 1 by default.

Editing the Responses Outline

In the Responses outline, note the following:

- Double-click a response to edit the response name.
- Click the goal to change it.
- Click on a limit or importance weight to change it.
- For multiple responses, you might want to enter values for the importance weights.

Response Limits Column Property

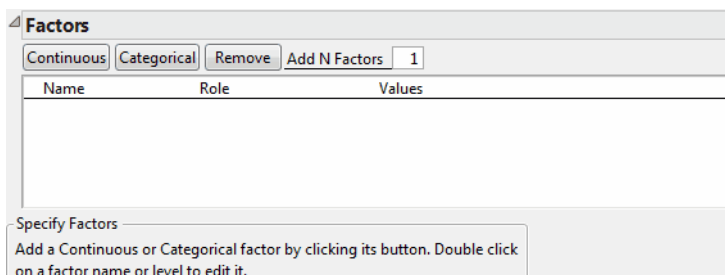
The Goal, Lower Limit, Upper Limit, and Importance that you specify when you enter a response are used in finding optimal factor settings. For each response, the information is saved in the generated design data table as a Response Limits column property. JMP uses this information to define the desirability function. The desirability function is used in the Prediction Profiler to find optimal factor settings. For further details about the Response Limits column property and examples of its use, see [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

If you do not specify a Lower Limit and Upper Limit, JMP uses the range of the observed data for the response to define the limits for the desirability function. Specifying the Lower Limit and Upper Limit gives you control over the specification of the desirability function. For more details about the construction of the desirability function, see the Profiler chapter in *Profilers*.

Factors

Add factors in the Factors outline.

Figure 7.18 Factors Outline



The factors outline contains the following buttons.

Continuous Enters the number of continuous factors specified in **Add N Factors**.

Categorical Enters the number of nominal factors specified in **Add N Factors**.

Remove Removes the selected factors.

Add N Factors Adds multiple factors of a given type. Enter the number of factors to add and click Continuous or Categorical. Repeat **Add N Factors** to add multiple factors of different types.

Tip: When you have completed your Factors panel, select Save Factors from the red triangle menu. This saves the factor names and values in a data table that you can later reload. See [“Definitive Screening Design Options”](#) on page 263.

The Factors outline contains the following columns:

Name The name of the factor. When added, a factor is given a default name of X1, X2, and so on. To change this name, double-click it and enter the desired name.

Role Specifies the Design Role of the factor. The Design Role column property for the factor is saved to the data table. This property ensures that the factor type is modeled appropriately.

Values The experimental settings for the factors. To insert Values, click the default values and type the desired values.

Editing the Factors Outline

In the Factors outline, note the following:

- To edit a factor name, double-click the factor name.
- To edit a value, click the value in the Values column.

Factor Types

Continuous Numeric data types only. A continuous factor is a factor that you can conceptually set to any value between the lower and upper limits you supply, given the limitations of your measurement system.

Categorical Either numeric or character data types with two levels. For a categorical factor, the value ordering is the order of the values, as entered from left to right. This ordering is saved in the Value Order column property after the design data table is created.

Factor Column Properties

For each factor, various column properties are saved to the data table.

Design Role Each factor is assigned the Design Role column property. The Role that you specify in defining the factor determines the value of its Design Role column property. When you add a block under Design Options, that factor is assigned the Blocking value. The Design Role property reflects how the factor is intended to be used in modeling the experimental data. Design Role values are used in the Augment Design platform.

Factor Changes Each factor is assigned the Factor Changes column property with a setting of Easy. In definitive screening designs, it is assumed that factor levels can be changed for each experimental run. Factor Changes values are used in the Evaluate Design and Augment Design platforms.

Coding If the Design Role is Continuous, the Coding column property for the factor is saved. This property transforms the factor values so that the low and high values correspond to –

1 and +1, respectively. The estimates and tests in the Fit Least Squares report are based on the transformed values.

Value Order If the Design Role is Categorical or Blocking, the Value Order column property for the factor is saved. This property determines the order in which levels of the factor appear.

RunsPerBlock Indicates the number of runs in each block. When a Block is selected in the Design Options outline and you then click Make Design, a factor with the default name Block is added to the Factors list. The RunsPerBlock column property is saved for that factor.

Design Options

The Design Options outline enables you to specify the blocking structure, the number of blocks, and the number of extra runs. Block effects are orthogonal to the main effects. Block sizes need not be equal.

The outline contains the following options:

No Blocks Required Indicates that the design will not contain a blocking factor. This is the default selection.

Add Blocks with Center Runs to Estimate Quadratic Effects Adds the number of blocks specified in the Number of Blocks text box. Constructs a design where block effects are orthogonal to main effects and where the model consisting of all main and quadratic effects is estimable. See [“Add Blocks with Center Runs to Estimate Quadratic Effects”](#) on page 259

Add Blocks without Extra Center Runs Adds the number of blocks specified in the Number of Blocks text box. Adds only as many center runs as required by the design structure. Constructs a design where block effects are orthogonal to main effects, but the model consisting of all main effects and quadratic effects might *not* be estimable. See [“Add Blocks without Extra Center Runs”](#) on page 260.

Note: Use the **Add Blocks without Extra Center Runs** option only if you can assume that not all quadratic effects are important.

Number of Blocks Indicates the number of blocks to add. The number of blocks that you can add ranges from two to the number of factors.

Number of Extra Runs Adds non-center runs that enable you to conduct *effective model selection*. See [“Extra Runs”](#) on page 269 and [“Effective Model Selection for DSDs”](#) on page 275 in the “The Fit Definitive Screening Platform” chapter.

Tip: Adding runs to your design with the Extra Runs option enhances your ability to detect effects in the presence of many active effects. The recommended number of Extra Runs is four, which dramatically improves the power of the design to identify active second-order effects.

Make Design Generates the design, presents it in the Design outline, and provides evaluation information in the Design Evaluation outline. The Output Options panel also appears, enabling you to create the design table.

Blocking in Definitive Screening Designs

This section describes the two blocking options:

- “[Add Blocks with Center Runs to Estimate Quadratic Effects](#)” on page 259
- “[Add Blocks without Extra Center Runs](#)” on page 260

If a design contains a categorical factor, a *center run* is a run where all continuous factors are set at their middle values. If all factors are categorical, the two blocking options are available. Both options produce designs whose blocks are orthogonal to main effects.

Add Blocks with Center Runs to Estimate Quadratic Effects

Note: For more information about the construction and properties of blocked designs that estimate quadratic effects, see Jones and Nachtsheim (2016). The paper also contains information about treating the blocks as random effects.

The **Add Blocks with Center Runs to Estimate Quadratic Effects** option constructs a design with these properties:

- Block effects are orthogonal to main effects.
- The model consisting of all main and quadratic effects is estimable.

If a design contains only continuous factors, a blocked design for k factors having these properties can be constructed as follows:

- Remove the center run from the DSD design for k factors.
- Assign conference matrix foldover pairs to the same block.
- Add one center run to each block.

When some factors are categorical, the **Add Blocks with Center Runs to Estimate Quadratic Effects** option adds pairs of center runs within certain blocks. This structure ensures orthogonality and the ability to estimate all main and quadratic effects.

Because the only requirement on block size is that a block contains a foldover pair, the number of blocks can range from 2 to k , if k is even and from 2 to $k+1$, if k is odd. See [“Conference Matrices and the Number of Runs”](#) on page 269. JMP attempts to construct blocks of equal size.

Add Blocks without Extra Center Runs

The **Add Blocks without Extra Center Runs** option constructs a design that has a single center run when all factors are continuous and two center runs when some factors are categorical. The resulting design has these properties:

- Block effects are orthogonal to main effects.
- Block effects might be confounded with a linear combination of quadratic effects. This implies that the model consisting of all main and quadratic effects might *not* be estimable.

For this reason, use this option only if you can assume that some quadratic effects are negligible.

A blocked design for k factors without extra center runs is constructed as follows:

- Assign conference matrix foldover pairs to the same block.
- If all factors are continuous, assign the single center run to a single block.
- If there are categorical factors, the unblocked definitive screening design requires the addition of two center runs to the foldover pairs defined by the conference matrix. See [“Conference Matrices and the Number of Runs”](#) on page 269. To construct the blocked design without extra center runs, these two center runs are added to a single block.

Because the only requirement on block size is that a block contains a foldover pair, the number of blocks can range from 2 to k , if k is even and from 2 to $k+1$, if k is odd. See [“Conference Matrices and the Number of Runs”](#) on page 269. JMP attempts to construct blocks of equal size.

Design

The Design outline shows the runs for the definitive screening design. The runs are given in a standard order. To change the run order for your design table, you can select Run Order options in the Output Options panel before generating the table.

Note: Definitive screening designs for four or fewer factors are based on a five-factor design. See [“Definitive Screening Designs for Four or Fewer Factors”](#) on page 270.

Design Evaluation

Note: The Design Evaluation outline is not shown for Cotter designs.

The Design Evaluation outline provides a number of ways to evaluate the properties of the generated design. Open the Design Evaluation outline to see the following options:

Power Analysis Enables you to explore your ability to detect effects of given sizes.

Prediction Variance Profile Shows the prediction variance over the range of factor settings.

Fraction of Design Space Plot Shows how much of the model prediction variance lies below (or above) a given value.

Prediction Variance Surface Shows a surface plot of the prediction variance for any two continuous factors.

Estimation Efficiency For each parameter, gives the fractional increase in the length of a confidence interval compared to that of an ideal (orthogonal) design, which might not exist. Also gives the relative standard error of the parameters.

Alias Matrix Gives coefficients that indicate the degree by which the model parameters are biased by effects that are potentially active, but not in the model.

Color Map on Correlations Shows the absolute correlation between effects on a plot using an intensity scale.

Design Diagnostics Indicates the optimality criterion used to construct the design. Also gives efficiency measures for your design.

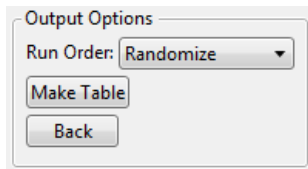
Note: The model used for the design diagnostics contains all main effects and two-factor interactions when all two-factor interactions are estimable. Otherwise, the model contains all main effects.

For more details about the Design Evaluation panel, see [“Design Evaluation”](#) on page 449 in the “Evaluate Designs” chapter.

Output Options

Use the Output Options panel to perform the following tasks:

- specify the order for the runs in the design data table
- construct the design table
- return to a previous point in the Definitive Screening Design window

Figure 7.19 Output Options Panel


The Output Options panel contains these options:

Run Order

The **Run Order** options determine the order of the runs in the design table. Choices include the following:

Keep the Same Rows in the design table are in the same order as in the Design outline.

Sort Left to Right Columns in the design table are sorted from left to right.

Randomize Rows in the design table are in random order.

Sort Right to Left Columns in the design table are sorted from right to left.

Randomize within Blocks Rows in the design table are in random order within the blocks.

Make Table

Click **Make Table** to construct the Definitive Screening Design data table.

In the Definitive Screening Design table, the Table panel (in the upper left) contains the following scripts. To run a script, click the green triangle next to the script name.

Fit Definitive Screening Runs the DOE > Definitive Screening > Fit Definitive Screening platform. See [Chapter 8, “The Fit Definitive Screening Platform”](#).

Evaluate Design Runs the DOE > Design Diagnostics > Evaluate Design platform. See [Chapter 15, “Evaluate Designs”](#).

DOE Dialog Re-creates the Definitive Screening Design window that you used to generate the design table.

Figure 7.20 Definitive Screening Design Table Showing Scripts

▼ Definitive Screening...										
Design Definitive Screening										
▶ Fit Definitive Screening										
▶ Evaluate Design										
▶ DOE Dialog										
▼ Columns (8/0)										
■ Block *										
■ X1 *										
■ X2 *										
■ X3 *										
■ X4 *										
■ X5 *										
■ X6 *										
■ Y *										
▼ Rows										
All rows		18								
Selected		0								
Excluded		0								
Hidden		0								

Back

The Back button takes you back to where you were before clicking Make Design. You can make changes to the previous outlines and regenerate the design.

Definitive Screening Design Options

The red triangle menu in the Definitive Screening Design platform contains these options:

Save Responses Saves the information in the Responses panel to a new data table. You can then quickly load the responses and their associated information into most DOE windows. This option is helpful if you anticipate reusing the responses.

Load Responses Loads responses that you saved using the Save Responses option.

Save Factors Saves the information in the Factors panel to a new data table. Each factor's column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Note: It is possible to create a factors table by entering data into an empty table, but remember to assign each column an appropriate Design Role. Do this by right-clicking on the column name in the data grid and selecting **Column Properties > Design Role**. In the Design Role area, select the appropriate role.

Load Factors Loads factors that you saved using the Save Factors option.

Save Constraints (Unavailable for some platforms) Saves factor constraints that you defined in the Define Factor Constraints or Linear Constraints outline into a data table, with a column for each constraint. You can then quickly load the constraints into most DOE windows.

In the constraint table, the first rows contain the coefficients for each factor. The last row contains the inequality bound. Each constraint's column contains a column property called `ConstraintState` that identifies the constraint as a "less than" or a "greater than" constraint. See "[ConstraintState](#)" on page 766 in the "Column Properties" appendix.

Load Constraints (Unavailable for some platforms) Loads factor constraints that you saved using the Save Constraints option.

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction
- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.
- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click Apply, a formula for the simulated response values is saved in a

new column called <Y> Simulated, where Y is the name of the response. Clicking Apply again updates the formula and values in <Y> Simulated.

For additional details, see [“Simulate Responses”](#) on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Save X Matrix Saves the Moments Matrix and Model Matrix to table scripts in the design data table. These scripts contain the moments and design matrices. See [“Save X Matrix”](#) on page 112 in the “Custom Designs” chapter.

Caution: For a design with nominal factors, the Model Matrix saved by the Save X Matrix option is *not* the coding matrix used in fitting the linear model. You can obtain the coding matrix used for fitting the model by selecting the option Save Columns > Save Coding Table in the Fit Model report that you obtain when you run the Model script

Advanced Options > Set Delta for Power Specify the difference in the mean response that you want to detect for model effects. See [“Evaluate Design Options”](#) on page 467 in the “Evaluate Designs” chapter.

Save Script to Script Window Creates the script for the design that you specified in the Definitive Screening Design window and saves it in an open script window.

Simulate Responses

When you click Make Table to create your design table, the Simulate Responses option does the following for each response:

- It adds random response values to the response column in your design table.
- It adds a new column containing a simulation model formula to the design table. The formula and values are based on a main effects model.

A Model window opens where you can add and remove effects to define a model, specify parameter values, and select a response distribution for simulation. When you click Apply in the Model window, each column containing a simulation model formula is updated.

Control Window

Figure 7.21 shows the Model window for a design with two continuous factors (X1 and X2) and one two-level categorical factor (X3).

Figure 7.21 Simulate Responses Control Window

Factors

X1
X2
X3
optional

Interactions ▼ RSM Powers ▼

Simulate Responses

Effects	Y
Intercept	1
X1	1
X2	1
X3 1	1

Reset Coefficients Remove Term

Distribution

☒ Normal Error σ : 1
☐ Binomial
☐ Poisson

Apply

The window has three outlines:

- Factors
- Simulate Responses
- Distribution

The initial Simulate Responses outline shows terms for a main effects model with values of 1 for all coefficients. The Distribution outline shows a Normal distribution with error standard deviation equal to 1. If you have set Anticipated Coefficients as part of Power Analysis under Design Evaluation in the DOE window, then the initial values in the Simulate Responses outline are the values that you specified as Anticipated Coefficients and Anticipated RMSE (Error Std) in the Power Analysis outline.

Factors

Add terms to the simulation model using the Factors outline.

Interactions Select factors in the list. Then select the interaction order from the Interactions menu. Those interactions are added to the list of Effects in the Simulate Responses outline.

RSM Adds all possible response surface terms to the list of Effects in the Simulate Responses outline.

Powers Select factors in the list. Then select the order from the Powers menu. Those powers are added to the list of Effects in the Simulate Responses outline.

Simulate Responses

To specify a model for simulated values, do the following:

1. For each term in the list of Effects, enter coefficients for the linear model used to simulate the response values. These define a linear function, $L(\mathbf{x}, \beta) = \mathbf{x}'\beta$. See the Simulate Responses outline in Figure 7.21:
 - The vector $\mathbf{x}$ consists of the terms that define the effects listed under Effects.
 - The vector β is the vector of model coefficients that you specify under Y.
2. Under Distribution, select a response distribution.
3. Click **Apply**. A $\langle Y \rangle$ Simulated column containing simulated values and their formula is added to the design table, where Y is the name of the response column.

Reset coefficients Sets all coefficients to 0.

Remove Term Remove terms from the list of Effects. Select the effects to remove and click Remove Term. Note that you cannot remove main effects.

Distribution

Choose from one of the following distributions in the Simulate Responses window:

Normal Simulates values from a normal distribution. Enter a value for Error σ , the standard deviation of the normal error distribution. If you have designated factors to have Changes of Hard in the Factors outline, you can enter a value for Whole Plots σ , the whole plot error. If you have designated factors to have Changes of Hard and Very Hard, you can enter values for both the subplot and whole plot errors. When you click Apply, random values and a formula containing a random response vector based on the model are entered in the column $\langle Y \rangle$ Simulated.

Binomial Simulates values from a binomial distribution. Enter a value for N, the number of trials. Random integer values are generated according to a binomial distribution based on N trials with probability of success $1/(1 + \exp(-L(\mathbf{x}, \beta)))$. When you click Apply, random values and their formula are entered in the column $\langle Y \rangle$ Simulated. A column called N Trials that contains the value N is also added to the data table.

Poisson Simulates random integer values according to a Poisson distribution with parameter $\exp(L(\mathbf{x}, \beta))$. When you click Apply, random values and their formula are entered in the column $\langle Y \rangle$ Simulated.

Note: You can set a preference to simulate responses every time you click Make Table. To do so, select **File > Preferences > Platforms > DOE**. Select **Simulate Responses**.

Technical Details

- [“Structure of Definitive Screening Designs”](#)
- [“Analysis of Experimental Data”](#)

Structure of Definitive Screening Designs

Figure 7.22 shows an example of a definitive design with eight continuous factors, and four Extra Runs that correspond to fake factors. Notice the following:

- Each pair of rows is a foldover pair; each even-numbered row is -1 times the previous row. The foldover aspect of the design removes the confounding of two-factor interactions and main effects.
- Each factor is set at its center value for three runs; this, together with the design’s construction, makes all quadratic effects estimable.
- Rows 17 through 20 are the Extra Runs that correspond to the fake factors.
- Adding the center run in the last row enables you to fit a model that includes an intercept and all main and quadratic effects.

This structure is typical of definitive screening designs for continuous factors.

Figure 7.22 Definitive Screening Design for Eight Continuous Factors

Design								
Run	X1	X2	X3	X4	X5	X6	X7	X8
1	0	1	1	1	1	1	1	1
2	0	-1	-1	-1	-1	-1	-1	-1
3	1	0	-1	-1	-1	-1	1	1
4	-1	0	1	1	1	1	-1	-1
5	1	-1	0	-1	1	1	-1	-1
6	-1	1	0	1	-1	-1	1	1
7	1	-1	-1	0	1	1	1	1
8	-1	1	1	0	-1	-1	-1	-1
9	1	-1	1	1	0	-1	-1	1
10	-1	1	-1	-1	0	1	1	-1
11	1	-1	1	1	-1	0	1	-1
12	-1	1	-1	-1	1	0	-1	1
13	1	1	-1	1	-1	1	0	-1
14	-1	-1	1	-1	1	-1	0	1
15	1	1	-1	1	1	-1	-1	0
16	-1	-1	1	-1	-1	1	1	0
17	1	1	1	-1	-1	1	-1	1
18	-1	-1	-1	1	1	-1	1	-1
19	1	1	1	-1	1	-1	1	-1
20	-1	-1	-1	1	-1	1	-1	1
21	0	0	0	0	0	0	0	0

Conference Matrices and the Number of Runs

Definitive screening designs in JMP are constructed using *conference matrices* (Xiao et al., 2012). A conference matrix is an $m \times m$ matrix C where m is even. The matrix C has 0s on the diagonal, off-diagonal entries equal to 1 or -1 , and satisfies

$$C'C = (m-1)I_{m \times m}.$$

Note: For certain even values of m , it is not known if a conference matrix exists.

Suppose that the number of factors, k , is five or larger. For the case of $k \leq 4$ factors, see “[Definitive Screening Designs for Four or Fewer Factors](#)” on page 270.

Consider the case of k continuous factors and suppose that a conference matrix is available.

- When k is even, the $k \times k$ conference matrix is used to define k runs of the design. Its negative, $-C$, defines the foldover runs. A center point is added to the design to ensure that a model containing an intercept, main effects, and quadratic effects is estimable. So, for k even, the minimum number of runs in the definitive screening design is $2k + 1$.
- When k is odd, a $(k+1) \times (k+1)$ conference matrix is used, with its last column deleted. A center point is added. Thus, for k odd, the minimum number of runs in the screening design is $2k + 3$.

A similar procedure is used when some factors are categorical and a conference matrix is available. See Jones and Nachtsheim (2013).

- Instead of a single center point, two additional runs are required. These two runs are center runs where all continuous factors are set at their middle values.
- When there are k factors and k is even, the number of runs in the design is $2k + 2$.
- When k is odd, the number of runs is $2k + 4$.

For those values of m for which a conference matrix is not available, a definitive screening design can be constructed using the next largest conference matrix. As a result, the required number of runs might exceed $2k + 3$, in the continuous case, and $2k + 4$, in the categorical case.

Extra Runs

Extra runs are constructed using fictitious, or *fake*, factors. Adding k_1 fake factors to a design results in 2^{k_1} additional runs.

Denote the number of factors in your experimental study by k . Extra runs are constructed by creating a design for $k + k_1$ factors, as described in “[Conference Matrices and the Number of Runs](#)” on page 269, and then dropping the last k_1 columns. If $k_1 = 2$, four extra runs are added. If $k_1 = 3$, eight extra runs are added. As few as four extra runs can be highly beneficial in model selection.

For information about how extra runs are used, see [“Effective Model Selection for DSDs”](#) on page 275 in the “The Fit Definitive Screening Platform” chapter.

Definitive Screening Designs for Four or Fewer Factors

Definitive screening designs for four or fewer factors are constructed using the five-factor definitive screening design as a base. This is because designs for $k \leq 4$ factors constructed strictly according to the conference matrix approach have undesirable properties. In particular, it is difficult to separate second-order effects.

If you specify $k \leq 4$ factors, a definitive screening design for five factors is constructed and unnecessary columns are dropped. For this reason, the number of runs for an unblocked design with $k \leq 4$ factors is 13 if all factors are continuous or 14 if some factors are categorical.

Analysis of Experimental Data

In general, you want to fit a model that permits the possibility that two-way interactions are active. You also might want to include pure quadratic terms in your model. You might want to postulate a full second-order model, or you might want to specify an a priori model containing only certain second-order terms.

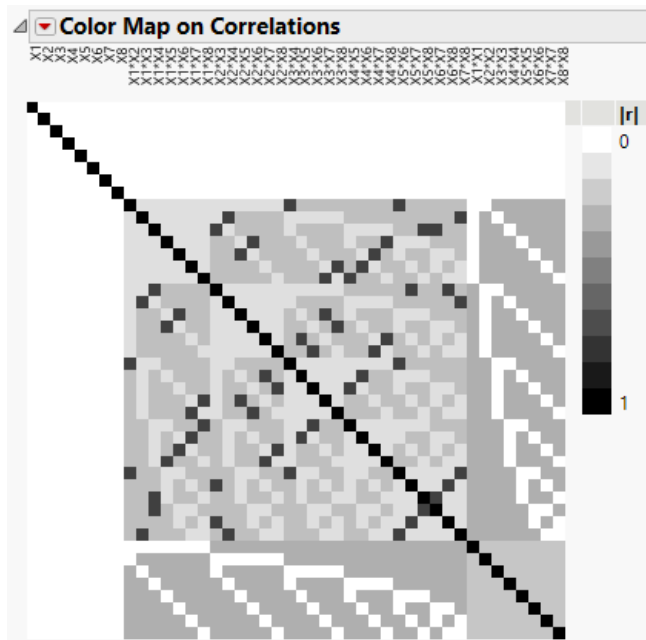
Two-Way Interactions

In fitting such a model, you need to be mindful of two facts:

- two-way interaction effects and quadratic effects are often correlated
- two-way interaction and quadratic effects cannot *all* be estimated simultaneously

Figure 7.23 shows a Color Map on Correlations for the design with eight continuous factors shown in Figure 7.22. The color map is for a full quadratic design. The eight pure quadratic effects are listed to the far right. You can construct this plot by using DOE > Design Diagnostics > Evaluate Design and entering the appropriate terms into the Alias Terms list. See [“Alias Terms”](#) on page 448 in the “Evaluate Designs” chapter.

Figure 7.23 Color Map on Correlations for Full Quadratic Model



Use your cursor to hold your mouse pointer over the cells of the color map in order to see the absolute correlations between effects. You see that main effects are uncorrelated with all two-way interaction and pure quadratic effects. You also see that none of the effects are completely confounded with other effects because the only red cells are on the main diagonal. But note that some of the absolute correlations between two-factor interactions are substantial, with some at 0.75. Note also that absolute correlations between two-factor interactions and pure quadratic effects are either 0 or 0.3118.

If only main and pure quadratic effects are active, you can fit a saturated model that contains main effects and quadratic effects. This model will result in effect estimates that are unbiased, assuming no active three-way or higher order effects.

Because of the correlations involving second-order effects, you must be careful in fitting a model with two-way interactions. Analysis methodologies include the following, where the first is preferred:

- The method of Efficient Model Selection performs well, especially if many effects are likely to be active. See [“Effective Model Selection for DSDs”](#) on page 275 in the “The Fit Definitive Screening Platform” chapter.
- Forward stepwise or all possible subsets regression performs adequately if the following conditions hold:
 - The number of active effects is no more than half the number of runs.

- There are at most two active two-way interactions or at most one active quadratic effect.

See [“Forward Stepwise Regression or All Possible Subsets Regression”](#) on page 272.

Forward Stepwise Regression or All Possible Subsets Regression

This method consists of first specifying a full response surface model. Then do one of the following:

- Use forward stepwise regression with the Stopping Rule set to Minimum AICc and the Rules set to Combine to ensure model heredity.
- Use All Possible Models regression, where you select the option that imposes the heredity restriction and use the AICc criterion for model selection.

You cannot fit the full response surface model because the number of runs is less than the number of parameters. So your analysis depends on the assumption of effect sparsity, where you assume that the number of active effects is less than the number of runs. This approach has some limitations:

- If the number of active effects exceeds half the number of runs, both stepwise and all possible models regression have difficulty finding the correct model.
- The power of tests to detect moderate quadratic effects is low. A quadratic effect must exceed three times the error standard deviation for the power to exceed 0.9.
- Because of effect confounding, several models might be equivalent. Additional runs will be necessary to resolve the confounding.

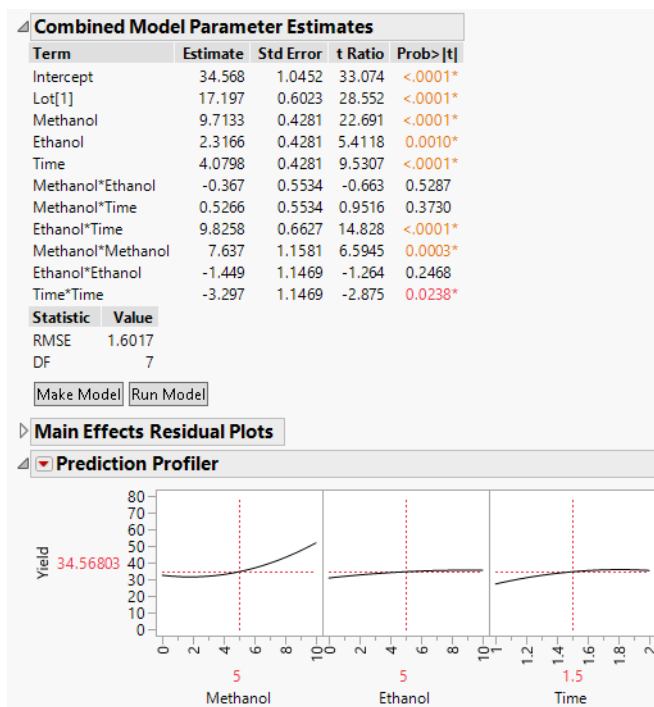
The Fit Definitive Screening Platform

Analyze Data from Definitive Screening Experiments

Use the Fit Definitive Screening platform to analyze definitive screening designs (DSDs) using a methodology called *Effective Model Selection for DSDs*. This methodology takes advantage of the unique structure of definitive screening designs.

Standard model selection methods applied to DSDs can fail to identify active effects. To identify active main effects and second-order effects, the Effective Model Selection for DSDs algorithm leverages the structure of DSDs.

Figure 8.1 Fit Definitive Screening Results



Contents

Overview of the Fit Definitive Screening Platform	275
Identification of Active Effects in DSDs	275
Effective Model Selection for DSDs	275
Example of the Fit Definitive Screening Platform	276
Fit the Model	276
Examine Results	277
Reduce the Model	279
Launch the Fit Definitive Screening Platform	280
Fit Definitive Screening Report	281
Stage 1 - Main Effect Estimates	281
Stage 2 - Even Order Effect Estimates	282
Combined Model Parameter Estimates	283
Main Effects Residual Plots	284
Prediction Profiler	284
Fit Definitive Screening Platform Options	284
Technical Details for the Fit Definitive Screening Platform	285
The Effective Model Selection for DSDs Algorithm	285

Overview of the Fit Definitive Screening Platform

The Fit Definitive Screening platform analyzes definitive screening designs (DSDs) using a methodology that takes advantage of their special structure. The methodology is called *Effective Model Selection for DSDs*. If you created your design in JMP, the design table contains a script called Fit Definitive Screening that automatically runs an analysis using the Effective Model Selection for DSDs methodology.

Identification of Active Effects in DSDs

DSDs are three-level designs that are valuable for identifying main effects and second-order effects in a single experiment. A minimum run-size DSD is capable of correctly identifying active terms with high probability if the number of active effects is less than about half the number of runs and if the effects sizes exceed twice the standard deviation.

However, by augmenting a minimum run-size DSD with four or more properly selected runs, you can identify substantially more effects with high probability. These runs are called *Extra Runs*, and correspond to fictitious inactive factors, called *fake factors*. For information about Extra Runs, see [“Structure of Definitive Screening Designs”](#) on page 268 in the “Definitive Screening Designs” chapter.

Extra Runs substantially increase the design’s ability to detect second-order effects. For this reason, Jones and Nachtsheim (2016) strongly encourage the inclusion of at least four Extra Runs.

Effective Model Selection for DSDs

When standard model selection methods are applied to DSDs, they can fail to identify active effects. See Errore et al. (2017). Also, standard selection methods do not leverage the structure of DSDs. The Fit Definitive Screening platform uses the Effective Model Selection for DSDs approach, which takes full advantage of the structure of the DSD.

Jones and Nachtsheim (2016) report on simulation studies using Effect Model Selection for DSDs as well as standard approaches. Denote by c the sum of the number of factors and the number of fake factors in a DSD. In many situations, if the number of active main effects exceeds three, then up to $c/2$ active second-order effects can be reliably identified. Assuming strong effect heredity, if there are three or fewer active main effects, then all active second-order effects can be reliably identified. *Reliable identification* means that the ratio of the absolute value of the coefficient to the error standard deviation exceeds three and that the power to detect the effect exceeds 0.80.

The Fit Definitive Screening platform default settings assume *strong effect heredity*. Strong effect heredity means that the A*B interaction can only be considered for inclusion in the model if both A and B have been included. Strong effect heredity requires that all lower-order components of a model effect be included in the model. In identifying active second-order effects, the algorithm uses strong effect heredity and the results cited earlier about how many active second-order effects can be reliably identified.

In a DSD, main effects and second-order effects are orthogonal to each other. The Effective Model Selection for DSDs approach takes advantage of this fact. The linear space of the response is separated into the subspace spanned by the main effects and the orthogonal complement of this subspace. Miller and Sitter (2005) refer to the linear subspace spanned by the main effects as the *odd space*, because it contains all the information about odd effects: main effects, 3-factor effects, 5-factors effects, and so on. They refer to its orthogonal complement as the *even space*, because it contains all the information about even effects: the intercept, 2-factor effects, 4-factor effects, and so on.

Fit Definitive Screening follows this thinking. The subspace spanned by the main effects is the odd space. Its orthogonal complement, the even space, contains the second-order effects and the block variable, if one exists. For more information about the algorithm, see [“The Effective Model Selection for DSDs Algorithm”](#) on page 285 and Jones and Nachtsheim (2016).

Example of the Fit Definitive Screening Platform

The design in the sample data table Extraction 3 Data.jmp is a definitive screening design for six factors in two blocks. The **Add Blocks with Center Runs to Estimate Quadratic Effects** option was selected and 4 Extra Runs were added in the design generation. The resulting design has 18 runs. This design is used to explore the analysis of a definitive screening design.

- [“Fit the Model”](#)
- [“Examine Results”](#)
- [“Reduce the Model”](#)

Note: For an example of a design with continuous and categorical factors, open the sample data file Peanut Data.jmp and run the Fit Definitive Screening table script.

Fit the Model

1. Select **Help > Sample Data Library** and open Design Experiment/Extraction 3 Data.jmp.
2. Select **DOE > Definitive Screening > Fit Definitive Screening**.
3. Select Yield and click Y.

4. Select Lot through Time and click X.
5. Click OK.

The fit performs a two-stage analysis. For more information about the algorithm, see [“Technical Details for the Fit Definitive Screening Platform”](#) on page 285.

Examine Results

Stage 1: Main Effect Estimates

Stage 1 determines which main effects are likely to be active.

Figure 8.2 Stage 1 Report for Main Effects

Stage 1 - Main Effect Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Methanol	9.7133	0.3674	26.438	<.0001*
Ethanol	2.3166	0.3674	6.3055	0.0015*
Time	4.0798	0.3674	11.104	0.0001*
Statistic	Value			
RMSE	1.3747			
DF	5			
☒ Quadratic Terms Obey Strong Heredity				
☒ Interactions Obey Strong Heredity				

Note: The fake factors do not appear in the design or as factors in the analysis.

A two-degree-of-freedom error sum of squares is computed from the four runs corresponding to the two fake factors. Because the fake factors are, by construction, inactive, this estimate of error variance is unbiased. For each main effect, the main effects response Y_{ME} is tested against this estimate. In this example, three factors, Methanol, Ethanol, and Time, have p -values smaller than the threshold value and are retained as active. For more information about the threshold values, see [“Stage 1 Methodology”](#) on page 286.

The variability from the three inactive factors, Propanol, Butanol, and pH, is pooled with the fake factor sum of squares to produce the five-degree-of-freedom RMSE statistic shown in Figure 8.2.

Stage 2: Even Order Effect Estimates

Stage 2 uses guided subset selection to arrive at a list of second-order effects that are likely to be active. Interactions and quadratic terms are second-order or even order effects.

Figure 8.3 Stage 2 Report for Even-Order Effects

Stage 2 - Even Order Effect Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Intercept	34.568	1.3459	25.683	0.0015*
Lot[1]	17.197	0.7757	22.171	0.0020*
Methanol*Ethanol	-0.367	0.7127	-0.515	0.6581
Methanol*Time	0.5266	0.7127	0.7389	0.5369
Ethanol*Time	9.8258	0.8534	11.514	0.0075*
Methanol*Methanol	7.637	1.4914	5.1208	0.0361*
Ethanol*Ethanol	-1.449	1.477	-0.981	0.4299
Time*Time	-3.297	1.477	-2.232	0.1552
Statistic	Value			
RMSE	2.0626			
DF	2			

Because three main effects are identified as active in Stage 1, the guided subset selection procedure for active second-order effects can continue until all second-order effects are included. Because all six second-order effects are reported in Stage 2, it follows that the Stage 2 RMSE remained larger than the Stage 1 RMSE. See “[Stage 2 Methodology](#)” on page 287.

The two-degree-of-freedom RMSE given in the Stage 2 report is the error estimate obtained from the final subset of all six second-order effects.

Combined Results

The effects selected for the model are listed in the Combined Model Parameter Estimates report.

Figure 8.4 Combined Model Parameter Estimates Report

Combined Model Parameter Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Intercept	34.568	1.0452	33.074	<.0001*
Lot[1]	17.197	0.6023	28.552	<.0001*
Methanol	9.7133	0.4281	22.691	<.0001*
Ethanol	2.3166	0.4281	5.4118	0.0010*
Time	4.0798	0.4281	9.5307	<.0001*
Methanol*Ethanol	-0.367	0.5534	-0.663	0.5287
Methanol*Time	0.5266	0.5534	0.9516	0.3730
Ethanol*Time	9.8258	0.6627	14.828	<.0001*
Methanol*Methanol	7.637	1.1581	6.5945	0.0003*
Ethanol*Ethanol	-1.449	1.1469	-1.264	0.2468
Time*Time	-3.297	1.1469	-2.875	0.0238*
Statistic	Value			
RMSE	1.6017			
DF	7			

The RMSE and degrees of freedom given at the bottom of the report are the usual standard least squares quantities. Use these effects as potential factors for your final model.

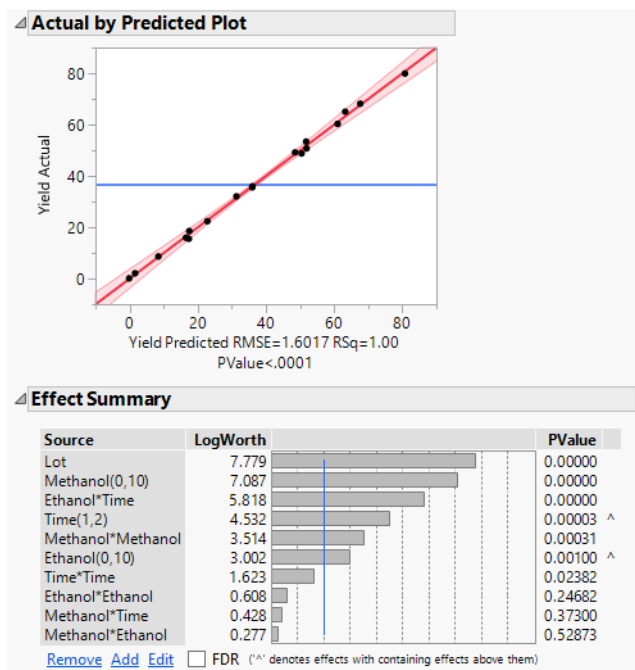
Reduce the Model

The Make Model button enters the model for the listed terms in a Fit Model specification window. To run the model directly using standard least squares, click the Run Model button.

1. Click **Run Model**.

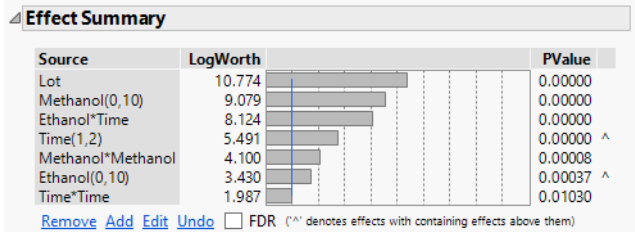
The Actual by Predicted Plot shows no lack of fit. The Effect Summary report suggests that you can reduce the model further.

Figure 8.5 Actual by Predicted Plot and Effect Summary Report



2. Select Methanol*Ethanol in the Effect Summary report and click **Remove**.
Methanol*Time has p -value 0.33750. Remove it next.
3. Select Methanol*Time in the Effect Summary report and click **Remove**.
Ethanol*Ethanol has p -value 0.15885. Remove it next.
4. Select Ethanol*Ethanol in the Effect Summary report and click **Remove**.

Figure 8.6 Effect Summary Report Showing Effects in Final Model



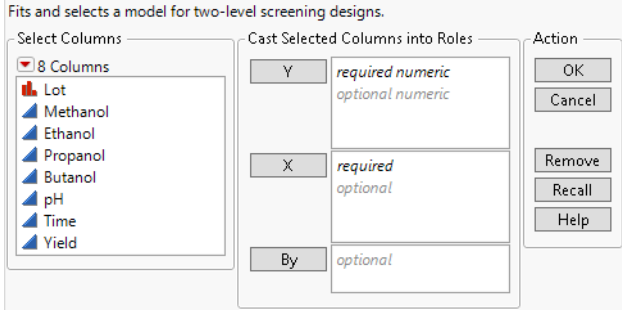
The remaining effects are significant. You conclude that these are the active effects.

Launch the Fit Definitive Screening Platform

To launch the Fit Definitive Screening platform, select **DOE > Definitive Screening > Fit Definitive Screening**. The launch window in Figure 8.7 uses Extraction3 Data.jmp.

Note: If you created your design in JMP, the design table contains a script called Fit Definitive Screening. Run this script to run the analysis directly.

Figure 8.7 Fit Definitive Screening Launch Window



- Y** One or more numeric response variables.
- X** Continuous or two-level categorical factors. Because the platform uses the unique features of a DSD in performing the analysis, these factors must define a DSD or a fold-over design.
- By** A column whose levels define separate analyses. For each level of the specified column, the corresponding rows are analyzed. The results appear in separate reports. If more than one By variable is assigned, a separate analysis is produced for each possible combination of the levels of the By variables.

For more information about the options in the Select Columns red triangle menu, see the Get Started chapter in *Using JMP*.

Fit Definitive Screening Report

- [“Stage 1 - Main Effect Estimates”](#)
- [“Stage 2 - Even Order Effect Estimates”](#)
- [“Combined Model Parameter Estimates”](#)
- [“Main Effects Residual Plots”](#)
- [“Prediction Profiler”](#)

Stage 1 - Main Effect Estimates

The Main Effect Estimates report lists main effects that are identified as active. Main effects with p -values less than the threshold p -value are considered active.

- If fake factors or center point replicates are available, an estimator of error variance that is independent of the model is constructed. The main effects are tested against this estimate.
- If no fake factors or center point replicates are available, subsets of main effects are tested sequentially against an estimate of error variance constructed from the inactive main effects. For this procedure to be viable, at least one of the main effects must be inactive.

In either case, variability from the inactive main effects is pooled into the error variance used to test the main effects.

Figure 8.8 Stage 1 Report

Stage 1 - Main Effect Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Methanol	9.7133	0.3674	26.438	<.0001*
Ethanol	2.3166	0.3674	6.3055	0.0015*
Time	4.0798	0.3674	11.104	0.0001*
Statistic	Value			
RMSE	1.3747			
DF	5			
☒ Quadratic Terms Obey Strong Heredity				
☒ Interactions Obey Strong Heredity				

Term Main effects identified as active. These effects have p -values less than the threshold when tested as described in [“Stage 1 Methodology”](#) on page 286.

Estimate Parameter estimate for a regression fit of Y on the main effects.

Std Error The standard error of the estimate, computed using the Stage 1 RMSE.

t Ratio The Estimate divided by its Std Error.

Prob>|t| The p -value computed using the t Ratio and the degrees of freedom for error (DF).

RMSE The square root of the mean square error that results from the Stage 1 analysis.

- If fake factors or centerpoint replicates are available, the mean square error is the estimate of variance from fake factors and centerpoints pooled with the variance estimate constructed from the main effects that are not identified as active.
- If no fake factors or centerpoint replicates are available, the mean square error is the estimate of variance constructed from the main effects that are not identified as active.

DF The degrees of freedom associated with the error estimate used to construct RMSE.

- If fake factors or centerpoint replicates are available, DF is the sum of the number of fake factors, centerpoint replicates, and main effects not identified as active.
- If no fake factors or centerpoint replicates are available, DF is the number of main effects that are not identified as active.

Quadratic Terms Obey Strong Heredity Check or uncheck to control heredity rules applied to quadratic terms in Stage 2.

Interactions Obey Strong Heredity Check or uncheck to control heredity rules applied to interactions in Stage 2.

Note: Strong effect heredity means that the A*B interaction can only be considered for inclusion in the model if both A and B have been included. Strong effect heredity requires that all lower-order components of a model effect be included in the model.

Stage 2 - Even Order Effect Estimates

The Even Order Effect Estimates report lists second-order effects that are identified as active. Active second-order effects are identified using the guided variable selection procedure or forward selection as described in “[Stage 2 Methodology](#)” on page 287. The block effect (if one is included) is also listed, whether it is significant or not.

Figure 8.9 Stage 2 Report

Stage 2 - Even Order Effect Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Intercept	34.568	1.3459	25.683	0.0015*
Lot[1]	17.197	0.7757	22.171	0.0020*
Methanol*Ethanol	-0.367	0.7127	-0.515	0.6581
Methanol*Time	0.5266	0.7127	0.7389	0.5369
Ethanol*Time	9.8258	0.8534	11.514	0.0075*
Methanol*Methanol	7.637	1.4914	5.1208	0.0361*
Ethanol*Ethanol	-1.449	1.477	-0.981	0.4299
Time*Time	-3.297	1.477	-2.232	0.1552
Statistic	Value			
RMSE	2.0626			
DF	2			

Term The block factor and second-order effects identified as active.

Estimate Parameter estimates for a regression fit of Y on the Stage 2 second-order effects defined by Y_{2nd} . See [“Decomposition of Response”](#) on page 285.

Std Error The standard error of the estimate, computed using the Stage 2 RMSE.

t Ratio The Estimate divided by its Std Error.

Prob>|t| The *p*-value computed using the t Ratio and the degrees of freedom for error (DF).

RMSE The square root of the mean square error that results from the Stage 2 analysis. RMSE is estimated as the residual variance from Y_{2nd} after fitting the second order effects identified as active. See [“Decomposition of Response”](#) on page 285.

DF The degrees of freedom associated with the error estimate used to construct RMSE.

Combined Model Parameter Estimates

The Combined Model Parameter Estimates report lists the terms in the final model and their usual standard least squares estimates, standard errors, t ratios, *p*-values, RMSE, and model degrees of freedom.

Below the report are buttons that construct or run the combined model.

Figure 8.10 Combined Model Parameter Estimates Report

Combined Model Parameter Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Intercept	34.568	1.0452	33.074	<.0001*
Lot[1]	17.197	0.6023	28.552	<.0001*
Methanol	9.7133	0.4281	22.691	<.0001*
Ethanol	2.3166	0.4281	5.4118	0.0010*
Time	4.0798	0.4281	9.5307	<.0001*
Methanol*Ethanol	-0.367	0.5534	-0.663	0.5287
Methanol*Time	0.5266	0.5534	0.9516	0.3730
Ethanol*Time	9.8258	0.6627	14.828	<.0001*
Methanol*Methanol	7.637	1.1581	6.5945	0.0003*
Ethanol*Ethanol	-1.449	1.1469	-1.264	0.2468
Time*Time	-3.297	1.1469	-2.875	0.0238*
Statistic	Value			
RMSE	1.6017			
DF	7			
Make ModelRun Model				

Make Model Creates a model for the Fit Model window containing the model terms in the Combined Model Parameter Estimates report and the response specified for the Fit Definitive Screening analysis. The Standard Least Squares personality is specified.

Run Model Runs a standard least squares fit for the model terms in the Combined Model Parameter Estimates report and the response specified for the Fit Definitive Screening analysis.

Main Effects Residual Plots

Shows a plot of the residuals from a main effects model fit to all other factors versus the levels of the plotted factor. A factor with residuals that differ across levels indicates an important main effect.

Note: There is not a plot for a block factor.

Prediction Profiler

Shows a prediction profiler for the combined model. For more information about the prediction profiler, see the Profiler chapter in *Profilers*.

Fit Definitive Screening Platform Options

Set Stage 1 p value Set the *p*-value threshold used for Stage 1. See “[Stage 1 Methodology](#)” on page 286.

Set Stage 2 ratio Specify the ratio of the Stage 2 MSE to the Stage 1 MSE. See “[Stage 2 Methodology](#)” on page 287.

See the JMP Reports chapter in *Using JMP* for more information about the following options:

Local Data Filter Shows or hides the local data filter that enables you to filter the data used in a specific report.

Redo Contains options that enable you to repeat or relaunch the analysis. In platforms that support the feature, the Automatic Recalc option immediately reflects the changes that you make to the data table in the corresponding report window.

Save Script Contains options that enable you to save a script that reproduces the report to several destinations.

Technical Details for the Fit Definitive Screening Platform

The Effective Model Selection for DSDs Algorithm

This section provides a summary of the algorithm used in the Fit Definitive Screening platform. See Jones and Nachtsheim (2016).

Decomposition of Response

The Effective Model Selection algorithm expresses the response, Y , in terms of two responses Y_{ME} and Y_{2nd} , so that $Y = Y_{ME} + Y_{2nd}$.

- Y_{ME} is the predicted value obtained from a regression of Y on the main effects and fake factors.

There is no need to include the block factor in Y_{ME} because of the fold-over structure of the design. The block factor is included in Y_{2nd} .

- Y_{2nd} is given by $Y_{2nd} = Y - Y_{ME}$.

Note: In a DSD, the columns Y_{ME} and Y_{2nd} are orthogonal.

The analysis proceeds in two stages:

- Stage 1: The response Y_{ME} is used to identify main effects. Stage 1 identifies the main effects that are considered active.

- Stage 2: The response Y_{2nd} is used to identify second-order effects. Stage 2 considers all second-order terms in the active main effects from Stage 1 and determines a subset of these containing effects considered to be active.

Note: If there is a blocking factor, it is included in the Stage 2 list of effects even if it is not significant.

Stage 1 Methodology

The Stage 1 methodology depends on whether the design contains fake factors or centerpoint replicates.

Case 1: Fake Factors or Centerpoint Replicates Available

1. Using the fake factors or center point replicates, an estimator of error variance that is independent of the model is constructed. Assuming that there are no active third or higher odd order effects, this estimate is unbiased.
2. Using Y_{ME} , main effects are tested against this estimate. Main effects with p -values less than a threshold p -value are considered active. The threshold values are the following:
 - For one error degree of freedom, the threshold value is 0.20.
 - For two error degrees of freedom, the threshold value is 0.10.
 - For more than two error degrees of freedom, the threshold value is 0.05.
 - User specified p -value is the threshold.

Note: To specify a different p -value threshold select **Set Stage 1 p value** from the Fit Definitive Screening red triangle menu.

3. If no main effect has a p -value less than the threshold value, conclude that there are no active main effects and no active two-factor effects. The procedure terminates.
4. If active main effects are found, then variability from the inactive main effects is pooled into the error variance constructed in (1).

Note: If Categorical factors are in the design, the estimated coefficients are recalculated each time a main effect is chosen as active.

Case 2: No Fake Factors or Centerpoint Replicates Available

In this case, there is no model-independent estimator of error variance available. Subsets of main effects are tested sequentially against an estimate of error variance constructed from the inactive main effects. Suppose that there are m main effects.

1. The absolute values of the estimated effects, using Y_{ME} as the response, are ordered from largest to smallest.
2. For each $1 \leq i < m$, the effect with the i^{th} largest absolute value is tested against the adjusted residual sum of squares for the model containing that effect and all effects with larger absolute values.
3. The effects in the model with the smallest p -value are considered to be the active effects.
4. If active main effects are found, then variability from the inactive main effects is used to construct an estimate of error variance, using Y_{ME} as the response.

Note: For the Fit Definitive Screening procedure to work properly in Case 2, at least one of the main effects must be active and at least one must be inactive. If no main effects are active, or if all main effects are active, the procedure will identify a set of main effects, but the procedure for arriving at that subset is compromised.

Stage 2 Methodology

In Stage 2, the factors considered depend on the Strong Heredity options. When strong heredity is selected, only second-order effects involving the factors whose main effects are identified as active in Stage 1 are considered. The Stage 2 methodology depends on the number of active main effects identified in Stage 1.

Case 1: Seven or Fewer Active Main Effects

Stage 2 uses a guided subset selection procedure. The goal is to continue to add second-order effects to the model as long as the ratio of the RMSE from Stage 2 to the RMSE from Stage 1 is greater than the specified threshold. When the ratio is less than or equal to the threshold, this indicates that there are no additional second-order effects to add to the model. The default threshold is 1. Smaller thresholds increases the number of terms likely to identified as active as compared to larger thresholds.

Note: To specify a RMSE ratio threshold other than one, select **Set Stage 2 ratio** from the Fit Definitive Screening red triangle menu.

For Stage 2:

- For one error degree of freedom, the threshold value is 0.20.
- For two error degrees of freedom, the threshold value is 0.10.

- For more than two error degrees of freedom, the threshold value is 0.05.
 - User specified p -value is the threshold.
1. The variability for Y_{2nd} is tested against the error estimate from Stage 1 to determine if there is additional variability due to second-order effects.
 - If the p -value for this test exceeds the threshold value the procedure terminates and no active second-order effects are identified.
 2. If the p -value for this test is less than or equal to the threshold value, then subsets of size k , $k = 1, 2, 3, \dots$ are successively tested, starting with $k = 1$.
 3. For each k , the residual sum of squares for each subset of that size is tested against the error estimate from Stage 1. The subset with the smallest RMSE is identified.
 4. The procedure continues until a k is found with a ratio of RMSE to the Stage 1 RMSE smaller than the Stage 2 ratio.
 5. The effects in the subset preceding the one that corresponds to the terminal value of k are considered to be the active two-factor effects.

Case 2: Eight or More Active Main Effects

Stage 2 uses forward selection for second order terms when eight or more active main effects are identified in Stage 1.

Screening Designs

Screening designs are among the most popular designs for industrial experimentation. Typically used in the initial stages of experimentation, they examine many factors in order to identify those factors that have the greatest effect on the response. The factors that are identified are then studied using more sensitive designs. Because screening designs generally require fewer experimental runs than other designs, they are a relatively inexpensive and efficient way to begin improving a process.

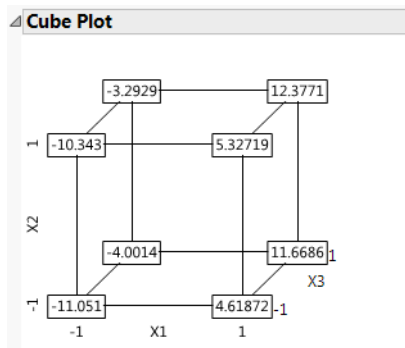
If a standard screening design exists for your experimental situation, you can choose from several standard screening designs. The list includes blocked designs when applicable. Your factors can be two-level continuous factors, three-level categorical factors, or continuous factors that can assume only discrete values (*discrete numeric* factors).

However, there are situations where standard screening designs are not available. In these cases, the Screening Design platform constructs a *main effects screening design*. A main effects screening design is either orthogonal or near orthogonal. It focuses on estimating main effects in the presence of negligible interactions.

Note that JMP also provides two compelling alternatives to screening designs:

- Definitive screening designs are particularly useful if you suspect active two-factor interactions or if you suspect that a plot of a continuous factor's effect on the response might exhibit strong curvature. See [Chapter 7, "Definitive Screening Designs"](#).
- Custom designs are highly flexible and often more cost-effective than a design obtained using alternative methods. See [Chapter 4, "Custom Designs"](#).

Figure 9.1 Results from a Fractional Factorial Design



Contents

Overview of Screening Designs	291
Underlying Principles.....	291
Analysis of Screening Design Results	292
Examples of Screening Designs	292
Compare a Fractional Factorial Design and a Main Effects Screening Design	292
Main Effects Screening Design where No Standard Design Exists	300
Screening Design Window.....	302
Responses	304
Factors.....	306
Choose Screening Type	308
Choose from a List of Fractional Factorial Designs	308
Design Type	309
Resolution as a Measure of Confounding	312
Display and Modify Design.....	313
Main Effects Screening Designs	316
Design Generation	317
Design	317
Design Evaluation.....	317
Output Options	318
Make Table	319
Screening Design Options	320
Additional Examples of Screening Designs	323
Modify Generating Rules in a Fractional Factorial Design.....	323
Plackett-Burman Design.....	329

Overview of Screening Designs

Screening experiments tend to be small and are aimed at identifying the factors that affect a response. Because identification is the goal (rather than sophisticated modeling), continuous factors in a screening design are typically set at only two levels. However, a screening situation might also involve discrete numeric or categorical factors, in which case classical screening designs might not fit your situation. The Screening Design platform can handle all three types of factors: two-level continuous factors, categorical factors, and discrete numeric factors.

There are two types of designs:

- Classical designs: For situations where standard screening designs exist, you can choose from a list that includes fractional factorial designs, Plackett-Burman, Cotter, and mixed-level designs.
- Main effects screening designs: Whether a standard design is available, you can ask JMP to construct a main effects screening design. These designs are orthogonal or near orthogonal and focus on estimating main effects in the presence of negligible interactions. See [“Main Effects Screening Designs”](#) on page 316.

Underlying Principles

The emphasis on studying main effects early on in the experimentation process is supported by the empirical principle of effect hierarchy. This principle maintains that lower order effects are more likely to be important than higher order effects. For this reason, screening designs focus on identifying active main effects. In cases where higher order interactions *are* of interest, screening designs assume that two-factor interactions are more important than three-factor interactions, and so on. See [“Effect Hierarchy”](#) on page 62 in the “Starting Out with DOE” chapter and Wu and Hamada (2009).

The efficiency of screening designs also depends on the principle of effect sparsity. Effect sparsity asserts that most of the variation in the response is explained by a relatively small number of effects. See [“Effect Sparsity”](#) on page 63 in the “Starting Out with DOE” chapter.

To appreciate the importance of effect sparsity, consider an example where you have seven two-level factors. Contrast a full factorial design to a screening design:

- A full factorial design consists of all combinations of the levels of the factors. The number of runs is the product of the numbers of levels for each factor. In this example, a full factorial design has $2^7 = 128$ runs.
- In contrast, a screening design requires only a fraction of the runs in the full factorial design. The main effects of the seven factors can be studied in an eight-run screening design.

Analysis of Screening Design Results

Screening designs are often used to test a large number of factors or interactions. When there are degrees of freedom for error, allowing construction of an error estimate, the experimental results can be analyzed using the usual regression techniques (Analyze > Fit Model).

However, sometimes there are no degrees of freedom for error. In this case, assuming effect sparsity, the Screening platform (DOE > Classical > Two Level Screening > Fit Two Level Screening) provides a way to analyze the results of a two-level design. The Screening platform accepts multiple responses and multiple factors. It automatically shows significant effects with plots and statistics. See [“The Fit Two Level Screening Platform”](#) chapter on page 335. For an examples in the current chapter, see [“Modify Generating Rules in a Fractional Factorial Design”](#) on page 323 and [“Plackett-Burman Design”](#) on page 329.

Examples of Screening Designs

- [“Compare a Fractional Factorial Design and a Main Effects Screening Design”](#)
- [“Main Effects Screening Design where No Standard Design Exists”](#)

Compare a Fractional Factorial Design and a Main Effects Screening Design

In this example, suppose an engineer wants to investigate a process that uses an electron beam welding machine to join two parts. The engineer fits the two parts into a welding fixture that holds them snugly together. A voltage applied to a beam generator creates a stream of electrons that heats the two parts, causing them to fuse. The ideal depth of the fused region is 0.17 inches. The engineer wants to study the welding process to determine the best settings for the beam generator to produce the desired depth in the fused region.

For this study, the engineer wants to explore the following seven factors:

- Operator is the technician operating the welding machine. Two technicians typically operate the machine.
- Speed (in rpm) is the speed at which the part rotates under the beam.
- Current (in amps) is a current that affects the intensity of the beam.
- Mode is the welding method used.
- Wall Size (in mm) is the thickness of the part wall.
- Geometry indicates whether the joint is a single-bevel joint or a double-bevel joint.
- Material is the type of material being welded.

Notice that three of these factors are continuous: Speed, Current, and Wall Size. Four are categorical: Operator, Mode, Geometry, and Material. Each of these categorical factors has two levels.

After each processing run, the engineer cuts the part in half. This reveals an area where the two parts have fused. The length of this fused area, measured in inches, is the depth of penetration of the weld. The depth of penetration is the response for the study.

The goals of the study are the following:

- Find which factors affect the depth of the weld.
- Quantify those effects.
- Find specific factor settings that predict a weld depth of 0.17 inches with a tolerance of ± 0.05 inches.

Your experimental budget allows you at most 12 runs. Construct and compare two designs for your experimental situation. The first is a classical fractional factorial design using eight runs. The second is a main effects screening design using 12 runs.

Constructing a Standard Screening Design

In this section, construct a standard screening design for this experimental situation.

Specify the Response

1. Select **DOE > Classical > Two Level Screening > Screening Design**.
2. In the Responses panel, double-click Y under Response Name and type **Depth**.
Note that the default Goal is Maximize. Your goal is to find factor settings that enable you to obtain a target depth of 0.17 inches with limits of 0.12 and 0.22.
3. Click the default Goal of Maximize and change it to **Match Target**.
4. Click under **Lower Limit** and type 0.12.
5. Click under **Upper Limit** and type 0.22.
6. Leave the area under **Importance** blank.

Because there is only one response, that response is given Importance 1 by default.

The completed Responses outline appears in Figure 9.2. Now, specify the factors.

Specify Factors

You can enter the factors manually or automatically:

- To enter the factors manually, see [“Specify Factors Manually”](#) on page 294.
- To enter the factors automatically, use the Weld Factors.jmp data table:

1. Select **Help > Sample Data Library** and open Design Experiment/Weld Factors.jmp.
2. Click the Screening Design red triangle and select **Load Factors**. Proceed to [“Choose a Design”](#) on page 295.

Specify Factors Manually

1. Type **3** in the **Add N Factors** box and click **Continuous**.
2. Double-click X1 and type Speed.
3. Use the Tab key to move through the rest of the values and factors. Make the following changes:
 - a. Change the Speed values to 3 and 5.
 - b. Change X2 to Current, with values of 150 and 165.
 - c. Change X3 to Wall Size, with values of 20 and 30.
4. Type **4** in the **Add N Factors** box and select **Categorical > 2 Level**.
5. Double-click X4 and type Operator.
6. Use the Tab key to move through the rest of the values and factors. Make the following changes:
 - a. Change the Operator values to John and Mary.
 - b. Change X5 to Mode, with values of Conductance and Keyhole.
 - c. Change X6 to Geometry, with values of Double and Single.
 - d. Change X7 to Material, with values of Aluminum and Magnesium.

Figure 9.2 Responses and Factors Outlines for Weld Experiment

Screening Design

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Depth <i>optional item</i>	Match Target	0.12	0.22	.

Factors

Continuous Discrete Numeric Categorical Remove Add N Factors 1

Name	Role	Values
Speed	Continuous	3 5
Current	Continuous	150 165
Wall Size	Continuous	20 30
Operator	Categorical	John Mary
Mode	Categorical	Conductance Keyhole
Geometry	Categorical	Double Single
Material	Categorical	Aluminum Magnesium

Specify Factors

Add a Continuous or Categorical factor by clicking its button. Double click on a factor name or level to edit it.

Continue

Choose a Design

1. Click **Continue**.

Because the combination of factors and levels that you have specified can be accommodated by a standard fractional factorial design, the Choose Screening Type panel appears. You can either select a standard design from a list or construct a main effects design.

Note: Setting the Random Seed in the next step reproduces the results shown in this example. When you are constructing a design on your own, this step is not necessary.

2. (Optional) Click the Screening Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
3. Accept the default selection to **Choose from a list of fractional factorial designs** and click **Continue**.
4. Select the first Fractional Factorial design.

Figure 9.3 Design List for Three Continuous Factors and Four Categorical Factors

Design List

Choose a design by clicking on its row in the list.

Number Of Runs	Block Size	Design Type	Resolution - what is estimable
8		Fractional Factorial	3 - Main Effects Only
12		Plackett-Burman	3 - Main Effects Only
16		Fractional Factorial	4 - Some 2-factor interactions
16	8	Fractional Factorial	4 - Some 2-factor interactions
16	4	Fractional Factorial	4 - Some 2-factor interactions
16	2	Fractional Factorial	4 - Some 2-factor interactions
16		Cotter	Sums of Even/Odd Terms
32		Fractional Factorial	4 - Some 2-factor interactions
32	16	Fractional Factorial	4 - Some 2-factor interactions
32	8	Fractional Factorial	4 - Some 2-factor interactions
32	4	Fractional Factorial	4 - Some 2-factor interactions
32	2	Fractional Factorial	4 - Some 2-factor interactions
64		Fractional Factorial	5+ - All 2-factor interactions
64	32	Fractional Factorial	5+ - All 2-factor interactions

This specifies an eight-run Resolution 3 fractional factorial design. For information about resolution, see [“Resolution as a Measure of Confounding”](#) on page 312.

5. Click **Continue**.

In the Output Options outline, note that **Run Order** is set to **Randomize**. This means that the design runs will appear in random order. This is the order you should use to conduct your experimental runs.

Figure 9.4 Completed Screening Design Window

Screening Design

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Depth	Match Target	0.12	0.22	.

optional item

Factors

Continuous Discrete Numeric Categorical Remove Add N Factors 1

Name	Role	Values
Speed	Continuous	3 5
Current	Continuous	150 165
Wall Size	Continuous	20 30
Operator	Categorical	John Mary
Mode	Categorical	Conductance Keyhole
Geometry	Categorical	Double Single
Material	Categorical	Aluminum Magnesium

Display and Modify Design

Change Generating Rules

Aliasing of Effects

Coded Design

Design Evaluation

Output Options

Run Order: Randomize

Make JMP Table from design plus

Number of Center Points: 0

Number of Replicates: 0

Make Table

Back

- Open the Aliasing of Effects outline under Display and Modify Design.

Figure 9.5 Aliasing for an Eight-Run Fractional Factorial Design

Aliasing of Effects

Effects	Aliases
Speed	= Current*Material = Wall Size*Geometry = Operator*Mode
Current	= Speed*Material = Wall Size*Mode = Operator*Geometry
Wall Size	= Speed*Geometry = Current*Mode = Operator*Material
Operator	= Speed*Mode = Current*Geometry = Wall Size*Material
Mode	= Speed*Operator = Current*Wall Size = Geometry*Material
Geometry	= Speed*Wall Size = Current*Operator = Mode*Material
Material	= Speed*Current = Wall Size*Operator = Mode*Geometry

Recall that you selected a Resolution 3 design (Figure 9.3). In a Resolution 3 design, some main effects are confounded with two-way interactions. The Aliasing of Effects outline indicates that, for this Resolution 3 design, every main effect is completely confounded with three two-way interactions. If you suspect that two-way interactions are active, this is a poor design. For a description of confounding, see [“Two-Level Regular Fractional Factorial”](#) on page 310.

- Click **Make Table**.

Figure 9.6 The Design Data Table

	Pattern	Speed	Current	Wall Size	Operator	Mode	Geometry	Material	Depth
1	++++++	5	165	30	Mary	Keyhole	Single	Magnesium	•
2	-----	3	165	20	Mary	Conductance	Single	Aluminum	•
3	++++--	5	150	30	John	Conductance	Single	Aluminum	•
4	-----	3	150	20	John	Keyhole	Single	Magnesium	•
5	-----	3	165	30	John	Keyhole	Double	Aluminum	•
6	++++--	5	165	20	John	Conductance	Double	Magnesium	•
7	++++--	5	150	20	Mary	Keyhole	Double	Aluminum	•
8	++++--	3	150	30	Mary	Conductance	Double	Magnesium	•

Notice the following:

- The table uses the names for the responses, factors, and levels that you specified.
- The Pattern column shows the assignment of high and low settings for the design runs.
- This fractional factorial design is a Resolution 3 design. It enables you to study the main effects of seven factors in eight runs.

Constructing a Main Effects Screening Design

Main effects screening designs are orthogonal or near orthogonal designs. In this section, construct a main effects screening design for your seven factors.

1. Open your Screening Design window. If you have closed it, then run the **DOE Dialog** script in the Design Data table.
2. Click **Back**.
3. Click **Continue**.

Note: Setting the Random Seed and the Number of Starts in the next two steps reproduces the exact results shown in this example. When constructing a design on your own, these steps are not necessary.

4. (Optional) Click the Screening Design red triangle, select **Set Random Seed**, type 12345, and click **OK**.
5. (Optional) Click the Screening Design red triangle, select **Number of Starts**, type 50, and click **OK**.
6. In the Choose Screening Type panel, select the **Construct a main effects screening design** option.
7. Click **Continue**.

Under Number of Runs, the selected option is **Default** with the number of runs set to 12. Keep this setting.

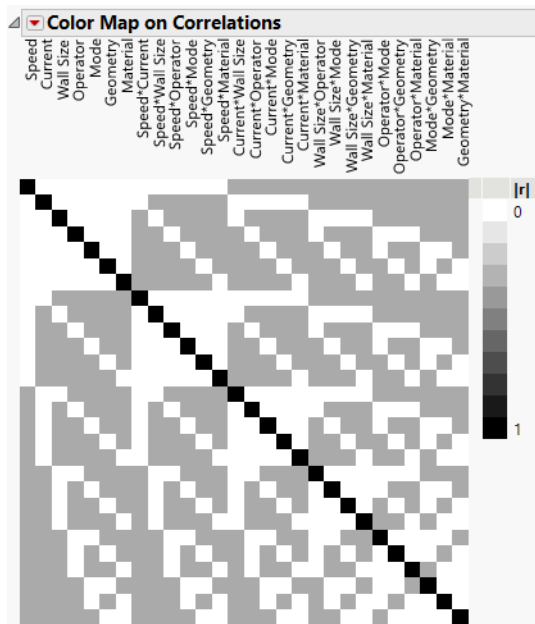
8. Click **Make Design**.

Figure 9.7 Main Effects Screening Design

Design							
Run	Speed	Current	Wall Size	Operator	Mode	Geometry	Material
1	3	165	20	John	Conductance	Single	Aluminum
2	5	150	30	Mary	Conductance	Double	Aluminum
3	3	150	30	John	Keyhole	Double	Aluminum
4	5	165	20	Mary	Keyhole	Double	Magnesium
5	3	165	20	Mary	Conductance	Double	Aluminum
6	5	150	20	John	Conductance	Double	Magnesium
7	3	150	30	Mary	Conductance	Single	Magnesium
8	5	150	20	John	Keyhole	Single	Aluminum
9	3	165	30	John	Keyhole	Double	Magnesium
10	5	165	30	John	Conductance	Single	Magnesium
11	3	150	20	Mary	Keyhole	Single	Magnesium
12	5	165	30	Mary	Keyhole	Single	Aluminum

- Open the Design Evaluation outline and then open the Color Map on Correlations outline.

Figure 9.8 Color Map on Correlations for 12-Run Main Effects Screening Design



The Color Map on Correlations shows that the main effects are uncorrelated with each other. This is indicated by the white off-diagonal cells in the upper left corner of the color map. Each main effect is partially aliased with some two-way interactions, indicated by the gray cells. You can see, by hovering your cursor above one of the gray cells, that the absolute correlations are 0.333.

In this case, the 12-run main effects screening design is a Plackett-Burman design, which you could have obtained in the Design List. However, in many design situations, the partial aliasing that occurs in a main effects design is preferable to the complete

confounding that occurs in a fractional factorial design that you adapt to your experimental situation.

The next section shows an example of a situation where no standard design exists. In this case, JMP constructs a main effects screening design.

Main Effects Screening Design where No Standard Design Exists

Main effects screening designs are orthogonal or near orthogonal designs for the main effects. You can use them in place of standard designs and in situations where standard designs do not exist. Main effects screening designs are excellent for estimating main effects when interactions are negligible.

In the following experimental situation, no standard design exists. You need a design to study 13 factors: 2 are categorical, one with 4 levels and one with 6 levels, and 11 are continuous.

1. Select **DOE > Classical > Two Level Screening > Screening Design**.

In the Responses panel, there is a single default response called Y. Keep this as the default response.

2. In the Factors panel, click **Categorical** and select **4 Level**.

This adds the variable X1 with levels L1 through L4.

3. Click **Categorical** and select **6 Level**.

This adds the variable X2, with levels L1 through L6.

4. Enter 11 next to **Add N Factors**.

5. Click **Continuous**.

This adds 11 factors, X3 to X13, each at two levels, -1 and 1.

6. Click **Continue**.

The Design Generation panel appears.

There is no option to select a design from the Design List since there are no available standard designs in this situation.

Keep the default number of runs, which is 24.

Figure 9.9 Screening Design Window for 13-Factor Design with Design Generation Panel

Screening Design

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.

optional item

Factors

Continuous Discrete Numeric Categorical Remove Add N Factors 1

Name	Role	Values
X1	Categorical	L1 L2 L3 L4
X2	Categorical	L1 L2 L3 L4 L5 L6
X3	Continuous	-1 1
X4	Continuous	-1 1
X5	Continuous	-1 1

Design Generation

Number of Runs:

☐ Minimum 20

☒ Default 24

☐ User Specified 24

Make Design

Back

- Click **Make Design**.

A Design and a Design Evaluation outline appear.

- Open the Design outline to see the randomized design.

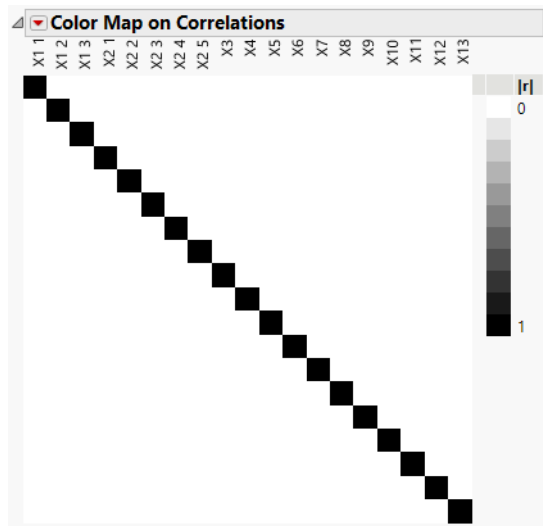
Note: The algorithm that generates the design uses a random starting design. To reproduce this design, save the script with the random seed by selecting **Save Script to Script Window** from the red triangle menu next to the report title.

Next, examine the Color Map on Correlations to see that this specific design is orthogonal.

- Open the **Design Evaluation > Color Map on Correlations** outline.

The color map (Figure 9.10) shows black entries (using JMP default colors) on the main diagonal, indicating correlations of one. This is because each diagonal cell corresponds to the correlation of a term with itself, which is one. Off-diagonal correlations are all white, indicating that correlations between distinct terms are zero. Hold your mouse pointer over any cell to see the relevant terms and their absolute correlation.

Figure 9.10 Color Map on Correlations



10. Click **Make Table** to construct the design table.

The table contains the runs for your experiment in random order. Conduct the experiment in this randomized order and insert the results in column Y. Run the **Model** script in the data table to analyze your results.

Screening Design Window

The Screening Design window updates as you work through the design steps. See [“The DOE Workflow: Describe, Specify, Design”](#) on page 57. The outlines that appear are separated by buttons that update the window. These follow the flow in the figures below.

Figure 9.11 Screening Design Flow when a Standard Design Exists

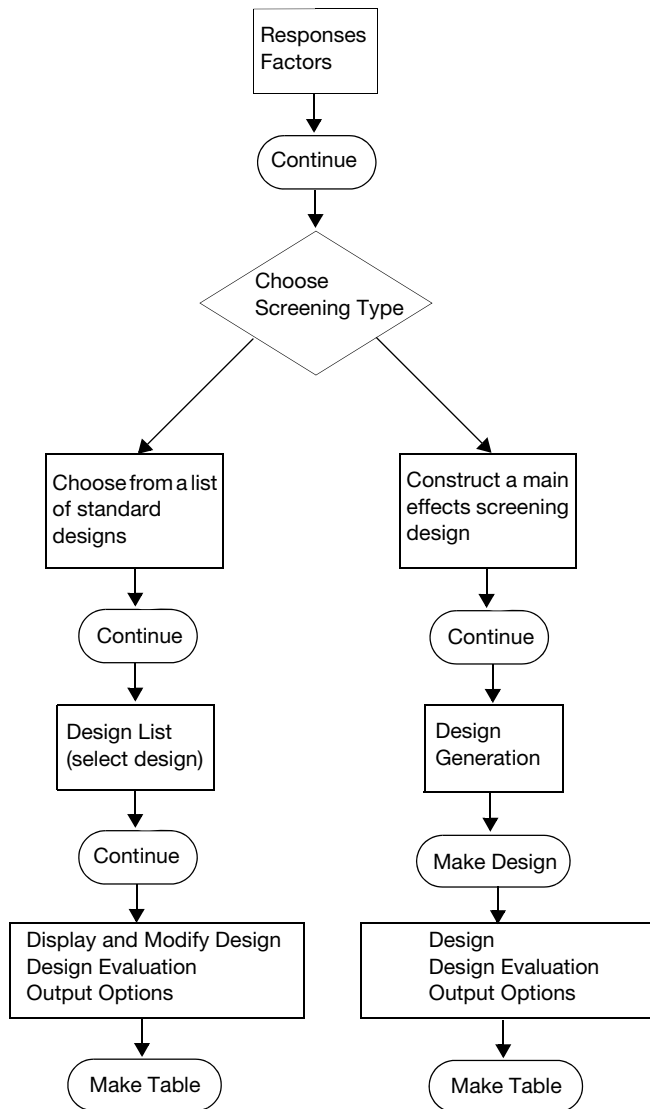
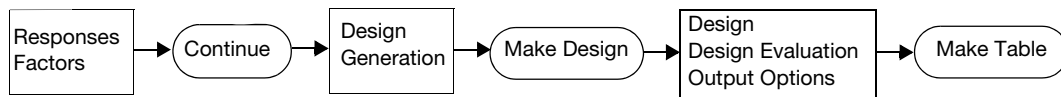


Figure 9.12 Screening Design Flow when No Standard Design Exists

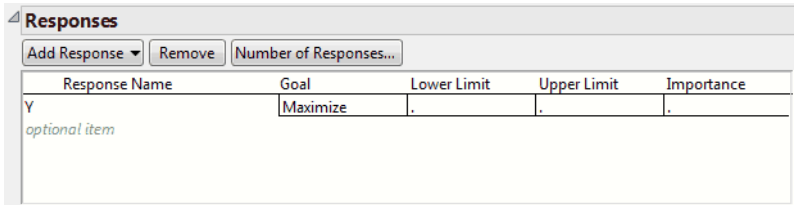


Responses

Use the Responses outline to specify one or more responses.

Tip: When you have completed the Responses outline, consider selecting **Save Responses** from the red triangle menu. This option saves the response names, goals, limits, and importance values in a data table that you can later reload in DOE platforms.

Figure 9.13 Responses Outline



Add Response Enters a single response with a goal type of Maximize, Match Target, Minimize, or None. If you select Match Target, enter limits for your target value. If you select Maximize or Minimize, entering limits is not required but can be useful if you intend to use desirability functions.

Remove Removes the selected responses.

Number of Responses Enters additional responses so that the number that you enter is the total number of responses. If you have entered a response other than the default Y, the Goal for each of the additional responses is the Goal associated with the last response entered. Otherwise, the Goal defaults to Match Target. Click the Goal type in the table to change it.

The Responses outline contains the following columns:

Response Name The name of the response. When added, a response is given a default name of Y, Y2, and so on. To change this name, double-click it and enter the desired name.

Goal, Lower Limit, Upper Limit The Goal tells JMP whether you want to maximize your response, minimize your response, match a target, or that you have no response goal. JMP assigns a Response Limits column property, based on these specifications, to each response column in the design table. It uses this information to define a desirability function for each response. The Profiler and Contour Profiler use these desirability functions to find optimal factor settings. For further details, see the Profiler chapter in *Profiler* and [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

- A Goal of Maximize indicates that the best value is the largest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.

- A Goal of Minimize indicates that the best value is the smallest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Match Target indicates that the best value is a specific target value. The default target value is assumed to be midway between the Lower Limit and Upper Limit.
- A Goal of None indicates that there is no goal in terms of optimization. No desirability function is constructed.

Note: If your target response is not midway between the Lower Limit and the Upper Limit, you can change the target after you generate your design table. In the data table, open the Column Info window for the response column (**Cols > Column Info**) and enter the desired target value.

Importance When you have several responses, the Importance values that you specify are used to compute an overall desirability function. These values are treated as weights for the responses. If there is only one response, then specifying the Importance is unnecessary because it is set to 1 by default.

Editing the Responses Outline

In the Responses outline, note the following:

- Double-click a response to edit the response name.
- Click the goal to change it.
- Click on a limit or importance weight to change it.
- For multiple responses, you might want to enter values for the importance weights.

Response Limits Column Property

The Goal, Lower Limit, Upper Limit, and Importance that you specify when you enter a response are used in finding optimal factor settings. For each response, the information is saved in the generated design data table as a Response Limits column property. JMP uses this information to define the desirability function. The desirability function is used in the Prediction Profiler to find optimal factor settings. For further details about the Response Limits column property and examples of its use, see [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

If you do not specify a Lower Limit and Upper Limit, JMP uses the range of the observed data for the response to define the limits for the desirability function. Specifying the Lower Limit and Upper Limit gives you control over the specification of the desirability function. For more details about the construction of the desirability function, see the Profiler chapter in *Profilers*.

Factors

Add factors in the Factors outline.

Tip: When you have completed the Factors outline, consider selecting **Save Factors** from the red triangle menu. This option saves the factor names, roles, changes, and values in a data table that you can later reload in DOE platforms.

Figure 9.14 Factors Outline

Name	Role	Values
X1	Continuous	-1 1

Specify Factors

Add a Continuous or Categorical factor by clicking its button. Double click on a factor name or level to edit it.

Continue

Continuous Adds a Continuous factor. The data type in the resulting data table is Numeric. A continuous factor is a factor that you can conceptually set to any value between the lower and upper limits you supply, given the limitations of your process and measurement system.

Discrete Numeric Adds a Discrete Numeric factor. A discrete numeric factor can assume only a discrete number of numeric values. These values have an implied order. The data type in the resulting data table is Numeric.

A screening design includes all levels of a discrete numeric factor and attempts to balance the levels. Fit Model treats a discrete numeric factor as a continuous predictor.

The default values for a discrete numeric factor with k levels, where $k > 2$, are the integers 1, 2, ..., k . The default values for a discrete numeric factor with $k = 2$ levels are -1 and 1. Replace the default values with the settings that you plan to use in your experiment.

Categorical Adds a Categorical factor. Click to select or specify the number of levels. The data type in the resulting data table is Character. The value ordering of the levels is the order of the values, as entered from left to right. This ordering is saved in the Value Order column property after the design data table is created.

The default values for a categorical factor are L1, L2, ..., L k , where k is the number of levels that you specify. Replace the default values with level names that are relevant for your experiment.

Remove Removes the selected factors.

Add N Factors Adds multiple factors. Enter the number of factors to add, click **Add Factor**, and then select the factor type. Repeat **Add N Factors** to add multiple factors of different types.

Factors Outline

The Factors outline contains the following columns:

Name The name of the factor. When added, a factor is given a default name of X1, X2, and so on. To change this name, double-click it and enter the desired name.

Role The Design Role of the factor. The Design Role column property for the factor is saved to the data table. This property ensures that the factor type is modeled appropriately. The Role of the factor determines other factor properties that are saved to the data table. See [“Factor Column Properties”](#) on page 307.

Values The experimental settings for the factors.

Editing the Factors Outline

In the Factors outline, note the following:

- To edit a factor name, double-click the factor name.
- To edit a value, click the value in the Values column.

Factor Column Properties

For each factor, various column properties are saved to the design table after you create the design by selecting Make Table in the Screening Design window. These properties are also saved automatically to the data table that is created when you select the Save Factors option. You can find more information about these column properties and related examples in [Appendix A, “Column Properties”](#).

Coding If the Role is Continuous or Discrete Numeric, the Coding column property for the factor is saved. This property transforms the factor values so that the low and high values correspond to -1 and +1, respectively. See [“Coding”](#) on page 742 in the “Column Properties” appendix.

Value Order If the Role is Categorical or if a Block variable is constructed, the Value Order column property for the factor is saved. This property determines the order in which levels of the factor appear. See [“Value Order”](#) on page 759 in the “Column Properties” appendix.

Design Role Each factor is given the Design Role column property. The Role that you specify in defining the factor determines the value of its Design Role column property. When you select a design with a block, that Block factor is assigned the Blocking value. The Design

Role property reflects how the factor is intended to be used in modeling the experimental data. Design Role values are used in the Augment Design platform. See [“Design Role”](#) on page 740 in the “Column Properties” appendix.

Factor Changes Each factor is assigned the Factor Changes column property with the value of Easy. The Factor Changes property reflects how the factor is used in modeling the experimental data. Factor Changes values are used in the Augment Design and Evaluate Design platforms. See [“Factor Changes”](#) on page 755 in the “Column Properties” appendix.

RunsPerBlock For a blocking factor, indicates the maximum allowable number of runs in each block. When a Blocking factor is specified in the Factors outline, the RunsPerBlock column property is saved for that factor. See [“RunsPerBlock”](#) on page 765 in the “Column Properties” appendix.

Choose Screening Type

After you enter your responses and factors and click **Continue**, one of the following results occurs:

- If a standard design can accommodate your factors and levels, two options appear in the Choose Screening Type panel. See [“Choose Screening Type Options”](#) on page 308.
- If no listed standard design exists for your factors and levels, then the Choose Screening Type panel does not appear. The Design Generation outline for constructing a main effects screening design opens. See [“Design Generation”](#) on page 317.

Choose Screening Type Options

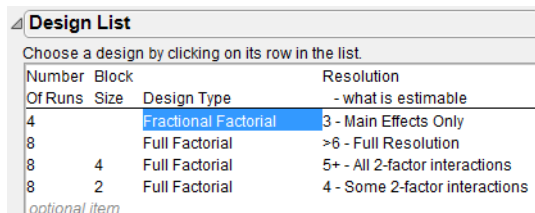
Choose from a list of fractional factorial designs Enables you to select from a list of designs. This option is the default. See [“Choose from a List of Fractional Factorial Designs”](#) on page 308.

Construct a main effects screening design Opens the Design Generation outline where you can specify the number of runs in the main effects screening design. For more information about main effects screening designs, see [“Main Effects Screening Designs”](#) on page 316.

Choose from a List of Fractional Factorial Designs

The list of screening designs that you can choose from includes designs that group the experimental runs into blocks of equal sizes where the size is a power of two. Select the type of screening design that you want to use and click **Continue**.

Figure 9.15 Choosing a Type of Fractional Factorial Design



Design List			
Choose a design by clicking on its row in the list.			
Number Of Runs	Block Size	Design Type	Resolution - what is estimable
4		Fractional Factorial	3 - Main Effects Only
8		Full Factorial	>6 - Full Resolution
8	4	Full Factorial	5+ - All 2-factor interactions
8	2	Full Factorial	4 - Some 2-factor interactions

optional item

The Design List contains the following columns:

Number of Runs Total number of runs in the design.

Block Size Number of runs in a block. The number of blocks is the Number of Runs divided by Block Size.

Design Type Description of the type of design. See [“Design Type”](#) on page 309.

Resolution Gives the resolution of the design and a brief description of the type of aliasing. See [“Resolution as a Measure of Confounding”](#) on page 312.

Design Type

The Design List provides the following types of designs:

- [“Two-Level Full Factorial”](#) on page 309
- [“Two-Level Regular Fractional Factorial”](#) on page 310
- [“Plackett-Burman Designs”](#) on page 310
- [“Mixed-Level Designs”](#) on page 310
- [“Cotter Designs”](#) on page 311

Two-Level Full Factorial

A full factorial design has runs for all combinations of the levels of the factors. The sample size is the product of the levels of the factors. For two-level designs, this is 2^k where k is the number of factors.

Full factorial designs are orthogonal for all effects. It follows that estimates of the effects are uncorrelated. Also, if you remove an effect from the analysis, the values of the other estimates do not change. Their p -values change slightly, because the estimate of the error variance and the degrees of freedom are different.

Full factorial designs allow the estimation of interactions of all orders up to the number of factors. However, most empirical modeling involves only first- or second-order approximations to the true functional relationship between the factors and the responses. From this perspective, full factorial designs are an inefficient use of experimental runs.

Two-Level Regular Fractional Factorial

A regular fractional factorial design also has a sample size that is a power of two. For two-level designs, if k is the number of factors, the number of runs in a regular fractional factorial design is 2^{k-p} where $p < k$. A 2^{k-p} fractional factorial design is a 2^p fraction of the k -factor full factorial design. Like full factorial designs, regular fractional factorial designs are orthogonal.

A full factorial design for k factors provides estimates of all interaction effects up to degree k . But because experimental runs are typically expensive, smaller designs are preferred. In a smaller design, some of the higher-order effects are *confounded* with other effects, meaning that the effects cannot be distinguished from each other. Although a linear combination of the confounded effects is estimable, it is not possible to attribute the variation to a specific effect or effects.

In fact, fractional factorials are designed by deciding in advance which interaction effects are confounded with other interaction effects. Experimenters are usually not concerned with interactions involving more than two factors. Three-way and higher-order interaction effects are often assumed to be negligible.

Plackett-Burman Designs

Plackett-Burman designs are an alternative to regular fractional factorials for screening. The number of runs in a Plackett-Burman design is a multiple of four rather than a power of two. There are no two-level fractional factorial designs with run sizes between 16 and 32. However, there are 20-run, 24-run, and 28-run Plackett-Burman designs.

In a Plackett-Burman design, main effects are orthogonal and two-factor interactions are only partially confounded with main effects. By contrast, in a regular Resolution 3 fractional factorial design, some two-factor interactions are indistinguishable from main effects. Plackett-Burman designs are useful when you are interested in detecting large main effects among many factors and where interactions are considered negligible.

Mixed-Level Designs

For most designs that involve categorical or discrete numeric factors at three or more levels, standard designs do not exist. In such cases, the screening platform generates *main effects screening designs*. These designs are orthogonal or near orthogonal for main effects.

For cases where standard mixed-level designs exist, the possible designs are given in the Design List. The Design List provides fractional factorial designs for pure three-level factorials with up to 13 factors. For mixed two-level and three-level designs, the Design list includes the complete factorials and the orthogonal-array designs listed in Table 9.1.

If your number of factors does not exceed the number for a design listed in the table, you can adapt that design by using an appropriate subset of its columns.

Table 9.1 Table of Mixed-Level Designs

Design	Number of Factors	
	Two-Level	Three-Level
L18 John and L18 Taguchi	1	7
L18 Chakravarty	3	6
L18 Hunter	8	4
L36 Taguchi	11	12

Cotter Designs

Note: By default, Cotter designs are not included in the Design List. To include Cotter designs, deselect **Suppress Cotter Designs** in the Screening Design red triangle menu. To always show Cotter designs, select **File > Preferences > Platforms > DOE** and deselect **Suppress Cotter Designs**.

Cotter designs are useful when you must test many factors, some of which might interact, in a very small number of runs. Cotter designs rely on the principle of effect sparsity. They assume that the sum of effects shows an effect if one of the components of the sum has an active effect. The drawback is that several active effects with mixed signs might sum to near zero, thereby failing to signal an effect. Because of this false-negative risk, many statisticians discourage their use.

For k factors, a Cotter design has $2k + 2$ runs. The design structure is similar to the “vary one factor at a time” approach.

The Cotter design is constructed as follows:

- A run is defined with all factors set to their high level.
- For each of the next k runs, one factor in turn is set at its low level and the others high.
- The next run sets all factors at their low level.

- For each of the next k runs, one factor in turn is set at its high level and the others low.
- The runs are randomized.

When you construct a Cotter design, the design data table includes a set of columns to use as regressors. The column names are of the form <factor name> Odd and <factor name> Even. They are constructed by summing the odd-order and even-order interaction terms, respectively, that contain the given factor.

For example, suppose that there are three factors, A, B, and C. Table 9.2 shows how the values in the regressor columns are calculated.

Table 9.2 Cotter Design Table

Effects Summed for Odd and Even Regressor Columns	
AOdd = A + ABC	AEven = AB + AC
BOdd = B + ABC	BEven = AB + BC
COdd = C + ABC	CEven = BC + AC

The Odd and Even columns define an orthogonal transformation. For this reason, tests for the parameters of the odd and even columns are equivalent to testing the combinations on the original effects.

Resolution as a Measure of Confounding

The *resolution* of a design is a measure of the degree of confounding in the design. The trade-off in screening designs is between the number of runs and the resolution of the design.

Experiments are classified by *resolution number* into these groups:

- Resolution 3 means that some main effects are confounded with one or more two-factor interactions. In order for the main effects to be meaningful, these interactions must be assumed to be negligible.
- Resolution 4 means that main effects are not confounded with other main effects or two-factor interactions. However, some two-factor interactions are confounded with other two-factor interactions.
- Resolution 5 means that there is no confounding between main effects, between main effects and two-factor interactions, and between pairs of two-factor interactions. Some two-factor interactions are confounded with three-factor interactions.
- Resolution 5+ means that the design has resolution greater than 5 but is not a full factorial design.

- Resolution 6 means that there is no confounding between effects of any order. The design is a full factorial design.

A *minimum aberration* design is one that minimizes the number of confoundings for a given resolution. A minimum aberration design of a given resolution minimizes the number of *words* in the defining relation that are of minimum length. For a description of *words*, see [“Change Generating Rules”](#) on page 313. For a discussion of minimum aberration designs, see Fries and Hunter (1984).

Display and Modify Design

In the Design List, if you select a fractional factorial design with all continuous or two-level categorical factors, and possibly a blocking factor, the Display and Modify Design outline opens after you click Continue. Modify your design using the reports in this outline. See [“Modify Generating Rules in a Fractional Factorial Design”](#) on page 323 for an example of changing the generating rules to construct a design.

Note: The Change Generating Rules and Aliasing of Effects outlines do not appear for Plackett-Burman designs or Cotter designs, because interactions are not identically equal to main effects.

Change Generating Rules Specify the defining relation for the design. The defining relation determines which fraction of the full factorial design that JMP provides. See [“Change Generating Rules”](#) on page 313.

Aliasing of Effects Shows the confounding pattern for the fractional factorial design. Click the red arrow at the bottom of the panel to see interactions to a specified order. The interactions and their aliases are presented in a data table.

Coded Design Shows the pattern of high and low values for the factors in each run.

Note: For Cotter designs, the Change Generating Rules and Aliasing of Effects outlines do not apply and are not shown.

Change Generating Rules

The generating rules define the relation used to construct a specific fractional factorial design. The default generator in the screening platform results in a minimum aberration design. A minimum aberration design minimizes the aliasing of lower-order effects. Your experimental situation might require that you define a fraction of the design that provides a coding or aliasing structure that is different from the standard fraction. You can do this by changing the generating rules in the Display and Modify Design outline. For more information about defining relations and generating rules, see Montgomery (2009).

Tip: If you want to create a design to estimate particular effects, consider using the Custom Designer. For more information about custom designs see the “[Custom Designs](#)” chapter on page 65.

The defining relation for a design is determined by the *words* in the generating rules. A *word* is represented by a product of factors, but it is interpreted as the element wise product of the entries in the design matrix for those columns. A defining relation consists of words whose product is a column of ones, called the *identity*.

Figure 9.16 shows the default-generating rules for a 2^{5-3} design (five factors and eight runs).

Figure 9.16 Generating Rules for the Standard 2^{5-3} Design

The screenshot shows the 'Display and Modify Design' window. The 'Change Generating Rules' panel has three columns: Factors, Temperature, and Concentration. The 'Factors' column lists Feed Rate, Catalyst, Stir Rate, and +/-, all with checked boxes. The 'Temperature' column has a checked box for the +/- factor. The 'Concentration' column has checked boxes for Catalyst, Stir Rate, and +/-, but an unchecked box for Feed Rate. An 'Apply' button is below the columns. The 'Aliasing of Effects' panel shows a table of effects and their aliases. The 'Coded Design' panel is partially visible at the bottom.

Factors	Temperature	Concentration
Feed Rate ☒	☐	☐
Catalyst ☒	☒	☒
Stir Rate ☒	☒	☒
+/- ☒	☒	☒

Apply

Effects	Aliases
Feed Rate	= Temperature*Concentration
Catalyst	= Stir Rate*Concentration
Stir Rate	= Catalyst*Concentration
Temperature	= Feed Rate*Concentration
Concentration	= Feed Rate*Temperature = Catalyst*Stir Rate
Feed Rate*Catalyst	= Stir Rate*Temperature
Feed Rate*Stir Rate	= Catalyst*Temperature

Coded Design

In each column of the Change Generating Rules panel, the factor listed at the top and the factors in the column whose boxes are selected form a word in the defining relation. For example, the first column indicates that Temperature = Feed Rate*Catalyst*Stir Rate is a word in the defining relation.

- If the +/- box is selected, the sign associated with the generating rule is positive and the corresponding word equals the identity.
- If the +/- box is not selected, the sign associated with the generating rule is negative and the corresponding word equals minus the identity.

See “[Obtain the Defining Relations in the \$2^{5-3}\$ Design](#)” on page 315.

The *principal fraction* of a full factorial design is the fractional factorial design obtained by setting all the defining relations equal to the identity. By default, the factorial design that JMP provides is the principal fraction. Notice that the +/- box is selected by default for all generating rules, so that each word in the defining relation equals the identity.

Generating rules determine the coding and aliasing of effects for the design. In some cases, you might want to use a fraction that results in a coding or an aliasing structure that differs from that of the standard fraction.

- To change the generating rules, select the appropriate boxes.
- To see the effect of your selections on the Aliasing of Effects results and on the Coded Design, click **Apply**.

For an example, see [“Change the Generating Rules to Obtain a Different Fraction”](#) on page 326.

Obtain the Defining Relations in the 2^{5-3} Design

Figure 9.16 shows two columns of check boxes:

- The first column represents the word $\text{Temperature} = \text{Feed Rate} * \text{Catalyst} * \text{Stir Rate}$.
- The second column represents the word $\text{Concentration} = \text{Catalyst} * \text{Stir Rate}$.

Define I to represent a column consisting of the values $+1$. Because all factor levels are -1 or $+1$, the word in the first column is equivalent to $\text{Temperature} * \text{Feed Rate} * \text{Catalyst} * \text{Stir Rate} = I$. The word in the second column is equivalent to $\text{Concentration} * \text{Catalyst} * \text{Stir Rate} = I$. Together, these give the defining relations for the 2^{5-3} design:

$$I = \text{Temperature} * \text{Feed Rate} * \text{Catalyst} * \text{Stir Rate} = \text{Concentration} * \text{Catalyst} * \text{Stir Rate}$$

Obtain the Aliasing of Effects Relations in the 2^{5-3} Design

The aliasing structure in the Aliasing of Effects outline is determined by the defining relations and the fact that factor levels are $+1$ and -1 . Recall that the first generating rule is $\text{Temperature} = \text{Feed Rate} * \text{Catalyst} * \text{Stir Rate}$ and the second is $\text{Concentration} = \text{Catalyst} * \text{Stir Rate}$.

To obtain the first relation in the Aliasing of Effects outline, notice that applying these two generating rules gives the expression:

$$\text{Temperature} = \text{Feed Rate} * \text{Catalyst} * \text{Stir Rate} = \text{Feed Rate} * \text{Concentration}$$

- The second equality follows from replacing $\text{Catalyst} * \text{Stir Rate}$ by Concentration using the second generating rule.

Now, post-multiply the first and third expressions by Concentration to obtain the following expression:

$$\text{Temperature} * \text{Concentration} = \text{Feed Rate} * \text{Concentration} * \text{Concentration}$$

Because the column for Concentration in the design matrix contains values of -1 and $+1$, the term $\text{Concentration} * \text{Concentration}$ represents a column of $+1$ values. The expression becomes the first alias relation shown in the Aliasing of Effects outline:

$$\text{Temperature} * \text{Concentration} = \text{Feed Rate} * I = \text{Feed Rate}$$

The other alias relations can be obtained using similar calculations.

Main Effects Screening Designs

If an experiment involves categorical or discrete numeric factors, or if the number of runs is constrained, it might not be possible to construct an orthogonal design for screening main effects. However, a main effects screening design can be constructed. See Lekivetz et al. (2015).

A main effects screening design is a design with good balance properties as described by a Chi-square criterion. See “Chi-Square Efficiency” on page 316. Such designs have desirable statistical properties for main effect models.

The algorithm used to generate the design attempts to construct an orthogonal array of strength two. Strength-two orthogonal arrays permit orthogonal estimation of main effects when interactions are negligible. These arrays are ideal for screening designs. Regular fractional factorial designs of Resolution 3 and Plackett-Burman designs are examples of strength-two orthogonal arrays.

Consider all possible pairs of levels for factors in the design. The algorithm attempts to balance the number of pairs of levels as far as possible. Given that a fixed number of columns has been generated, a new balanced column is randomly constructed. A measure is defined that reflects the degree of balance achieved for pairs that involve the new column. The algorithm attempts to minimize this measure by interchanging levels within the new column.

Chi-Square Efficiency

Suppose that a design has n runs and p factors corresponding to the columns of the design matrix.

- Denote the levels of factors k and l by $a = 0, 1, \dots, s_k - 1$ and $b = 0, 1, \dots, s_l - 1$, respectively.
- Denote the number of times that the combination of levels (a, b) appears in columns k and l by $n_{kl}(a, b)$.

A measure of the lack of orthogonality evidenced by columns k and l is given by the following expression:

$$\chi_{kl}^2 = \sum_{a=0}^{s_k-1} \sum_{b=0}^{s_l-1} \frac{[n_{kl}(a, b) - n/(s_k s_l)]^2}{n/(s_k s_l)}$$

A measure of the average non-orthogonality of the design is given by this expression:

$$\chi^2 = \sum_{1 \leq k < l \leq p} \chi_{kl}^2 / [p(p-1)/2]$$

The maximum possible value of χ^2 , denoted χ_{max}^2 , is obtained. The chi-square efficiency of a design is defined as follows:

$$\text{Chi-Square Efficiency} = 100(1 - \chi^2 / \chi_{max}^2)$$

Chi-square efficiency indicates how close χ^2 is to zero, relative to a design in which pairs of levels show extreme lack of balance.

Design Generation

When you construct a main effects screening design, the Design Generation outline enables you to specify the number of runs. To generate the design, click **Make Design**.

Minimum A lower bound on the number of runs necessary to avoid failures in design generation. When you select Minimum, the resulting design is saturated. There are no degrees of freedom for error.

Note: If you select the Minimum number of runs, there is no error term for testing. You cannot test parameter estimates. This choice is appropriate only when the cost of additional runs is prohibitive.

Default Suggests the number of runs. This value is based on heuristics for creating a balanced design with at least four runs more than the Minimum number of runs.

User Specified Specify the number of runs that you want. Enter that value into the Number of Runs text box. This option enables you to balance the cost of additional runs against the potential gain in information.

Design

The Design outline shows the runs for the main effects screening design. To change the run order for your design table, you can select Run Order options in the Output Options panel before generating the table.

Design Evaluation

Note: The Design Evaluation outline is not shown for Cotter designs.

The Design Evaluation outline provides a number of ways to evaluate the properties of the generated design. Open the Design Evaluation outline to see the following options:

Power Analysis Enables you to explore your ability to detect effects of given sizes.

Prediction Variance Profile Shows the prediction variance over the range of factor settings.

Fraction of Design Space Plot Shows how much of the model prediction variance lies below (or above) a given value.

Prediction Variance Surface Shows a surface plot of the prediction variance for any two continuous factors.

Estimation Efficiency For each parameter, gives the fractional increase in the length of a confidence interval compared to that of an ideal (orthogonal) design, which might not exist. Also gives the relative standard error of the parameters.

Alias Matrix Gives coefficients that indicate the degree by which the model parameters are biased by effects that are potentially active, but not in the model.

Color Map on Correlations Shows the absolute correlation between effects on a plot using an intensity scale.

Design Diagnostics Indicates the optimality criterion used to construct the design. Also gives efficiency measures for your design.

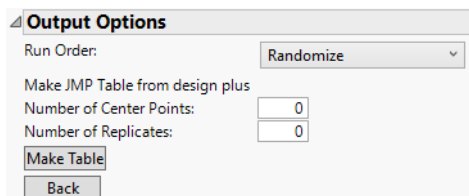
Note: The model used for the design diagnostics contains all main effects and two-factor interactions when all two-factor interactions are estimable. Otherwise, the model contains all main effects.

For more details about the Design Evaluation panel, see [“Design Evaluation”](#) on page 449 in the “Evaluate Designs” chapter.

Output Options

Specify details for the output data table in the Output Options panel. When you have finished, click **Make Table** to construct the data table for the design. Figure 9.17 shows the Output Options panel for a standard design selected from the Design List. For a main effects screening design, only Run Order is available.

Figure 9.17 Select the Output Options



Run Order

The **Run Order** options determine the order of the runs in the design table. Choices include the following:

Keep the Same Rows in the design table are in the same order as in the Coded Design or Design outlines.

Sort Left to Right Columns in the design table are sorted from left to right.

Randomize Rows in the design table are in random order.

Sort Right to Left Columns in the design table are sorted from right to left.

Randomize within Blocks Rows in the design table are in random order within the blocks. (Not available if you select **Construct a main effects screening design**.)

Center Points and Replicates

Number of Center Points Specifies how many additional runs to add as center points to the design. A center point is a run where every continuous factor is set at the center of the factor's range. This option is not available if you select **Construct a main effects screening design**.

Suppose that your design includes both continuous and categorical factors. If you request center points in the Output Options panel, the center points are distributed as follows:

1. The settings for the categorical factors are ordered using the value ordering specified in the Factors outline.
2. One center point is assigned to each combination of the settings of the categorical factors in order, and this is repeated, until all center points are assigned.

Number of Replicates For designs in the Design List, specify the number of times to replicate the entire design, including center points. One replicate doubles the number of runs. This option is not available if you select **Construct a main effects screening design**.

Note: If you request center points or replicates and click Make Table repeatedly, these actions are applied to the most recently constructed design table.

Make Table

Click **Make Table** to create a data table that contains the runs for your experiment. In the table, the high and low values that you specified appear for each run.

Figure 9.18 The Design Data Table

Design Fractional Factorial	Pattern	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration	Percent Reacted
Screening	1 -+++-	10	2	120	140	6	•
Model	2 +-+--	15	1	120	140	3	•
Evaluate Design	3 -+++-	10	2	100	180	3	•
DOE Dialog	4 +++++	15	2	120	180	6	•
Columns (7/0)	5 ---+-	10	1	120	180	3	•
Pattern *	6 +++++	15	1	100	180	6	•
Feed Rate *	7 -----	10	1	100	140	6	•
Catalyst *	8 +-----	15	2	100	140	3	•
Stir Rate *							
Temperature *							
Concentration *							
Percent Reacted *							

The name of the table is the design type that generated it.

The design table includes the following scripts:

Screening Runs the DOE > Classical > Two Level Screening > Fit Two Level Screening platform. Only provided when all factors are at two levels.

Model Runs the Analyze > Fit Model platform.

Evaluate Design Runs the DOE > Design Diagnostics > Evaluate Design platform.

DOE Dialog Re-creates the Screening Design window that you used to generate the design table. The script also contains the random seed used to generate your design.

Run the **Screening** or **Model** scripts to analyze the data.

If the design was selected from the Design List, the design table contains a **Pattern** column.

The **Pattern** column contains entries that summarize the run in the given row. Low settings are denoted by “-”, high settings by “+”, and center points by “0”. **Pattern** can be useful as a label variable in plots.

Screening Design Options

The Screening Design red triangle menu contains the following options:

Save Responses Saves the information in the Responses panel to a new data table. You can then quickly load the responses and their associated information into most DOE windows. This option is helpful if you anticipate reusing the responses.

Load Responses Loads responses that you saved using the Save Responses option.

Save Factors Saves the information in the Factors panel to a new data table. Each factor's column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Note: It is possible to create a factors table by entering data into an empty table, but remember to assign each column an appropriate Design Role. Do this by right-clicking on the column name in the data grid and selecting **Column Properties > Design Role**. In the Design Role area, select the appropriate role.

Load Factors Loads factors that you saved using the Save Factors option.

Save Constraints (Unavailable for some platforms) Saves factor constraints that you defined in the Define Factor Constraints or Linear Constraints outline into a data table, with a column for each constraint. You can then quickly load the constraints into most DOE windows.

In the constraint table, the first rows contain the coefficients for each factor. The last row contains the inequality bound. Each constraint's column contains a column property called `ConstraintState` that identifies the constraint as a "less than" or a "greater than" constraint. See "[ConstraintState](#)" on page 766 in the "Column Properties" appendix.

Load Constraints (Unavailable for some platforms) Loads factor constraints that you saved using the Save Constraints option.

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction
- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.

- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.
- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click Apply, a formula for the simulated response values is saved in a new column called <Y> Simulated, where Y is the name of the response. Clicking Apply again updates the formula and values in <Y> Simulated.

For additional details, see [“Simulate Responses”](#) on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Save X Matrix Saves the Moments Matrix and Model Matrix to table scripts in the design data table. These scripts contain the moments and design matrices. See [“Save X Matrix”](#) on page 112 in the “Custom Designs” chapter.

Caution: For a design with nominal factors, the Model Matrix saved by the Save X Matrix option is *not* the coding matrix used in fitting the linear model. You can obtain the coding matrix used for fitting the model by selecting the option Save Columns > Save Coding Table in the Fit Model report that you obtain when you run the Model script

Suppress Cotter Designs Excludes Cotter designs from the Design List. This option is selected by default. Deselect it to show Cotter designs in the Design List.

Note: You can set a preference to always show Cotter designs. Select File > Preferences > Platforms > DOE and deselect Suppress Cotter Designs.

Number of Starts (Available only for Main Effects Screening Designs.) Specify the maximum number of times that the algorithm regenerates entire designs from scratch, attempting to optimize the final design.

Design Search Time (Available only for Main Effects Screening Designs.) Specify the maximum number of seconds spent searching for a design. The default search time is 15 seconds.

If the iterations of the algorithm require more than a few seconds, a Computing Design progress window appears. The progress bar displays *Chi2 Efficiency*. See [“Chi-Square](#)

[Efficiency](#)” on page 316. If you click **Cancel** in the progress window, the calculation stops and gives the best design found at that point.

Note: You can set a preference for Design Search Time. Select **File > Preferences > Platforms > DOE**. Select **Design Search Time** and enter the maximum number of seconds. If an orthogonal array is found, the search terminates. In certain situations where more time is required, JMP automatically extends the search time.

Number of Column Starts (Available only for Main Effects Screening Designs.) Specify the maximum number of times that the algorithm attempts to optimize a given column before moving on to constructing the next column. The default number of column starts is 50. See [“Main Effects Screening Designs”](#) on page 316.

Advanced Options > Set Delta for Power Specify the difference in the mean response that you want to detect for model effects. See [“Set Delta for Power”](#) on page 116 in the “Custom Designs” chapter.

Save Script to Script Window Creates the script for the design that you specified in the Screening Design window and places it in an open script window.

Additional Examples of Screening Designs

- [“Modify Generating Rules in a Fractional Factorial Design”](#)
- [“Plackett-Burman Design”](#)

Modify Generating Rules in a Fractional Factorial Design

The following example, adapted from Meyer, *et al.* (1996), shows how to use the Screening Design platform when you have many factors. In this example, a chemical engineer investigates the effects of five factors on the percent reaction of a chemical process. The factors are:

- Feed Rate - the amount of raw material added to the reaction chamber in liters per minute
- Catalyst (as a percent)
- Stir Rate - the RPMs of a propeller in the chamber
- Temperature (in degrees Celsius)
- Concentration of reactant

Production constraints limit the size of the experiment to no more than twelve runs. You decide to consider the 8-run fractional factorial design and the 12-run Plackett-Burman design. Also, you suspect the following statements to be true:

- The Temperature*Concentration interaction is active, so you want a design that does not alias this interaction with a main effect.
- The Catalyst*Temperature* interaction is not likely to be active.
- The Stir Rate*Concentration interaction is not likely to be active.

Use this information in constructing your design.

Create the Standard Fractional Factorial Design

To create the standard fractional factorial design, do the following:

- [“Specify the Response”](#) on page 324
- [“Specify the Factors”](#) on page 324 or [“Specify Factors Manually”](#) on page 325
- [“Choose a Design”](#) on page 325

Specify the Response

1. Select **DOE > Classical > Two Level Screening > Screening Design**.
2. Double-click Y under Response Name and type Percent Reacted.

Note that the default Goal is Maximize. The Goal is to maximize the response, but the minimum acceptable reaction percentage is 90 (Lower Limit) and the upper limit is 100 (Upper Limit).

3. Click under Lower Limit and type 90.
4. Click under Upper Limit and type 100.
5. Leave the area under Importance blank.

Because there is only one response, that response is given Importance 1 by default.

See Figure 9.19 for the completed Responses outline. Now, specify the factors.

Specify the Factors

You can enter the factors manually or automatically:

- To enter the factors automatically, use the Reactor Factors.jmp data table:
 1. Select **Help > Sample Data Library** and open Design Experiment/Reactor Factors.jmp.
 2. Click the Screening Design red triangle and select **Load Factors**. Proceed to [“Choose a Design”](#) on page 325.
- To enter the factors manually, follow the steps below.

Specify Factors Manually

1. Add five continuous factors by entering 5 in the **Add N Factors** box and clicking **Continuous**.
2. Change the default factor names (X1-X5) to Feed Rate, Catalyst, Stir Rate, Temperature, and Concentration.
3. Enter the low and high values, as follows:
 - Feed Rate: 10, 15
 - Catalyst: 1, 2
 - Stir Rate: 100, 120
 - Temperature: 140, 180
 - Concentration: 3, 6

Figure 9.19 Responses and Factors Outlines

Screening Design

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Percent Reacted <i>optional item</i>	Maximize	90	100	1

Factors

Continuous Discrete Numeric Categorical Remove Add N Factors 1

Name	Role	Values
Feed Rate	Continuous	10 15
Catalyst	Continuous	1 2
Stir Rate	Continuous	100 120
Temperature	Continuous	140 180
Concentration	Continuous	3 6

Specify Factors
Add a Continuous or Categorical factor by clicking its button. Double click on a factor name or level to edit it.

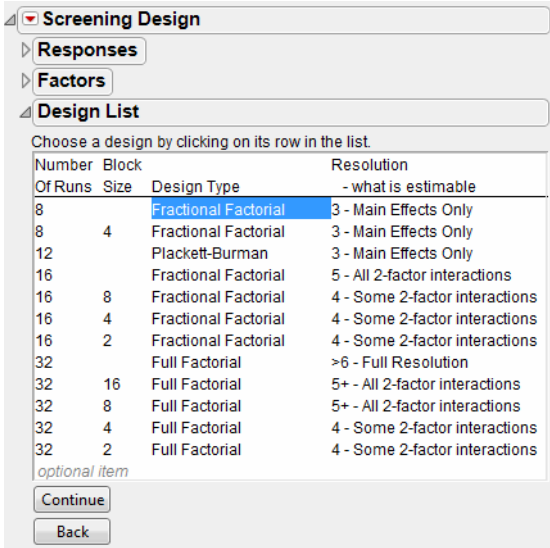
Continue

Choose a Design

1. Click **Continue**.
2. From the Choose Screening Type panel, accept the default selection to **Choose from a list of fractional factorial designs** and click **Continue**.

Designs for the factors and levels that you specified are listed in the Design List (Figure 9.20).

Figure 9.20 Fractional Factorial Designs for Five Continuous Factors



3. The design that you want is the first in the list and happens to be selected by default (Figure 9.20). Accept that selection and click **Continue**.

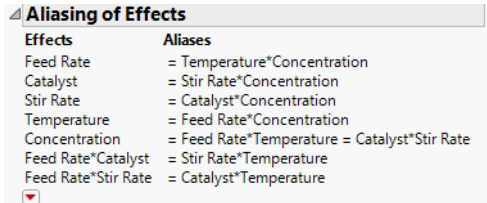
Because you are limited to eight runs and have no blocking factor, your best design option is the 8-run fractional factorial design with no blocks. This design is a 2^{5-2} fractional factorial design. It is one quarter of the full factorial design for five factors.

Change the Generating Rules to Obtain a Different Fraction

In this example, you want to know whether the Temperature*Concentration interaction is confounded with a main effect. Use the Display and Modify Design outline to view the aliasing structure for the design that you selected and to change it, if appropriate.

1. Open the Aliasing of Effects outline.

Figure 9.21 Aliasing of Effects Outline



The Temperature*Concentration interaction, which you suspect is active, is confounded with Feed Rate, a main effect. You want to change the generating rules to construct a

design where Feed Rate is aliased with effects that you suspect are inactive, and where the Temperature*Concentration interaction is not aliased with a main effect.

2. Open the Change Generating Rules outline.

The default-generating rules give you the standard (or *principal*) one-quarter fraction of the full factorial design. Recall that you suspect that the Catalyst*Temperature and Stir Rate*Concentration interactions are not likely to be active. Redefine the generating rules so that these two interactions are confounded with Feed Rate. The redefined generating rules give you a different one-quarter fraction of the full factorial design.

3. Do the following:

- Deselect Stir Rate in the Temperature column.
- Deselect Catalyst in the Concentration column.
- Select Feed Rate in the Concentration column.

Figure 9.22 New Generating Rules

Factors	Temperature	Concentration
Feed Rate ☒		☒
Catalyst ☒		☐
Stir Rate ☐		☒
+/- ☒	☒	☒

Apply

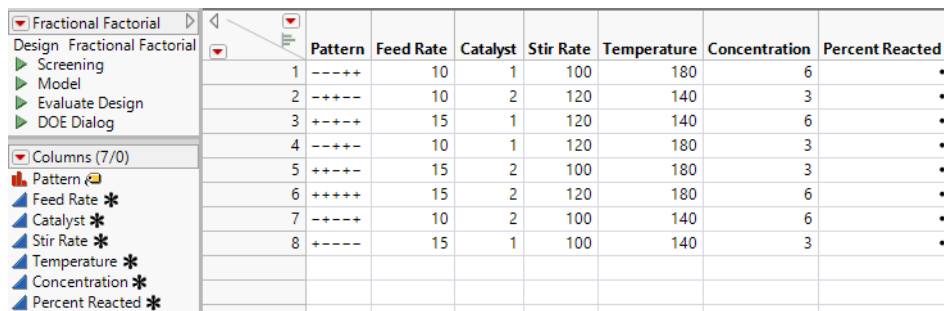
4. Click Apply.

Figure 9.23 Aliasing of Effects Outline for Modified Generating Rules

Effects	Aliases
Feed Rate	= Catalyst*Temperature = Stir Rate*Concentration
Catalyst	= Feed Rate*Temperature
Stir Rate	= Feed Rate*Concentration
Temperature	= Feed Rate*Catalyst
Concentration	= Feed Rate*Stir Rate
Catalyst*Stir Rate	= Temperature*Concentration
Catalyst*Concentration	= Stir Rate*Temperature

In the design that you have defined, Feed Rate is confounded with Catalyst*Temperature and Stir Rate*Concentration. Also, the Temperature*Concentration interaction is now confounded with the two-way interaction Catalyst*Stir Rate.

5. In the Output Options outline, accept the default Run Order setting of **Randomize** and click **Make Table**.

Figure 9.24 Eight-Run Fractional Factorial Design Table


	Pattern	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration	Percent Reacted
1	-----	10	1	100	180	6	•
2	--++--	10	2	120	140	3	•
3	+-++--	15	1	120	140	6	•
4	---+-	10	1	120	180	3	•
5	++--	15	2	100	180	3	•
6	++++	15	2	120	180	6	•
7	-----	10	2	100	140	6	•
8	+-++--	15	1	100	140	3	•

The design table shows the design that you constructed. Notice that the table contains a column for the response that you defined in the Screening window, **Percent Reacted**, where you can record your experimental results.

The **Screening**, **Model**, and **DOE Dialog** scripts are also included. For more information about these scripts, see [“Make Table”](#) on page 319.

Analyze the Results

Next you conduct the experiment, record your data, and proceed to analyze the results.

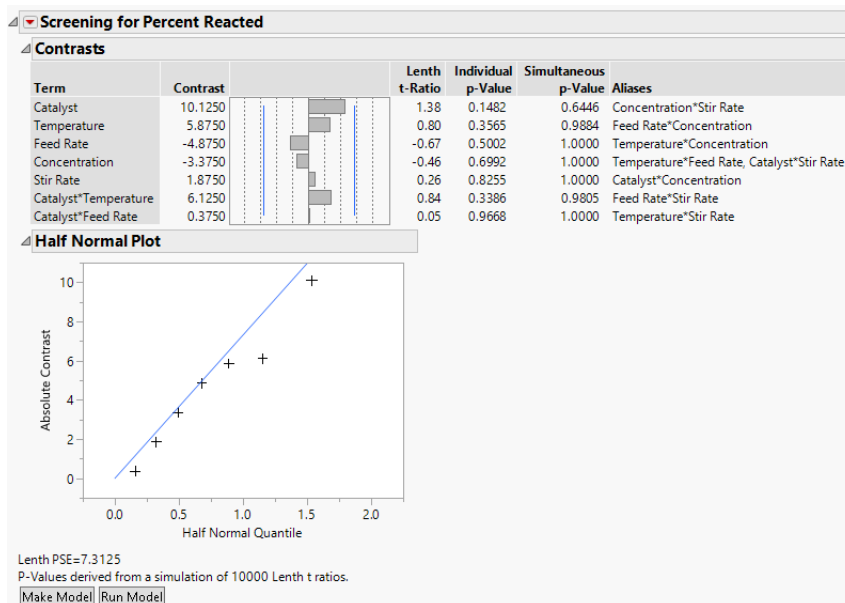
1. Select **Help > Sample Data Library** and open **Design Experiment/Reactor 8 Runs.jmp**.

You can estimate seven effects with your eight runs. Of these, you expect only a few to be active. Because you want to estimate seven effects, there are no degrees of freedom for error. For these reasons, you use the Screening platform to analyze the results.

2. Run the **Screening** script in the data table.

The **Screening** script launches the Screening platform (**DOE > Classical > Two Level Screening > Fit Two Level Screening**) for your response and factors.

Figure 9.25 Report for Screening Example



Note: Since the p -values are obtained using a simulation-based technique, your p -values might not precisely match those shown here.

The report shows both Individual and Simultaneous p -values based on Lenth t -ratios. None of the effects are significant, even with respect to the Individual p -values. The Half Normal Plot suggests that the effects reflect only random noise.

For more information about the Screening report, see [“The Screening Report”](#) on page 341 in the [“The Fit Two Level Screening Platform”](#) chapter.

Plackett-Burman Design

The Fractional Factorial example shows an 8-run fractional factorial design for five continuous factors. But suppose you can afford 4 additional runs. In this example, construct a 12-run Plackett-Burman design. To facilitate completing the Screening window, use the Load Responses and Load Factors commands.

Create the Plackett-Burman Design

1. Select **DOE > Classical > Two Level Screening > Screening Design**.
2. Select **Help > Sample Data Library** and open **Design Experiment/Reactor Response.jmp**.
3. Click the Screening Design red triangle and select **Load Responses**.

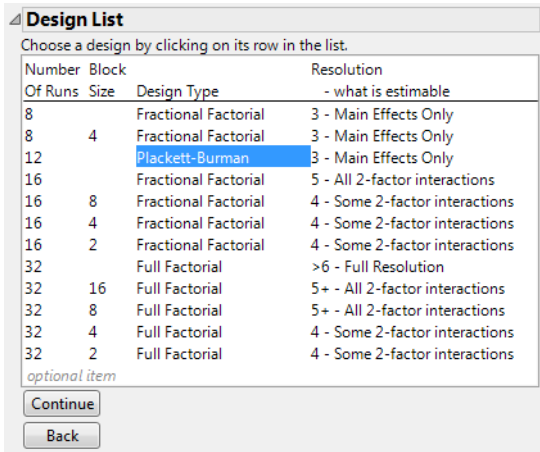
- 4. Select **Help > Sample Data Library** and open **Design Experiment/Reactor Factors.jmp**.
- 5. Click the Screening Design red triangle and select **Load Factors**.
The **Load Responses** and **Load Factors** commands fill in the Responses and Factors outlines with the response and factor names, goal and limits for the response, and values for the factors. See Figure 9.19 for the completed Responses and Factors outlines.
- 6. Click **Continue**.

Note: Setting the random seed in the next step reproduces the run order shown in this example. In constructing a design on your own, this step is not necessary.

- 7. (Optional) Click the Screening Design red triangle, select **Set Random Seed**, type 34567, and click **OK**.
- 8. From the Choose Screening Type panel, accept the default selection to **Choose from a list of fractional factorial designs** and click **Continue**.
- 9. Select the Plackett-Burman design.

Plackett-Burman designs with run sizes that are not a power of two tend to have complex aliasing structures. In particular, main effects can be partially aliased with several two-way interactions. See [“Evaluate the Design”](#) on page 331. Notice that the 12-run Plackett-Burman design is designated as having Resolution 3.

Figure 9.26 Design List Showing Plackett-Burman Screening Design



- 10. Click **Continue**.
- 11. Click **Make Table**.

Figure 9.27 Design Table for Plackett-Burman Design

	Pattern	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration	Percent Reacted
1	+++++	15	2	120	180	6	•
2	+-----	15	1	100	140	6	•
3	++----	15	1	120	180	6	•
4	---++-	10	2	120	180	3	•
5	++----	15	2	100	140	3	•
6	-----	10	1	100	180	3	•
7	--+++-	10	2	100	180	6	•
8	++++--	15	2	120	140	3	•
9	---++-	10	1	120	140	6	•
10	-----	10	1	120	140	3	•
11	-----	10	2	100	140	6	•
12	++++--	15	1	100	180	3	•

A column called Percent Reacted is included in the design table. You should conduct your experimental runs in the order shown in the table, recording your results in the Percent Reacted column.

Evaluate the Design

1. Return to your Screening Design window. If you have closed this window, run the DOE Dialog script in your design table.
2. Open the **Design Evaluation > Color Map on Correlations** outline.

Figure 9.28 Color Map for Absolute Correlations



The diagonal cells have correlations of one, as expected. White cells correspond to effects that have correlations equal to 0. The gray and black cells correspond to effects that have correlations greater than zero. Place your cursor over a few of these cells with your cursor to see the effects involved and their absolute correlations. For example, notice that **Feed Rate** is correlated with several two-way and three-way interactions.

3. Open the Alias Matrix outline.

Figure 9.29 Alias Matrix - Partial View Showing Up to Two-Way Interactions

Alias Matrix												
Effect	Feed Rate*Catalyst	Feed Rate*Stir Rate	Feed Rate*Temperature	Feed Rate*Concentration	Catalyst*Stir Rate	Catalyst*Temperature	Catalyst*Concentration	Stir Rate*Temperature	Stir Rate*Concentration	Temperature*Concentration		
Intercept	0	0	0	0	0	0	0	0	0	0		
Feed Rate	0	0	0	0	0.333	-0.33	-0.33	0.333	0.333	0.333	0	0
Catalyst	0	0.333	-0.33	-0.33	0	0	0	0.333	-0.33	0	0.333	0.333
Stir Rate	0.333	0	0.333	0.333	0	0.333	-0.33	0	0	0.333	0	0.333
Temperature	-0.33	0.333	0	0.333	0.333	0	0.333	0	0.333	0	0.333	0
Concentration	-0.33	0.333	0.333	0	-0.33	0.333	0	0.333	0	0.333	0	0

Because the design is orthogonal for the main effects, the Alias Matrix gives the numerical values of the correlations between effects. See “[Alias Matrix](#)” on page 462 in the “Evaluate Designs” chapter. For example, notice that **Feed Rate** is partially aliased with six two-way interactions and with four three-way interactions. These are the interactions corresponding to the entries of 0.333 and -0.33 in the row for **Feed Rate**.

Analyze the Results

The data table Plackett-Burman.jmp contains the results of the designed experiment. Recall that you suspect that the Temperature*Concentration interaction is active. You proceed under the assumption that this is the only potentially active interaction.

1. Select **Help > Sample Data Library** and open Design Experiment/Plackett-Burman.jmp.
2. Run the **Model** script by clicking the icon to its left.
3. Select Temperature in the Select Columns list and Concentration in the Construct Model Effects list.
4. Click **Cross**.
5. Click **Run**.

Figure 9.30 Parameter Estimates for Full Model

Parameter Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Intercept	66.166667	2.639444	25.07	<.0001*
Feed Rate(10,15)	-2.333333	2.850926	-0.82	0.4503
Catalyst(1,2)	10.5	2.850926	3.68	0.0142*
Stir Rate(100,120)	1.333333	2.850926	0.47	0.6597
Temperature(140,180)	5	2.639444	1.89	0.1167
Concentration(3,6)	-0.5	2.639444	-0.19	0.8572
Concentration*Temperature	-6.5	3.232646	-2.01	0.1006

The Actual by Predicted Plot indicates no lack of model fit. The Parameter Estimates report shows that Catalyst is significant at the 0.05 level and that the Concentration*Temperature interaction is almost significant at the 0.10 level.

Reduce the Model

You want to identify those effects that have the most impact on the response. To see these active effects more clearly, remove insignificant effects using the Effect Summary outline.

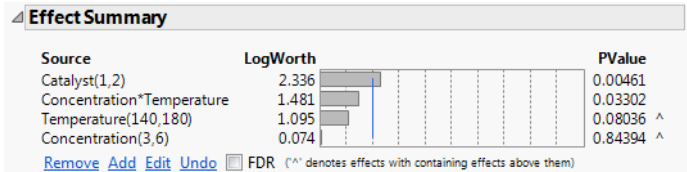
Figure 9.31 Effect Summary Outline for Full Model

Effect Summary			
Source	LogWorth		PValue
Catalyst(1,2)	1.846		0.01425
Concentration*Temperature	0.998		0.10055
Temperature(140,180)	0.933		0.11671 ^
Feed Rate(10,15)	0.346		0.45033
Stir Rate(100,120)	0.181		0.65969
Concentration(3,6)	0.067		0.85720 ^
Remove Add Edit ☐ FDR (^ denotes effects with containing effects above them)			

Although Concentration is the least significant effect, it is involved in a higher-order interaction (Concentration*Temperature), as indicated by the caret to the right of its PValue. Based on the principle of effect heredity, Concentration should not be removed from the model while the Concentration*Temperature interaction remains in the model. See “Effect Heredity” on page 62 in the “Starting Out with DOE” chapter. The next least significant effect is Stir Rate.

- 1. In the Effect Summary outline, select Stir Rate and click **Remove**.
Feed Rate is the next least significant effect that can be removed.
- 2. In the Effect Summary outline, select Feed Rate and click **Remove**.

Figure 9.32 Effect Summary Outline for Reduced Model



The PValue column indicates that the Catalyst main effect and the Concentration*Temperature interaction are both significant at the 0.05 level. The model should not be reduced any further. If all other interactions are inactive or negligible, then you can conclude that Catalyst and the Concentration*Temperature interaction are active effects.

Chapter 10

The Fit Two Level Screening Platform

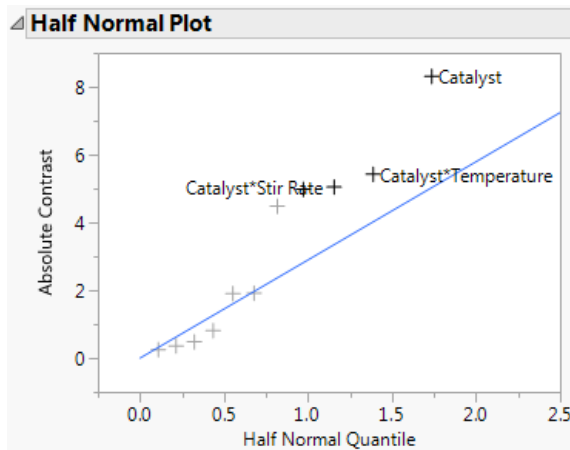
Analyze Data from Screening Experiments

The Fit Two Level Screening platform is a modeling platform that you can use to analyze experimental data results from a screening design. The Fit Two Level Screening platform helps you identify effects that have a large impact on the response.

The Fit Two Level Screening platform is based on the principle of effect sparsity (Box and Meyer 1986). This principle asserts that relatively few of the effects that you study in a screening design are active. Most are inactive, meaning that their true effects are negligible and that their estimates can be treated as random error.

A screening design often provides no degrees of freedom for error when the model of interest includes interaction terms. Consequently, classical tests for effects are not available. In such cases, the Fit Two Level Screening platform is particularly useful.

Figure 10.1 Half Normal Plot from Fit Two Level Screening Report



Contents

Overview of the Fit Two Level Screening Platform	337
Example of Fit Two Level Screening Platform	337
Launch the Fit Two Level Screening Platform	340
The Screening Report	341
Contrasts	342
Half Normal Plot	342
Make or Run Model	343
Additional Fit Two Level Screening Analysis Examples	343
Analyze a Plackett-Burman Design	343
Effect Heredity Example	344
Analyze a Supersaturated Design	346
Technical Details for the Fit Two Level Screening Platform	348
Order of Effect Entry	348
Fit Two Level Screening as an Orthogonal Rotation	348
Lenth's Pseudo-Standard Error	349
Lenth t-Ratios	349

Overview of the Fit Two Level Screening Platform

The analysis of screening designs depends on the principle of *effect sparsity*, where most of the variation in the response is explained by a small number of effects. Under this principle, effects with small estimates are used to estimate the error in the model. This then allows one to test whether the larger effects are active.

You can analyze data from a screening experiment using Fit Model (Analyze > Fit Model) or Fit Two Level Screening (DOE > Classical > Two Level Screening > Fit Two Level Screening). Use the following guidelines to select the appropriate modeling platform:

- If your factors are all two-level and orthogonal, all of the statistics in the Fit Two Level Screening platform are appropriate.
- If you have data from a highly supersaturated main effect design, the Fit Two Level Screening platform is effective in selecting active factors, but it is not effective at estimating the error or the significance. The Monte Carlo simulation to produce p -values uses assumptions that are not valid for this case.
- If you have a categorical or a discrete numeric factor with more than two levels, the Fit Two Level Screening platform is not appropriate. JMP treats the associated model terms as continuous. For such factor, the variation is scattered across main effects and polynomial effects. In this situation, it is recommended that you use the Fit Model platform.
- If your data are not orthogonal, the constructed estimates in the Fit Two Level Screening Platform are different from standard regression estimates. JMP can identify large effects, but it does not effectively test each effect. This is because effects are artificially orthogonalized as they are entered into the model, making effects that enter early in the model appear less significant than in standard regression. See [“Order of Effect Entry”](#) on page 348.
- For mixture designs, the Fit Two Level Screening platform is not appropriate. See instead [“Fitting Mixture Designs”](#) on page 413 in the “Mixture Designs” chapter.

Example of Fit Two Level Screening Platform

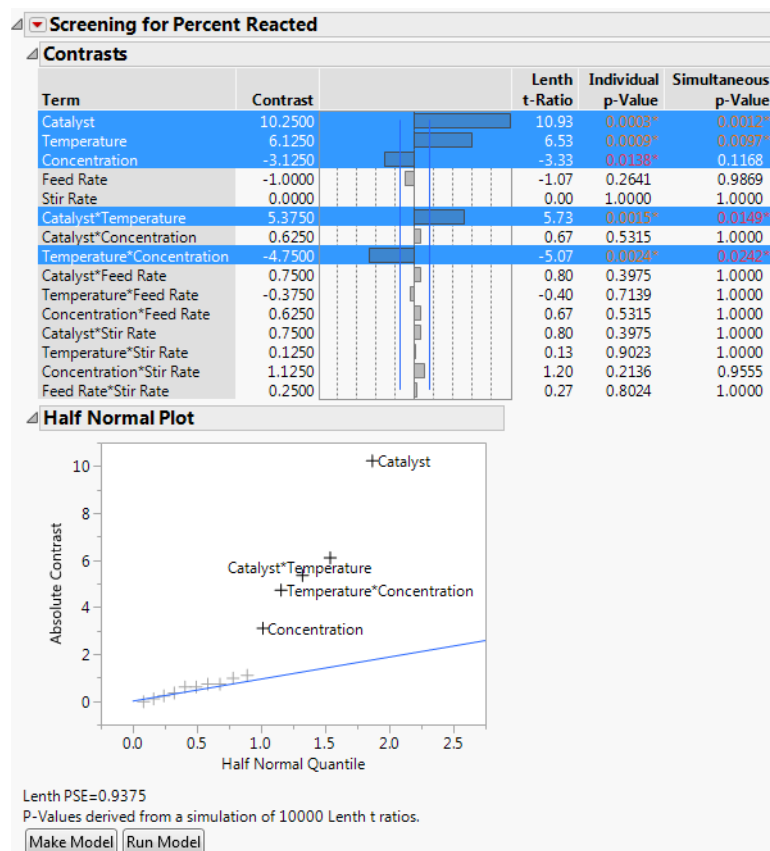
The Reactor Half Fraction.jmp sample data table contains the results of an experiment derived from a design discussed in Box et al. (1978). You are interested in identifying significant effects in a model that contains main effects and two-way interactions. This example uses a model with fifteen parameters for a design with sixteen runs. The example illustrates the analysis both with the Fit Two Level Screening Platform and the Fit Model Platform.

Fit Two Level Screening

1. Select **Help > Sample Data Library** and open Reactor Half Fraction.jmp.
2. Select **DOE > Classical > Two Level Screening > Fit Two Level Screening**.
3. Select Percent Reacted and click **Y**.
4. Select Feed Rate through Concentration, click **X**, and then click **OK**.

Note: The Fit Two Level Screening Platform automatically constructs a model with interaction terms. This is in contrast to the Fit Model Platform, where you manually specify the interactions that you want to include in your model.

Figure 10.2 Reactor Half Fraction.jmp Fit Two Level Screening Design Report



Note the following features of the Screening report:

- The effects that have an individual p -value less than 0.10 are selected.

- A t -ratio is calculated using Lenth's PSE (pseudo-standard error). The Lenth PSE value is shown below the Half Normal Plot.
- Both individual and simultaneous p -values are shown. Those that are less than 0.05 are shown with an asterisk.
- The Half Normal Plot enables you to quickly examine the effects. The effects that are initially selected in the effects list are also labeled in this plot.

In this example, Catalyst, Temperature, and Concentration, along with two of their two-factor interactions, are selected. Alternatively, you can fit the same model in the Fit Model Platform.

Fit Model

1. From the Reactor Half Fraction.jmp table, select **Analyze > Fit Model**.
2. Select Percent Reacted, click **Y**.
3. Select Feed Rate through Concentration, and then select **Macros > Factorial to Degree**.
4. Click **Run**.
5. Open the Parameter Estimates outline.

Figure 10.3 Saturated Reactor Half Fraction.jmp Design Parameter Estimates

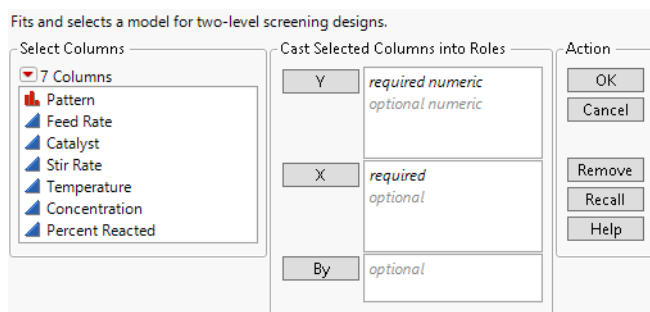
Response Percent Reacted				
Effect Summary				
Parameter Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
Intercept	65.25	.	.	.
Feed Rate(10,15)	-1	.	.	.
Catalyst(1,2)	10.25	.	.	.
Stir Rate(100,120)	0	.	.	.
Temperature(140,180)	6.125	.	.	.
Concentration(3,6)	-3.125	.	.	.
Feed Rate*Catalyst	0.75	.	.	.
Feed Rate*Stir Rate	0.25	.	.	.
Feed Rate*Temperature	-0.375	.	.	.
Feed Rate*Concentration	0.625	.	.	.
Catalyst*Stir Rate	0.75	.	.	.
Catalyst*Temperature	5.375	.	.	.
Catalyst*Concentration	0.625	.	.	.
Stir Rate*Temperature	0.125	.	.	.
Stir Rate*Concentration	1.125	.	.	.
Temperature*Concentration	-4.75	.	.	.

Since there are 16 observations and 16 model terms, there are not enough observations to estimate an error term. Without an estimate of error, it is not possible to conduct standard tests. Parameter estimates are provided, but because there are no degrees of freedom for error, standard errors, t -ratios, and p -values are all missing. This illustrates the strength of the Fit Two Level Screening platform for getting the most information out of a screening design.

Launch the Fit Two Level Screening Platform

Launch the Fit Two Level Screening platform by selecting **DOE > Classical > Two Level Screening > Fit Two Level Screening**.

Figure 10.4 Launch Window for the Fit Two Level Screening Platform



Y The columns to be analyzed. These must have a Numeric data type.

X Continuous or discrete two-level orthogonal factors.

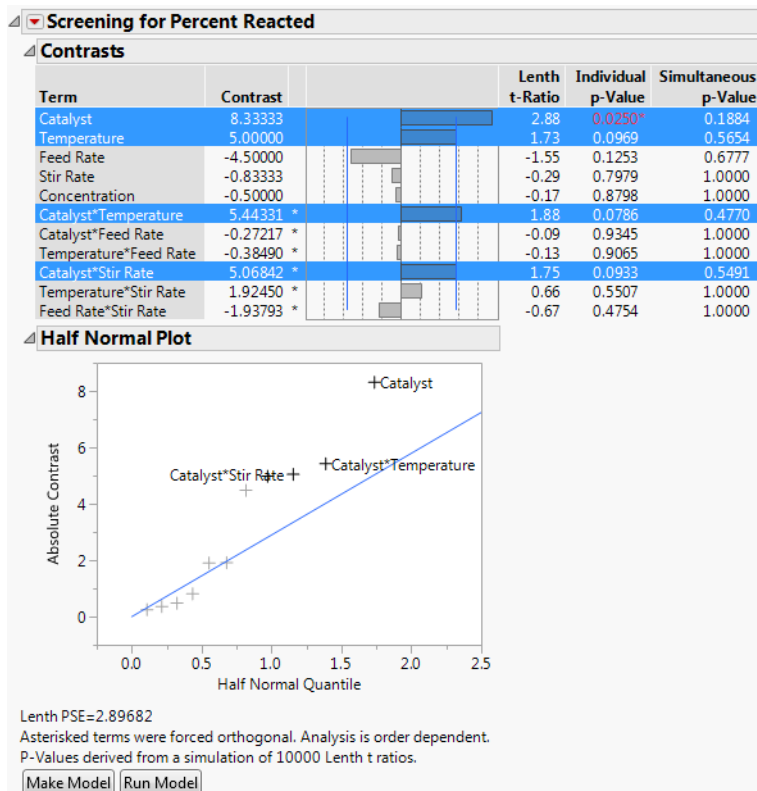
By A column whose level define separate analyses. For each level of the specified column, the corresponding rows are analyzed. The results appear in separate reports. If more than one By variable is assigned, a separate analysis is produced for each possible combination of the levels of the By variables.

Note: If you created your design in the Screening Design Platform from a list of fractional factorial designs, the design table contains a script called **Screening**. Run this script to run the analysis directly.

The Screening Report

The Fit Two Level Screening report shows a summary table of the design contrasts (candidate model terms) and a half normal plot. You can launch or run the selected model from the report.

Figure 10.5 Screening Report



- [“Contrasts”](#) on page 342
- [“Half Normal Plot”](#) on page 342
- [“Make or Run Model”](#) on page 343

Contrasts

The Contrasts report lists model effects, a contrast value, Lenth t -ratios, individual and simultaneous p -values, and aliases (if there are any). Effects are entered into the analysis following a hierarchical ordering. See [“Order of Effect Entry”](#) on page 348. Effects that have an Individual p -Value less than 0.10 are selected.

Term The name of the factor.

Contrast Estimate for the factor. For orthogonal designs, this number is the same as the regression parameter estimate. This is not the case for non-orthogonal designs. An asterisk might appear next to the contrast, indicating a lack of orthogonality.

Bar Chart Shows the Lenth t -ratios with blue vertical lines that indicate a value that is significant at the 0.10 level.

Lenth t -Ratio Lenth’s t -ratio, calculated as $Contrast/PSE$, where PSE is Lenth’s Pseudo-Standard Error. See [“Lenth’s Pseudo-Standard Error”](#) on page 349.

Individual p -Value Analogous to the standard p -values for a linear model. Small individual p -values indicate a significant effect. See [“Lenth’s Pseudo-Standard Error”](#) on page 349.

Note: Do not expect the p -values to be exactly the same if the analysis is re-run. The Monte Carlo method should give similar, but not identical, values if the same analysis is repeated.

Simultaneous p -Value A multiple-comparison adjusted p -value.

Aliases Appears only if there are exact aliases of later effects to earlier effects.

Half Normal Plot

The Half Normal Plot shows the absolute value of the contrasts plotted against the absolute value of quantiles for the half-normal distribution. The blue line passes through the origin with a slope of the Lenth’s estimate of σ . Effects that are small are considered error terms and assumed to have a normal distribution with mean zero and standard deviation σ . These terms fall on the blue line. Significant effects are those that have nonzero means and do not fall on the blue line.

The Half Normal Plot is interactive. You can select model effects by dragging a rectangle around the effects of interest. You can also hold down the Ctrl button and click an effect name(s) in the report.

Make or Run Model

The **Make Model** button launches a Fit Model window that is populated with the selected effects.

The **Run Model** button runs an effect screening model for the selected effects.

Note: A JMP alert appears if your selected model does not follow effect heredity. You can **Cancel**, use the **Make Model** button to launch the Fit Model window, and add the missing terms. Or, you can **Continue** and build a model that does not follow effect heredity rules. See [“Effect Heredity”](#) on page 62 in the “Starting Out with DOE” chapter.

Additional Fit Two Level Screening Analysis Examples

- [“Analyze a Plackett-Burman Design”](#)
- [“Effect Heredity Example”](#)
- [“Analyze a Supersaturated Design”](#)

Analyze a Plackett-Burman Design

Plackett-Burman designs are an alternative to fractional-factorial screening designs. Two-level fractional factorial designs must, by their nature, have a number of runs that are a power of two. However, Plackett-Burman designs exist for 12-, 24-, and 28-run designs.

The Weld-Repaired Castings.jmp sample data table uses a Plackett-Burman design, and is found in Box et al. (1978). Seven factors are thought to be influential on weld quality. The seven factors include Initial Structure, Bead Size, Pressure Treatment, Heat Treatment, Cooling Rate, Polish, and Final Treatment. A Plackett-Burman design with 12 runs is used to investigate the importance of the seven factors. The response is $100 \times \log(\text{lifetime})$. (The sample data table also contains four terms that were used to model error, but those terms are not used in this analysis.)

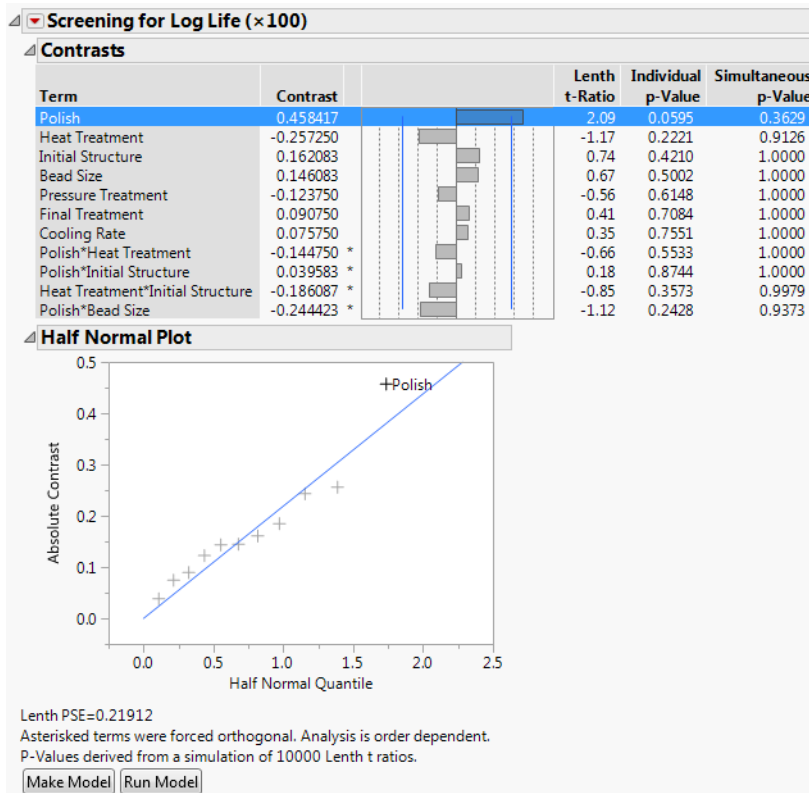
1. Select **Help > Sample Data Library** and open Weld-Repaired Castings.jmp.
2. Select **DOE > Classical > Two Level Screening > Fit Two Level Screening**.

The launch window is populated based on the column properties. Log Life ($\times 100$) is the response **Y**.

The seven factors Initial Structure, Bead Size, Pressure Treatment, Heat Treatment, Cooling Rate, Polish, and Final Treatment are specified for **X**.

3. Scroll to the bottom of the **X** window. Select ϵ_1 , ϵ_2 , ϵ_3 , and ϵ_4 , and click **Remove**.
4. Click **OK**.

Figure 10.6 Screening Report for Weld-Repaired Castings.jmp



The only significant effect identified is Polish. Note that asterisks mark four terms, indicating that they are not orthogonal to the effects preceding them. The contrast values obtained for these effects are after orthogonalization. As a result, these estimates do not match estimates obtained from a corresponding regression analysis. You can use the **Run Model** button to fit the single-factor model.

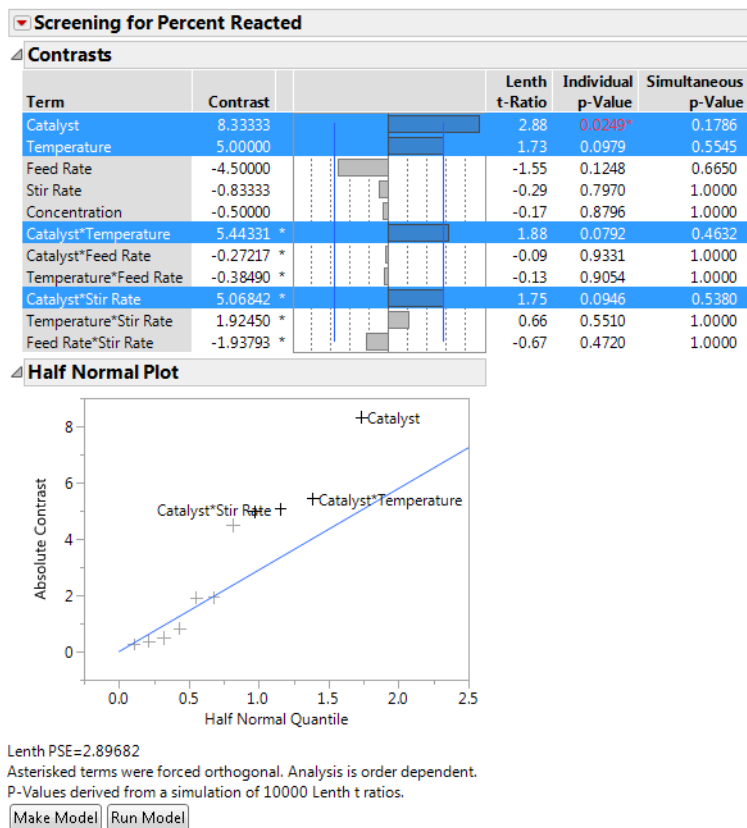
Effect Heredity Example

This example illustrates building a model when the Fit Two Level Screening platform identifies a model that does not follow effect heredity.

1. Select **Help > Sample Data Library** and open **Design Experiment/Plackett-Burman.jmp**.
2. Select **DOE > Classical > Two Level Screening > Fit Two Level Screening**.

3. Select Percent Reacted and click Y.
4. Select Feed Rate through Concentration and click X.
5. Click OK.

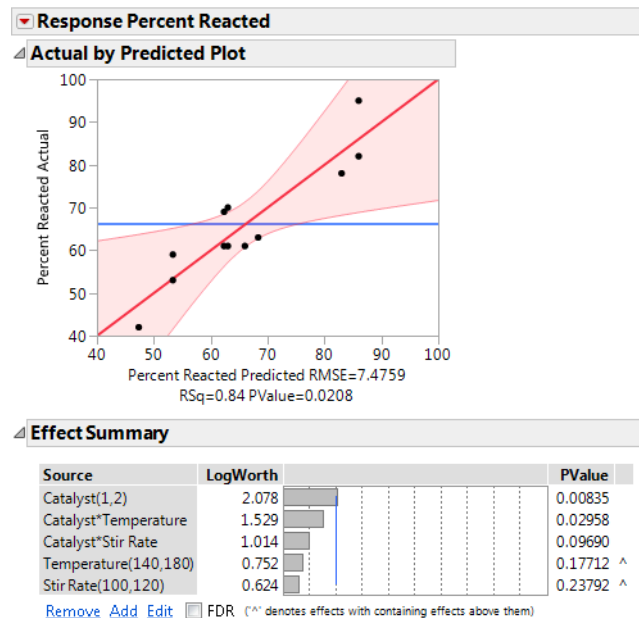
Figure 10.7 Model Contrasts



The Catalyst*Stir Rate interaction is highlighted, but the Stir Rate main effect is not. In accordance with the principle of Effect Heredity, you should add the Stir Rate main effect to the model. See “Effect Heredity” on page 62 in the “Starting Out with DOE” chapter.

6. Click **Make Model**.
7. Select Stir Rate and click **Add** in the Construct Model Effects section.
8. Click **Run**.

Figure 10.8 Partial Fit Model Report



The ^ symbol to the right of the PValue for Stir Rate and Temperature show that while not significant, these effects are contained in significant higher-order effect.

Analyze a Supersaturated Design

Supersaturated designs have more factors than runs. They rely heavily on effect sparsity for their analysis, so the Fit Two Level Screening platform is ideal for their analysis. The objective is to determine which effects are active.

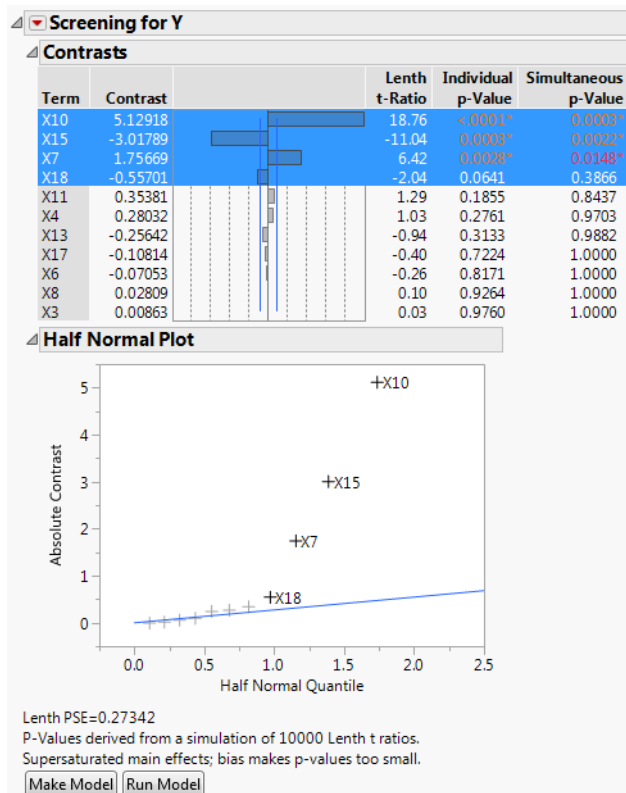
In this example, a simulated design with 18 factors but only 12 runs is considered. Y is generated by:

$$Y = 2(X7) + 5(X10) - 3(X15) + \varepsilon$$

where $\varepsilon \sim N(0,1)$. So, Y has been constructed with three active factors.

1. Select **Help > Sample Data Library** and open Supersaturated.jmp.
2. Select **DOE > Classical > Two Level Screening > Fit Two Level Screening**.
3. Select Y and click Y.
4. Select X1 through X18 and click X.
5. Click **OK**.

Figure 10.9 Screening Report for Supersaturated.jmp



Note that four factors have been highlighted. Factors X10, X15, and X7 are active based on a 0.05 critical value criteria. In the Half Normal Plot, X18 is close to the blue line, which indicates that it is close to the 0.1 cutoff value for significance. The 0.1 critical value is generous in its selection of factors so that you do not miss those that are possibly active.

The contrast estimates for X10, X15, and X7 of 5.1, -3, and 1.8 are close to their simulated values (5, -3, 2).

The p -values are not entirely valid statistically, since they are based on a simulation that assumes orthogonal designs, which is not the case for supersaturated designs. However, as shown, they can be useful in identifying effects for further investigation.

Technical Details for the Fit Two Level Screening Platform

- [“Order of Effect Entry”](#)
- [“Fit Two Level Screening as an Orthogonal Rotation”](#)
- [“Lenth’s Pseudo-Standard Error”](#)
- [“Lenth t-Ratios”](#)

Order of Effect Entry

The Fit Two Level Screening platform has a carefully defined order of operations:

1. First, the main effect terms enter the model according to the absolute size of their contrast. All effects are orthogonalized to the effects preceding them in the model. This method assures that their order is the same as it would be in a forward stepwise regression. Ordering by main effects also helps in selecting preferred aliased terms later in the process.
2. After main effects, all second-order interactions enter the model, followed by third-order interactions, and so on. The second-order interactions cross with all earlier terms before bringing in a new term because B and C are both larger than D. The complete order of entry is AB, AC, BC, AD, BD, and then CD.
3. An effect that is an exact alias for an effect already in the model appears in the Alias column. Effects that are a linear combination of several previous effects do not appear. If there is partial aliasing (due to a lack of orthogonality), the effects involved are marked with an asterisk.
4. The process continues until n effects are obtained, where n is the number of rows in the data table, thus fully saturating the model. If complete saturation is not possible with the factors, JMP generates random orthogonalized effects to absorb the rest of the variation. They are labeled Null n where n is a number. For example, this situation occurs if there are exact replicate rows in the design.

Fit Two Level Screening as an Orthogonal Rotation

Mathematically, the Fit Two Level Screening platform takes the n values in the response vector and rotates them into n new values. The rotated values are then mapped by the space of the factors and their interactions.

$$\text{Contrasts} = \mathbf{T}' \times \text{Responses}$$

In the above equation, $\mathbf{T}$ is an n by n orthonormalized set of values starting with the intercept, main effects of factors, two-way interactions, three-way interactions, and so on, until n values have been obtained. Since the first column of $\mathbf{T}$ is an intercept, and all the other columns are orthogonal to it, these other columns are all contrasts, that is, they sum to zero. Since $\mathbf{T}$ is orthogonal, it can serve as $\mathbf{X}$ in a linear model. It does not need inversion, since $\mathbf{T}'$ is equivalent to $\mathbf{T}^{-1}$ and $(\mathbf{T}'\mathbf{T})\mathbf{T}'$. The contrasts are the parameters estimated in a linear model.

If no effect in the model is active other than the intercept, the contrasts are just an orthogonal rotation of random independent variates into different random independent variates. These orthogonally-rotated random variates have the same variance as the original random independent variates. To the extent that some effects are active, the inactive effects still represent the same variation as the error in the model. The hope is that the effects and the design are strong enough to separate the active effects from the random error effects.

Lenth's Pseudo-Standard Error

Lenth's method (Lenth 1989), known as the *Lenth Pseudo Standard Error (PSE)*, constructs an estimate of the residual standard error using effects that appear to be inactive. Lenth's PSE can be used to estimate the standard error for experiments where contrasts are independent and have a common variance.

If there are n rows, the platform constructs $n - 1$ contrasts. Denote these contrasts by $\hat{C}_i$, where $i = 1, \dots, n - 1$.

To obtain Lenth's PSE, first calculate the following:

$$v = 1.5[\text{median}_{i=1, \dots, n-1} |\hat{C}_i|]$$

Lenth's PSE is based on the effects that are likely to be inactive and is defined as follows:

$$PSE = 1.5 \left[\text{median}_{|\hat{C}_i| < 2.5v} |\hat{C}_i| \right]$$

The value for Lenth's PSE is shown at the bottom of the Screening report.

Lenth t-Ratios

For each contrast, a t -ratio is computed by dividing the contrast by the PSE. The reference distribution of these t -ratios under the null hypothesis is not computationally tractable. Therefore, it is obtained by simulation. The method, described below, is based on a discussion in Ye and Hamada (2000).

Denote the t -ratio for the i^{th} contrast by t_i :

$$t_i = \hat{C}_i / (PSE)$$

Of primary importance in screening experiments is the *individual error rate*, namely the probability of declaring a given effect as active when it is not. For the i^{th} effect, this occurs when $|t_i|$ is large and falls into the upper tail of its reference distribution.

Because the platform constructs a relatively large number of effects, the *experimentwise error rate* is also of importance. The experimentwise error rate is the probability of declaring *any* effect as active when no effects are active. An experimentwise error occurs when no effects are active and the maximum of the absolute t -ratios, $\max |t_i|$, is large and falls into the upper tail of its reference distribution.

The Fit Two Level Screening platform obtains reference distributions for both types of error rates using Monte Carlo simulation. Consider a set of $n - 1$ values that is simulated from a normal distribution with mean 0 and standard deviation equal to PSE. This set of values represents potential contrast values for the experiment under the null hypothesis of no active effects. In all, 10,000 sets of $n - 1$ random contrast values are generated.

Individual p-Values

Consider the i^{th} contrast. Lenth t -ratios are constructed using each of the $10,000 \cdot (n - 1)$ simulated values. The reference distribution for the individual error rate is approximated by the absolute values of these t -ratios. The Individual p -value is the interpolated fractional position of the observed absolute Lenth t -Ratio among the $10,000 \cdot (n - 1)$ simulated absolute t -ratios arranged in descending order. This approximates the area to the right of the absolute value of the observed absolute Lenth t -Ratio with respect to the reference distribution.

Simultaneous p-Values

An experimentwise error occurs if any t -ratio leads to rejecting the null hypothesis when all effects are inactive. Equivalently, an experimentwise error occurs if the maximum of the absolute t -ratios, $\max |t_i|$, leads to rejecting the null hypothesis.

To obtain a reference distribution in this case, consider the maximum of the computed absolute t -ratios in each of the 10,000 simulations. These 10,000 maximum values form the reference distribution. The Simultaneous p -value is the interpolated fractional position of the observed absolute Lenth t -Ratio among the 10,000 simulated maximum absolute t -ratios arranged in descending order. This approximates the area to the right of the absolute value of the absolute Lenth t -Ratio with respect to the reference distribution based on the simulated maximum absolute t -ratios.

Monte Carlo Simulation Options

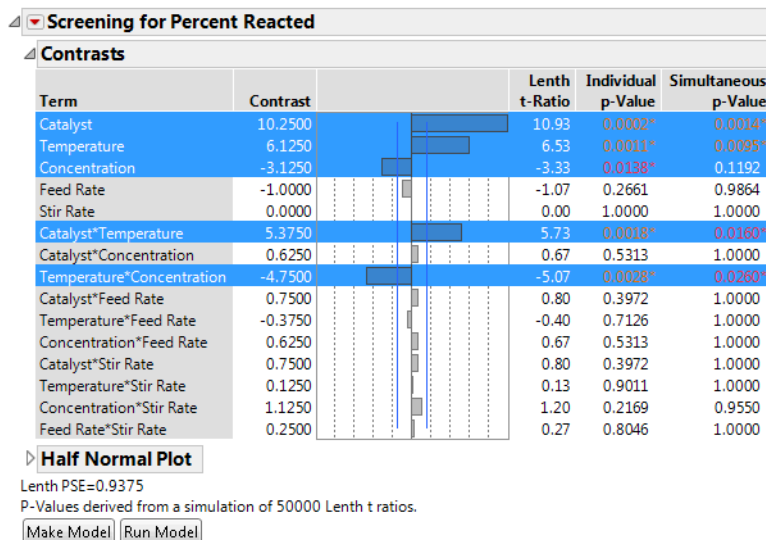
To change the number of default sets of simulations from 10,000, you must assign a value to a global JSL variable named `LenthSimN`. The number of simulations used to derive the *p*-values is reported in the report window.

The following example sets the number of simulations to 50,000.

1. Select **Help > Sample Data Library** and open **Reactor Half Fraction.jmp**.
2. Select **DOE > Classical > Two Level Screening > Fit Two Level Screening**.
3. Select **Percent Reacted** and click **Y**.
4. Select **Feed Rate** through **Concentration**, click **X**, and then click **OK**.
5. Click the red triangle next to **Screening for Percent Reacted** and select **Save Script > To Script Window**.
6. Insert `LenthSimN=50000`; as the first line of the script.


```
LenthSimN=50000;
Fit Two Level Screening(
    Y( :Percent Reacted),
    X( :Feed Rate, :Catalyst, :Stir Rate, :Temperature, :Concentration )
);
```
7. Right-click in the script window and select **Run Script**.

Figure 10.10 Fit Two Level with N = 50,000 Simulations for *p*-Value Derivations



The number of simulations is stated in the report window.

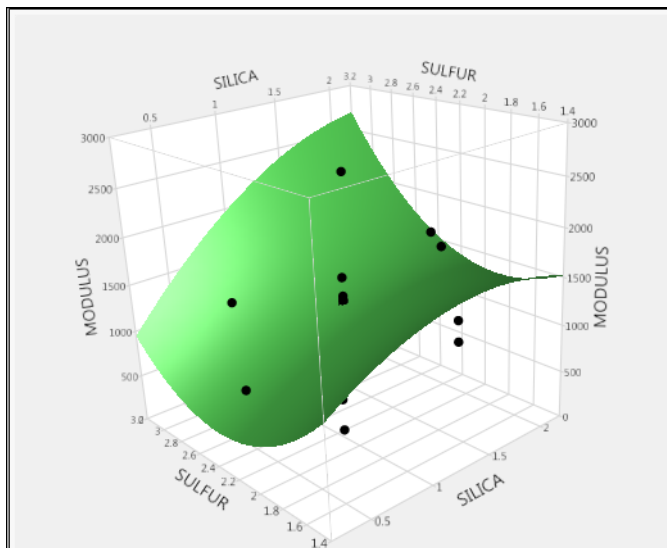
Note: If $LenthSimN=0$, the standard t -distribution is used and simultaneous p -values are not provided. This approach is not recommended.

Response Surface Designs

Response surface designs are useful for modeling quadratic surfaces. A response surface model can identify a point where a minimum or maximum value of the response occurs, if one exists inside the design region. Three distinct values for each factor are necessary to fit a quadratic function, so standard two-level designs are not appropriate for fitting curved surfaces.

Response surface designs are capable of fitting a second-order prediction equation for the response. The quadratic terms in these equations model the curvature in the true response function. If a maximum or minimum exists inside the design region, the point where that value occurs can be estimated.

Figure 11.1 Model for Response Surface Design Results



Contents

- Overview of Response Surface Designs..... 355
- Example of a Response Surface Design..... 357
 - Construct a Box-Behnken Design..... 358
 - Analyze the Experimental Data..... 359
 - Explore Optimal Settings..... 361
- Response Surface Design Window..... 364
 - Responses..... 365
 - Factors..... 367
 - Choose a Design..... 368
 - Specify Output Options..... 370
 - Make Table..... 371
- Response Surface Design Options..... 372

Overview of Response Surface Designs

The Response Surface Design platform provides the classical central composite and Box-Behnken designs, including blocked versions of these designs. For central composite designs, you can control the placement of axial points and other aspects of the design. Response surface designs are available for continuous factors only and are provided for up to eight factors.

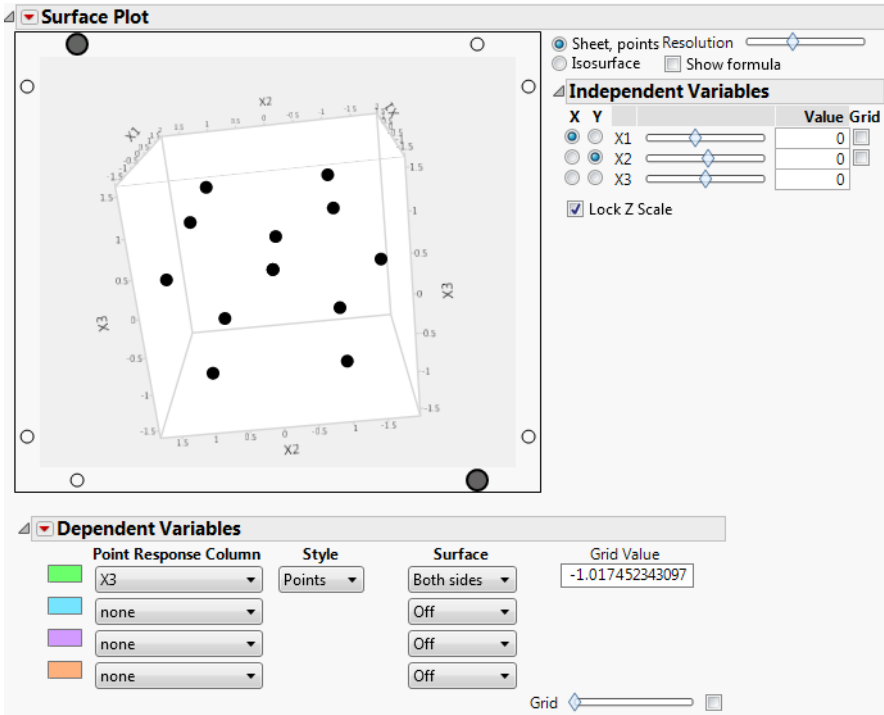
Tip: You can use DOE > Custom Design to construct optimal response surface designs that accommodate your specific experimental situation. Custom Design constructs response surface designs that are much more flexible than classical response surface designs. In particular, you can use the Custom Design platform to create response surface designs that involve categorical factors or more than eight continuous factors. You can also specify the number of runs and restrictions on the design space. For examples, see [“Response Surface Experiments”](#) on page 156 in the “Examples of Custom Designs” chapter.

A *central composite design* (Figure 11.2) combines a two-level fractional factorial design and two other types of points:

- *Center points*, where all the factor values are set to the midrange value.
- *Axial points*, where one factor is set to a high or low value (an *axial* value) and all other factors are set to the midrange value.

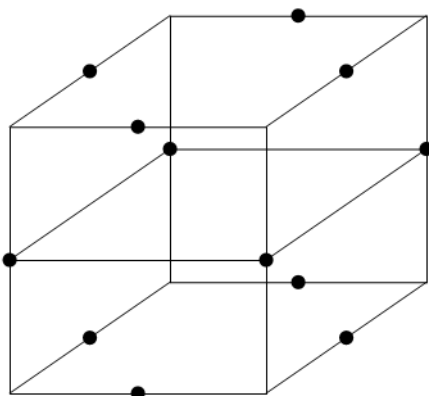
Depending on your selections relative to axial points, a central composite design can have as many as five distinct settings for each factor and the axial points can extend beyond the specified range of the factors.

Figure 11.2 Central Composite Design for Three Factors



A *Box-Behnken design* (Figure 11.3) has only three levels per factor and has no design points at the vertices of the cube defined by the ranges of the factors. This type of design can be useful when you must avoid these points due to engineering considerations. But, the lack of design points at the vertices of the cube means that a Box-Behnken design has higher prediction variance, and so less precision, near the vertices compared to a central composite design.

Figure 11.3 Box-Behnken Design for Three Factors



In JMP, you can construct a response surface design in two ways:

- Using the Response Surface Design platform (for up to eight continuous factors)
- Using the Custom Design platform (and clicking the RSM button in the Model outline)

In both cases, the design table contains a Model script that you can run to fit a model. The Model script applies the Response Surface Effect attribute to each main effect, so that the main effects appear with a **&RS** suffix in the Fit Model window. This attribute ensures that the Fit Least Squares report contains a Response Surface report. For more information about this report, see the Standard Least Squares Report and Options chapter in *Fitting Linear Models*.

Note: The Response Surface outline in the Standard Least Squares report is not shown for response surface designs that contain more than 20 continuous factors.

Example of a Response Surface Design

In this example, you construct a Box-Behnken design for a tire tread experiment. Your objective is to match a target value of 450 for a measure of elongation (Stretch). The stretch varies as a function of the amounts of Silica, Silane, and Sulfur used to manufacture the tire tread compound. You want to experiment over a wide range of factor settings to find the settings that achieve the target.

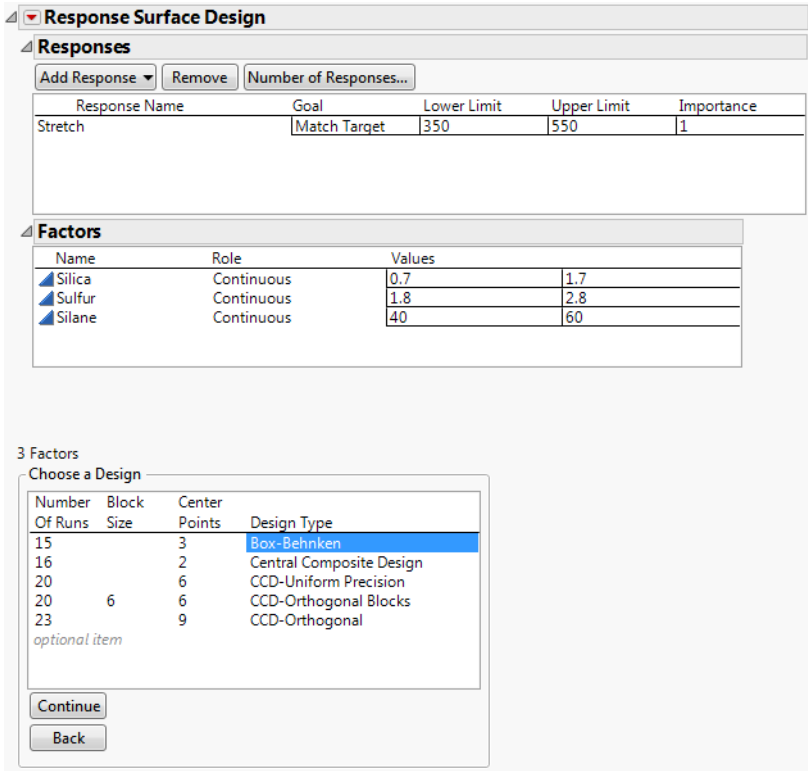
- [“Construct a Box-Behnken Design”](#)
- [“Analyze the Experimental Data”](#)
- [“Explore Optimal Settings”](#)

Construct a Box-Behnken Design

In this example, for convenience, you load the responses and factors from existing tables. When designing a new experiment on your own, enter the responses and factors manually. See “Responses” on page 365 and “Factors” on page 367.

- 1. Select DOE > Classical > Response Surface Design.
- 2. Select Help > Sample Data Library and open Design Experiment/Bounce Response.jmp.
- 3. Click the Response Surface Design red triangle and select Load Responses.
- 4. Select Help > Sample Data Library and open Design Experiment/Bounce Factors.jmp.
- 5. Click the Response Surface Design red triangle and select Load Factors.

Figure 11.4 Responses and Factors Outlines for Tire Tread Design



In the Responses outline, notice that the Goal for Stretch is set to Match Target.

In the Choose a Design panel, possible designs appear.

Note: Setting the Random Seed in step 6 reproduces the exact results shown in this example. In constructing a design on your own, this step is not necessary.

6. (Optional) Click the Response Surface Design red triangle and select **Set Random Seed**. Type 12345 and click **OK**.
7. Click **Continue** to retain the Box-Behnken design selection.
8. Click **Make Table**.

Figure 11.5 Box-Behnken Design Table

	Pattern	Silica	Sulfur	Silane	Stretch
1	-0+	0.7	2.3	60	•
2	0+–	1.2	2.8	40	•
3	000	1.2	2.3	50	•
4	000	1.2	2.3	50	•
5	+0–	1.7	2.3	40	•
6	+0+	1.7	2.3	60	•
7	0–+	1.2	1.8	60	•
8	0++	1.2	2.8	60	•
9	–0–	0.7	2.3	40	•
10	–+0	0.7	2.8	50	•
11	000	1.2	2.3	50	•
12	––0	0.7	1.8	50	•
13	0––	1.2	1.8	40	•
14	+–0	1.7	1.8	50	•
15	++0	1.7	2.8	50	•

At this point, conduct the experiment and enter the responses into the data table.

Analyze the Experimental Data

1. Select **Help > Sample Data Library** and open Design Experiment/Bounce Data.jmp.
The file Bounce Data.jmp contains your experiment results.
2. Run the **Model** script.
Notice that the main effects in the Construct Model Effects list are followed by the **& RS** suffix. This suffix indicates that these are response surface effects, which produce a Response Surface report in the Standard Least Squares report.
3. Click **Run**.

Figure 11.6 Lack of Fit and Effect Tests Reports

Lack Of Fit					
Source	DF	Sum of Squares	Mean Square	F Ratio	
Lack Of Fit	3	11.750000	3.91667	0.9792	
Pure Error	2	8.000000	4.00000		Prob > F
Total Error	5	19.750000		0.5411	
				Max RSq	
				0.9999	

Box-Cox Transformations					
Parameter Estimates					

Effect Tests					
Source	Nparm	DF	Sum of Squares	F Ratio	Prob > F
Silica(0,7,1,7)	1	1	12880.125	3260.791	<.0001*
Sulfur(1,8,2,8)	1	1	5778.125	1462.816	<.0001*
Silane(40,60)	1	1	968.000	245.0633	<.0001*
Silica*Sulfur	1	1	52441.000	13276.20	<.0001*
Silica*Silane	1	1	0.250	0.0633	0.8114
Sulfur*Silane	1	1	8556.250	2166.139	<.0001*
Silica*Silica	1	1	2592.923	656.4362	<.0001*
Sulfur*Sulfur	1	1	4653.231	1178.033	<.0001*
Silane*Silane	1	1	0.231	0.0584	0.8186

There is no indication of lack of fit and the Effect Tests report indicates that all but two higher-order terms (Silica*Silane and Silane*Silane) have p -values below 0.0001. See the Standard Least Squares Report and Options chapter in *Fitting Linear Models* for more information about interpretation of the tables in Figure 11.6.

4. Click the disclosure icon next to Response Surface to open the report.
5. Click the disclosure icon next to Canonical Curvature.

Figure 11.7 Response Surface Report

Response Surface				
Coef				
	Silica(0,7,1,7)	Sulfur(1,8,2,8)	Silane(40,60)	Stretch
Silica(0,7,1,7)	26.5	114.5	-0.25	-40.125
Sulfur(1,8,2,8)	.	-35.5	46.25	-26.875
Silane(40,60)	.	.	0.25	11

Solution	
Variable	Critical Value
Silica(0,7,1,7)	1.7912411
Sulfur(1,8,2,8)	2.1986422
Silane(40,60)	23.424426
Solution is a SaddlePoint	
Critical values outside data range	
Predicted Value at Solution 360.38388	

Canonical Curvature				
Eigenvalues and Eigenvectors				
Eigenvalue	62.9095	3.2989	-74.9584	
Silica(0,7,1,7)	0.82779	-0.29879	-0.47486	
Sulfur(1,8,2,8)	0.52687	0.12315	0.84097	
Silane(40,60)	0.19280	0.94634	-0.25937	

The Coef table shown as the first part of the report gives a concise summary of the estimated model parameters. The first three columns give the coefficients of the second-order terms. The last column gives the coefficients of the linear terms. To see the prediction expression in its entirety, select **Estimates > Show Prediction Expression** from the Response Stretch red triangle.

The Solution report gives the coordinates of the point where the single critical value occurs. In this instance, that point is a saddle point (a point where neither a maximum nor a minimum occurs) and falls outside the range of the design space.

The Canonical Curvature report shows eigenvalues and eigenvectors of the effects. These give information about the nature and direction of the surface's curvature. The large positive eigenvalue of 62.9095 indicates positive curvature and the eigenvector values indicate that the curvature is primarily in the Silica direction. The large negative eigenvalue of -74.9584 indicates negative curvature and the eigenvector values indicate that the curvature is primarily in the Sulfur direction.

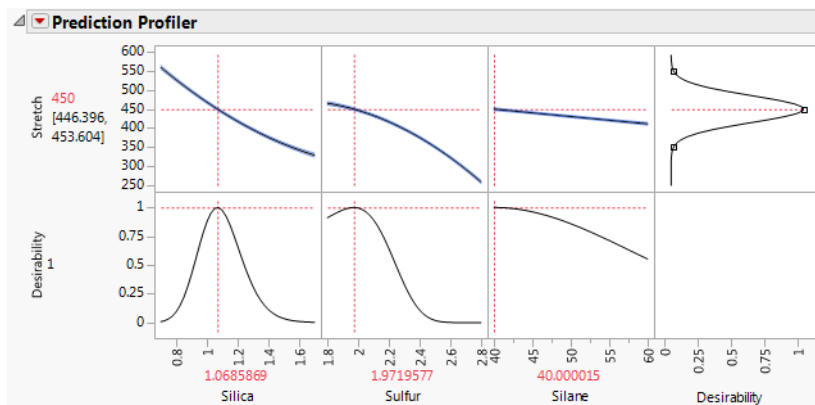
See the Standard Least Squares Report and Options chapter in *Fitting Linear Models* for more information about the response surface analysis tables in Figure 11.7.

Next, use the prediction profiler and the contour profiler to find optimal settings.

Explore Optimal Settings

1. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

Figure 11.8 Prediction Profiler for Bounce Data.jmp with Desirability Maximized



Note: Your optimal settings might differ. This is because there are many points for which the predicted Stretch is 450.

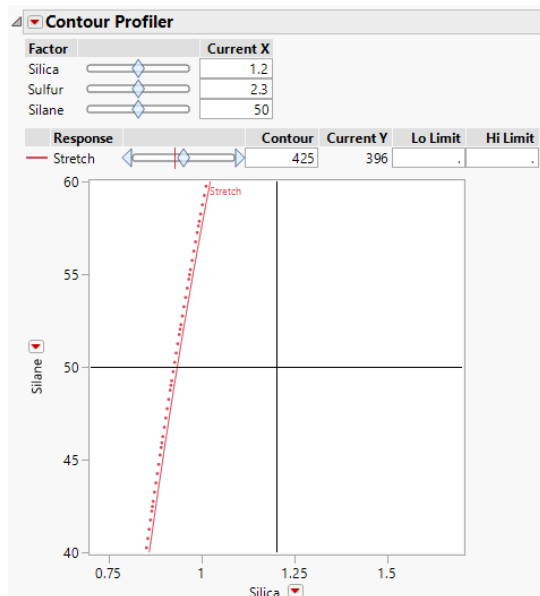
When you specified the response, the goal was set to match a target of 450, with lower and upper limits of 350 and 550. This goal was carried over to the design table and these limits were put in a Response Limits column property for Stretch. A desirability function is constructed from these response limits (top right cell in Figure 11.8). See [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

When you maximize the desirability function, JMP identifies one combination of factor level settings (usually out of many possible combinations) that results in a predicted Stretch of 450. Figure 11.8 shows these settings as Silica = 1.069, Sulfur = 1.972, and Silane = 40.000. Next, you use the Contour Profiler to identify other points that maximize the desirability function.

For more information about the Prediction Profiler, see the Profiler chapter in *Profilers*.

2. Click the Response Stretch red triangle and select **Factor Profiling > Contour Profiler**. Suppose that you want to achieve your target while setting Sulfur to the value 2.0. Also, you want to make sure that the settings that you choose for Silane and Silica maintain predicted Stretch within 5 units of 450.
3. On the vertical axis, click the red triangle next to Sulfur and select Silane.

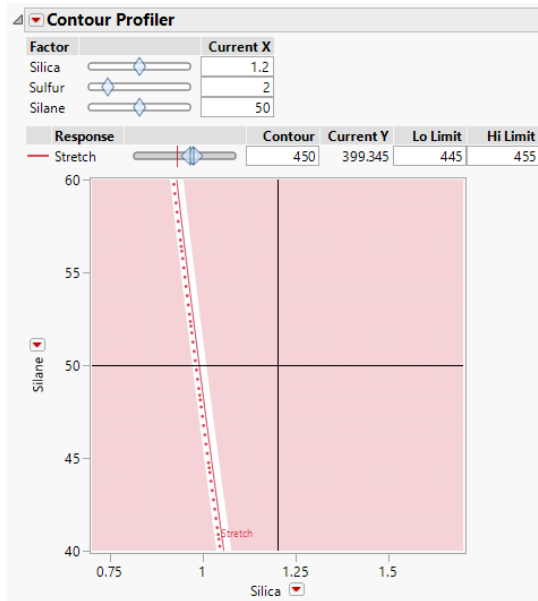
Figure 11.9 Contour Profiler for Bounce Data.jmp



The plot shows the contour of values of Silane and Silica for Stretch at 425 and Sulfur at 2.3.

4. Set the **Current X** for Sulfur to 2.
5. Set the **Contour** for Stretch to 450.
6. Set the **Lo Limit** and **Hi Limit** for Stretch to 445 and 455, respectively. Press Enter.

Figure 11.10 Contour Profiler Showing Optimal Settings for Silica and Silane

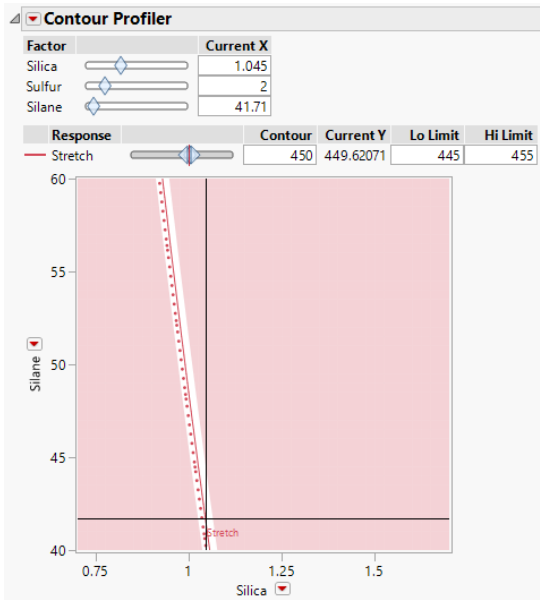


The unshaded band of Silica and Silane values gives predicted Stretch between 445 and 455 when Sulfur is set at 2.0. The values on the solid red curve give predicted Stretch of 450.

7. Drag the crosshairs that appear in the plot to the unshaded band to find settings for Silica and Silane that are best for your process from a practical perspective.

Suppose that your process is known to be more robust at low levels of Silane than at high levels. Then you might consider the settings in Figure 11.11.

Figure 11.11 Contour Profiler Showing Specific Settings for Silica and Silane



For Sulfur = 2.0, the factor settings identified by the crosshairs are Silica = 1.045 and Silane = 41.71. These settings are shown under Current X. At these settings, the predicted Stretch is 449.62071, shown next to Current Y.

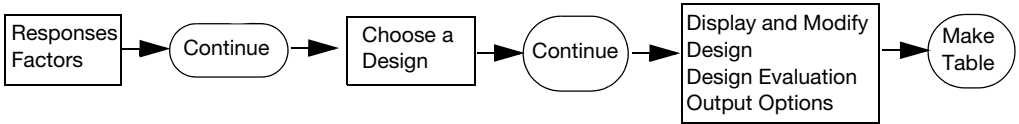
For further information about the Contour Profiler, see the Contour Profiler chapter in *Profilers*.

Response Surface Design Window

The Response Surface Design window walks you through the steps to construct a design for modeling a quadratic surface. You can select a central composite design, a Box-Behnken design, or a blocked version of one of these design types. If you select a central composite design, you can adjust the axial points.

The Response Surface Design window updates as you work through the design steps. The outlines, separated by buttons that update the outlines, follow the flow in Figure 11.12.

Figure 11.12 Response Surface Design Flow



Responses

Use the Responses outline to specify one or more responses.

Tip: When you have completed the Responses outline, consider selecting **Save Responses** from the red triangle menu. This option saves the response names, goals, limits, and importance values in a data table that you can later reload in DOE platforms.

Figure 11.13 Responses Outline

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.
optional item				

Add Response Enters a single response with a goal type of Maximize, Match Target, Minimize, or None. If you select Match Target, enter limits for your target value. If you select Maximize or Minimize, entering limits is not required but can be useful if you intend to use desirability functions.

Remove Removes the selected responses.

Number of Responses Enters additional responses so that the number that you enter is the total number of responses. If you have entered a response other than the default Y, the Goal for each of the additional responses is the Goal associated with the last response entered. Otherwise, the Goal defaults to Match Target. Click the Goal type in the table to change it.

The Responses outline contains the following columns:

Response Name The name of the response. When added, a response is given a default name of Y, Y2, and so on. To change this name, double-click it and enter the desired name.

Goal, Lower Limit, Upper Limit The Goal tells JMP whether you want to maximize your response, minimize your response, match a target, or that you have no response goal. JMP assigns a Response Limits column property, based on these specifications, to each response column in the design table. It uses this information to define a desirability function for each response. The Profiler and Contour Profiler use these desirability functions to find optimal factor settings. For further details, see the Profiler chapter in *Profilers* and [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

- A Goal of Maximize indicates that the best value is the largest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.

- A Goal of Minimize indicates that the best value is the smallest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Match Target indicates that the best value is a specific target value. The default target value is assumed to be midway between the Lower Limit and Upper Limit.
- A Goal of None indicates that there is no goal in terms of optimization. No desirability function is constructed.

Note: If your target response is not midway between the Lower Limit and the Upper Limit, you can change the target after you generate your design table. In the data table, open the Column Info window for the response column (**Cols > Column Info**) and enter the desired target value.

Importance When you have several responses, the Importance values that you specify are used to compute an overall desirability function. These values are treated as weights for the responses. If there is only one response, then specifying the Importance is unnecessary because it is set to 1 by default.

Editing the Responses Outline

In the Responses outline, note the following:

- Double-click a response to edit the response name.
- Click the goal to change it.
- Click on a limit or importance weight to change it.
- For multiple responses, you might want to enter values for the importance weights.

Response Limits Column Property

The Goal, Lower Limit, Upper Limit, and Importance that you specify when you enter a response are used in finding optimal factor settings. For each response, the information is saved in the generated design data table as a Response Limits column property. JMP uses this information to define the desirability function. The desirability function is used in the Prediction Profiler to find optimal factor settings. For further details about the Response Limits column property and examples of its use, see [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

If you do not specify a Lower Limit and Upper Limit, JMP uses the range of the observed data for the response to define the limits for the desirability function. Specifying the Lower Limit and Upper Limit gives you control over the specification of the desirability function. For more details about the construction of the desirability function, see the Profiler chapter in *Profilers*.

Factors

Factors in a response surface design can only be continuous.

Tip: Use **DOE > Custom Design** to create response surface designs that involve categorical factors.

The initial Factors panel for a response surface design appears with two continuous factors.

Figure 11.14 Factors Outline

Factors

Add 1 Continuous

Remove Selected

Name	Role	Values
X1	Continuous	-1 1
X2	Continuous	-1 1

Specify Factors

Specify desired number of factors. Double click on a factor name or setting to edit it.

Continue

The factors outline contains the following buttons.

Add Enters the number of continuous factors specified.

Remove Selected Removes the selected factors.

Tip: When you have completed your Factors panel, select Save Factors from the red triangle menu. This saves the factor names and values in a data table that you can later reload. See [“Response Surface Design Options”](#) on page 372.

The Factors outline contains the following columns:

Name The name of the factor. When added, a factor is given a default name of X1, X2, and so on. To change this name, double-click it and enter the desired name.

Role Specifies the Design Role of the factor as Continuous. The Design Role column property for the factor is saved to the data table. This property ensures that the factor type is modeled appropriately.

Values The experimental settings for the factors. To insert Values, click the default values and enter the desired values.

Factor Column Properties

For each factor, various column properties are saved to the data table for the completed design.

Design Role Each factor is assigned the Design Role column property. The Role that you specify in defining the factor determines the value of its Design Role column property. When you select a design with a block, that factor is assigned the Blocking value. The Design Role property reflects how the factor is intended to be used in modeling the experimental data. Design Role values are used in the Augment Design platform.

Factor Changes Each factor is assigned the Factor Changes column property with a setting of Easy. In the Response Surface Design platform, it is assumed that factor levels can be changed for each experimental run. Factor Changes values are used in the Evaluate Design and Augment Design platforms.

Coding If the Design Role is Continuous, the Coding column property for the factor is saved. This property transforms the factor values so that the low and high values correspond to -1 and $+1$, respectively. The estimates and tests in the Fit Least Squares report are based on the transformed values.

RunsPerBlock Indicates the number of runs in each block. When you select a design with a block and then click Make Table, a factor with the default name Block is added to the Factors list. The RunsPerBlock column property is saved for that factor.

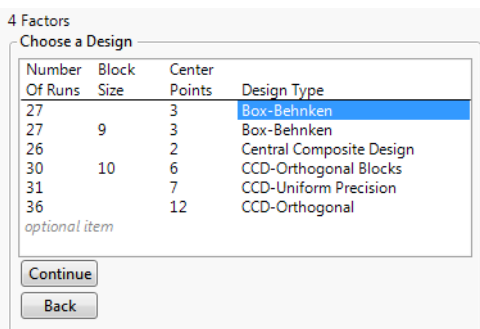
Choose a Design

After you enter your responses and factors and click **Continue**, you select from a list of designs. The designs include two types:

- [“Box-Behnken Designs”](#) on page 369
- [“Central Composite Designs”](#) on page 369

Select the design that you want to use and click **Continue**.

Figure 11.15 Choose a Design Panel for Four Factors



Box-Behnken Designs

Box-Behnken designs have only three levels for each factor and have no design points at the vertices of the cube defined by the ranges of the factors. These designs can be useful when it is desirable to avoid extreme settings for engineering considerations. However, these designs result in higher prediction variance near the vertices than do central composite designs.

Central Composite Designs

Central composite designs have center points and axial points. An *axial point* is a point where one factor is set to a high or low value (an *axial* value) and all other factors are set to the midrange, or center, value.

A central composite design can have axial points that fall beyond the faces of the hypercube defined by the specified factor ranges. This means that each factor might require five distinct settings, including two that fall beyond the range of values specified in the Factors outline. However, JMP enables you to place design points on the face.

The following types of central composite designs are available:

Central Composite Design The usual central composite design for the specified number of factors.

CCD-Uniform Precision The number of center points is chosen so that the prediction variance near the center of the design space is very flat.

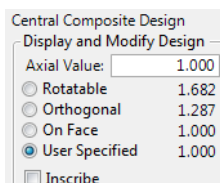
CCD-Orthogonal The number of center points and the axial values are chosen so that the second-order parameter estimates are minimally correlated with the other parameter estimates.

CCD-Orthogonal Blocks The second-order parameter estimates and block effects are minimally correlated with the other parameter estimates.

Axial Value

When you select a central composite design and then click **Continue**, you have the option to provide axial scaling information. In placing axial values, the values shown are used to multiply half of the specified range of a factor. If you specify a value of 1.0 next to Axial Value, then axial points in the resulting design are placed on the faces of the cube defined by the factors. You can set the axial value according to the following options:

Figure 11.16 Axial Value Panel



Rotatable The prediction variance depends only on the scaled distance from the center of the design. The axial points are more extreme than the factor ranges. If this factor range cannot be practically achieved, select **On Face** or specify your own value.

Orthogonal The effects are orthogonal. The axial points are more extreme than the factor ranges. If this factor range cannot be practically achieved, select **On Face** or specify your own value.

On Face Places the axial points at the extremes of the specified factor ranges.

User Specified Places the axial points at a distance specified by the value that you enter in the Axial Value text box.

Inscribe Rescales the design so that the axial points are at the low and high ends of the factor range. The factorial design points are shrunk based on that scaling.

Specify Output Options

You can specify details for the output data table in the Output Options panel. When you finish, click **Make Table** to construct the data table for the design.

Run Order

The **Run Order** options determine the order of the runs in the design table. Choices include the following:

Keep the Same Rows in the design table are in the same order as in the Design and Anticipated Coefficients outline.

Sort Left to Right Columns in the design table are sorted from left to right.

Randomize Rows in the design table are in random order.

Sort Right to Left Columns in the design table are sorted from right to left.

Randomize within Blocks Rows in the design table are in random order within the blocks.

Center Points and Replicates

Number of Center Points The number of center points that appear in the design. A center point is a run where every continuous factor is set at the center of the factor's range. The initial value shown is the number of center points in the design that you selected.

Number of Replicates The number of times to replicate the entire design, including center points. One replicate doubles the number of runs.

Make Table

Click **Make Table** to create a design table that contains the runs for your experiment.

Figure 11.17 Orthogonal Central Composite Design for Bounce Factors and Response

▼ Central Composite D... ▾		▼	▼					
Design: Central Composite D		▼	▼	Pattern	Silica	Sulfur	Silane	Stretch
▶ Model				1	000	1.2	2.3	50
▶ Evaluate Design				2	000	1.2	2.3	50
▶ DOE Dialog				3	---	0.7	2.8	60
				4	A00	2.0340158846	2.3	50
				5	+++	1.7	1.8	60
				6	000	1.2	2.3	50
				7	a00	0.3659841154	2.3	50
				8	000	1.2	2.3	50
▼ Columns (5/0)				9	---	0.7	1.8	40
■ Pattern				10	00a *	1.2	2.3	33.319682308
■ Silica *				11	+++	1.7	2.8	40
■ Sulfur *				12	---	1.7	1.8	40
■ Silane *				13	000	1.2	2.3	50
■ Stretch *				14	000	1.2	2.3	50
				15	000	1.2	2.3	50
				16	00A	1.2	2.3	66.680317692
				17	0a0	1.2	1.4659841154	50
▼ Rows				18	--	0.7	2.8	40
All rows		23		19	---	0.7	1.8	60
Selected		0		20	000	1.2	2.3	50
Excluded		0		21	0A0	1.2	3.1340158846	50
Hidden		0		22	+++	1.7	2.8	60
Labelled		0		23	000	1.2	2.3	50

The **Design** note in the Table panel at the upper left gives the design type that generated the table (Central Composite Design). This information can be helpful if you are comparing multiple designs.

Pattern Column

A Pattern column gives a symbolic description of the run in each row in terms of the factor values.

Tip: Pattern can be a useful label variable in plots.

Table 11.1 Pattern Column Description

-	Low value
+	High value
0	Midrange (center) value
a	Low axial value
A	High axial value

Design Table Scripts

The design table includes the following scripts:

Model Runs the **Analyze > Fit Model** platform.

Evaluate Design Runs the **DOE > Design Diagnostics > Evaluate Design** platform.

DOE Dialog Re-creates the Response Surface Design window that you used to generate the design table. The script also contains the random seed used to generate your design.

Response Surface Design Options

The red triangle menu in the Response Surface Design platform contains these options:

Save Responses Saves the information in the Responses panel to a new data table. You can then quickly load the responses and their associated information into most DOE windows. This option is helpful if you anticipate reusing the responses.

Load Responses Loads responses that you saved using the Save Responses option.

Save Factors Saves the information in the Factors panel to a new data table. Each factor’s column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Note: It is possible to create a factors table by entering data into an empty table, but remember to assign each column an appropriate Design Role. Do this by right-clicking on the column name in the data grid and selecting **Column Properties > Design Role**. In the Design Role area, select the appropriate role.

Load Factors Loads factors that you saved using the Save Factors option.

Save Constraints (Unavailable for some platforms) Saves factor constraints that you defined in the Define Factor Constraints or Linear Constraints outline into a data table, with a column for each constraint. You can then quickly load the constraints into most DOE windows.

In the constraint table, the first rows contain the coefficients for each factor. The last row contains the inequality bound. Each constraint's column contains a column property called `ConstraintState` that identifies the constraint as a "less than" or a "greater than" constraint. See "[ConstraintState](#)" on page 766 in the "Column Properties" appendix.

Load Constraints (Unavailable for some platforms) Loads factor constraints that you saved using the Save Constraints option.

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction
- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.

- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click **Apply**, a formula for the simulated response values is saved in a new column called **<Y> Simulated**, where Y is the name of the response. Clicking **Apply** again updates the formula and values in **<Y> Simulated**.

For additional details, see [“Simulate Responses”](#) on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Save X Matrix Saves the Moments Matrix and Model Matrix to table scripts in the design data table. These scripts contain the moments and design matrices. See [“Save X Matrix”](#) on page 112 in the “Custom Designs” chapter.

Caution: For a design with nominal factors, the Model Matrix saved by the Save X Matrix option is *not* the coding matrix used in fitting the linear model. You can obtain the coding matrix used for fitting the model by selecting the option **Save Columns > Save Coding Table** in the Fit Model report that you obtain when you run the Model script

Advanced Options > Set Delta for Power Specifies the difference in the mean response that you want to detect for model effects. See [“Evaluate Design Options”](#) on page 467 in the “Evaluate Designs” chapter.

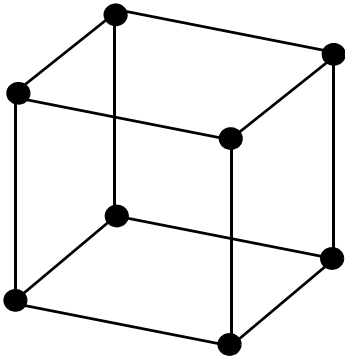
Save Script to Script Window Creates the script for the design that you specified in the Response Surface Design window and saves it in an open script window.

Chapter 12

Full Factorial Designs

A full factorial design defines an experiment where trials are run at all possible combinations of factor settings. A full factorial design allows the estimation of all possible interactions. Full factorial designs are large compared to screening designs, and since high-level interactions are often not active, they can be inefficient. They are typically used when you have a small number of factors and levels and want information about all possible interactions. For example, full factorial designs often form the basis for a measurement systems analysis (MSA).

Figure 12.1 Full Factorial Design for Three Two-Level Factors



Contents

Overview of Full Factorial Design 377

Example of a Full Factorial Design..... 377

 Construct the Design 377

 Analyze the Experimental Data 379

Full Factorial Design Window..... 383

 Responses 383

 Factors..... 385

 Select Output Options 388

 Make Table 389

Full Factorial Design Options 390

Overview of Full Factorial Design

In a full factorial design, you perform an experimental run at every combination of the factor levels. The sample size is the product of the numbers of levels of the factors. For example, a factorial experiment with a two-level factor, a three-level factor, and a four-level factor has $2 \times 3 \times 4 = 24$ runs.

The Full Factorial Design platform supports both continuous factors and categorical factors with arbitrary numbers of levels. It is assumed that you can run the trials in a completely random fashion.

Full factorial designs are the most conservative of all design types. Unfortunately, because the sample size grows exponentially with the number of factors, full factorial designs are often too expensive to run. Custom designs, definitive screening designs, and screening designs are less conservative but more efficient and cost-effective.

Example of a Full Factorial Design

In this example, you construct a full factorial design to study the effects of five two-level factors (Feed Rate, Catalyst, Stir Rate, Temperature, and Concentration) on the yield of a reactor. Because there are five factors, each at two levels, the full factorial design includes at least $2^5 = 32$ runs. For smaller screening designs involving this experimental situation, see [“Additional Examples of Screening Designs”](#) on page 323 in the “Screening Designs” chapter.

In this example, you load the responses and factors from existing tables. When designing a new experiment on your own, enter the responses and factors manually. See [“Responses”](#) on page 383 and [“Factors”](#) on page 385.

Construct the Design

1. Select **DOE > Classical > Full Factorial Design**.
2. Select **Help > Sample Data Library** and open Design Experiment/Reactor Response.jmp.
3. Click the Full Factorial Design red triangle and select **Load Responses**.
4. Select **Help > Sample Data Library** and open Design Experiment/Reactor Factors.jmp.
5. Click the Full Factorial Design red triangle and select **Load Factors**.
6. Click **Continue**.

Figure 12.2 Full Factorial Example Response and Factors Panels

Full Factorial Design

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Percent Reacted	Maximize	90	100	1

Factors

Continuous Categorical Remove Add N Factors 1

Name	Role	Values
Feed Rate	Continuous	10 15
Catalyst	Continuous	1 2
Stir Rate	Continuous	100 120
Temperature	Continuous	140 180
Concentration	Continuous	3 6

2x2x2x2x2 Factorial

Output Options

Run Order: Randomize

Number of Runs: 32

Number of Center Points: 0

Number of Replicates: 0

Make Table

Back

Note: Setting the Random Seed in step 7 ensures that the runs in your design table appear in the same order as in this example. In constructing a design on your own, this step is not necessary.

- (Optional) Click the Full Factorial Design red triangle and select **Set Random Seed**. Type 12345 and click **OK**.

The Run Order in the Output Options panel is set to Randomize. The order of runs in the design table will be random, as determined by the random seed.

The Number of Runs is set to 32. This is the number of all possible factor level combinations.

- Click **Make Table**.

The first column in the design data table shows the factor level combination for each run in terms of + and - signs, indicating high and low factor settings. The table also has an empty Y column named Percent Reacted for entering response values as you conduct the experiment.

Figure 12.3 Full Factorial Design for Reactor Experiment

	Pattern	Feed Rate	Catalyst	Stir Rate	Temperature	Concentration	Percent Reacted
1	++++-	15	2	100	180	3	•
2	+++++	15	2	120	180	6	•
3	---+-	10	2	120	140	3	•
4	---++	10	2	100	140	6	•
5	---++	10	2	120	180	6	•
6	---+-	10	2	100	140	3	•
7	+++++	15	2	100	140	6	•
8	---++	10	1	120	140	6	•
9	-----	10	1	100	140	3	•
10	+++++	15	2	100	180	6	•
11	++++-	15	1	120	180	3	•
12	---++	10	2	120	180	3	•
13	++++-	15	1	100	180	3	•
14	---+-	10	1	120	180	3	•
15	---++	10	2	120	140	6	•
16	+++++	15	2	120	140	6	•
17	+++++	15	1	100	180	6	•
18	++++-	15	2	120	180	3	•
19	-----	10	1	100	180	3	•
20	+++++	15	1	100	140	6	•
21	---++	10	2	100	180	6	•
22	---+-	10	1	100	180	6	•
23	++++-	15	2	100	140	3	•
24	+++++	15	1	120	140	6	•
25	---++	10	1	120	180	6	•
26	++++-	15	2	120	140	3	•
27	++++-	15	1	120	140	3	•
28	---+-	10	1	120	140	3	•
29	-----	10	1	100	140	6	•
30	---+-	10	2	100	180	3	•
31	++++-	15	1	100	140	3	•
32	+++++	15	1	120	180	6	•

Analyze the Experimental Data

Next, proceed to analyze the data from the completed experiment. You will use two methods to analyze the results: Screening and Stepwise Regression. Then you will find optimal settings using the Prediction Profiler.

Analysis Using Screening Platform

1. Select **Help > Sample Data Library** and open Design Experiment/Reactor 32 Runs.jmp.
2. Run the **Screening** script.

The Screening report shows a Contrasts report and a Half Normal Plot. The Contrasts report shows estimates for all 31 potential effects, up to the five-way interaction.

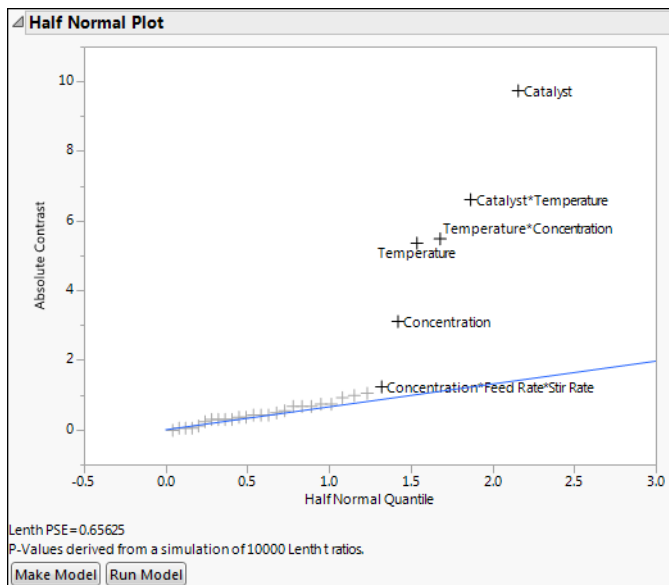
Figure 12.4 Contrasts Report for Reactor 32 Runs.jmp

Contrasts					
Term	Contrast		Lenth t-Ratio	Individual p-Value	Simultaneous p-Value
Catalyst	9.75000		14.86	<.0001*	<.0001*
Temperature	5.37500		8.19	<.0001*	0.00009*
Concentration	-3.12500		-4.76	<.0001*	0.0152*
Feed Rate	-0.68750		-1.05	0.2883	0.9998
Stir Rate	-0.31250		-0.48	0.6494	1.0000
Catalyst*Temperature	6.62500		10.10	<.0001*	0.00002*
Catalyst*Concentration	1.00000		1.52	0.1335	0.9553
Temperature*Concentration	-5.50000		-8.38	<.0001*	0.00004*
Catalyst*Feed Rate	0.68750		1.05	0.2883	0.9998
Temperature*Feed Rate	-0.43750		-0.67	0.4941	1.0000
Concentration*Feed Rate	0.06250		0.10	0.9266	1.0000
Catalyst*Stir Rate	0.43750		0.67	0.5212	1.0000
Temperature*Stir Rate	1.06250		1.62	0.1139	0.9205
Concentration*Stir Rate	0.43750		0.67	0.4941	1.0000
Feed Rate*Stir Rate	0.37500		0.57	0.5859	1.0000
Catalyst*Temperature*Concentration	-0.12500		-0.19	0.8557	1.0000
Catalyst*Temperature*Feed Rate	0.68750		1.05	0.2883	0.9998
Catalyst*Concentration*Feed Rate	-0.93750		-1.43	0.1587	0.9784
Temperature*Concentration*Feed Rate	0.31250		0.48	0.6494	1.0000
Catalyst*Temperature*Stir Rate	0.56250		0.86	0.3772	1.0000
Catalyst*Concentration*Stir Rate	0.06250		0.10	0.9266	1.0000
Temperature*Concentration*Stir Rate	0.06250		0.10	0.9266	1.0000
Catalyst*Feed Rate*Stir Rate	0.75000		1.14	0.2478	0.9989
Temperature*Feed Rate*Stir Rate	-0.37500		-0.57	0.5859	1.0000
Concentration*Feed Rate*Stir Rate	-1.25000		-1.90	0.0705	0.7592
Catalyst*Temperature*Concentration*Feed Rate	0.31250		0.48	0.6494	1.0000
Catalyst*Temperature*Concentration*Stir Rate	-0.31250		-0.48	0.6494	1.0000
Catalyst*Temperature*Feed Rate*Stir Rate	0.00000		0.00	1.0000	1.0000
Catalyst*Concentration*Feed Rate*Stir Rate	0.75000		1.14	0.2478	0.9989
Temperature*Concentration*Feed Rate*Stir Rate	0.50000		0.76	0.4328	1.0000
Catalyst*Temperature*Concentration*Feed Rate*Stir Rate	-0.25000		-0.38	0.7150	1.0000

Note: Because the p -values in the Contrasts report are obtained using a Monte Carlo simulation, you might not obtain the same values as shown in Figure 12.4. See “[Lenth’s Pseudo-Standard Error](#)” on page 349 in the “The Fit Two Level Screening Platform” chapter.

The six highlighted effects in the Contrasts outline are labeled in the Half Normal Plot.

Figure 12.5 Half Normal Plot for Reactor 32 Runs.jmp



The Half Normal Plot provides strong evidence that at least five of the labeled effects are larger than would be expected if they were the result of random variation. This suggests that these effects are active. The plot does *not* provide a clear indication that the three-way Concentration*Feed Rate*Stir Rate interaction is active.

In the Contrasts outline in Figure 12.4, the Individual p-Value for the three-way Concentration*Feed Rate*Stir Rate interaction is 0.0705 and its Simultaneous p-Value is 0.7592. Because the effect does not stand out on the Half Normal Plot and because its *p*-values are large, you decide not to include this effect in your model.

3. In the Half Normal Plot, drag a rectangle to select all labeled effects except for Concentration*Feed Rate*Stir Rate.
4. Click **Make Model** to open a Fit Model window containing the five effects.
5. Click **Run**.

The Actual by Predicted plot shows no evidence of lack of fit and the Effect Summary outline shows that all five effects are significant.

Analysis Using Stepwise Regression

1. Return to the Reactor 32 Runs.jmp data table, or reopen it by selecting **Help > Sample Data Library** and opening Design Experiment/Reactor 32 Runs.jmp.
2. Run the **Model** script.

The Construct Model Effects list contains only up to two-way interactions. You want to consider all interactions.

3. From the Select Columns list, select Feed Rate through Concentration.

4. Click **Macros > Full Factorial**.

All possible effects are added in the Construct Model Effects list.

5. Change the **Personality** to **Stepwise**.

6. Click **Run**.

7. Change the **Stopping Rule** to **Minimum AICc**.

For designed experiments, AICc is preferred to BIC. This is because BIC is typically a more lenient stopping rule than AICc as it tends to allow inactive effects into the model.

8. Click **Go**.

The Stepwise procedure selects six effects as potentially active.

9. Click **Run Model**.

This fits a model using the six effects. The Effect Summary outline indicates that Catalyst*Concentration is not significant at the 0.05 significance level as the p -value is 0.0896.

10. Select Catalyst*Concentration in the Effect Summary outline and click **Remove**.

The five remaining effects are all highly significant. These are the same five effects that you identified using the Screening platform ("[Analysis Using Screening Platform](#)" on page 379).

Optimal Settings Using the Prediction Profiler

Now, find optimal settings for the three active factors involved in the five significant effects that you retained in your model.

1. In the Reactor 32 Runs.jmp data table, run the **Reduced Model** script.

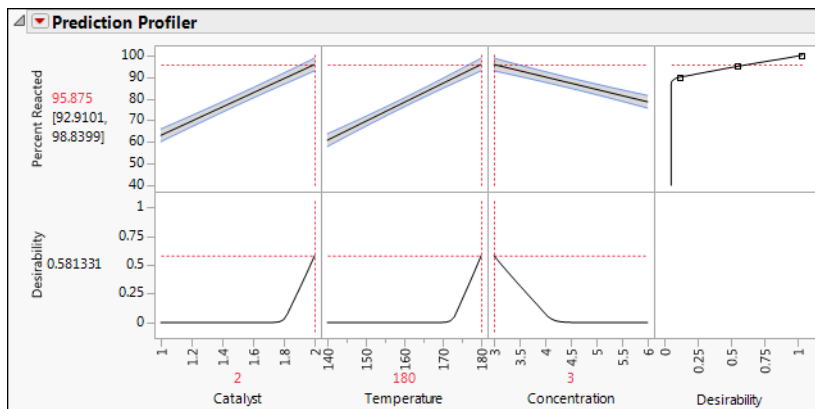
The Reduced Model script opens a Fit Model window for the five-effect model that you identified in "[Analysis Using Screening Platform](#)" on page 379 and "[Analysis Using Stepwise Regression](#)" on page 381.

2. Click **Run**.

The Prediction Profiler report displays Desirability because in the Full Factorial window, you specified a Goal of Maximize when you defined your response. The desirability function shown in the rightmost cell in the top row of the profiler shows that a value of 100 is most desirable and a value of 90 or below is least desirable.

3. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

Figure 12.6 Prediction Profiler Showing Settings That Optimize Desirability

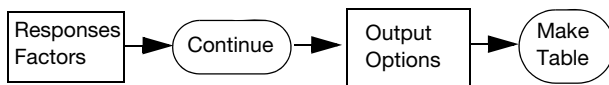


The predicted mean Percent Reacted at the settings that are shown is 95.875, with a confidence interval of 92.91 to 98.84. Note that, for all three factors, the settings that are identified are at the extremes of the ranges used in the experiment. In a future experiment, you should explore the process behavior beyond these settings.

Full Factorial Design Window

The Full Factorial design window updates as you work through the design steps. See [“The DOE Workflow: Describe, Specify, Design”](#) on page 57. The outlines, separated by buttons that update the outlines, follow the flow in Figure 12.7.

Figure 12.7 Full Factorial Design Flow



Responses

Use the Responses outline to specify one or more responses.

Tip: When you have completed the Responses outline, consider selecting **Save Responses** from the red triangle menu. This option saves the response names, goals, limits, and importance values in a data table that you can later reload in DOE platforms.

Figure 12.8 Responses Outline

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.
optional item				

Add Response Enters a single response with a goal type of Maximize, Match Target, Minimize, or None. If you select Match Target, enter limits for your target value. If you select Maximize or Minimize, entering limits is not required but can be useful if you intend to use desirability functions.

Remove Removes the selected responses.

Number of Responses Enters additional responses so that the number that you enter is the total number of responses. If you have entered a response other than the default Y, the Goal for each of the additional responses is the Goal associated with the last response entered. Otherwise, the Goal defaults to Match Target. Click the Goal type in the table to change it.

The Responses outline contains the following columns:

Response Name The name of the response. When added, a response is given a default name of Y, Y2, and so on. To change this name, double-click it and enter the desired name.

Goal, Lower Limit, Upper Limit The Goal tells JMP whether you want to maximize your response, minimize your response, match a target, or that you have no response goal. JMP assigns a Response Limits column property, based on these specifications, to each response column in the design table. It uses this information to define a desirability function for each response. The Profiler and Contour Profiler use these desirability functions to find optimal factor settings. For further details, see the Profiler chapter in *Profiler* and [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

- A Goal of Maximize indicates that the best value is the largest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Minimize indicates that the best value is the smallest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Match Target indicates that the best value is a specific target value. The default target value is assumed to be midway between the Lower Limit and Upper Limit.

- A Goal of None indicates that there is no goal in terms of optimization. No desirability function is constructed.

Note: If your target response is not midway between the Lower Limit and the Upper Limit, you can change the target after you generate your design table. In the data table, open the Column Info window for the response column (**Cols > Column Info**) and enter the desired target value.

Importance When you have several responses, the Importance values that you specify are used to compute an overall desirability function. These values are treated as weights for the responses. If there is only one response, then specifying the Importance is unnecessary because it is set to 1 by default.

Editing the Responses Outline

In the Responses outline, note the following:

- Double-click a response to edit the response name.
- Click the goal to change it.
- Click on a limit or importance weight to change it.
- For multiple responses, you might want to enter values for the importance weights.

Response Limits Column Property

The Goal, Lower Limit, Upper Limit, and Importance that you specify when you enter a response are used in finding optimal factor settings. For each response, the information is saved in the generated design data table as a Response Limits column property. JMP uses this information to define the desirability function. The desirability function is used in the Prediction Profiler to find optimal factor settings. For further details about the Response Limits column property and examples of its use, see [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

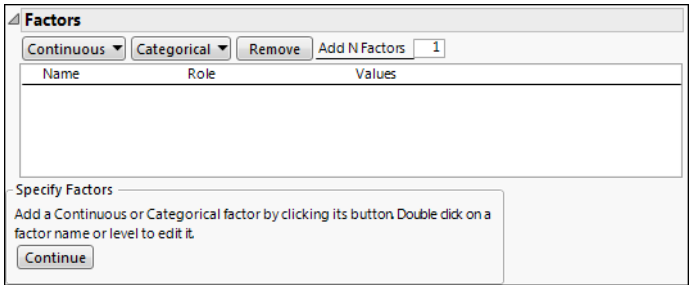
If you do not specify a Lower Limit and Upper Limit, JMP uses the range of the observed data for the response to define the limits for the desirability function. Specifying the Lower Limit and Upper Limit gives you control over the specification of the desirability function. For more details about the construction of the desirability function, see the Profiler chapter in *Profilers*.

Factors

Factors in a full factorial design can be continuous or categorical.

Tip: When you have completed the Factors outline, consider selecting **Save Factors** from the red triangle menu. This option saves the factor names, roles, changes, and values in a data table that you can later reload in DOE platforms.

Figure 12.9 Factors Outline



Continuous Adds a Continuous factor. The data type in the resulting data table is Numeric. A continuous factor is a factor that you can conceptually set to any value between the lower and upper limits you supply, given the limitations of your process and measurement system.

Categorical Adds a Categorical factor. Click to select or specify the number of levels. The data type in the resulting data table is Character. The value ordering of the levels is the order of the values, as entered from left to right. This ordering is saved in the Value Order column property after the design data table is created.

The default values for a categorical factor are L1, L2,..., Lk, where *k* is the number of levels that you specify. Replace the default values with level names that are relevant for your experiment.

Remove Removes the selected factors.

Add N Factors Adds multiple factors. Enter the number of factors to add, click **Add Factor**, and then select the factor type. Repeat **Add N Factors** to add multiple factors of different types.

Factors Outline

The Factors outline contains the following columns:

Name The name of the factor. When added, a factor is given a default name of X1, X2, and so on. To change this name, double-click it and enter the desired name.

Role The Design Role of the factor. The Design Role column property for the factor is saved to the data table. This property ensures that the factor type is modeled appropriately. The

Role of the factor determines other factor properties that are saved to the data table. See [“Factor Column Properties”](#) on page 387.

Values The experimental settings for the factors.

Editing the Factors Outline

In the Factors outline, note the following:

- To edit a factor name, double-click the factor name.
- To edit a value, click the value in the Values column.

Factor Column Properties

For each factor, various column properties are saved to the design table after you create the design by selecting Make Table in the Screening Design window. These properties are also saved automatically to the data table that is created when you select the Save Factors option. You can find more information about these column properties and related examples in [Appendix A, “Column Properties”](#).

Coding If the Role is Continuous, the Coding column property for the factor is saved. This property transforms the factor values so that the low and high values correspond to -1 and $+1$, respectively. See [“Coding”](#) on page 742 in the “Column Properties” appendix.

Value Order If the Role is Categorical or if a Block variable is constructed, the Value Order column property for the factor is saved. This property determines the order in which levels of the factor appear. See [“Value Order”](#) on page 759 in the “Column Properties” appendix.

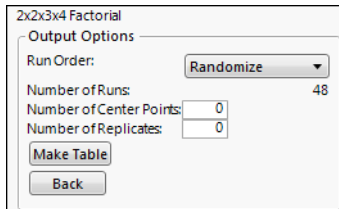
Design Role Each factor is given the Design Role column property. The Role that you specify in defining the factor determines the value of its Design Role column property. The Design Role property reflects how the factor is intended to be used in modeling the experimental data. Design Role values are used in the Augment Design platform. See [“Design Role”](#) on page 740 in the “Column Properties” appendix.

Factor Changes Each factor is assigned the Factor Changes column property with the value of Easy. The Factor Changes property reflects how the factor is used in modeling the experimental data. Factor Changes values are used in the Augment Design and Evaluate Design platforms. See [“Factor Changes”](#) on page 755 in the “Column Properties” appendix.

Select Output Options

After you enter your responses and factors and click **Continue**, you can make selections for your design table in the Output Options outline. The structure of the full factorial design appears at the top of the outline.

Figure 12.10 Output Options Panel



Run Order

The **Run Order** options determine the order of the runs in the design table. Choices include the following:

Keep the Same Rows in the design table are sorted from left to right.

Sort Left to Right Columns in the design table are sorted from left to right.

Randomize Rows in the design table are in random order.

Sort Right to Left Columns in the design table are sorted from right to left.

Center Points and Replicates

Number of Runs Shows the number of runs in the design before you add center points or replicates.

Number of Center Points Specifies how many additional runs to add as center points to the design. A center point is a run where every continuous factor is set at the center of the factor's range.

Suppose that your design includes both continuous and categorical factors. If you request center points in the Output Options panel, the center points are distributed as follows:

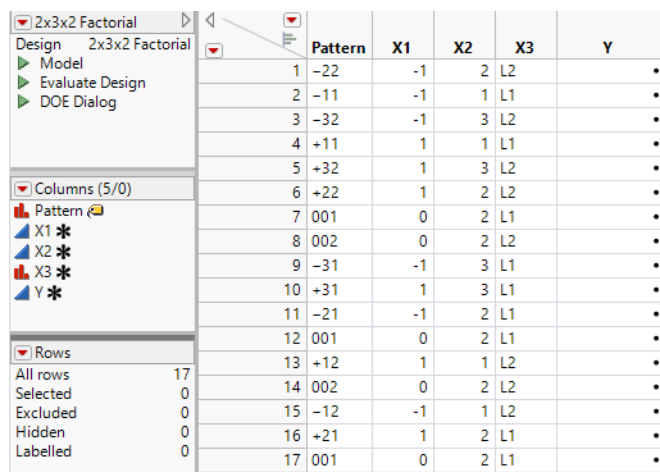
1. The settings for the categorical factors are ordered using the value ordering specified in the Factors outline.
2. One center point is assigned to each combination of the settings of the categorical factors in order, and this is repeated until all center points are assigned.

Number of Replicates The number of times to replicate the entire design, including center points. One replicate doubles the number of runs.

Make Table

Clicking Make Table creates a data table that contains the runs for your experiment. The example in Figure 12.11 shows a full factorial design with five center points for three factors: X1 (a two-level continuous factor), X2 (a three-level continuous factor), and X3 (a two-level categorical factor). The design uses the default values for the factor levels. The center points are in rows 7, 8, 12, 14, and 17. See [“Pattern Column”](#) on page 389.

Figure 12.11 Design Data Table



	Pattern	X1	X2	X3	Y
1	-22	-1	2	L2	•
2	-11	-1	1	L1	•
3	-32	-1	3	L2	•
4	+11	1	1	L1	•
5	+32	1	3	L2	•
6	+22	1	2	L2	•
7	001	0	2	L1	•
8	002	0	2	L2	•
9	-31	-1	3	L1	•
10	+31	1	3	L1	•
11	-21	-1	2	L1	•
12	001	0	2	L1	•
13	+12	1	1	L2	•
14	002	0	2	L2	•
15	-12	-1	1	L2	•
16	+21	1	2	L1	•
17	001	0	2	L1	•

The name of the table, shown in the upper left corner, is the design type that generated it.

Design Table Scripts

The design table includes the following scripts:

Model Runs the Analyze > Fit Model platform.

Evaluate Design Runs the DOE > Design Diagnostics > Evaluate Design platform.

DOE Dialog Re-creates the Full Factorial Design window that you used to generate the design table. The script also contains the random seed used to generate your design.

Run the **Screening** or **Model** scripts to analyze the data.

Pattern Column

The Pattern column contains entries that summarize the run in the given row. You can use Pattern as a label variable in plots.

- For a two-level continuous factor, the low setting is denoted by “–”, the high setting by “+”, and a center point by “0”.
- For a continuous factor with more than two levels:
 - For a non-center point, the factor setting is denoted by an integer that corresponds to the value level for the run.
 - For a center point, the factor setting is denoted by a “0”.
- For a categorical factor, the factor setting is denoted by an integer that corresponds to the value level for the run.

Full Factorial Design Options

The red triangle menu in the Full Factorial Design platform contains these options:

Save Responses Saves the information in the Responses panel to a new data table. You can then quickly load the responses and their associated information into most DOE windows. This option is helpful if you anticipate reusing the responses.

Load Responses Loads responses that you saved using the Save Responses option.

Save Factors Saves the information in the Factors panel to a new data table. Each factor’s column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Note: It is possible to create a factors table by entering data into an empty table, but remember to assign each column an appropriate Design Role. Do this by right-clicking on the column name in the data grid and selecting **Column Properties > Design Role**. In the Design Role area, select the appropriate role.

Load Factors Loads factors that you saved using the Save Factors option.

Save Constraints (Unavailable for some platforms) Saves factor constraints that you defined in the Define Factor Constraints or Linear Constraints outline into a data table, with a column for each constraint. You can then quickly load the constraints into most DOE windows.

In the constraint table, the first rows contain the coefficients for each factor. The last row contains the inequality bound. Each constraint’s column contains a column property called **ConstraintState** that identifies the constraint as a “less than” or a “greater than” constraint. See [“ConstraintState”](#) on page 766 in the “Column Properties” appendix.

Load Constraints (Unavailable for some platforms) Loads factor constraints that you saved using the Save Constraints option.

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction
- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.
- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click Apply, a formula for the simulated response values is saved in a new column called <Y> Simulated, where Y is the name of the response. Clicking Apply again updates the formula and values in <Y> Simulated.

For additional details, see [“Simulate Responses”](#) on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Advanced Options None available.

Save Script to Script Window Creates the script for the design that you specified in the Full Factorial Design window and saves it in an open script window.

Chapter 13

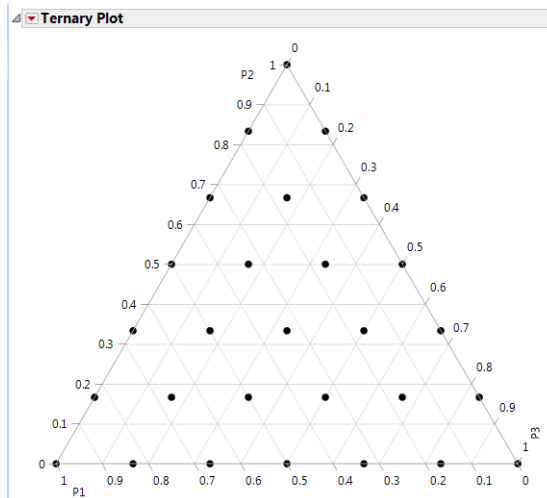
Mixture Designs Designs for Formulations

Use the Mixture Design platform to build experiments with factors that are components in a mixture. Choose from several classical mixture design approaches, such as simplex, extreme vertices, and lattice. For the extreme vertices approach, you can supply a set of linear inequality constraints limiting the geometry of the mixture factor space.

In mixture experiments, a factor's value is its proportion of the mixture, which falls between zero and one. Mixture experiments have three or more factors with the sum of the factor proportions equal to one (100%). Mixture experiments differ from other experimental types in that you cannot vary factors independently of one another. When you change the proportion of one factor, the proportion of one or more other factors must also change. This simple fact has a profound effect on every aspect of experimentation with mixtures: the factor space, the design properties, and the interpretation of the results.

Because the mixture components are proportions that sum to one, the feasible region of a mixture design takes the form of a simplex. For example, three factor designs can be visualized with a triangular 2-D graph and four factor designs can be visualized with a 3-D tetrahedron.

Figure 13.1 Ternary Plot



Contents

Overview of Mixture Designs	395
Mixture Design Window	395
Responses	396
Factors	398
Ternary Plot Overview	399
Examples of Mixture Design Types	400
Optimal Mixture Design	400
Simplex Centroid Design	402
Simplex Lattice Design	403
Extreme Vertices Design	405
ABCD Design	409
Space Filling Design	410
Fitting Mixture Designs	413
Whole Model Tests and Analysis of Variance Reports	414
Understanding Response Surface Reports	414
A Chemical Mixture Example	414
Create the Design	414
Analyze the Mixture Model	417
The Prediction Profiler	417
The Mixture Profiler	418

Overview of Mixture Designs

You can choose from the following types of mixture designs:

Optimal Generates an optimal mixture design using the custom designer. See [“Optimal Mixture Design”](#) on page 400.

Simplex Centroid Generates a design of pure mixtures and blends of factors up to a specified degree. In a pure mixture, one component is at 100% and all other components are at 0%. See [“Simplex Centroid Design”](#) on page 402.

Simplex Lattice Generates a design of pure mixtures and blends in a space filling triangular grid of runs. You can specify the number of grid levels. See [“Simplex Lattice Design”](#) on page 403.

Extreme Vertices Generates a design based on the vertices and combinations of the vertices of a constrained factor space. Available when you have linear constraints or restricted upper and lower bounds on one or more of your factors. See [“Extreme Vertices Design”](#) on page 405.

ABCD Design Generates a screening design for mixtures devised by Snee (1975). See [“ABCD Design”](#) on page 409.

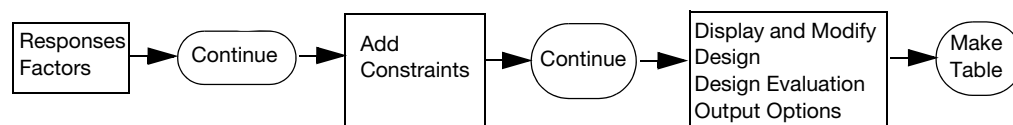
Space Filling Generates a space filling design that accommodates linear constraints. See [“Space Filling Design”](#) on page 410.

Note: The optimal, extreme vertices, and space filling designs can incorporate linear constraints and restrictions on components.

Mixture Design Window

The mixture design window updates as you work through the design steps. See [“The DOE Workflow: Describe, Specify, Design”](#) on page 57. The workflow and outlines vary depending on design type.

Figure 13.2 Mixture Design Workflow



Responses

Use the Responses outline to specify one or more responses.

Tip: When you have completed the Responses outline, consider selecting **Save Responses** from the red triangle menu. This option saves the response names, goals, limits, and importance values in a data table that you can later reload in DOE platforms.

Figure 13.3 Responses Outline

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.

optional item

Add Response Enters a single response with a goal type of Maximize, Match Target, Minimize, or None. If you select Match Target, enter limits for your target value. If you select Maximize or Minimize, entering limits is not required but can be useful if you intend to use desirability functions.

Remove Removes the selected responses.

Number of Responses Enters additional responses so that the number that you enter is the total number of responses. If you have entered a response other than the default Y, the Goal for each of the additional responses is the Goal associated with the last response entered. Otherwise, the Goal defaults to Match Target. Click the Goal type in the table to change it.

The Responses outline contains the following columns:

Response Name The name of the response. When added, a response is given a default name of Y, Y2, and so on. To change this name, double-click it and enter the desired name.

Goal, Lower Limit, Upper Limit The Goal tells JMP whether you want to maximize your response, minimize your response, match a target, or that you have no response goal. JMP assigns a Response Limits column property, based on these specifications, to each response column in the design table. It uses this information to define a desirability function for each response. The Profiler and Contour Profiler use these desirability functions to find optimal factor settings. For further details, see the Profiler chapter in *Profiler* and [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

- A Goal of Maximize indicates that the best value is the largest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.

- A Goal of Minimize indicates that the best value is the smallest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Match Target indicates that the best value is a specific target value. The default target value is assumed to be midway between the Lower Limit and Upper Limit.
- A Goal of None indicates that there is no goal in terms of optimization. No desirability function is constructed.

Note: If your target response is not midway between the Lower Limit and the Upper Limit, you can change the target after you generate your design table. In the data table, open the Column Info window for the response column (**Cols > Column Info**) and enter the desired target value.

Importance When you have several responses, the Importance values that you specify are used to compute an overall desirability function. These values are treated as weights for the responses. If there is only one response, then specifying the Importance is unnecessary because it is set to 1 by default.

Editing the Responses Outline

In the Responses outline, note the following:

- Double-click a response to edit the response name.
- Click the goal to change it.
- Click on a limit or importance weight to change it.
- For multiple responses, you might want to enter values for the importance weights.

Response Limits Column Property

The Goal, Lower Limit, Upper Limit, and Importance that you specify when you enter a response are used in finding optimal factor settings. For each response, the information is saved in the generated design data table as a Response Limits column property. JMP uses this information to define the desirability function. The desirability function is used in the Prediction Profiler to find optimal factor settings. For further details about the Response Limits column property and examples of its use, see [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

If you do not specify a Lower Limit and Upper Limit, JMP uses the range of the observed data for the response to define the limits for the desirability function. Specifying the Lower Limit and Upper Limit gives you control over the specification of the desirability function. For more details about the construction of the desirability function, see the Profiler chapter in *Profilers*.

Factors

Add factors in the Factors outline.

Tip: When you have completed the Factors outline, consider selecting **Save Factors** from the red triangle menu. This saves the factor names, roles, changes, and values in a data table that you can later reload.

Figure 13.4 Factors Outline

Factors

Add 1 Mixture

Remove Selected

Name	Role	Values

Specify Factors

Specify desired number of factors. Double click on a factor name or setting to edit it.

Continue

Add Enter the number of factors to add and click Add.

Remove Selected Removes the selected factors.

Factors List

The Factors list contains the following columns:

Name The name of the factor. When added, a factor is given a default name of X1, X2, and so on. To change this name, double-click it and enter the desired name.

Role Specifies the Design Role of the factor. The Design Role column property for the factor is saved to the data table. This property ensures that the factor type is modeled appropriately.

Values The experimental settings for the factors. To insert Values, click the default values and enter the desired values.

Editing the Factors List

In the Factors list, do the following:

- To edit a factor name, double-click the factor name.
- To edit a value, click the value in the Values column.

Linear Constraints

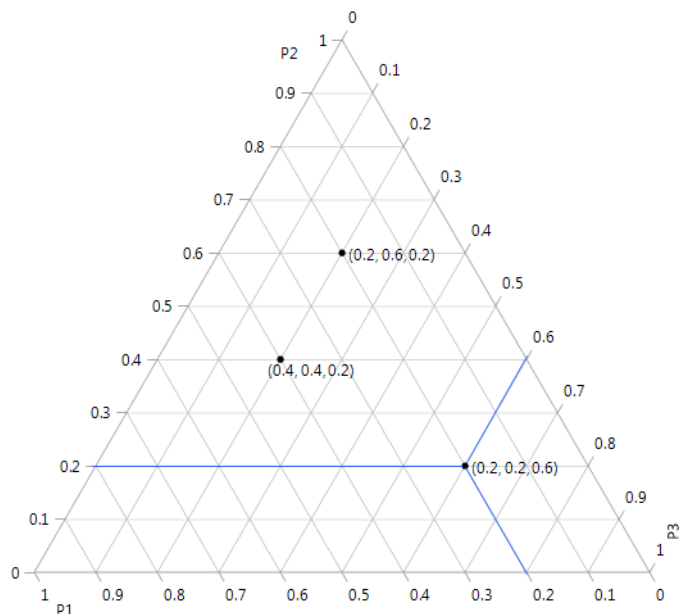
Click the **Linear Constraint** button to enter one or more linear inequality constraints. A template appears for a linear expression involving all the continuous factors in your design. Enter coefficient values for the factors and select the direction of the inequality to reflect your linear constraint. Specify the constraining value in the box to the right of the inequality. To add more constraints, click **Linear Constraint** again.

Note: When you save a script for a design that involves a linear constraint, the script expresses the linear constraint as a *less than or equal to* inequality ($\leq$).

Ternary Plot Overview

A ternary plot is a two-dimensional representation of three mixture components that sum to a constant. The plot is an equilateral triangle with an edge for each component. When unconstrained, each vertex of the triangle corresponds to a pure blend where one component is 1 (100%) and all other components are 0. When components are constrained, the feasible mixtures are represented by a portion of the ternary plot. Shading is used to exclude infeasible portions of the plot.

Figure 13.5 displays three components P1, P2, and P3. The three components are unconstrained with the range of each proportion ranging from 0 to 1 (100%). Three points are labeled with their coordinates (P1, P2, P3). One point (0.2, 0.2, 0.6) includes blue lines along the axis grid lines for each mixture component to guide you in reading the axes.

Figure 13.5 Ternary Plot Example

Examples of Mixture Design Types

- “Optimal Mixture Design”
- “Simplex Centroid Design”
- “Simplex Lattice Design”
- “Extreme Vertices Design”
- “ABCD Design”
- “Space Filling Design”

Optimal Mixture Design

Optimal mixture designs can be generated using either the Mixture Design Platform or the Custom Design platform.

Here you use the Mixture Platform to create an example optimal mixture design with three factors:

1. Select **DOE > Classical > Mixture Design**.

2. (Optional) To match the output of this example, click the Mixture Design red triangle and select Set Random Seed, and then enter 1409.
3. (Optional) To match the output of this example, click the Mixture Design red triangle and select Number of Starts, and then enter 2.
4. Use three factors for this example. No changes need to be made to the Factors section.
5. Click **Continue**.
6. Click **Optimal** on the Choose Mixture Design Type panel.
7. In the **Define Factor Constraints** panel, click **Specify Linear Constraints**, and then click **Add**.
8. Enter a 1 in the X1 and X2 boxes and enter 0.8 in the constraint box.
This constrains $X1+X2$ to less than 80% of the mixture.
9. In the **Model** panel, click **Interactions > 2nd**.
This adds interaction effects to the model.
10. In the **Design Generation** panel, enter 2 for the **Number of Center Points**.
11. Click **Make Design**.

Figure 13.6 12-Run Optimal Design Runs

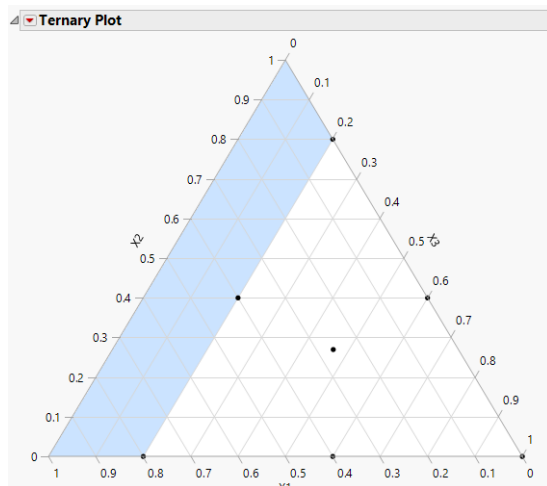
Design			
Run	X1	X2	X3
1	0.8	0	0.2
2	0	0	1
3	0	0.8	0.2
4	0	0	1
5	0	0.4	0.6
6	0.4	0	0.6
7	0	0.8	0.2
8	0.4	0	0.6
9	0.4	0.4	0.2
10	0.8	0	0.2
11	0.264183	0.269687	0.46613
12	0.264183	0.269687	0.46613

12. Click **Make Table**.

Note: In practice, you would round the mixtures amounts to the appropriate level of precision.

Visualize the Design

1. From the design table, select **Graph > Ternary Plot**.
2. Select X1, X2, and X3 and click **X, Plotting**, and then click **OK**.

Figure 13.7 Ternary Plot for Optimal Design


The ternary plot shows the 7 unique mixtures that make up the design. The shading indicates the constrained region. For more information about ternary plots, see [“Ternary Plot Overview”](#) on page 399.

Simplex Centroid Design

A simplex centroid design of degree k with n factors consists of mixtures with the following characteristics:

- all one factor mixtures or pure blends
- all blends of two factors at equal levels
- additional blends, up to k factors at a time blended at k equal levels

A center point run with equal amounts of all the ingredients is always included.

Creating the Design

To create an example simplex centroid design:

1. Select **DOE > Classical > Mixture Design**.
2. Use three factors for this example. No changes need to be made to the Factors section.
3. Click **Continue**.
4. Click **Simplex Centroid** to generate the design

Note: The default setting of $k = 2$ was used. There is a text box that enables you to adjust the value of k .

Figure 13.8 Three Factor Simplex Centroid Design of Degree 2.

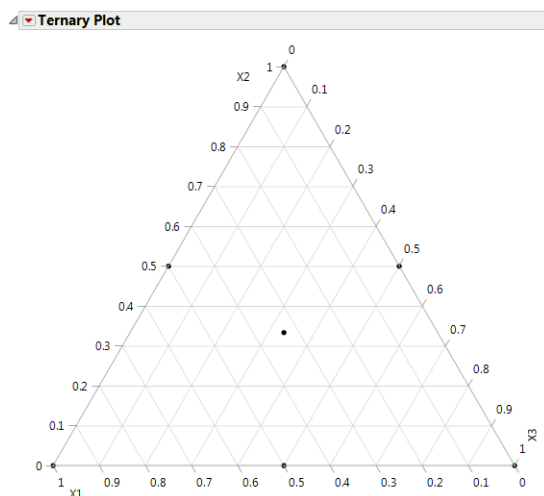
Design			
Run	X1	X2	X3
1	1.00000	0.00000	0.00000
2	0.00000	1.00000	0.00000
3	0.00000	0.00000	1.00000
4	0.50000	0.50000	0.00000
5	0.50000	0.00000	0.50000
6	0.00000	0.50000	0.50000
7	0.33333	0.33333	0.33333

5. Click **Make Table**.

Visualize the Design

- From design table, select **Graph>Ternary Plot**.
- Select X1, X2, and X3 and click **X, Plotting**, and then click **OK**.

Figure 13.9 Ternary Plot for Simplex Centroid Design



For more information about ternary plots, see [“Ternary Plot Overview”](#) on page 399.

Simplex Lattice Design

The simplex lattice design is a space filling design that creates a triangular grid of runs. The design is the set of blends where the factors' values are i/m , where i varies from 0 to m and the sum of the factors is 1. For a three-level design, a factor can be set at 0, 0.33, 0.66, or 1.

To create an example simplex lattice design:

- Select **DOE > Classical > Mixture Design**.

2. Use three factors for this example. No changes need to be made to the Factors section.
3. Click **Continue**.
4. Click **Simplex Lattice** to generate the design.

Note: The default setting of 5 levels was used. There is a text box that enables you to adjust the number of levels

Figure 13.10 Three Factor Five-Level Simplex Lattice Design

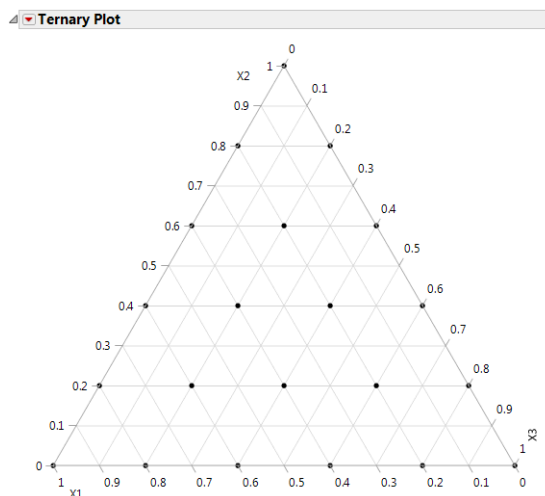
Design			
Run	X1	X2	X3
1	0.00000	0.00000	1.00000
2	0.00000	0.20000	0.80000
3	0.00000	0.40000	0.60000
4	0.00000	0.60000	0.40000
5	0.00000	0.80000	0.20000
6	0.00000	1.00000	0.00000
7	0.20000	0.00000	0.80000
8	0.20000	0.20000	0.60000
9	0.20000	0.40000	0.40000
10	0.20000	0.60000	0.20000
11	0.20000	0.80000	0.00000
12	0.40000	0.00000	0.60000
13	0.40000	0.20000	0.40000
14	0.40000	0.40000	0.20000
15	0.40000	0.60000	0.00000
16	0.60000	0.00000	0.40000
17	0.60000	0.20000	0.20000
18	0.60000	0.40000	0.00000
19	0.80000	0.00000	0.20000
20	0.80000	0.20000	0.00000
21	1.00000	0.00000	0.00000

5. Click **Make Table**.

Visualize the Design

6. From the design table, select **Graph >Ternary Plot**.
7. Select X1, X2, and X3 and click **X, Plotting**, and then click **OK**.

Figure 13.11 Ternary Plot for Simplex Lattice Design



For more information about ternary plots, see [“Ternary Plot Overview”](#) on page 399.

As the number of levels increases, the number of design points increases, decreasing the space between design points. In contrast to the simplex centroid design, the simplex lattice design does not necessarily include the centroid.

Extreme Vertices Design

An extreme vertices design requires restricted ranges on one or more of the factors, as defined by limits in the Factors panel or by linear constraints. An extreme vertices design consists of mixtures at the vertices of the factor space and at the averages of vertices up to a specified degree. The centroid point for two neighboring vertices joined by a line is a second-degree centroid. The centroid point for vertices sharing a plane is a third-degree centroid.

An Extreme Vertices Example with Range Constraints

To create an example extreme vertices design:

1. Select **DOE > Classical > Mixture Design**.
2. Under Factors, enter 2 for the number of additional factors to add and click **Add**.
3. Set the ranges for the five factors as shown in Figure 13.12 and click **Continue**.

Figure 13.12 Ranges for Five-factors

Name	Role	Values	
▲ X1	Mixture	0.05	0.25
▲ X2	Mixture	0.1	0.3
▲ X3	Mixture	0.1	0.3
▲ X4	Mixture	0.1	0.4
▲ X5	Mixture	0.05	0.25

- Enter 4 in the **Degree** text box and click **Extreme Vertices**.

Figure 13.13 Display and Modify Panel for Extreme Vertices Example

Display and Modify Design

Output Options

Run Order: Randomize

D Optimal Subset

Choose desired sample size: 116

Make Table Find Subset Back

The default number of runs for this five-factor, degree 4 extreme vertices design is 116. This is the number of extreme vertices and averages up to degree 4 for these five factors given the constraints on their ranges. JMP uses the default set of mixtures as a candidate set to generate a smaller D Optimal design.

- (Optional) To match the output of this example, click the Mixture Design red triangle and select Set Random Seed, and then enter 1409.
- Enter 10 in the Choose desired sample size box and click **Find Subset** to generate the design.

Note: The **Find Subset** option uses the row exchange method (not coordinate exchange) to find the optimal subset of rows.

Figure 13.14 Ten Run D-optimal Extreme Vertices Design

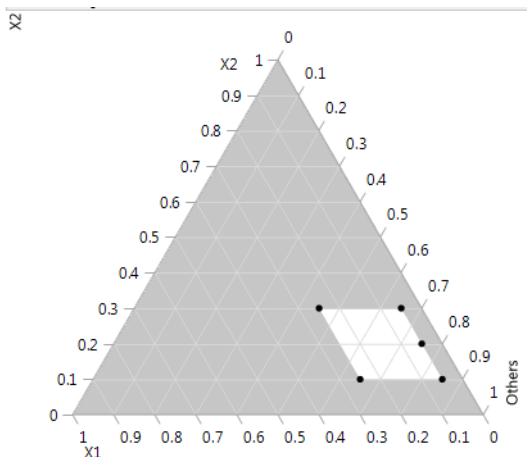
Design					
Run	X1	X2	X3	X4	X5
1	0.05000	0.30000	0.20000	0.40000	0.05000
2	0.05000	0.10000	0.30000	0.40000	0.15000
3	0.25000	0.30000	0.10000	0.10000	0.25000
4	0.05000	0.30000	0.30000	0.10000	0.25000
5	0.25000	0.10000	0.30000	0.10000	0.25000
6	0.25000	0.30000	0.30000	0.10000	0.05000
7	0.25000	0.10000	0.30000	0.30000	0.05000
8	0.05000	0.20000	0.10000	0.40000	0.25000
9	0.25000	0.30000	0.10000	0.30000	0.05000
10	0.25000	0.10000	0.10000	0.30000	0.25000

- Click **Make Table**.

Visualize the Design

8. From the design table, select **Graph > Ternary Plot**.
9. Select X1, X2, X3, X4, and X5 and click **X, Plotting**, and then click **OK**.

Figure 13.15 Partial Output of Ternary Plot for Five-Factor Design



When you have more than three mixture components, the ternary plot shows a series of plots. Each plot has two component axes and the third axis is the sum of all other components. The plots are shaded if you have a constrained region. The unshaded portion represents the feasible region. For more information about ternary plots, see [“Ternary Plot Overview”](#) on page 399.

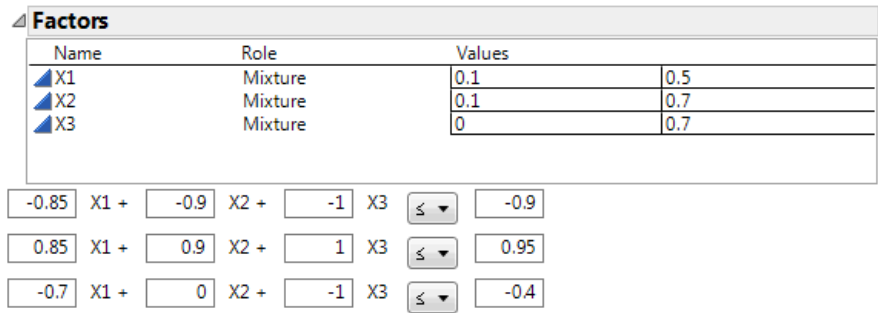
An Extreme Vertices Example with Linear Constraints

Consider the classic example presented by Snee (1979) and Piepel (1988). This example has three factors, X1, X2, and X3, with factor bounds and three linear constraints.

To create an extreme vertices design for this example:

1. Select **DOE > Classical > Mixture Design**.
2. Enter the values from Figure 13.16 for X1, X2, and X3 and click **Continue**.

Figure 13.16 Values and Linear Constraints for the Snee and Piepel Example

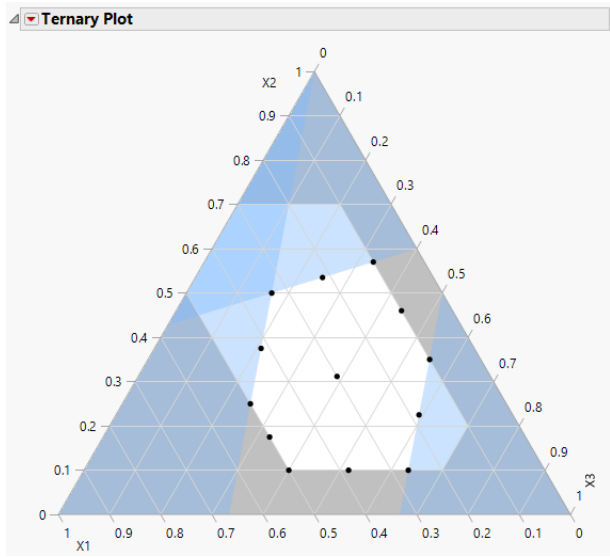


- 3. Click **Linear Constraint** three times. Enter the constraints as shown in Figure 13.16.
- 4. Click the **Extreme Vertices** button.
- 5. Click **Make Table**.

Visualize the Design

- 6. From the design table, select **Graph >Ternary Plot**.
- 7. Select X1, X2, and X3 and click **X, Plotting**, and then click **OK**.

Figure 13.17 Ternary Plot Showing Piepel Example with Constraints



The ternary plot shows the feasible region as defined by the factor limits and the linear constraints. The design points are at the six vertices of the feasible region, at the six edge mid-points, and the overall centroid. For more information about ternary plots, see [“Ternary Plot Overview”](#) on page 399.

Statistical Details for Extreme Vertices Method

If there are linear constraints, JMP uses the CONSIM algorithm developed by R.E. Wheeler, described in Snee (1979) and presented by Piepel (1988) as CONVRT. The method is also described in Cornell (1990, Appendix 10a). The method combines constraints and checks to see whether vertices violate them. If so, it drops the vertices and calculates new ones. The method named CONAEV for doing centroid points is by Piepel (1988).

If there are no linear constraints (only range constraints), the extreme vertices design is constructed using the XVERT method developed by Snee and Marquardt (1974) and Snee (1975). After the vertices are found, a simplex centroid method generates combinations of vertices up to a specified order.

The XVERT method first creates a full 2^{nf-1} design using the given low and high values of the $nf-1$ factors with smallest range. Then, it computes the value of the one factor left out based on the restriction that the factors' values must sum to one. It keeps points that are not in factor's range. If not, it increments or decrements the value to bring it within range, and decrements or increments each of the other factors in turn by the same amount. This method keeps the points that still satisfy the initial restrictions.

The above algorithm creates the vertices of the feasible region in the simplex defined by the factor constraints. However, Snee (1975) has shown that it can also be useful to have the centroids of the edges and faces of the feasible region. A generalized n -dimensional face of the feasible region is defined by $nf-n$ of the boundaries and the centroid of a face defined to be the average of the vertices lying on it. The algorithm generates all possible combinations of the boundary conditions and then averages over the vertices generated on the first step.

ABCD Design

The ABCD Design is a screening design for mixtures. See Snee (1975).

To create an example ABCD design:

1. Select **DOE > Classical > Mixture Design**.
2. Under Factors, enter 3 for the number of additional factors to add and click **Add**.
3. Click **Continue**.
4. Click **ABCD Design**.

Figure 13.18 14 of the 28 Run 6 Factor ABCD Design

Design						
Run	X1	X2	X3	X4	X5	X6
1	1.00000	0.00000	0.00000	0.00000	0.00000	0.00000
2	0.00000	1.00000	0.00000	0.00000	0.00000	0.00000
3	0.00000	0.00000	1.00000	0.00000	0.00000	0.00000
4	0.00000	0.00000	0.00000	1.00000	0.00000	0.00000
5	0.00000	0.00000	0.00000	0.00000	1.00000	0.00000
6	0.00000	0.00000	0.00000	0.00000	0.00000	1.00000
7	0.50000	0.50000	0.00000	0.00000	0.00000	0.00000
8	0.50000	0.00000	0.50000	0.00000	0.00000	0.00000
9	0.50000	0.00000	0.00000	0.50000	0.00000	0.00000
10	0.50000	0.00000	0.00000	0.00000	0.50000	0.00000
11	0.50000	0.00000	0.00000	0.00000	0.00000	0.50000
12	0.00000	0.50000	0.50000	0.00000	0.00000	0.00000
13	0.00000	0.50000	0.00000	0.50000	0.00000	0.00000
14	0.00000	0.50000	0.00000	0.00000	0.50000	0.00000

5. Click **Make Table**.

Space Filling Design

The Space Filling mixture designs spread design points throughout the design region. It accommodates linear constraints. The design is generated in a fashion similar to the Fast Flexible Filling design method found under DOE > Special Purpose > Space Filling Design ("[Fast Flexible Filling Designs](#)" on page 636).

Two Mixture Design red triangle options relate specifically to Space Filling Designs:

FFF Optimality Criterion For the Fast Flexible Filling mixture design type, enables you to select between the MaxPro criterion (the default) and the Centroid criterion. See "[FFF Optimality Criterion](#)" on page 411.

Advanced Options > Set Average Cluster Size For the Fast Flexible Filling mixture design type, enables you to specify the average number of randomly generated points used to define each cluster or, equivalently, each design point. See "[Set Average Cluster Size](#)" on page 412.

Space Filling Example

To create an example space filling design:

1. Select **DOE > Classical > Mixture Design**.
2. (Optional) To match the output of this example click the Mixture Design red triangle and select Set Random Seed and enter 1409.
3. Use three factors for this example. No changes need to be made to the Factors section.
4. Click **Continue**.
5. Click **Linear Constraints**

6. Enter 0.7 in the X1 box, 1 in the X3 box, change the direction of the inequality to $\geq$ and enter 0.4 in the constraint box.

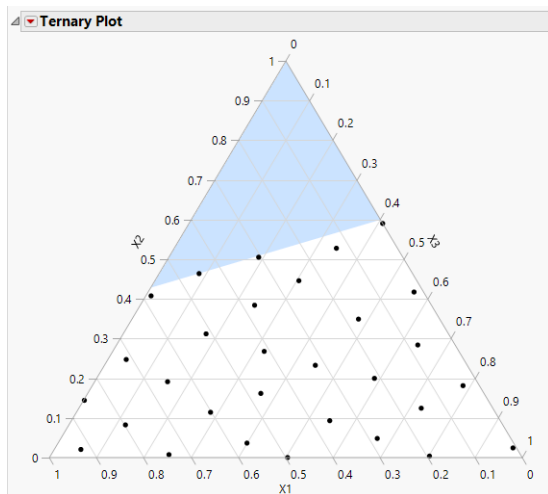
This constrains $0.7X1+X3$ to at least 40% of the mixture.

7. Enter 30 in the **Runs** box.
8. Click **Space Filling**.
9. Click **Make Table**.

Visualize the Design

10. From the design table, select **Graph >Ternary Plot**.
11. Select X1, X2, and X3 and click **X, Plotting**, and then click **OK**.

Figure 13.19 Space Filling Design with One Linear Constraint



This design is constructed using the Centroid FFF Optimality Criterion. Note that the points fall in the constrained design region and are fairly well spread throughout this region. For more information about ternary plots, see [“Ternary Plot Overview”](#) on page 399.

Statistical Details for Mixture Space Filling Designs

FFF Optimality Criterion

The algorithms for Fast Flexible Filling designs begin by generating a large number of random points within the specified design region. These points are then clustered using a Fast Ward algorithm into a number of clusters that equals the Number of Runs that you specified.

The final design points can be obtained by using the default MaxPro (*maximum projection*) optimality criterion or by selecting the Centroid criterion. You can find these options under FFF Optimality Criterion in the report's red triangle menu.

MaxPro For p factors and n equal to the specified Number of Runs, the MaxPro criterion strives to find points in the clusters that minimize the following criterion:

$$C_{MaxPro} = \sum_i^{n-1} \sum_{j=i+1}^n \left[1 / \prod_{k=1}^p (x_{ik} - x_{jk})^2 \right]$$

The MaxPro criterion maximizes the product of the distances between potential design points in a way that involves all factors. This supports the goal of providing good space-filling properties on projections of factors. See Joseph et al. (2015). The Max Pro option is the default.

Centroid This method places a design point at the centroid of each cluster. It has the property that the average distance from an arbitrary point in the design space to its closest neighboring design point is smaller than for other designs.

Note: You can set a preference to always use a given optimality criterion. Select **File > Preferences > Platforms > DOE**. Select FFF Optimality Criterion and select your preferred criterion.

Set Average Cluster Size

The **Set Average Cluster Size** option is found under Advanced Options in the Mixture Design red triangle menu. This option enables you to specify the average number of uniformly generated points used to define each cluster or, equivalently, each design point.

By default, if the number of Runs for the Space Filling design type is 200 or smaller, a total of 10,000 random uniformly generated points are used as the basis for the clustering algorithm. When the number of Runs exceeds 200, the default value is 50. Increasing this value can be particularly useful in designs with a large number of factors.

Note: Depending on the number of factors and the specified value for Runs, you might want to increase the average number of initial points per design point by selecting **Advanced Options > Set Average Cluster Size**.

Linear Constraints

The design region can be restricted by selecting the **Linear Constraint** option in the Linear Constraints outline.

When you specify linear constraints, the random points that form the basis for the clustering algorithm are randomly distributed within the constrained design region. The clustering algorithm uses these points.

Fitting Mixture Designs

When fitting a model for mixture designs, one must take into account that the factors sum to a constant. A traditional full linear model is not fully estimable.

An appropriate response surface model for mixture responses is the Scheffé polynomial (Scheffé 1958). See the discussion of Cox Mixtures and the Scheffé cubic macro in the Standard Least Squares Report and Options chapter in *Fitting Linear Models*. The Scheffé polynomial model does the following:

- suppresses the intercept
- includes all the linear main-effect terms
- excludes all the square terms (such as X_1^2)
- includes all the cross terms (such as X_1X_2)

In this model, the parameters are easy to interpret (Cornell 1990). The coefficients on the linear terms are the fitted response at the extreme points where the mixture consists of a single factor. The coefficients on the cross terms indicate the curvature across each edge of the factor space.

For a mixture model with third-degree polynomial terms, the Scheffé cubic model can be used. The Scheffé cubic model includes terms of the form $X_1X_2(X_1-X_2)$.

- To fit a Scheffé polynomial model use the Mixture Response Surface macro in the Fit Model platform.
- To fit the Scheffé cubic model use the **Scheffe Cubic** macro in the Fit Model platform.
- If you chose to fit a different mixture model the Fit Model platform includes options for a No Intercept model and Mixture Effects as an effect attribute.

Tip: The custom design model outline has an option for adding the Scheffé cubic model terms. The generated design table will include a script for the Scheffé cubic model.

Whole Model Tests and Analysis of Variance Reports

In a whole-model Analysis of Variance table, JMP traditionally tests that all the parameters are zero except for the intercept. In a mixture model without an intercept, JMP looks for a hidden intercept, in the sense that a linear combination of effects is a constant. If it finds a hidden intercept, it does the whole model test with respect to the intercept model rather than a zero-intercept model. This test is equivalent to testing that all the parameters are zero except the linear parameters, and testing that they are equal.

The hidden-intercept property also causes the R^2 to be reported with respect to the intercept model rather than reported as missing.

Understanding Response Surface Reports

When there are effects marked as response surface effects “&RS,” JMP creates additional reports that analyze the fitted response surface. These reports were originally designed for full response surfaces, not mixture models. However, JMP might encounter a no-intercept model and find a hidden intercept with linear response surface terms, but no square terms. Then it *folds* its calculations, collapsing on the last response surface term to calculate critical values for the optimum. This can be done for any combination that yields a constant and involves the last response surface term.

A Chemical Mixture Example

Three plasticizers (p1, p2, and p3) comprise 79.5% of the vinyl used for automobile seat covers (Cornell, 1990). Within this 79.5%, the individual plasticizers are restricted by the following constraints: $0.409 \leq x_1 \leq 0.849$, $0 \leq x_2 \leq 0.252$, and $0.151 \leq x_3 \leq 0.274$. Cornell uses a 14-run design to fit a quadratic mixture model to the collected responses. The design is a 3-degree extreme vertices design with replicate measurements taken at the four vertices and the overall centroid.

- “Create the Design”
- “Analyze the Mixture Model”
- “The Prediction Profiler”
- “The Mixture Profiler”

Create the Design

To create Cornell’s mixture design in JMP

1. Select **Help > Sample Data Library** and open **Plastifactors.jmp**.

2. Select **DOE > Classical > Mixture Design**.
3. Click the Mixture Design red triangle and select **Load Factors**.

Figure 13.20 Factors and Factor Constraints for the Plasticizer Experiment

Factors

Add Mixture

Remove Selected

Name	Role	Values
p1	Mixture	0.409 0.849
p2	Mixture	0 0.252
p3	Mixture	0.151 0.274

4. Click **Continue**.
5. Enter 3 in the **Degree** box.
6. Click **Extreme Vertices**.
7. Click **Make Table**. JMP uses the 9 factor settings to generate a JMP table.

Figure 13.21 Extreme Vertices Mixture Design

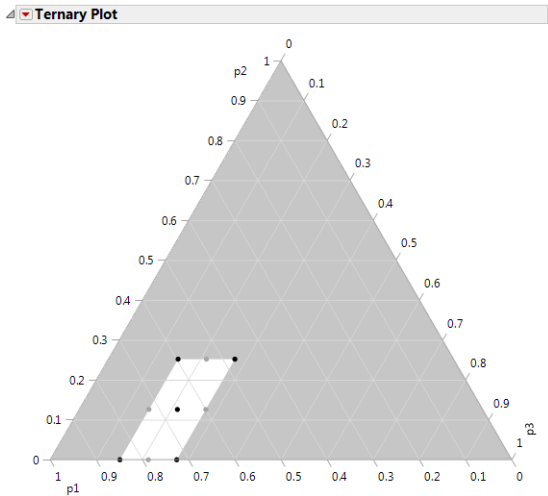
	p1	p2	p3	Y
1	0.597	0.252	0.151	•
2	0.726	0	0.274	•
3	0.6615	0.126	0.2125	•
4	0.6	0.126	0.274	•
5	0.7875	0	0.2125	•
6	0.849	0	0.151	•
7	0.5355	0.252	0.2125	•
8	0.474	0.252	0.274	•
9	0.723	0.126	0.151	•

Note: Your table might differ due to the random seed used to generate the design.

Next, you add an extra five design runs by duplicating the vertex points and overall centroid, to generate the 14 run design.

8. To identify the vertex points and the centroid use a ternary plot. From the design table, select **Graph > Ternary Plot**.
9. Select p1, p2, and p3 and click **X, Plotting**, and then click **OK**.

Figure 13.22 Ternary Plot for Design



10. Highlight the vertices and centroid points.

Figure 13.23 Identify Vertices and Center Point

	p1	p2	p3	Y
1	0.597	0.252	0.151	•
2	0.726	0	0.274	•
3	0.6615	0.126	0.2125	•
4	0.6	0.126	0.274	•
5	0.7875	0	0.2125	•
6	0.849	0	0.151	•
7	0.5355	0.252	0.2125	•
8	0.474	0.252	0.274	•
9	0.723	0.126	0.151	•

11. Select **Edit > Copy**, to copy the selected rows to the clipboard.
12. Click in the first cell of row 10 and select **Edit > Paste** to add the duplicate rows to the table.

Note: A similar design can be obtained using the custom designer with the quadratic model, 2 center points, 4 replicate designs, and 14 runs.

The Plasticizer design with the results (Y values) that Cornell obtained are available in the sample data file Plasticizer.jmp.

Analyze the Mixture Model

To analyze the Plasticizer experiment use the sample data table Plasticizer.jmp.

1. Select **Help > Sample Data Library** and open Plasticizer.jmp.
2. Click the green triangle next to the **Model** script in the upper left corner of the data table.

A completed Fit Model launch window appears.

Note: The model script is saved to the design table when you create a design.

3. Click **Run** to see the response surface analysis.
4. Plasticizer.jmp contains a column called Pred Formula Y. This column was added after the analysis by selecting **Save Columns > Prediction Formula** from the Response Y red triangle.
5. To see the prediction formula, click the plus symbol in the column list:

$$0 - 50.1465 * p1 - 282.198 * p2 - 911.648 * p3 + p2 * p1 * 317.363 + p3 * p1 * 1464.330 + p3 * p2 * 1846.218$$

Note: These results correct the coefficients reported in Cornell (1990).

The Response Surface Solution report shows that a maximum predicted value of 19.570299 occurs at point (0.63505, 0.15568, 0.20927).

Figure 13.24 Mixture Response Surface Analysis

Response Surface				
Coef				Y
(p1-0.474)/0.375	0	44.62923	205.92138	7.4923077
p2/0.375	.	0	259.62437	-1.491318
(p3-0.151)/0.375	.	.	0	-138.2039

Solution	
Variable	Critical Value
(p1-0.474)/0.375	0.635052
p2/0.375	0.1556806
(p3-0.151)/0.375	0.2092675

Solution is a Maximum
Assuming the following mixture sum: p1+p2+p3=1
Predicted Value at Solution 19.570299

The Prediction Profiler

The report contains a prediction profiler.

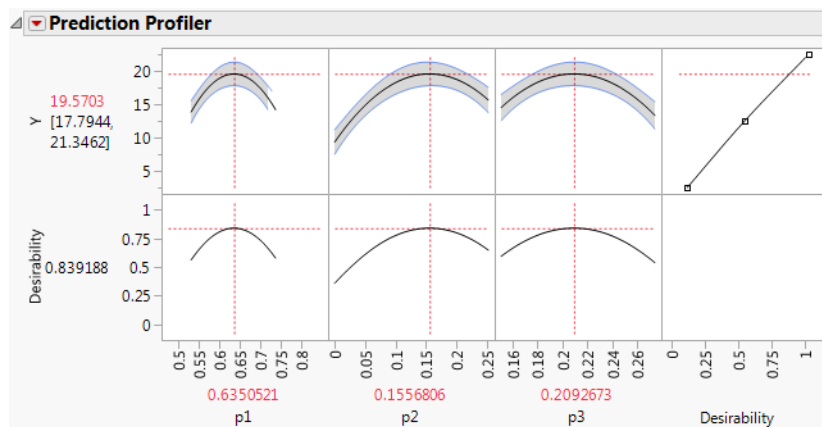
1. From the prediction profiler you can see how the crossed effects show as curvature in the prediction traces. As you change values of one component (drag the red lines) the values of the other two components move in the opposite direction to maintain their ratio and the overall mixture constraint (components must sum to 100%).

Note: The axes of prediction profiler traces range from the defined upper and lower bounds of the factors, p1, p2, and p3. When you explore the effect of a component and the limit of a second component is reached, it cannot move further and only the third variable changes to maintain the mixture constraint.

2. To the visible profile curves to bounds that take into account the levels of all three components, click the Prediction Profiler red triangle and select **Profile at Boundary > Stop at Boundaries**.
3. To optimize the mixture components click the Prediction Profiler red triangle and select **Optimization and Desirability > Desirability Functions**.
4. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability** to obtain the optimal factor settings for maximum Y.

The profiler displays optimal settings (rounded) of 0.6350 for p1, 0.1557 for p2, and 0.2093 for p3, which give an estimated response of 19.5703.

Figure 13.25 Maximum Desirability in Profiler for Mixture Analysis Example



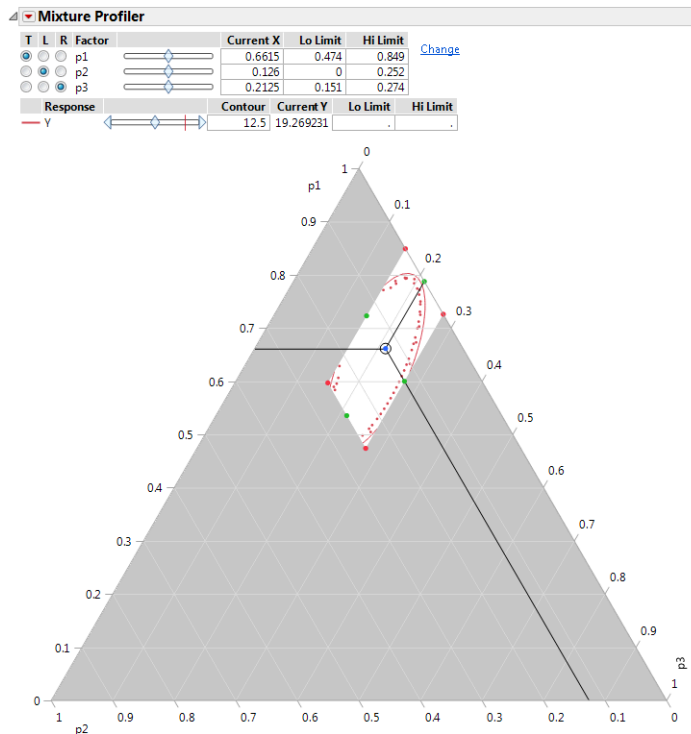
The Mixture Profiler

The Fit Model report also has a **Mixture Profiler** that is useful for visualizing and optimizing response surfaces from mixture experiments. Many of the features are the same as those of the Contour Profiler. However, some are unique to the Mixture Profiler:

- A ternary plot is used instead of a Cartesian plot, which enables you to view three mixture factors at a time.
- If you have more than three factors, radio buttons let you choose which factors to plot.
- If the factors have constraints, you can enter their low and high limits in the Lo Limit and Hi Limit columns. This setting shades non-feasible regions in the profiler.

Click the Response Y red triangle and elect **Factor Profiling** > **Mixture Profiler** to see the mixture profiler for the plasticizer data.

Figure 13.26 Mixture Profiler for Plasticizer Example



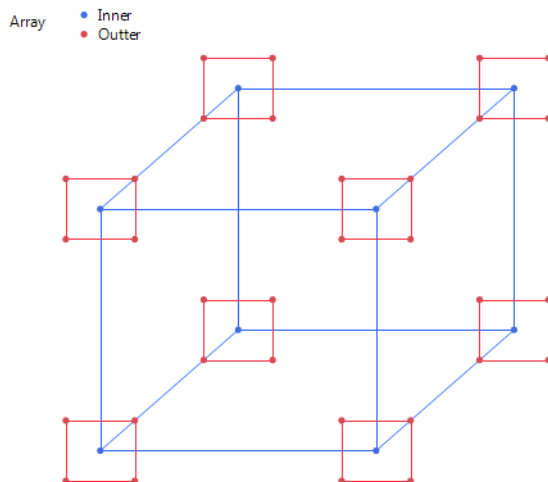
Chapter 14

Taguchi Designs Designs with Noise Factors

A Taguchi design uses crossed array to explore factors in the presence of noise. A Taguchi design has two parts: the inner array is a design for control factors, and the outer array is a design for the noise factors. The full design is the cross product of the two arrays.

Note that noise factors in production that are not easy or cost effective to control must be controlled during an experiment. Alternatives to Taguchi designs include combined arrays and mixed resolution designs (Borrer 2000). For an example of a design with noise factors using the custom designer see “[Experiments for Robust Process and Product Design](#)” on page 207 in the “Examples of Custom Designs” chapter.

Figure 14.1 Taguchi Design with Three Control Factors and Two Noise Factors



Contents

Overview of Taguchi Designs 423

Example of a Taguchi Design 423

Taguchi Design Window 427

 Response..... 428

 Factors..... 428

 Choose Inner and Outer Array Designs 429

 Display Coded Design 429

 Make the Design Table 430

Taguchi Design Options..... 430

Overview of Taguchi Designs

A Taguchi Design uses two types of factors: control factors and noise factors.

- An *inner* design is constructed for the control factors.
- An *outer* design is constructed for the noise factors.

The experiment is performed on all combinations of the inner and outer design runs. The noise factors have to be controlled during the experiment. The mean and standard deviation for each setting in the inner array across the noise settings are combined into a signal-to-noise measure for analysis. Table 14.1 lists the signal-to-noise performance statistics.

Table 14.1 Recommended Performance Statistics

Goal	S/N Ratio Formula
nominal is best	$\frac{S}{N} = 10\log\left(\frac{\bar{Y}^2}{s^2}\right)$
larger-is-better (maximize)	$\frac{S}{N} = -10\log\left(\frac{1}{n}\sum_i \frac{1}{Y_i^2}\right)$
smaller-is-better (minimize)	$\frac{S}{N} = -10\log\left(\frac{1}{n}\sum_i Y_i^2\right)$

Example of a Taguchi Design

The following example is an experiment described by Byrne and Taguchi (1986). The objective of the experiment is to find settings of control factors to maximize the adhesiveness (pull-off force) of nylon tubing. Four control factors are evaluated across three noise factors.

The signal factors for this example are as follows:

Interfer Tubing and connector interference. Signal factor with 3 levels.

Wall Wall thickness of the connector. Signal factor with 3 levels.

Depth Insertion depth of the tubing into the connector. Signal factor with 3 levels.

Adhesive Percent of adhesive. Signal factor with 3 levels.

The noise factors for this example are as follows:

Time Conditioning time. Noise factor with 2 levels.

Temperature Temperature. Noise factor with 2 levels.

Humidity Relative humidity. Noise factor with 2 levels.

Create the Design

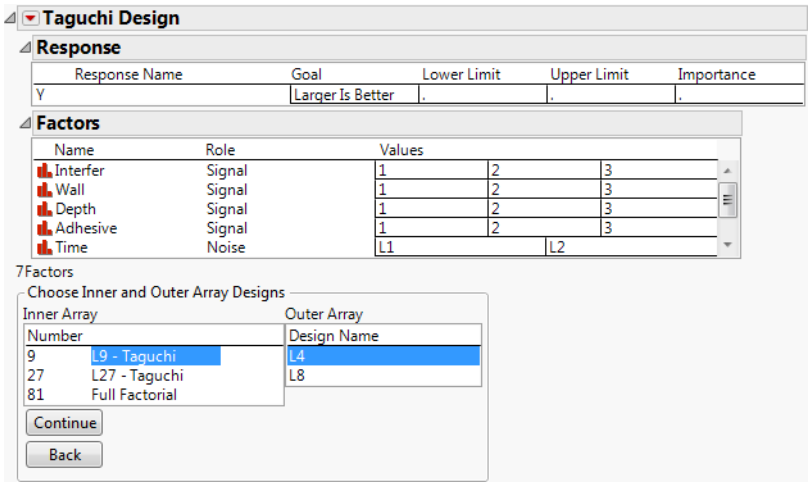
- 1. Select **DOE > Classical > Taguchi Arrays**.
- 2. Select **Help > Sample Data Library** and open Design Experiment/Byrne Taguchi Factors.jmp.
- 3. In the Taguchi Arrays window, click the Taguchi Design red triangle and select **Load Factors**.

The Factors panel shows the four three-level control (signal) factors and three noise factors.

Note: The Goal selected for the response determines the SN Ratio formula that is added to your design data table.

- 4. Ensure that **L9-Taguchi** is selected for the inner array.
- 5. Click **L8** for the outer array design.

Figure 14.2 Completed Taguchi Design Window



- 6. Click **Continue**.
 - 7. Click **Make Table** to create the design table shown in Figure 14.3.
- The inner array has nine runs for the four signal factors, each at three levels. The settings for the signal factors are provided in the first four columns of the table.

The outer design is a full factorial design for the three two-level noise factors. The outer array is provided in eight columns of the data table. The column names are the pattern of the outer array runs. For example, the column named “---” is for the results collected when all of the noise levels are at their low levels. Each of the nine trials are conducted at each of the eight combination of noise factors for a total of 72 experimental trials.

Figure 14.3 Taguchi Design before Data Entry

	Interfer	Wall	Depth	Adhesive	Pattern	---	--+	-+-	---+	+-	++	+++	Mean	SN Ratio
1	1	1	1	1	----	*	*	*	*	*	*	*	*	*
2	1	2	2	2	-000	*	*	*	*	*	*	*	*	*
3	1	3	3	3	--++	*	*	*	*	*	*	*	*	*
4	2	1	2	3	0-0+	*	*	*	*	*	*	*	*	*
5	2	2	3	1	00+-	*	*	*	*	*	*	*	*	*
6	2	3	1	2	0+-0	*	*	*	*	*	*	*	*	*
7	3	1	3	2	+-+0	*	*	*	*	*	*	*	*	*
8	3	2	1	3	+0-+	*	*	*	*	*	*	*	*	*
9	3	3	2	1	++0-	*	*	*	*	*	*	*	*	*

The design table allows for entry of the 72 experimental results.

8. Select **Help > Sample Data Library** and open Design Experiment/Byrne Taguchi Data.jmp.

Figure 14.4 Complete Taguchi Design Table (Byrne Taguchi Data.jmp)

	Interfer	Wall	Depth	Adhesive	Pattern	Y---	Y--+	Y+-	Y++	Y+-	Y++	Y++	Y++	Mean Y	SN Ratio Y
1	1	1	1	1	----	15.6	9.5	16.9	19.9	19.6	19.6	20	19.1	17.525	24.02534
2	1	2	2	2	-000	15	16.2	19.4	19.6	19.7	19.8	24.2	21.9	19.475	25.52164
3	1	3	3	3	--++	16.3	16.7	19.1	15.6	22.6	18.2	23.3	20.4	19.025	25.33476
4	2	1	2	3	0-0+	18.3	17.4	18.9	18.6	21	18.9	23.2	24.7	20.125	25.90425
5	2	2	3	1	00+-	19.7	18.6	19.4	25.1	25.6	21.4	27.5	25.3	22.825	26.90753
6	2	3	1	2	0+-0	16.2	16.3	20	19.8	14.7	19.6	22.5	24.7	19.225	25.32574
7	3	1	3	2	+-+0	16.4	19.1	18.4	23.6	16.8	18.6	24.3	21.6	19.85	25.71081
8	3	2	1	3	+0-+	14.2	15.6	15.1	16.8	17.8	19.6	23.2	24.4	18.3375	24.83231
9	3	3	2	1	++0-	16.1	19.9	19.3	17.3	23.1	22.7	22.6	28.6	21.2	26.15198

The SN Ratio Y column is the performance statistic (larger is better) to maximize the response. It is computed as -10 times the common logarithm of the average of the squared reciprocals of the responses:

$$-10 \log_{10} \left[\text{Mean} \left[\frac{1}{(Y_{--})^2}, \frac{1}{(Y_{-+})^2}, \frac{1}{(Y_{+-})^2}, \frac{1}{(Y_{++})^2}, \frac{1}{(Y_{+-})^2}, \frac{1}{(Y_{++})^2}, \frac{1}{(Y_{++})^2}, \frac{1}{(Y_{++})^2} \right] \right]$$

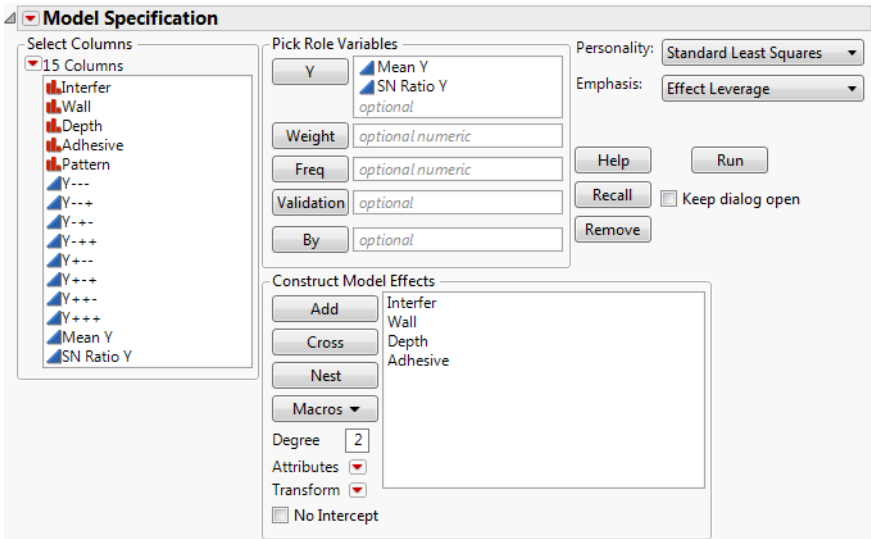
This expression is large when all of the individual y response values are large. That is, you are trying to find the signal settings that result in the largest response across the noise settings.

Analyze the Data

The data are now ready to analyze. The goal of the analysis is to find factor settings that maximize both the mean and the signal-to-noise ratio.

- 1. In the Byrne Taguchi Data.jmp data table, click the green arrow to run the Model script.

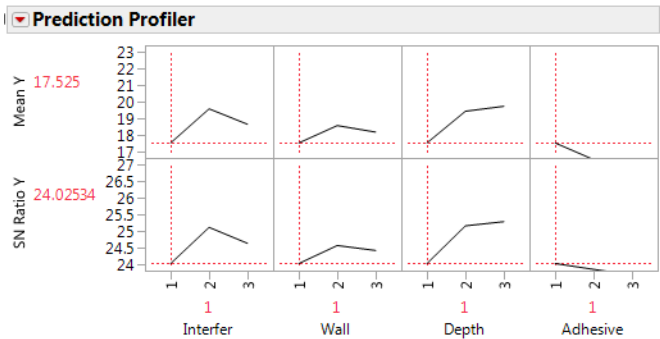
Figure 14.5 Fit Model Launch Window for Taguchi Data



- The script launches the Fit Model window. The model includes the main effects of the four signal factors to model the mean (Mean Y) and signal-to-noise ratio (SN Ratio Y) responses.
- 2. Click **Run**.

The Prediction Profiler at the bottom of the report is a quick way to find settings that give the highest signal-to-noise ratio for this experiment and the mean value at that maximum.

Figure 14.6 The Prediction Profiler

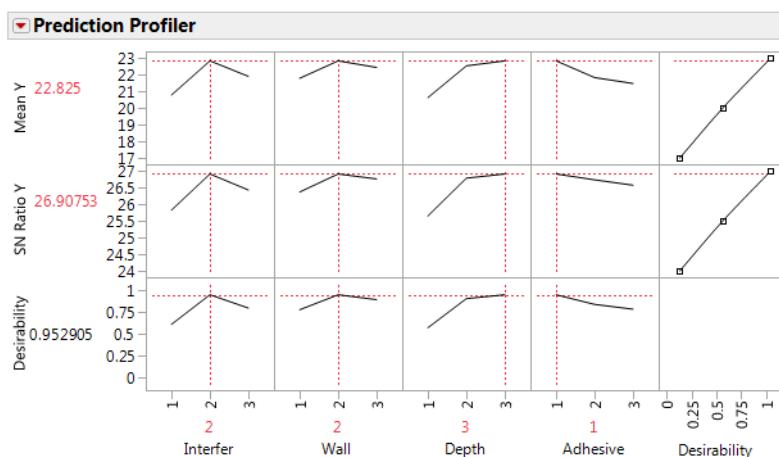


3. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Desirability Functions**.

This adds a row of traces and a column of function settings to the profiler, as shown in Figure 14.7. The default desirability functions are set to larger-is-better, which is what you want in this experiment. See the Profiler chapter in *Profilers* for more information about desirability functions in the prediction profiler.

4. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

Figure 14.7 Best Factor Settings for Byrne Taguchi Data

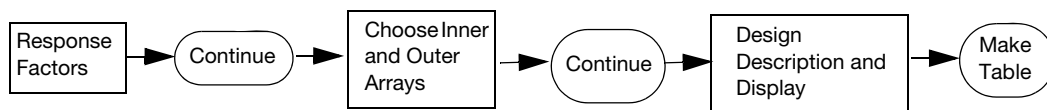


In this example, the optimal settings for Interfer and Wall are 2, Depth is 3, and Adhesive is 1. These settings result in a predicted Mean value of 22.8 and an SN Ratio of 26.9.

Taguchi Design Window

The Taguchi Design window updates as you work through the design steps. See [“The DOE Workflow: Describe, Specify, Design”](#) on page 57. The outlines, separated by buttons that update the outlines, follow the flow in Figure 14.8.

Figure 14.8 Taguchi Design Flow



Response

Use the Response outline to specify a response.

Figure 14.9 Response Outline

Response				
Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Larger Is Better	.	.	.

The Response outline contains the following columns:

Response Name The name of the response (the default is Y). To change this name, double-click it and enter a different name.

Goal, Lower Limit, Upper Limit The Goal determines the form of the signal-to-noise function (Table 14.1 on page 423). Select one of the following Goals: Larger Is Better, Nominal is Best, Smaller is Better, or None. When you create your design, JMP saves a formula for the SN Ratio to the data table that reflects your selected goal. To change a goal, click it and enter a different goal. Use the Lower and Upper Limits if you have limits on your response.

Importance When you have several responses, the Importance values that you specify are used to compute an overall desirability function. These values are treated as weights for the responses. In a Taguchi Design there is only one response, the Importance is set to 1 by default. No value is needed for the Importance.

Factors

Use the Factors outline to specify factors.

Signal Specify two or more 2- or 3-level signal factors. Signal factors are system control inputs. These are factors that you can control in production.

Noise Specify one or more noise factors. Noise factors are variables that are difficult or expensive to control in production. However, you must be able to control noise factors during the experiment. Click Noise to add a noise factor.

Remove Removes the selected factors.

Name The name of the factor. When added, a factor is given a default name of X1, X2, and so on. To change this name, double-click it and enter a different name.

Role Specifies the Design Role (Signal or Noise) of the factor. This is set when you add the factor.

Values The experimental settings for the factors. To insert Values, click the default values and enter the desired values.

Figure 14.10 Factors Outline

Factors

Signal Noise Remove

Name	Role	Values
X1	Signal	1 2
X2	Signal	1 2

Choose Inner and Outer Array Designs

The Choose Inner and Outer Array Design options appear after you click Continue in the Specify Factors panel. Based on your signal factors, you get one or more choices for orthogonal designs for the inner array. The outer array choices for orthogonal “L” designs depend on the number and type of factors in your design.

Display Coded Design

After you select a design type for the inner and outer arrays, the Coded Design is shown when your inner array has two-level factors.

Figure 14.11 Coding for Eight Factor L12 Design

[illegible]

The Coded Design shows the pattern of high and low values for the factors in each run.

Make the Design Table

When you click **Make Table**, a table similar to the one in Figure 14.12 appears. In the data table, each row represents an inner array run. In the values for the **Pattern** variable, plus signs designate high levels, and minus signs represent low levels. The table also contains a column for each factor and a column for each pattern that appears in the **Pattern** column. The columns for the patterns contain the responses to the experiment. If you selected **Simulate Responses** from the Taguchi Design red triangle menu, the pattern columns are filled with simulated response values.

Figure 14.12 Taguchi Design Table for an Eight Factor L12 Design

	X1	X2	X3	X4	X5	X6	X7	X8	Pattern	run 1	run 2	Mean	SN Ratio
1	1	1	1	1	1	1	1	1	-----	*	*	*	*
2	1	1	1	1	1	2	2	2	-----	*	*	*	*
3	1	1	2	2	2	1	1	1	-----	*	*	*	*
4	1	2	1	2	2	1	2	2	-----	*	*	*	*
5	1	2	2	1	2	2	1	2	-----	*	*	*	*
6	1	2	2	2	1	2	2	1	-----	*	*	*	*
7	2	1	2	2	1	1	2	2	-----	*	*	*	*
8	2	1	2	1	2	2	2	1	-----	*	*	*	*
9	2	1	1	2	2	2	1	2	-----	*	*	*	*
10	2	2	2	1	1	1	1	2	-----	*	*	*	*
11	2	2	1	2	1	2	1	1	-----	*	*	*	*
12	2	2	1	1	2	1	2	1	-----	*	*	*	*

Taguchi Design Options

Save Responses Saves the information in the Responses panel to a new data table. You can then quickly load the responses and their associated information into most DOE windows. This option is helpful if you anticipate reusing the responses.

Load Responses Loads responses that you saved using the Save Responses option.

Save Factors Saves the information in the Factors panel to a new data table. Each factor's column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Note: It is possible to create a factors table by entering data into an empty table, but remember to assign each column an appropriate Design Role. Do this by right-clicking on the column name in the data grid and selecting **Column Properties > Design Role**. In the Design Role area, select the appropriate role.

Load Factors Loads factors that you saved using the Save Factors option.

Save Constraints (Unavailable for some platforms) Saves factor constraints that you defined in the Define Factor Constraints or Linear Constraints outline into a data table, with a

column for each constraint. You can then quickly load the constraints into most DOE windows.

In the constraint table, the first rows contain the coefficients for each factor. The last row contains the inequality bound. Each constraint's column contains a column property called **ConstraintState** that identifies the constraint as a "less than" or a "greater than" constraint. See "[ConstraintState](#)" on page 766 in the "Column Properties" appendix.

Load Constraints (Unavailable for some platforms) Loads factor constraints that you saved using the Save Constraints option.

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction
- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.
- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click Apply, a formula for the simulated response values is saved in a new column called <Y> Simulated, where Y is the name of the response. Clicking Apply again updates the formula and values in <Y> Simulated.

For additional details, see "[Simulate Responses](#)" on page 110 in the "Custom Designs" chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Chapter 15

Evaluate Designs

Explore Properties of Your Design

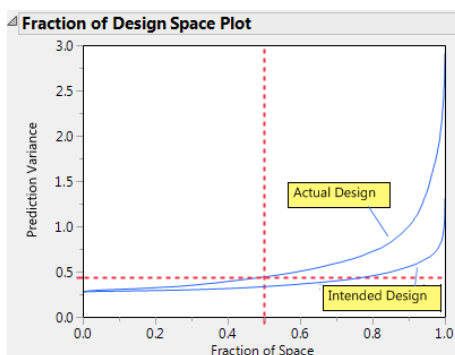
Using the Evaluate Design platform, you can:

- See the strengths and limitations of your existing experimental design.
- Determine your design's ability to detect effects associated with meaningful changes in the response.
- Address prediction variance and the precision of your estimates.
- Gain insight on aliasing.
- Obtain efficiency measures.

The Evaluate Design platform generates the results that appear in the Design Evaluation outlines provided by several DOE platforms. Diagnostics provided by both Evaluate Design and the Design Evaluation outlines include the following:

- power analysis
- a prediction variance profiler and surface plot
- a fraction of design space plot, showing how much of the design space has prediction variance above a given value
- estimation efficiency measures for parameter estimates
- the alias matrix, showing the bias structure for model effects
- a color map showing absolute correlations among effects
- design efficiency values

Figure 15.1 Comparison of Two Fraction of Design Space Plots



Contents

Overview of the Evaluate Design Platform	435
Example of Evaluate Design	436
Assessing the Impact of Lost Runs	436
Evaluating Power Relative to a Specified Model	443
Evaluate Design Launch Window	445
Evaluate Design Window	445
Factors	447
Model	447
Alias Terms	448
Design	449
Design Evaluation	449
Power Analysis	449
Prediction Variance Profile	456
Fraction of Design Space Plot	457
Prediction Variance Surface	458
Estimation Efficiency	460
Alias Matrix	462
Color Map on Correlations	463
Design Diagnostics	465
Evaluate Design Options	467

Overview of the Evaluate Design Platform

The Evaluate Design platform provides powerful tools that enable you to assess the properties of your design, whether it is created by JMP or another tool. You can use the platform before conducting an experiment to choose from several designs. You can also assess the impact of incorrect settings or lost runs in a design that you have conducted. You can modify the terms in your assumed model to see the impact of estimating a modified model. You can also modify the terms that appear in the Alias Terms outline to see the impact on the Alias Matrix.

You start by entering information about the design in the launch window. Then you can modify the assumed model and specify which effects not included in the model are of potential interest. Based on your specifications, the Design Evaluation platform then provides a number of ways to evaluate the properties of the generated design:

Power Analysis Enables you to explore your ability to detect effects of given sizes.

Prediction Variance Profile Shows the prediction variance over the range of factor settings.

Fraction of Design Space Plot Shows how much of the model prediction variance lies below or above a given value.

Prediction Variance Surface Shows a surface plot of the prediction variance for any two continuous factors.

Estimation Efficiency For each parameter, gives the fractional increase in the length of a confidence interval compared to that of an ideal (orthogonal) design, which might not exist. Also gives the relative standard error of the parameters.

Alias Matrix Gives coefficients that indicate the degree to which the model parameters are biased by effects that are potentially active, but not in the model.

Color Map on Correlations Shows the absolute correlations between effects on a plot using an intensity scale.

Design Diagnostics Gives efficiency measures for your design.

Note: In several DOE platforms, when you construct a design, a Design Evaluation outline appears. This outline shows results provided by the Evaluate Design platform. The platforms that provide a Design Evaluation outline are: Custom Design, Definitive Screening Design, Screening Design, Response Surface Design, and Mixture Design with Optimal design type.

Example of Evaluate Design

- “Assessing the Impact of Lost Runs”
- “Evaluating Power Relative to a Specified Model”

Assessing the Impact of Lost Runs

An experiment was conducted to explore the effect of three factors (Silica, Sulfur, and Silane) on tennis ball bounciness (Stretch). The goal of the experiment is to develop a predictive model for Stretch. A 15-run Box-Behnken design was selected using the Response Surface Design platform. After the experiment, the researcher learned that the two runs where Silica = 0.7 and Silane = 50 were not processed correctly. These runs could not be included in the analysis of the data.

Use the Evaluate Design platform to assess the impact of not including those two runs. Obtain diagnostics for the intended 15-run design and compare these to the actual 13-run design that is missing the two runs.

Construct the Intended and Actual Designs

Intended Design

1. Select **Help > Sample Data Library** and open Design Experiment/Bounce Data.jmp.
2. Select **DOE > Design Diagnostics > Evaluate Design**.
3. Select Silica, Sulfur, and Silane and click **X, Factor**.

You can add Stretch as **Y, Response** if you wish. But specifying the response has no effect on the properties of the design.

4. Click **OK**.

Leave your Evaluate Design window for the intended design open.

Tip: Place the Evaluate Design window for the intended design in the left area of your screen. After the next steps, you will place the corresponding window for the actual design to its right.

Actual Design with Missing Runs

In this section, you will exclude the two runs where Silica = 0.7 and Silane = 50. These are rows 3 and 7 in the data table.

1. In Bounce Data.jmp, select rows 3 and 7, right-click in the highlighted area, and select **Hide and Exclude**.
2. Select **DOE > Design Diagnostics > Evaluate Design**.
3. Click **Recall**.
4. Click **OK**.

Leave your Evaluate Design window for the actual design open.

Tip: Place the Evaluate Design window for the actual design to the *right* of the Evaluate Design window for the intended design to facilitate comparing the two designs.

Comparison of Intended and Actual Designs

You can now compare the two designs using these methods:

- [“Power Analysis”](#) on page 437
- [“Prediction Variance Profile”](#) on page 438
- [“Fraction of Design Space Plot”](#) on page 440
- [“Estimation Efficiency”](#) on page 441
- [“Color Map on Correlations”](#) on page 442
- [“Design Diagnostics”](#) on page 443

Power Analysis

In each window, do the following:

1. Open the Power Analysis outline.

The outline shows default values of 1 for all Anticipated Coefficients. These values correspond to detecting a change in the anticipated response of 2 units across the levels of main effect terms, assuming that the interaction and quadratic terms are not active.

The power calculations assume an error term (Anticipated RMSE) of 1. From previous studies, you believe that the RMSE is approximately 2.

2. Type 2 next to **Anticipated RMSE**.

When you click outside the text box, the power values are updated.

You are interested in detecting differences in the anticipated response that are on the order of 6 units across the levels of main effects, assuming that interaction and quadratic terms are not active. To set these uniformly, use a red triangle option.

3. Click the Evaluate Design red triangle and select **Advanced Options > Set Delta for Power**.
4. Type 6 as your value for delta.
5. Click **OK**.

Figure 15.2 shows both outlines, with the Design and Anticipated Responses outline closed.

Figure 15.2 Power Analysis Outlines, Intended Design (Left) and Actual Design (Right)

Power Analysis			Power Analysis		
Significance Level	0.05		Significance Level	0.05	
Anticipated RMSE	2		Anticipated RMSE	2	
Term	Anticipated Coefficient	Power	Term	Anticipated Coefficient	Power
Intercept	3	0.553	Intercept	3	0.432
Silica	3	0.919	Silica	3	0.533
Sulfur	3	0.919	Sulfur	3	0.533
Silane	3	0.919	Silane	3	0.798
Silica*Sulfur	3	0.672	Silica*Sulfur	3	0.231
Silica*Silane	3	0.672	Silica*Silane	3	0.533
Sulfur*Silane	3	0.672	Sulfur*Silane	3	0.533
Silica*Silica	3	0.639	Silica*Silica	3	0.379
Sulfur*Sulfur	3	0.639	Sulfur*Sulfur	3	0.379
Silane*Silane	3	0.639	Silane*Silane	3	0.379
Apply Changes to Anticipated Coefficients			Apply Changes to Anticipated Coefficients		
Design and Anticipated Responses			Design and Anticipated Responses		

The power values for the actual design are uniformly smaller than for the intended design. For Silica and Sulfur, the power of the tests in the intended design is almost twice the power in the actual design. For the Silica*Sulfur interaction, the power of the test in the actual design is 0.231, compared to 0.672 in the intended design. The actual design results in substantial loss of power in comparison with the intended design.

Prediction Variance Profile

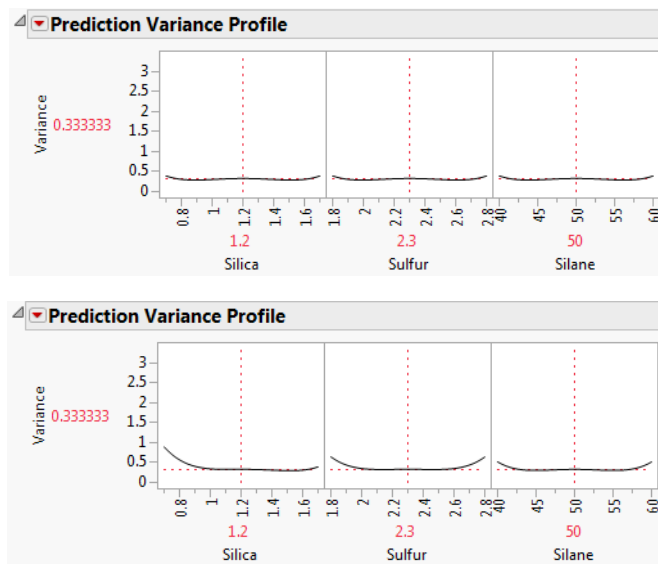
1. In each window, open the Prediction Variance Profile outline.
2. In the window for the actual design, place your cursor on the scale for the vertical axis. When your cursor becomes a hand, right-click. **Select Edit > Copy Axis Settings**.

This action creates a script containing the axis settings. Next, apply these axis settings to the Prediction Variance Profile plot for the intended design.

3. In the Evaluate Design window for the intended design, locate the Prediction Variance Profile outline. Place your cursor on the scale for the vertical axis and when your cursor becomes a hand, right-click. **Select Edit > Paste Axis Settings**.

The plots are shown in Figure 15.5, with the plot for the intended design at the top and for the actual design at the bottom.

Figure 15.3 Prediction Variance Profile, Intended Design (Top) and Actual Design (Bottom)



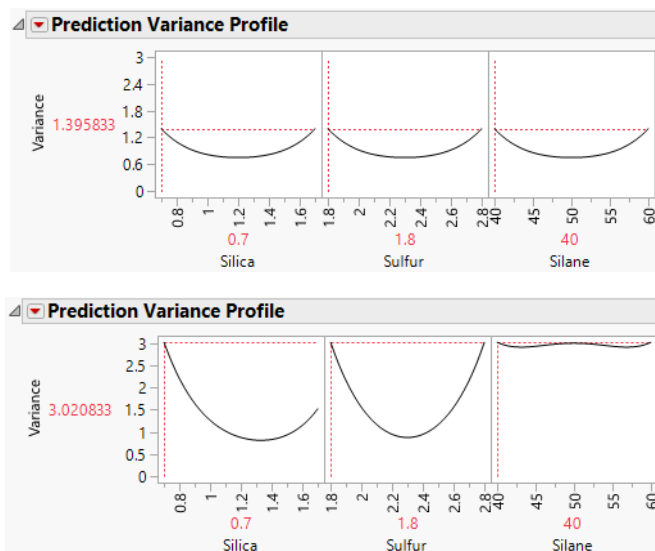
The Prediction Variance Profile plots are profiler views of the relative prediction variance. You can explore the relative prediction variance in various regions of design space.

Both plots show the same relative prediction variance in the center of the design space. However, the variance for points near the edges of the design space appears greater than for the same points in the intended design. Explore this phenomenon by moving all three vertical lines to points near the edges of the factor settings.

4. In both windows, click the Prediction Variance Profile red triangle and select **Maximize Variance**.

Figure 15.4 shows the maximum relative prediction variance for the intended and actual designs.

Figure 15.4 Prediction Variance Profile Maximized, Intended Design (Top) and Actual Design (Bottom)



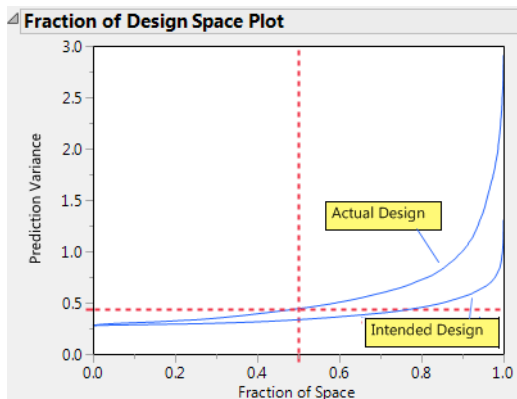
For both designs, the profilers identify the same point as one of the design points where the maximum prediction variance occurs: Silica=0.7, Sulfur=1.8, and Silane=40. The maximum prediction variance is 1.396 for the intended design, and 3.021 for the actual design. Note that there are other points where the prediction variance is maximized. The larger maximum prediction variance for the actual design means that predictions in parts of the design space are less accurate than they would have been had the intended design been conducted.

Fraction of Design Space Plot

1. In each window, open the Fraction of Design Space Plot outline.
2. In the window for the intended design, right-click in the plot and select **Edit > Copy Frame Contents**.
3. In the window for the actual design, locate the Fraction of Design Space Plot outline.
4. Right-click in the plot and select **Edit > Paste Frame Contents**

Figure 15.5 shows the plot with annotations. Each Fraction of Design Space Plot shows the proportion of the design space for which the relative prediction variance falls below a specific value.

Figure 15.5 Fraction of Design Space Plots



The relative prediction variance for the actual design is greater than that of the intended design over the entire design space. The discrepancy increases with larger design space coverage.

Estimation Efficiency

In each window, open the Estimation Efficiency outline.

Figure 15.6 Estimation Efficiency Outlines, Intended Design (Left) and Actual Design (Right)

Estimation Efficiency			Estimation Efficiency		
Term	Fractional Increase in CI Length	Relative Std Error of Estimate	Term	Fractional Increase in CI Length	Relative Std Error of Estimate
Intercept	1.236	0.577	Intercept	1.082	0.577
Silica	0.369	0.354	Silica	0.803	0.5
Sulfur	0.369	0.354	Sulfur	0.803	0.5
Silane	0.369	0.354	Silane	0.275	0.354
Silica*Sulfur	0.936	0.5	Silica*Sulfur	2.122	0.866
Silica*Silane	0.936	0.5	Silica*Silane	0.803	0.5
Sulfur*Silane	0.936	0.5	Sulfur*Silane	0.803	0.5
Silica*Silica	1.016	0.52	Silica*Silica	1.268	0.629
Sulfur*Sulfur	1.016	0.52	Sulfur*Sulfur	1.268	0.629
Silane*Silane	1.016	0.52	Silane*Silane	1.268	0.629

In the actual design (right), the relative standard errors for all parameters either exceed or equal the standard errors for the intended design (left). For all except three of the non-intercept parameters, the relative standard errors in the actual design exceed those in the intended design.

The Fractional Increase in CI Length compares the length of a parameter's confidence interval as given by the current design to the length of such an interval given by an ideal design of the same run size. The length of the confidence interval, and consequently the Fractional Increase in CI Length, is affected by the number of runs. See [“Fractional Increase in CI Length”](#) on page 460. Despite the reduction in run size, for the actual design, the terms Silane, Silica*Silane, and Sulfur*Silane have a smaller increase than for the intended design. This is because the two runs that were removed to define the actual design had Silane set to its center point. By removing these runs, the widths of the confidence intervals for these parameters more closely resemble those of an ideal orthogonal design, which has no center points.

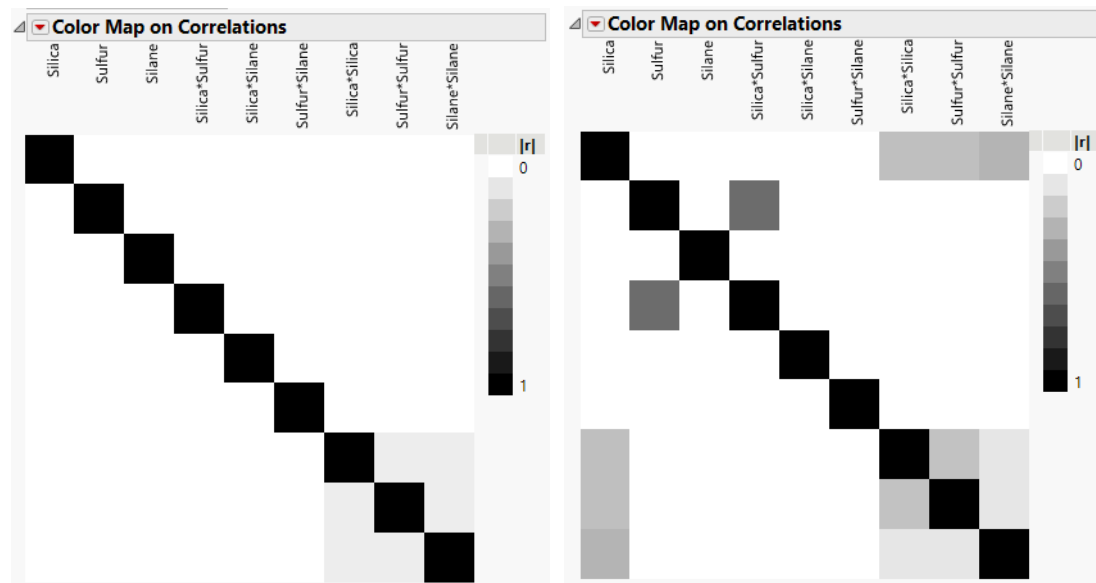
Color Map on Correlations

In each report, do the following:

1. Open the **Color Map on Correlations** outline.

The two color maps show the effects in the Model outline. Each plot shows the absolute correlations between effects colored using the JMP default white to black intensity scale. Ideally, you would like zero or very small correlations between effects.

Figure 15.7 Color Map on Correlations, Intended Design (Left) and Actual Design (Right)



The absolute values of the correlations range from 0 (white) to 1 (black). Hover over a cell to see the value of the absolute correlation. The color map for the actual design shows more absolute correlations that are large than does the color map for the intended design. For example, the correlation between Sulfur and Silica*Sulfur is $< .0001$ for the intended design, and 0.5774 for the actual design.

Design Diagnostics

In each report, open the Design Diagnostics outline.

Figure 15.8 Design Diagnostics, Intended Design (Left) and Actual Design (Right)

Design Diagnostics		Design Diagnostics	
D Efficiency	36.6429	D Efficiency	36.6429
G Efficiency	69.10947	G Efficiency	69.10947
A Efficiency	29.3578	A Efficiency	29.3578
Average Variance of Prediction	0.384722	Average Variance of Prediction	0.384722
Design Creation Time (seconds)	0	Design Creation Time (seconds)	0

The intended design (left) has higher efficiency values and a lower average prediction variance than the actual design (right). The results of the Design Evaluation analysis indicate that the two lost runs have had a negative impact on the design.

Note that both the number of runs and the model matrix factor into the calculation of efficiency measures. In particular, the *D*-, *G*-, and *A*- efficiencies are calculated relative to the ideal design for the run size of the given design. It is not necessarily true that larger designs are more efficient than smaller designs. However, for a given number of factors, larger designs tend to have smaller Average Variance of Prediction values than do smaller designs. For more information on how efficiency measures are defined, see [“Design Diagnostics”](#) on page 465.

Evaluating Power Relative to a Specified Model

For this example, you have constructed a definitive screening design to determine which of six factors have an effect on the yield of an extraction process. The data are given in the Extraction Data.jmp sample data table, located in the Design Experiment folder. Because the design is a *definitive* screening design, each factor has three levels. See the [“Definitive Screening Designs”](#) chapter on page 237.

You are interested in the power of tests to detect a strong quadratic effect. You consider a strong effect to be one whose magnitude is at least three times as large as the error variation.

Although the experiment studies six factors, effect sparsity suggests that only a small subset of factors is active. Consequently, you feel comfortable investigating power in a model based on a smaller number of factors. Also, past studies on a related process provide strong evidence to suggest that three of the factors, Propanol, Butanol, and pH, have negligible main effects, do not interact with other factors, and do not have quadratic effects. This leads you to believe that the likely model contains main, interaction, and quadratic effects only for Methanol, Ethanol, and Time. You decide to investigate power in the context of a three-factor response surface model.

Use the Evaluate Design platform to determine the power of your design to detect strong quadratic effects for Methanol, Ethanol, or Time.

1. Select **Help > Sample Data Library** and open Design Experiment/Extraction Data.jmp.
2. Select **DOE > Design Diagnostics > Evaluate Design**.
3. Select Methanol, Ethanol, and Time and click **X, Factor**.

You can add Yield as **Y, Response** if you wish. But specifying the response has no effect on the properties of the design.

4. Click **OK**.
5. In the Model outline, click **RSM**.

This adds the interaction and quadratic terms for the three factors.

6. Open the Power Analysis outline.

Note that the Anticipated RMSE is set to 1 by default. Although you have an estimate of the RMSE from past studies, you need not enter it. This is because the magnitude of the effect of interest is three times the error variation.

7. Under **Anticipated Coefficient**, type 3 next to Methanol*Methanol, Ethanol*Ethanol, and Time*Time.
8. Click **Apply Changes to Anticipated Coefficients**.

Figure 15.9 Power Analysis Outline after Applying Changes to Coefficients

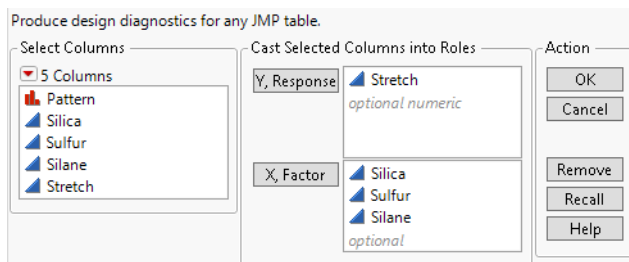
Power Analysis		
Significance Level	0.05	
Anticipated RMSE	1	
Term	Anticipated Coefficient	Power
Intercept	1	0.111
Methanol	1	0.573
Ethanol	1	0.573
Time	1	0.573
Methanol*Methanol	3	0.737
Methanol*Ethanol	-1	0.331
Ethanol*Ethanol	3	0.737
Methanol*Time	-1	0.331
Ethanol*Time	1	0.331
Time*Time	3	0.737
Apply Changes to Anticipated Coefficients		

The power of detecting a quadratic effect whose magnitude is three times the error variation is 0.737. This assumes a final model that is a response surface in three factors. It also assumes a 0.05 significance level for the test.

Evaluate Design Launch Window

To launch the Evaluate Design platform, open the data table of interest and select **DOE > Design Diagnostics > Evaluate Design**. The example in Figure 15.10 uses the Bounce Data.jmp sample data table, located in the Design Experiment folder.

Figure 15.10 Evaluate Design Launch Window



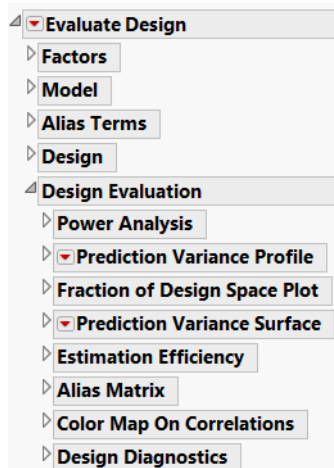
The launch window contains the following buttons:

- Y, Response** Enter the response column or columns. Entering a response is optional. Response values are not used in evaluating the design. Responses must be numeric.
- X, Factor** Enter the factor columns. Factors can be of any Data Type or Modeling Type.

Evaluate Design Window

The Evaluate Design window consists of two parts. See Figure 15.11, where all outline nodes are closed.

- The Factors, Model, Alias Terms, and Design outlines define the model and design.
- The Design Evaluation outline provides results that describe the properties of your design.

Figure 15.11 Evaluate Design Window Showing All Possible Outlines

The Factors, Model, Alias Terms, and Design outlines contain information that you enter about the factors, assumed model, potentially aliased effects of interest, and the actual design. JMP populates these outlines using your selections in the launch window and the design table. However, you can modify the effects in the Model and Alias Terms outlines. Once you have made your specifications, the Design Evaluation outlines are updated. You can open these outlines to see reports or control windows that provide information about your design.

- “Factors”
- “Model”
- “Alias Terms”
- “Design”
- “Design Evaluation”
- “Power Analysis”
- “Prediction Variance Profile”
- “Fraction of Design Space Plot”
- “Prediction Variance Surface”
- “Estimation Efficiency”
- “Alias Matrix”
- “Color Map on Correlations”
- “Design Diagnostics”

Factors

The factors outline lists the factors entered in the launch window. You can select factors to construct effects in the Model outline.

Figure 15.12 Factors Outline

Factors	
Name	Role
▲ Silica	Continuous
▲ Sulfur	Continuous
▲ Silane	Continuous

Model

What appears in the Model outline depends upon whether the associated data table contains a Model or Fit Model script:

- If there is a script, the Model outline contains the estimable effects specified in the script. Effects are entered into the model in the order specified in the script. If the model script contains more effects than are estimable, those that are not estimable are not included in the model for evaluation.
- If there is no script, the Model outline contains only main effects.

Figure 15.13 shows the Model outline for the **Bounce Data.jmp** data table, found in the Design Experiment folder. The Model script in the data table contains response surface effects for the three factors **Silica**, **Silane**, and **Sulfur**. Consequently, the Model outline contains the main effects, two-way interactions, and quadratic effects for these three factors.

Figure 15.13 Model Outline for Bounce Data.jmp

Model	
Main Effects	Interactions ▼
RSM	Cross
Powers ▼	Remove Term
Intercept	
Silica	
Silane	
Sulfur	
Silica*Silane	
Silica*Sulfur	
Sulfur*Silane	
Silica*Silica	
Silane*Silane	
Sulfur*Sulfur	

You can add effects to the Model outline using the following buttons:

Main Effects Adds main effects for all factors in the model.

Interactions Adds interaction effects. If no factors are selected in the Factors outline, select 2nd, 3rd, 4th, or 5th to add all appropriate interactions up to that order. Add interactions up to a given order for specific factors by selecting the factor names in the Factors outline, selecting Interactions, and then specifying the appropriate order. Interactions between non-mixture and mixture factors, and interactions with blocking and constant factors, are not added.

RSM Adds interaction and quadratic terms up to the second order (response surface model terms) for continuous factors. Categorical factors are not included in RSM terms. Main effects for non-mixture factors that interact with all the mixture factors are removed.

Cross Adds specific interaction terms. Select factor names in the Factors outline and effect names in the Model outline. Click Cross to add the crossed terms to the Model outline.

Powers Adds polynomial terms. If no factor names are selected in the Factors outline, adds polynomial terms for all continuous factors. If factor names are selected in the Factors outline, adds polynomial terms for only those factors. Select 2nd, 3rd, 4th, or 5th to add polynomial terms of that order.

Scheffé Cubic Adds Scheffé cubic terms for all mixture factors. These terms are used to specify a mixture model with third-degree polynomial terms.

Remove Term Removes selected effects.

Alias Terms

It is possible that effects *not* included in your assumed model are active. In the Alias Terms outline, list potentially active effects that are not in your assumed model but might bias the estimates of model terms. The Alias Matrix entries represent the degree of bias imparted to model parameters by the effects that you specified in the Alias Terms outline. See [“The Alias Matrix”](#) on page 771 in the “Technical Details” appendix.

By default, the Alias Terms outline includes all two-way interaction effects that are not in your Model outline (with the exception of terms involving blocking factors). Add or remove terms using the buttons. For a description of how to use these buttons to add effects to the Alias Terms table, see [“Model”](#) on page 447.

In the Evaluate Design platform, the Alias Matrix outline is immediately updated to reflect changes to Alias Matrix effects. In the Custom Design platform, you must click Make Design after modifying the effects in the Alias Terms outline. Within other DOE platforms that construct designs, there is no Alias Terms outline. However, the Alias Matrix outline, containing appropriate effects, appears under Design Evaluation after you construct the design.

Design

The Design outline shows the design runs for the factors that you have specified in the launch window. You can easily view the design as you explore its properties in the Design Evaluation outline.

Design Evaluation

Design Evaluation within the Evaluate Design platform is based on your design and the specifications that you make in the Model and Alias Terms outlines. Several DOE Design platforms provide a Design Evaluation outline: Custom, Definitive Screening, Screening, Response Surface, and Mixture with Optimal design type. Design Evaluation within these platforms is based on the design that you construct.

Power Analysis

The Power Analysis outline calculates the power of tests for the parameters in your model. Power is the probability of detecting an active effect of a given size. The Power Analysis outline helps you evaluate the ability of your design to detect effects of practical importance. The higher your power, the more likely you are to detect significant effects assuming your coefficient and RMSE assumptions are correct. Power depends on the number of runs, the significance level, and the estimated error variation. In particular, you can determine if additional runs are necessary.

This section covers the following topics:

- [“Power Analysis Overview”](#) on page 449
- [“Power Analysis Details”](#) on page 451
- [“Tests for Individual Parameters”](#) on page 451
- [“Tests for Categorical Effects with More Than Two Levels”](#) on page 452
- [“Design and Anticipated Responses Outline”](#) on page 453
- [“Power Analysis for Coffee Experiment”](#) on page 454

Power Analysis Overview

Power is calculated for the effects listed in the Model outline. These include continuous, discrete numeric, categorical, blocking, mixture, and covariate factors. The tests are for individual model parameters and for whole effects. For more information on how power is calculated, see [“Power Calculations”](#) on page 772 in the “Technical Details” appendix.

Power is the probability of rejecting the null hypothesis of no effect at specified values of the model parameters. In practice, your interest is not in the values of the model parameters, but in detecting differences in the mean response of practical importance. In the Power Analysis outline, you can compute Anticipated Responses for specified values of the Anticipated Coefficients. This helps you to determine the coefficient values associated with the differences you want to detect in the mean response.

Figure 15.14 shows the Power Analysis outline for the design in the Coffee Data.jmp sample data table, found in the Design Experiment folder. The model specified in the Model script is a main effects only model.

Figure 15.14 Power Analysis for Coffee Data.jmp

Power Analysis

Significance Level

Anticipated RMSE

Term	Anticipated Coefficient	Power
Intercept	1	0.789
Grind	-1	0.789
Temperature	1	0.789
Time	1	0.789
Charge	1	0.789
Station 1	1	0.507
Station 2	1	0.507

Effect **Power**

Station 0.888

Design and Anticipated Responses

Anticipated Response	Grind	Temperature	Time	Charge	Station
6 Medium		205	4	2.4	1
-2 Coarse		195	3	1.6	1
2 Medium		205	3	1.6	1
2 Coarse		195	4	2.4	1
2 Coarse		205	4	1.6	2
4 Medium		195	4	2.4	2
0 Medium		195	3	1.6	2
2 Coarse		205	3	2.4	2
-1 Coarse		205	4	1.6	3
-3 Coarse		195	3	2.4	3
-1 Medium		195	4	1.6	3
1 Medium		205	3	2.4	3

In the Power Analysis outline, you can:

- Specify coefficient values that reflect differences that you want to detect. You enter these as Anticipated Coefficients in the top part of the outline.
- Specify anticipated response values and apply these to determine the corresponding Anticipated Coefficients. You specify Anticipated Responses in the Design and Anticipated Responses panel.

Power Analysis Details

Specify values for the Significance Level and Anticipated RMSE. These are used to calculate the power of the tests for the model parameters.

Significance Level The probability of rejecting the hypothesis of no effect, if it is true. The power calculations update immediately when you enter a value.

Anticipated RMSE An estimate of the square root of the error variation. The power calculations update immediately when you enter a value.

The top portion of the Power Analysis report opens with default values for the Anticipated Coefficients (Figure 15.14). The default values are based on Delta. See [“Advanced Options > Set Delta for Power”](#) on page 467.

Note: If the design is supersaturated, meaning that the number of parameters to be estimated exceeds the number of runs, the anticipated coefficients are set to 0.

Figure 15.15 shows the top portion of the Power Analysis report where values have been specified for the Anticipated Coefficients. These values reflect the differences you want to detect.

Figure 15.15 Possible Specification of Anticipated Coefficients for Coffee Data.jmp

Power Analysis		
Significance Level	0.05	
Anticipated RMSE	0.1	
Term	Anticipated Coefficient	Power
Intercept	1.4	1
Grind	0.05	0.291
Temperature	0.05	0.291
Time	0.05	0.291
Charge	0.05	0.291
Station 1	0.1	0.507
Station 2	0.1	0.507
Apply Changes to Anticipated Coefficients		
Effect	Power	
Station	0.888	

Tests for Individual Parameters

The Term column contains a list of model terms. For each term, the Anticipated Coefficient column contains a value for that term. The value in the Power column is the power of a test that the coefficient for the term is 0 if the true value of the coefficient is given by the Anticipated Coefficient.

Term The model term associated with the coefficient being tested.

Note: The order in which model terms appear in the Power Analysis report may not be identical to their order in the Parameter Estimates report obtained using Standard Least Squares. This difference can only occur when the model contains an interaction with more than one degree of freedom.

Anticipated Coefficient A value for the coefficient associated with the model term. This value is used in the calculations for Power. These values are also used to calculate the Anticipated Response column in the Design and Anticipated Responses outline. When you set a new value in the Anticipated Coefficient column, click **Apply Changes to Anticipated Coefficients** to update the Power and Anticipated Response columns.

Note: The anticipated coefficients have default values of 1 for continuous effects. They have alternating values of 1 and -1 for categorical effects. You can specify a value for Delta by selecting **Advanced Options > Set Delta for Power** from the red triangle menu. If you change the value of Delta, the values of the anticipated coefficients are updated so that their absolute values are one-half of Delta. See [“Advanced Options > Set Delta for Power”](#) on page 467.

Power Probability of rejecting the null hypothesis of no effect when the true coefficient value is given by the specified Anticipated Coefficient. For a coefficient associated with a numeric factor, the change in the mean response (based on the model) is twice the coefficient value. For a coefficient associated with a categorical factor, the change in the mean response (based on the model) across the levels of the factor equals twice the absolute value of the anticipated coefficient.

Calculations use the specified Significance Level and Anticipated RMSE. For more information about the power calculation, see [“Power for a Single Parameter”](#) on page 772 in the “Technical Details” appendix.

Apply Changes to Anticipated Coefficients When you set a new value in the Anticipated Coefficient column, click **Apply Changes to Anticipated Coefficients** to update the Power and Anticipated Response columns.

Tests for Categorical Effects with More Than Two Levels

If your model contains a categorical effect with more than two levels, then the following columns appear below the Apply Changes to Anticipated Coefficients button:

Effect The categorical effect.

Power The power calculation for a test of no effect. The null hypothesis for the test is that all model parameters corresponding to the effect are zero. The difference to be detected is defined by the values in the Anticipated Coefficient column that correspond to the model

terms for the effect. The power calculation reflects the differences in response means determined by the anticipated coefficients.

Calculations use the specified Significance Level and Anticipated RMSE. For more information about the power calculation, see [“Power for a Categorical Effect”](#) on page 774 in the “Technical Details” appendix.

Design and Anticipated Responses Outline

The Design and Anticipated Responses outline shows the design preceded by an Anticipated Response column. Each entry in the first column is the Anticipated Response corresponding to the design settings. The Anticipated Response is calculated using the Anticipated Coefficients.

Figure 15.16 shows the Design and Anticipated Responses outline corresponding to the specification of Anticipated Coefficients given in Figure 15.15.

Figure 15.16 Anticipated Responses for Coffee Data.jmp

Design and Anticipated Responses						
Anticipated Response	Grind	Temperature	Time	Charge	Station	
1.6	Medium	205	4	2.4	1	
1.4	Coarse	195	3	1.6	1	
1.4	Medium	205	3	1.6	1	
1.6	Coarse	195	4	2.4	1	
1.6	Coarse	205	4	1.6	2	
1.5	Medium	195	4	2.4	2	
1.3	Medium	195	3	1.6	2	
1.6	Coarse	205	3	2.4	2	
1.3	Coarse	205	4	1.6	3	
1.2	Coarse	195	3	2.4	3	
1.1	Medium	195	4	1.6	3	
1.2	Medium	205	3	2.4	3	

Apply Changes to Anticipated Responses

In the Anticipated Response column, you can specify a value for each setting of the factors. These values reflect the differences you want to detect.

Click **Apply Changes to Anticipate Responses** to update both the Anticipated Coefficient and Power columns.

Anticipated Response The response value obtained using the Anticipated Coefficient values as coefficients in the model. When the outline first appears, the calculation of Anticipated Response values is based on the default values in the Anticipated Coefficient column. When you set new values in the Anticipated Response column, click **Apply Changes to Anticipated Responses** to update the Anticipated Coefficient and Power columns.

Design The columns to the right of the Anticipated Response column show the factor settings for all runs in your design.

Apply Changes to Anticipated Responses When you set new values in the Anticipated Response column, click **Apply Changes to Anticipated Responses** to update the Anticipated Coefficient and Power columns.

Power Analysis for Coffee Experiment

Consider the design in the Coffee Data.jmp data table. Suppose that you are interested in the power of your design to detect effects of various magnitudes on Strength. Recall that Grind is a two-level categorical factor, Temperature, Time, and Charge are continuous factors, and Station is a three-level categorical (blocking) factor.

In this example, ignore the role of Station as a blocking factor. You are interested in the effect of Station on Strength. Since Station is a three-level categorical factor, it is represented by two terms in the Parameters list: Station 1 and Station 2.

Specifically, you are interested the probability of detecting the following changes in the mean Strength:

- A change of 0.10 units as you vary Grind from Coarse to Medium.
- A change of 0.10 units or more as you vary Temperature, Time, and Charge from their low to high levels.
- An increase due to each of Stations 1 and 2 of 0.10 units beyond the overall anticipated mean. This corresponds to a decrease due to Station 3 of 0.20 units from the overall anticipated mean.

You set 0.05 as your Significance Level. Your estimate of the standard deviation of Strength for fixed design settings is 0.1 and you enter this as the Anticipated RMSE.

Figure 15.17 shows the Power Analysis node with these values entered. Specifically, you specify the Significance Level, Anticipated RMSE, and the value of each Anticipated Coefficient.

When you click **Apply Changes to Anticipated Coefficients**, the Anticipated Response values are updated to reflect the model you have specified.

Figure 15.17 Power Analysis Outline with User Specifications in Anticipated Coefficients Panel

Power Analysis

Significance Level 0.05
Anticipated RMSE 0.1

Term	Anticipated Coefficient	Power
Intercept	1.4	1
Grind	0.05	0.291
Temperature	0.05	0.291
Time	0.05	0.291
Charge	0.05	0.291
Station 1	0.1	0.507
Station 2	0.1	0.507

Apply Changes to Anticipated Coefficients

Effect Power
Station 0.888

Design and Anticipated Responses

Anticipated Response	Grind	Temperature	Time	Charge	Station
1.6	Medium	205	4	2.4	1
1.4	Coarse	195	3	1.6	1
1.4	Medium	205	3	1.6	1
1.6	Coarse	195	4	2.4	1
1.6	Coarse	205	4	1.6	2
1.5	Medium	195	4	2.4	2
1.3	Medium	195	3	1.6	2
1.6	Coarse	205	3	2.4	2
1.3	Coarse	205	4	1.6	3
1.2	Coarse	195	3	2.4	3
1.1	Medium	195	4	1.6	3
1.2	Medium	205	3	2.4	3

Apply Changes to Anticipated Responses

Recall that Temperature is a continuous factor with coded levels of -1 and 1. Consider the test whose null hypothesis is that Temperature has no effect on Strength. Figure 15.17 shows that the power of this test to detect a difference of 0.10 ($=2 \times 0.05$) units across the levels of Temperature is only 0.291.

Now consider the test for the whole Station effect, where Station is a three-level categorical factor. Consider the test whose null hypothesis is that Station has no effect on Strength. This is the usual F test for a categorical factor provided in the Effect Tests report when you run **Analyze > Fit Model**. See the Standard Least Squares Report and Options chapter in *Fitting Linear Models*.

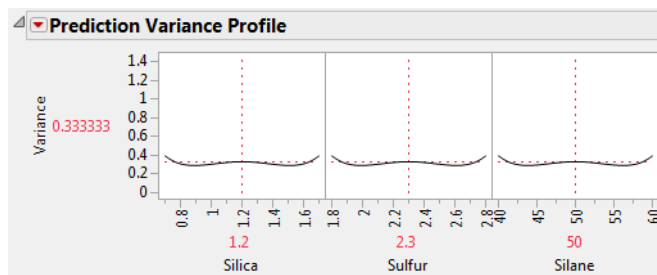
The Power of this test is shown directly beneath the Apply Changes to Anticipated Coefficients button. The entries under Anticipated Coefficients for the model terms Station 1 and Station 2 are both 0.10. These settings imply that the effect of both stations is to increase Strength by 0.10 units above the overall anticipated mean. For these settings of the Station 1 and Station 2 coefficients, the effect of Station 3 on Strength is to decrease it by 0.20 units from the overall anticipated mean. Figure 15.17 shows that the power of the test to detect a difference of at least this magnitude is 0.888.

Prediction Variance Profile

The Prediction Variance Profile helps you to understand where in the design space your predictions have more or less variability. Low prediction variance is desired. Use the **Maximize Variance** option to find the maximum variance. See [“Maximize Variance”](#) on page 457.

The Prediction Variance Profile plots the relative variance of prediction as a function of each factor at fixed values of the other factors. Figure 15.18 shows the Prediction Variance Profile for the Bounce Data.jmp data table, located in the Design Experiment folder.

Figure 15.18 Prediction Variance Profiler



Relative Prediction Variance

For given settings of the factors, the prediction variance is the product of the error variance and a quantity that depends on the design and the factor settings. Before you run your experiment, the error variance is unknown, so the prediction variance is also unknown. However, the ratio of the prediction variance to the error variance is not a function of the error variance. This ratio, called the *relative prediction variance*, depends only on the design and the factor settings. Consequently, the relative variance of prediction can be calculated before acquiring the data. See [“Relative Prediction Variance”](#) on page 775 in the “Technical Details” appendix.

After you run your experiment and fit a least squares model, you can estimate the error variance using the mean squared error (MSE) of the model fit. You can estimate the actual variance of prediction at any setting by multiplying the relative variance of prediction at that setting.

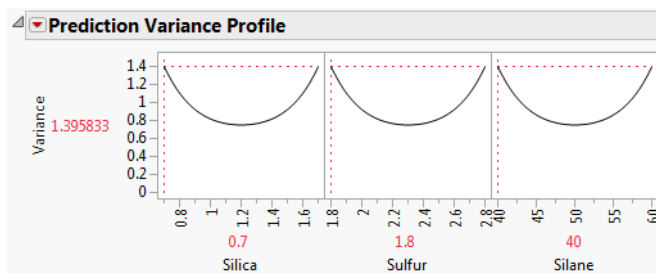
It is ideal for the prediction variance to be small throughout the design space. Generally, the error variance drops as the sample size increases. In comparing designs, you may want to place the prediction variance profilers for two designs side-by-side. A design with lower prediction variance on average is preferred.

Maximize Variance

You can evaluate a design or compare designs in terms of the maximum relative prediction variance. Select the **Maximize Variance** option from the red triangle menu next to Prediction Variance Profile. JMP uses a desirability function that maximizes the relative prediction variance. The value of the maximum variance in the Prediction Variance Profile is the worst (least desirable from a design point of view) value of the relative prediction variance. The maximum variance can occur at more than one combination of factor settings.

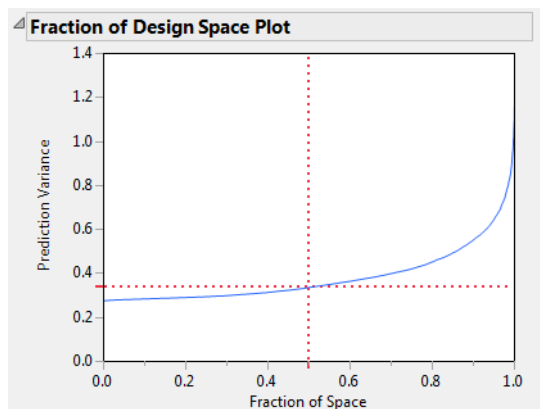
Figure 15.19 shows the Prediction Variance Profile with Maximize Variance selected. The plot is for the Bounce Data.jmp sample data table, located in the Design Experiment folder. The maximum value of the relative prediction variance is 1.3958 when Silica = 0.7, Sulfur = 1.8, and Silane = 40. However, keep in mind that several factor settings can have this same relative variance. The design point with Silica = 1.6, Sulfur = 2.8, and Silane = 60 also has relative prediction variance of 1.3958. To evaluate the prediction variance using a surface, see [“Prediction Variance Surface”](#) on page 458.

Figure 15.19 Prediction Variance Profile Showing Maximum Variance



Fraction of Design Space Plot

The Fraction of Design Space Plot shows the proportion of the design space over which the relative prediction variance lies below a given value. It is desirable to have a large proportion of the design space with low prediction variance values. Figure 15.20 shows the Fraction of Design Space plot for the Bounce Data.jmp sample data table, located in the Design Experiment folder.

Figure 15.20 Fraction of Design Space Plot

The X axis in the plot represents the proportion of the design space, ranging from 0 to 100%. The Y axis represents relative prediction variance values. For a point (x, y) that falls on the blue curve, the value x is the proportion of design space with variance less than or equal to y . Red dotted crosshairs mark the value that bounds the relative prediction variance for 50% of design space.

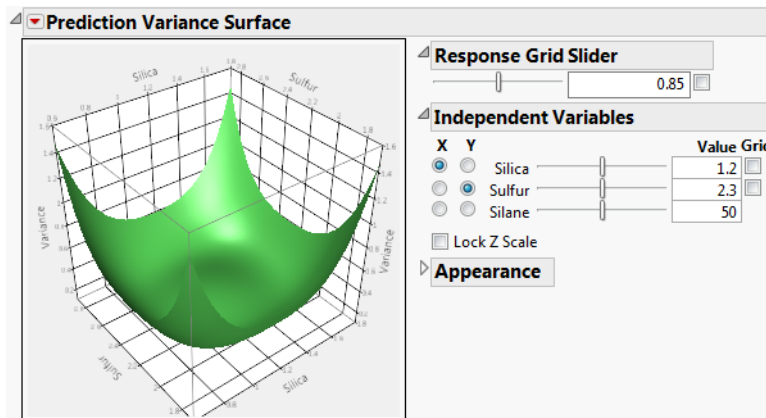
Figure 15.20 shows that the minimum relative prediction variance is slightly less than 0.3, while the maximum is below 1.4. (The actual maximum is 1.395833, as shown in Figure 15.19.) The red dotted crosshairs indicate that the relative prediction variance is less than 0.34 over about 50% of the design space. You can use the crosshairs tool to find the maximum relative prediction variance that corresponds to any Fraction of Space value. Use the crosshairs tool in Figure 15.20 to see that 90% of the prediction variance values are below approximately 0.55.

Note: Monte Carlo sampling of the design space is used in constructing the Fraction of Design Space Plot. Therefore, plots for the same design may vary slightly.

Prediction Variance Surface

The Prediction Variance Surface report plots the relative prediction variance surface as a function of any two design factors. Figure 15.21 shows the Prediction Variance Surface outline for the Bounce Data.jmp sample data table, located in the Design Experiment folder. Show or hide the controls by selecting **Control Panel** on the red triangle menu. See [“Control Panel”](#) on page 459.

Figure 15.21 Prediction Variance Surface



When there are two or more factors, the Prediction Variance Surface outline shows a plot of the relative prediction variance for any two variables. The Prediction Variance Surface outline plots the relative prediction variance formula. Drag on the plot to rotate and change the perspective.

Control Panel

The Control Panel consists of the following:

Response Grid Slider The **Grid** check box superimposes a grid that shows constant values of Variance. The value of the Variance is shown in the text box. The slider enables you to adjust the placement of the grid. Alternatively, you can enter a Variance value in the text box. Click outside the box to update the plot.

Independent Variables This panel enables you to select which two factors are used as axes for the plot and to specify the settings for factors not used as axes. Select a factor for each of the X and Y axes by clicking in the appropriate column. Use the sliders and text boxes to specify values for each factor *not* selected for an axis. The plot shows the three-dimensional slice of the surface at the specified values of the factors that are not used as axes in the plot. Move the sliders to see different slices.

Each grid check box activates a grid for the corresponding factor. Use the sliders to adjust the placement of each grid.

Lock Z Scale locks the z-axis to its current values. This is useful when moving the sliders that are not on an axis.

Appearance The **Resolution** slider affects how many points are evaluated for a formula. Too coarse a resolution means that a function with a sharp change might not be represented

very well. But setting the resolution high can make evaluating and displaying the surface slower.

The **Orthographic projection** check box shows a projection of the plot in two dimensions.

The **Contour** menu controls the placement of contour curves. A contour curve is a set of points whose Response values are constant. You can select to turn the contours Off (the default) or place them contours Below, Above, or On Surface.

Estimation Efficiency

This report gives the Fractional Increase in CI (Confidence Interval) Length and Relative Std (Standard) Error of Estimate for each parameter estimate in the model. Smaller is better for both of these values. Figure 15.22 shows the Estimation Efficiency outline for the Bounce Data.jmp sample data table, located in the Design Experiment folder.

Figure 15.22 Estimation Efficiency Outline

Estimation Efficiency		
Term	Fractional Increase in CI Length	Relative Std Error of Estimate
Intercept	1.236	0.577
Silica	0.369	0.354
Silane	0.369	0.354
Sulfur	0.369	0.354
Silica*Silane	0.936	0.5
Silica*Sulfur	0.936	0.5
Sulfur*Silane	0.936	0.5
Silica*Silica	1.016	0.52
Silane*Silane	1.016	0.52
Sulfur*Sulfur	1.016	0.52

Fractional Increase in CI Length

The Fractional Increase in CI Length compares the length of a parameter’s confidence interval as given by the current design to the length of such an interval given an ideal design:

- The length of the ideal confidence interval for the parameter is subtracted from the length of its actual confidence interval.
- This difference is then divided by the length of the ideal confidence interval.

For an orthogonal D-optimal design, the fractional increase is zero. In selecting a design, you would like the fractional increase in confidence interval length to be as small as possible.

The Ideal Design

The covariance matrix for the ordinary least squares estimator is $\sigma^2(\mathbf{X}'\mathbf{X})^{-1}$. The diagonal elements of $(\mathbf{X}'\mathbf{X})^{-1}$ are the *relative variances* (the variances divided by σ^2) of the parameter estimates. For two-level designs and using the effects coding convention (see “Coding” on page 742 in the “Column Properties” appendix), the minimum value of the relative variance for any parameter estimate is $1/n$, where n is the number of runs. This occurs when all the effects for the design are orthogonal and the design is D-optimal.

Let $\hat{\beta}$ denote the vector of parameter estimates. The *ideal design*, which may not exist, is a design whose covariance matrix is given as follows:

$$\text{Var}(\hat{\beta}) = (\sigma^2/n)\mathbf{I}_n$$

where $\mathbf{I}_n$ is the n by n identity matrix and σ is the standard deviation of the response.

If an orthogonal D-optimal design exists, it is the ideal design. However, the definition above extends the idea of an ideal design to situations where a design that is both orthogonal and D-optimal does not exist.

The definition is also appropriate for designs with multi-level categorical factors. The orthogonal coding used for categorical factors allows such designs to have the ideal covariance matrix. For a Custom Design, you can view the coding matrix by selecting Save X Matrix from the options in the Custom Design window, making the design table, and looking at the script Model Matrix that is saved to the design table.

Fractional Increase in Length of Confidence Interval

Note that, in the ideal design, the standard error for the parameter estimates would be given as follows:

$$SE_{Ideal}(\hat{\beta}) = (\sigma/\sqrt{n})\mathbf{I}_n$$

The length of a confidence interval is determined by the standard error. The Fractional Increase in Confidence Interval Length is the difference between the standard error of the given design and the standard error of the ideal design, divided by the standard error of the ideal design.

Specifically, for the i^{th} parameter estimate, the Fractional Increase in Confidence Interval Length is defined as follows:

$$FI = \frac{\sigma\sqrt{(\mathbf{X}'\mathbf{X}_{ii})^{-1}} - (\sigma/(\sqrt{n}))}{(\sigma/\sqrt{n})} = \sqrt{n(\mathbf{X}'\mathbf{X}_{ii})^{-1}} - 1$$

where

σ^2 is the unknown response variance,

$\mathbf{X}$ is the model matrix for the given design, defined in [“The Alias Matrix”](#) on page 771 in the [“Technical Details”](#) appendix,

$(\mathbf{X}'\mathbf{X})_{ii}^{-1}$ is the i^{th} diagonal entry of $(\mathbf{X}'\mathbf{X})^{-1}$, and

n is the number of runs.

Relative Std Error of Estimate

The Relative Std Error of Estimate gives the ratio of the standard deviation of a parameter's estimate to the error standard deviation. These values indicate how large the standard errors of the model's parameter estimates are, relative to the error standard deviation. For the i^{th} parameter estimate, the Relative Std Error of Estimate is defined as follows:

$$SE = \sqrt{(\mathbf{X}'\mathbf{X})_{ii}^{-1}}$$

where

$\mathbf{X}$ is the model matrix defined in [“The Alias Matrix”](#) on page 771 in the [“Technical Details”](#) appendix, and

$(\mathbf{X}'\mathbf{X})_{ii}^{-1}$ is the i^{th} diagonal entry of $(\mathbf{X}'\mathbf{X})^{-1}$.

Alias Matrix

The Alias Matrix addresses the issue of how terms that are not included in the model affect the estimation of the model terms, if they are indeed active. In the Alias Terms outline, you list potentially active effects that are not in your assumed model but that might bias the estimates of model terms. The Alias Matrix entries represent the degree of bias imparted to model parameters by the Alias Terms effects. See [“Alias Terms”](#) on page 448.

The rows of the Alias Matrix are the terms corresponding to the model effects listed in the Model outline. The columns are terms corresponding to effects listed in the Alias Terms outline. The entry in a given row and column indicates the degree to which the alias term affects the parameter estimate corresponding to the model term.

In evaluating your design, you ideally want one of two situations to occur relative to any entry in the Alias Matrix. Either the entry is small or, if it is not small, the effect of the alias term is small so that the bias will be small. If you suspect that the alias term may have a substantial effect, then that term should be included in the model or you should consider an alias optimal design.

For more information about the computation of the Alias Matrix, see [“The Alias Matrix”](#) on page 771 in the [“Technical Details”](#) appendix. See also Lekivetz, R. (2014).

Note the following:

- If the design is orthogonal for the assumed model, then the correlations in the Alias Matrix correspond to the absolute correlations in the Color Map on Correlations.
- Depending on the complexity of the design, it is possible to have alias matrix entries greater than 1 or less than -1.

Alias Matrix Example

Consider the *Coffee Data.jmp* sample data table, located in the Design Experiment folder. The design assumes a main effects model. You can see this by running the Model script in the data table. Consequently, in the Evaluate Design window's Model outline, only the Intercept and five main effects appear. The Alias Terms outline contains the two-way interactions.

Figure 15.23 Alias Matrix for *Coffee Data.jmp*

Alias Matrix							
Effect	Grind*Temperature	Grind*Time	Grind*Charge	Temperature*Time	Temperature*Charge	Time*Charge	
Intercept	0	0	0	0	0	0	
Grind	0	0	0	0.333	-0.33	-0.33	
Temperature	0	0.333	-0.33	0	0	-0.33	
Time	0.333	0	-0.33	0	-0.33	0	
Charge	-0.33	-0.33	0	-0.33	0	0	
Station 1	-0.41	0	0	0	0	0.816	
Station 2	0.707	0	0	0	0	0	

The Alias Matrix shows the Model terms in the first column defining the rows. The two-way interactions in the Alias Terms are listed across the top, defining the columns. Consider the model effect *Temperature* for example. If the *Grind*Time* interaction is the only active two-way interaction, the estimate for the coefficient of *Temperature* is biased by 0.333 times the true value of the *Grind*Time* effect. If other interactions are active, then the value in the Alias Matrix indicates the additional amount of bias incurred by the *Temperature* coefficient estimate.

Color Map on Correlations

The Color Map on Correlations shows the absolute value of the correlation between any two effects that appear in either the Model or the Alias Terms outline. The cells of the color map are identified above the map. There is a cell for each effect in the Model outline and a cell for each effect in the Alias Terms outline. Smaller values are desired.

By default, the absolute magnitudes of the correlations are represented by a white to gray to black intensity color theme. In general terms, the color map for a good design shows a lot of white off the diagonal, indicating orthogonality or small correlations between distinct terms. Large absolute correlations among effects inflate the standard errors of estimates.

To see the absolute value of the correlation between two effects, hover your cursor over the corresponding cell. To change the color theme for the entire plot, right-click in the plot and select Color Theme. To save a table of the correlations, right-click to the right of the plot below the legend and select **Table of Correlations**.

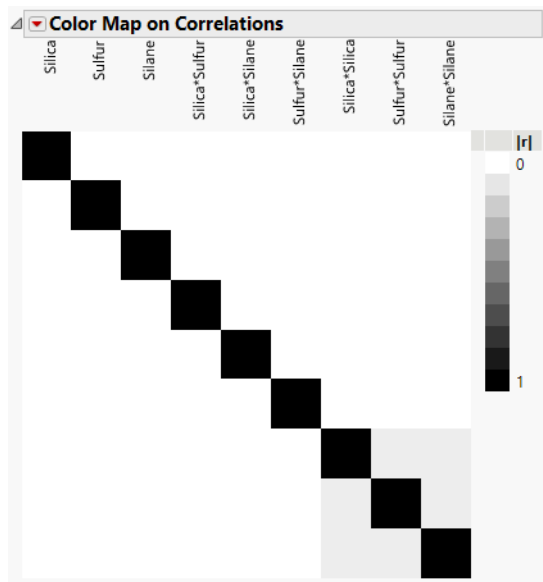
Color Map Example

Figure 15.24 shows the Color Map on Correlations for the Bounce Data.jmp sample data table, found in the Design Experiment folder. The black coloring indicates absolute correlations of one. Note that there are black cells on the diagonal, showing correlations of model terms with themselves.

All other cells are either black or gray. The gray squares correspond to correlations between quadratic terms. To see this, hover your cursor over each of the gray squares. The absolute correlations of quadratic terms with each other are small, 0.0714.

From the perspective of correlation, this is a good design. When effects are highly correlated, it is more difficult to determine which is responsible for an effect on the response.

Figure 15.24 Color Map on Correlations



Tip: To save a table of the correlations, right-click to the right of the plot below the legend and select **Table of Correlations**.

Design Diagnostics

The Design Diagnostics outline shows D -, G -, and A -efficiencies and the average variance of prediction. These diagnostics are not shown for designs that include factors with Changes set to Hard or Very Hard or effects with Estimability designated as If Possible.

When Design Diagnostics is accessed from a DOE platform other than Evaluate Design, the Design Creation Time gives the amount of time required to create the design. When Design Diagnostics is accessed from the Evaluate Design platform, Design Creation Time gives the amount of time required for the Evaluate Design platform to calculate results.

Figure 15.25 shows the Design Diagnostics outline for the Bounce Data.jmp sample data table, found in the Design Experiment folder.

Figure 15.25 Design Diagnostics Outline

Design Diagnostics	
D Efficiency	36.6429
G Efficiency	69.10947
A Efficiency	29.3578
Average Variance of Prediction	0.384722
Design Creation Time (seconds)	0

Caution: The efficiency measures should not be interpreted on their own. But they can be used to compare designs. Given two designs, the one with the higher efficiency measure is better. While the maximum efficiency is 100 for any criterion, an efficiency of 100% is impossible for many design problems.

Notation

The descriptions of the efficiency measures given below use the following notation:

- $\mathbf{X}$ is the model matrix
- n is the number of runs in the design
- p is the number of terms, including the intercept, in the model
- $Var(\hat{y}|\mathbf{x})$ is the relative prediction variance at the point $\mathbf{x}$. See [“Relative Prediction Variance”](#) on page 775 in the “Technical Details” appendix.
- $Var(\hat{y}|\mathbf{x})_{max}$ is the maximum relative prediction variance over the design region

D Efficiency

The efficiency of the design to that of an ideal orthogonal design in terms of the D-optimality criterion. A design is D-optimal if it minimizes the volume of the joint confidence region for the vector of regression coefficients:

$$D\text{-efficiency} = 100 \left(\frac{1}{n} |\mathbf{X}'\mathbf{X}|^{1/p} \right)$$

G Efficiency

The efficiency of the design to that of an ideal orthogonal design in terms of the G-optimality criterion. A design is G-optimal if it minimizes the maximum prediction variance over the design region:

$$G\text{-efficiency} = 100p / (n \text{Var}(\hat{y}|\mathbf{x})_{\max})$$

Letting D denote the design region,

$$\text{Var}(\hat{y}|\mathbf{x})_{\max} = \underset{\mathbf{x} \text{ in D}}{\text{maximum}} [\mathbf{x}'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}]$$

Note: G-Efficiency is calculated using Monte Carlo sampling of the design space. Therefore, calculations for the same design may vary slightly.

A Efficiency

The efficiency of the design to that of an ideal orthogonal design in terms of the A-optimality criterion. A design is A-optimal if it minimizes the sum of the variances of the regression coefficients:

$$A\text{-efficiency} = 100p / (n \text{Trace}(\mathbf{X}'\mathbf{X})^{-1})$$

Average Variance of Prediction

At a point $\mathbf{x}$ in the design space, the relative prediction variance is defined as:

$$\text{Var}(\hat{y}|\mathbf{x}) = \mathbf{x}'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}$$

Note that this is the prediction variance divided by the error variance. For more information about the calculation, see Section 4.3.5 in Goos and Jones (2011).

Design Creation Time

Design Creation Time gives the amount of time required for the Evaluate Design platform to calculate results.

Evaluate Design Options

The Evaluate Design red triangle menu contains the following options:

Advanced Options > Split Plot Variance Ratio Specify the ratio of the variance of the random whole plot and the subplot variance (if present) to the error variance. Before setting this value, you must define a hard-to-change factor for your split-plot design, or hard and very-hard-to-change factors for your split-split-plot design. Then you can enter one or two positive numbers for the variance ratios, depending on whether you have specified a split-plot or a split-split-plot design.

Advanced Options > Set Delta for Power Specify a value for the difference you want to detect that is applied to Anticipated Coefficients in the Power Analysis report. The Anticipated Coefficient values are set to $\Delta/2$ for continuous effects. For categorical effects, they are alternating values of $\Delta/2$ and $-\Delta/2$. For more information about power analysis, see [“Power Analysis”](#) on page 449.

By default, Delta is set to two. Consequently, the Anticipated Coefficient default values are 1 for continuous effects and alternating values of 1 and -1 for categorical effects. The default values that are entered as Anticipated Coefficients when Delta is 2 ensure these properties:

- The power calculation for a numeric effect assumes a change of Delta in the response mean due to linear main effects as the factor changes from the lowest setting to the highest setting in the design region.
- The power calculation for the parameter associated with a two-level categorical factor assumes a change of Delta in the response mean across the levels of the factor.
- The power calculation for a categorical effect with more than two levels is based on the multiple degree of freedom F test for the null hypothesis that all levels have the same response mean. Power is calculated at the values of the response means that are determined by the Anticipated Coefficients. Various configurations of the Anticipated Coefficients can define a difference in levels as large as Delta. However, the power values for such configurations will differ based on the Anticipated Coefficients for the other levels.

Save Script to Script Window Creates a script that reproduces the Evaluate Design window and places it in an open script window.

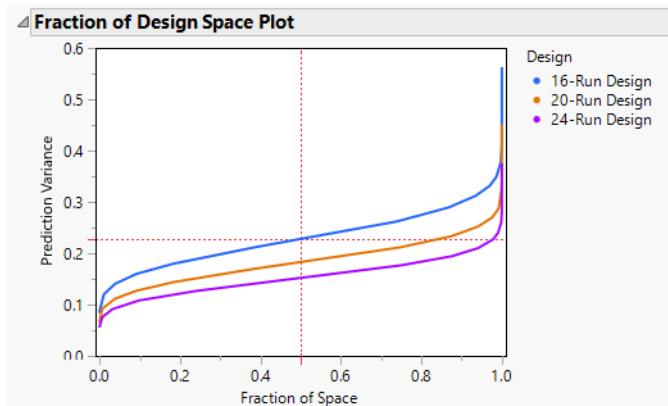
Chapter 16

Compare Designs

Compare and Evaluate Designs Simultaneously

The Compare Designs platform compares up to four designs to a reference design. Use to explore, evaluate, and compare design performance. Diagnostics show how the designs perform relative to each other and how they perform in an absolute sense. To compare designs relative to your specific needs, you can change the terms in the assumed model and in the alias terms list.

Figure 16.1 Comparing Three Designs with Different Run Sizes



Contents

Overview of the Compare Designs Platform	471
Examples of Comparing Designs	471
Designs of Same Run Size	471
Designs of Different Run Sizes	476
Split Plot Designs with Different Numbers of Whole Plots	480
Launch the Compare Designs Platform	484
Compare Designs Window: Specify Model and Alias Terms	486
Reference Design	486
Factors	486
Model	486
Alias Terms	487
Compare Designs Window: Design Evaluation	487
Power Analysis	488
Prediction Variance Profile	491
Fraction of Design Space Plot	493
Relative Estimation Efficiency	495
Alias Matrix Summary	496
Absolute Correlations	498
Design Diagnostics	501
Compare Designs Options	503

Overview of the Compare Designs Platform

The Compare Designs platform, which is an extension of the Evaluate Design platform, enables you to easily compare up to four designs to a reference design. To compare the performance of one or two designs relative to another, you select a reference design that is treated as the base design. You can specify effects in the Model outline, and effects of interest in the Alias Terms outline.

The Design Evaluation report shows diagnostic results and plots covering these areas:

- Power analysis
- Prediction variance
- Fraction of design space
- Relative estimation efficiency
- Alias matrix diagnostics
- Correlations among effects (including confounding)
- Relative efficiency measures for the overall designs

Examples of Comparing Designs

- [“Designs of Same Run Size”](#)
- [“Designs of Different Run Sizes”](#)
- [“Split Plot Designs with Different Numbers of Whole Plots”](#)

Designs of Same Run Size

In this example, you compare two designs for six factors each with 13 runs. One is a 12-run Plackett-Burman (PB) design augmented with a single center point. The other is a Definitive Screening Design (DSD).

Comparison in Terms of Main Effects Only

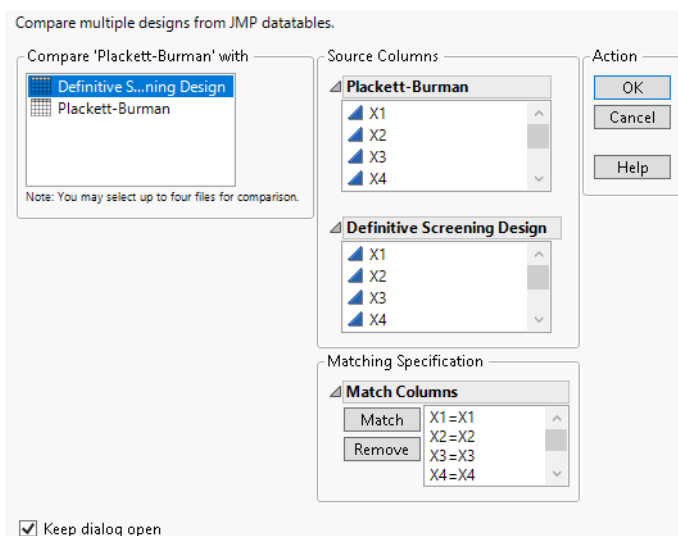
First, compare the two designs assuming that the model to be estimated contains only the main effects.

1. Select **Help > Sample Data**, click **Open the Sample Scripts Directory**, and select **Compare Same Run Size.jsl**.
2. Right-click in the script window and select **Run Script**.

Two 13-run design tables are constructed: Definitive Screening Design and Plackett-Burman. You want to compare these two designs. Because the Plackett-Burman table is active, it is the reference design to which you compare the DSD.

3. In the Plackett-Burman data table, select **DOE > Design Diagnostics > Compare Designs**.
4. Select Definitive Screening Design from the **Compare 'Plackett-Burman' with** list.
5. Select X1 through X6 in the Plackett-Burman panel and in the Definitive Screening Panel.
6. Open the Match Columns outline and click **Match**.

Figure 16.2 Launch Window with Matched Columns



This defines the correspondence between the factors in your two designs.

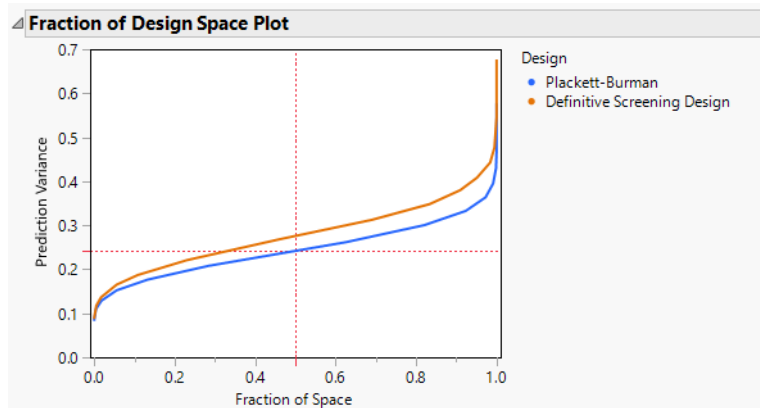
7. Click **OK**.

The reference design is the Plackett-Burman design. In the Design Evaluation outline, comparison metrics compare the PB to the DSD. The designs are compared relative to power, prediction variance, estimation efficiency, aliasing, and design efficiency measures.

Figure 16.3 Power Analysis for PB and DSD Comparison



In terms of power, prediction variance, and estimation efficiency, the PB design outperforms the DSD. Figure 16.3 shows the Power Analysis report with the default settings for the significance level, Anticipated RMSE, and coefficients. For tests for the main effects, the PB design has higher power than does the DSD.

Figure 16.4 Fraction of Design Space Plot for PB and DSD Comparison

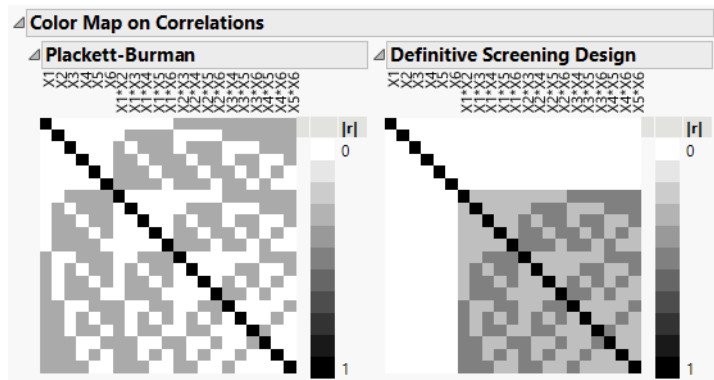
The Fraction of Design Space plot indicates that the PB design has smaller prediction variance than the DSD over the entire design space.

You conclude that, if you suspect that only main effects are active, the PB design is preferable.

Comparison in Terms of Two-Way Interactions

Now suppose you suspect that some two-way interactions might be active. The analysis below shows that if those two-way interactions are actually active, then the PB design might be less desirable than the DSD.

1. In the Absolute Correlations report, open the Color Map on Correlations report and the color map reports under it.

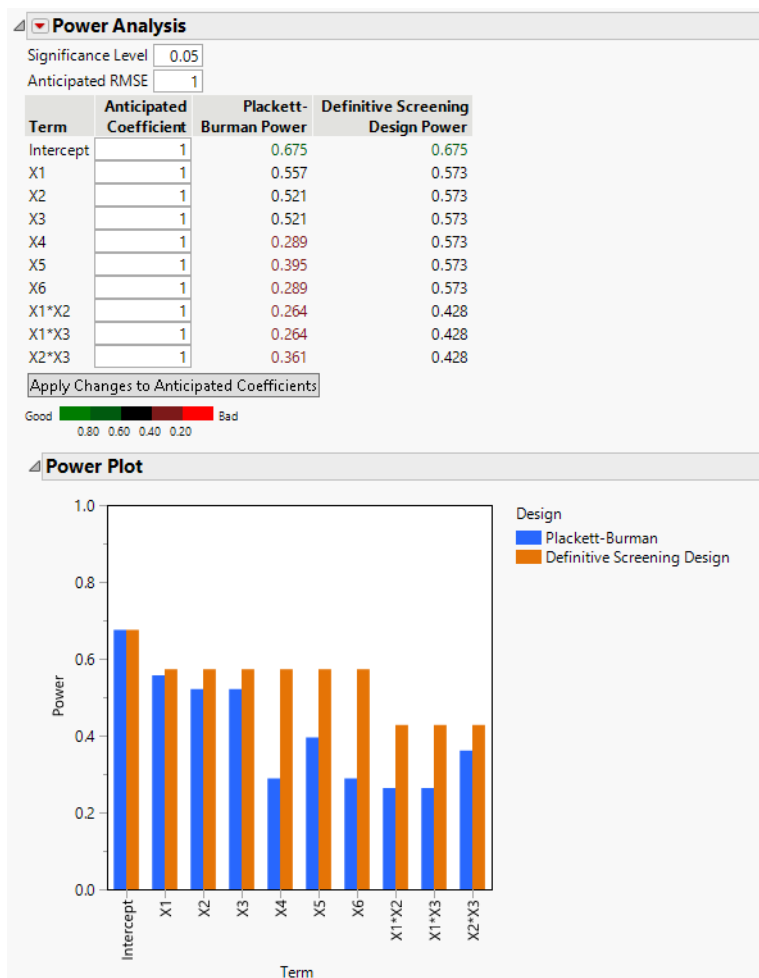
Figure 16.5 Color Maps for PB and DSD Comparison

The Color Map on Correlations plots show that the PB design aliases main effects with two-way interactions. In contrast, the DSD does not alias main effects with two-way interactions.

To gain more insight on how the designs compare if some two-way interactions are active, add two-way interactions in the Model outline.

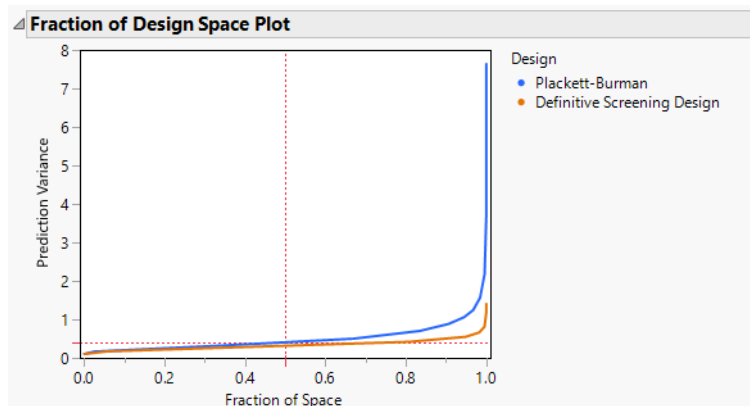
2. In the Factors outline, select X1 through X3.
3. In the Model outline, select **Interactions > 2nd**.

Figure 16.6 Power Analysis for PB and DSD Comparison with Interactions



The Term list shows the three two-way interactions. If these two-way interactions are active, then the DSD has better performance in terms of power across all effects than the PB.

Figure 16.7 Prediction Variance for PB and DSD Comparison with Interactions



The DSD also outperforms the PB design in terms of prediction variance with the three interactions in the model. You can explore the other reports to see that the DSD is preferred when there are potentially active interactions.

Designs of Different Run Sizes

In this example, compare three designs with run sizes 16, 20, and 24. The designs are constructed for main effect models. Use the Compare Designs platform to determine whether the potential benefits of using a larger run size are worth the additional cost in resources.

1. Select **Help > Sample Data**, click **Open the Sample Scripts Directory**, and select **Compare Three Run Sizes.jsl**.
2. Right-click in the script window and select **Run Script**.

Three design tables are constructed using Custom Design, with only main effects as entries in the Model outline:

- 16-Run Design
- 20-Run Design
- 24-Run Design

You want to compare these three designs. Notice that the 16-Run Design table is active.

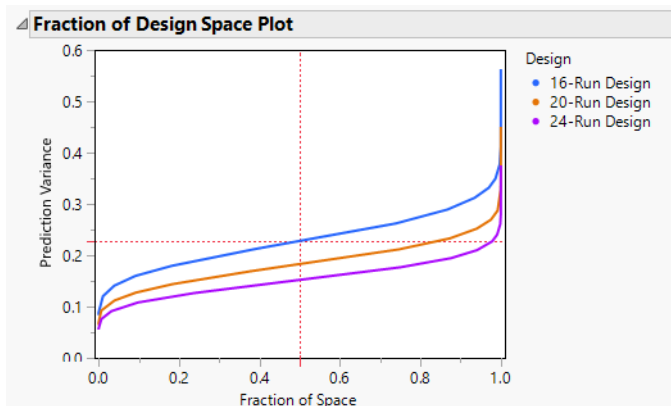
3. In the 16-Run Design table, select **DOE > Design Diagnostics > Compare Designs**.
4. From the **Compare '16-Run Design' with** list, select 20-Run Design and 24-Run Design.
Panels for each of these designs are added to the launch window. JMP automatically matches the columns in the order in which they appear in the three design tables.
5. Click **OK**.

Figure 16.8 Power Analysis Comparison



All three designs have high power for detecting main effects if the coefficients are on the order of the Anticipated RMSE.

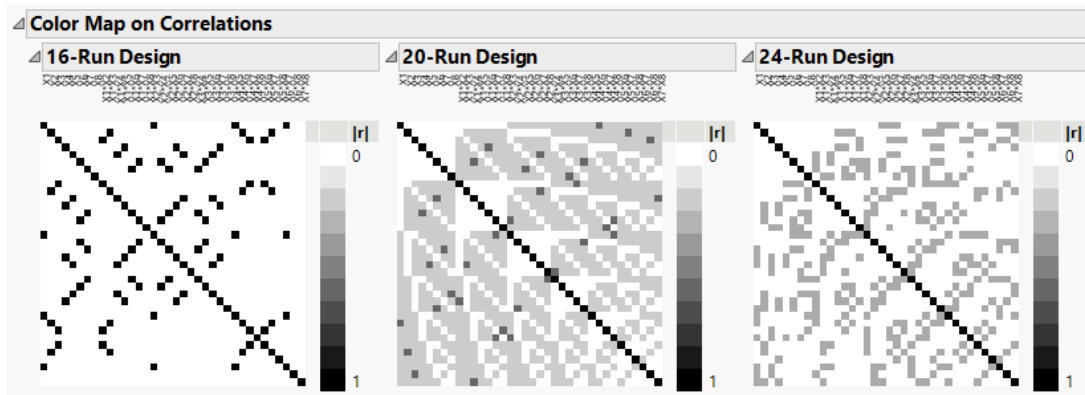
Figure 16.9 Fraction of Design Space Comparison



As expected, the 24-run design is superior to the other two designs in terms of prediction variance over the entire design space. The 20-run design is superior to the 16-run design.

6. In the Absolute Correlations report, open the Color Map on Correlations report and the three color map reports under it.

Figure 16.10 Color Map on Correlations Comparison



For the 16-run design, the Color Map on Correlations indicates that there is confounding of some main effects with some two-factor interactions, and confounding of two-factor interactions.

For the 20-run design, the Color Map on Correlations indicates that there are some large correlations between some main effects and some two-factor interactions, and between some two-factor interactions.

The 24-run design shows only moderate correlations between main effects and two-factor interactions, and between two-factor interactions.

Figure 16.11 Absolute Correlations Comparison

Absolute Correlations			
Model x Model	Average Correlation	Number of Confoundings	Number of Pairs
16-Run Design	0.000	0	28
20-Run Design	0.000	0	28
24-Run Design	0.000	0	28
Model x Alias	Average Correlation	Number of Confoundings	Number of Pairs
16-Run Design	0.094	21	224
20-Run Design	0.166	0	224
24-Run Design	0.098	0	224
Alias x Alias	Average Correlation	Number of Confoundings	Number of Pairs
16-Run Design	0.056	21	378
20-Run Design	0.121	0	378
24-Run Design	0.085	0	378

Good  Bad
0.10 0.30 0.50 0.70

The Absolute Correlations table summarizes the information shown in the Color Maps on Correlations. Recall that the model for all three designs consists of only main effects and the Alias Matrix contains two-factor interactions.

For the 16-run design, the Model x Alias portion of the table indicates that there are nine confoundings of main effects with two-factor interactions. The Alias x Alias portion indicates that six two-factor interactions are confounded.

Figure 16.12 Design Diagnostics Comparison

Design Diagnostics		
	Efficiency of 16-Run Design Relative to 20-Run Design	Efficiency of 16-Run Design Relative to 24-Run Design
D-efficiency	0.800	0.667
G-efficiency	0.800	0.667
A-efficiency	0.800	0.667
I-efficiency	0.800	0.667
Additional Run Size	-4	-8

Good  Bad
1.50 1.25 0.80 0.67

The Design Diagnostics report compares the efficiency of the 16-run design to both the 20-run and 24-run designs in terms of several efficiency measures. Relative efficiency values that exceed 1 indicate that the reference design is preferable for the given measure. Values less than 1 indicate that the design being compared to the reference design is preferable. The 16-run design has lower efficiency than the other two designs across all metrics, indicating that the larger designs are preferable.

7. In the Factors outline, select X1 through X3.
8. In the Model outline, select **Interactions > 2nd**.

An Inestimable Terms window appears, telling you that the 16-run design cannot fit one of the effects that you just added to the model (X1*X2).

9. Click **OK**.

The other two effects, $X1 \times X3$ and $X2 \times X3$, are added to the Compare Design report. You can examine the report to compare the designs if the two interactions are active.

Split Plot Designs with Different Numbers of Whole Plots

In this example, compare two split-plot designs with different numbers of whole plots. The designs are for three factors:

- A continuous hard-to-change factor
- A continuous easy-to-change factor
- A three-level categorical easy-to-change factor

The designs include all two-factor interactions in the assumed model. You can afford 20 runs and want to compare using 4 or 8 whole plots.

Launch Compare Designs

1. Select **Help > Sample Data**, click **Open the Sample Scripts Directory**, and select **Compare Split Plots.jsl**.
2. Right-click in the script window and select **Run Script**.

Two design tables are constructed using Custom Design:

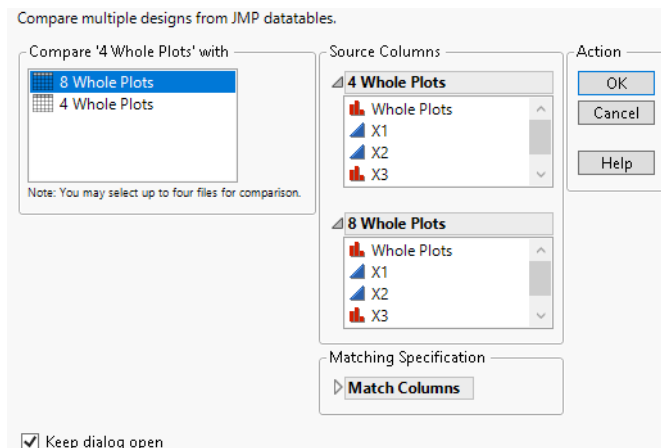
- 4 Whole Plots
- 8 Whole Plots

You want to compare these two designs. Notice that the 4 Whole Plots table is active.

3. In the 4 Whole Plots table, select **DOE > Design Diagnostics > Compare Designs**.
4. From the **Compare '4 Whole Plots' with** list, select 8 Whole Plots.

A panel for this design is added to the launch window. JMP automatically matches the columns in the order in which they appear in the two design tables.

Figure 16.13 Completed Launch Window



- Click OK.
- Open the **Matching Specification** outline under Reference Design: 20 run '4 Whole Plots'.

Figure 16.14 Matching Specification for Split-Plot Designs

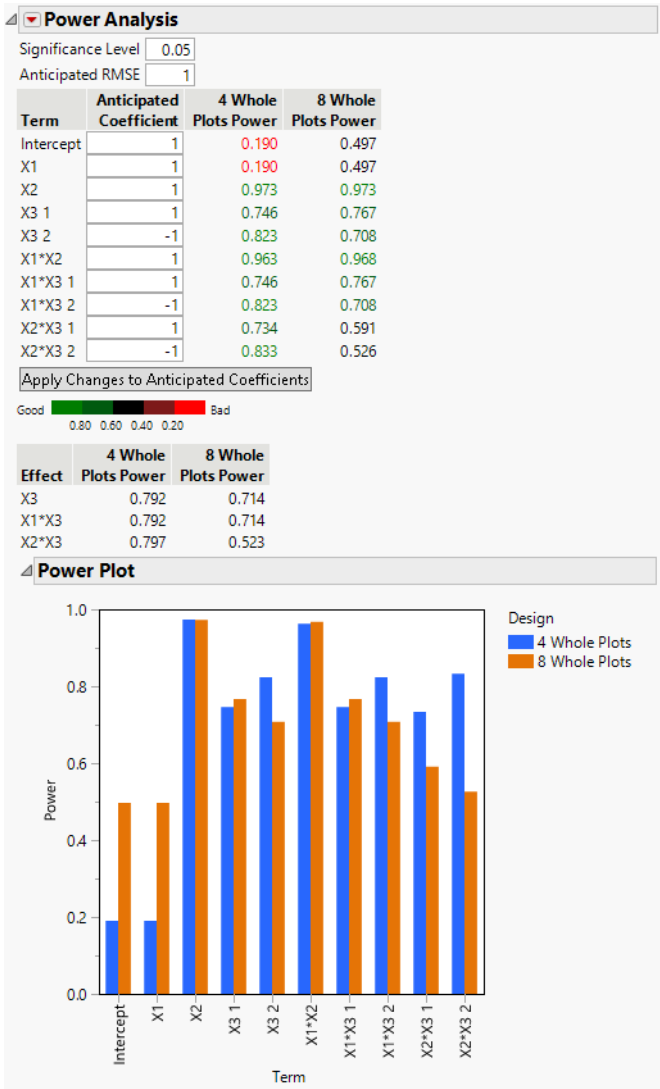


Notice that the Whole Plots column is entered as part of the design. This is necessary because Compare Designs needs to know the whole plot structure.

Examine the Report

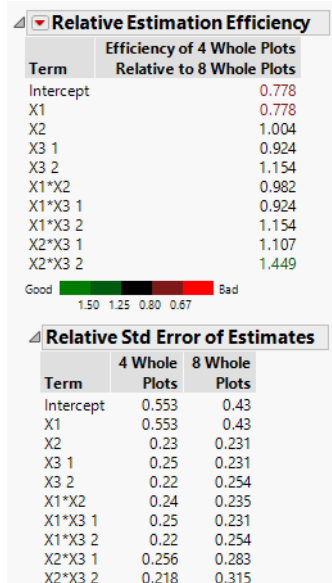
The Design Evaluation report provides various diagnostics that compare the two designs.

Figure 16.15 Power Analysis for Two Split-Plot Designs



The Power Analysis report shows that the power for the whole-plot factor, X1, is much smaller for the four whole-plot design (0.19) than for the eight whole-plot design (0.497). However, the four whole-plot design has higher power to detect split-plot effects, especially the interaction of the two split-plot factors, X2*X3 (0.797 compared to 0.523). Notice that the power for the combined effect X2*X3 is given under the color bar and legend.

Figure 16.16 Relative Estimation Efficiency Comparing Split-Plot Designs



The Relative Estimation Efficiency report shows the relative estimation efficiency for X1 to be 0.778. This indicates that the standard error for X1 is notably larger for the four whole-plot design than for the eight whole-plot design.

Open the Relative Std Error of Estimates report. You can see that the relative standard error for X1 in the four whole-plot design is 0.553, compared to the eight whole-plot error of 0.43.

In the Relative Estimation Efficiency report, the relative estimation efficiency for X2*X3 2 is 1.449, indicating that the standard error for the parameter associated with X2*X3 2 is notably larger for the eight whole-plot design than for the four whole-plot design.

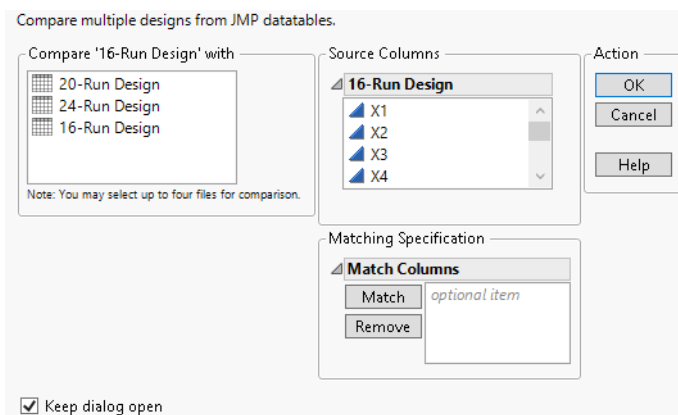
The Power Analysis and the Relative Estimation Efficiency reports indicate that the choice of designs revolves around the importance of detecting the whole plot effect X1. The eight whole-plots design gives you a better chance of detecting a whole plot effect. The four whole-plots design is somewhat better for detecting split-plot effects involving the categorical variable.

Launch the Compare Designs Platform

Launch the Compare Designs platform by selecting **DOE > Design Diagnostics > Compare Designs**. All open data tables appear in the list at the left. The active data table and its columns appear in a Source Columns panel. The design in the initial Source Columns panel is the *reference design*, namely, the design to which other designs are compared. When you add designs to compare to the reference design, their columns appear in panels under the reference design panel.

Figure 16.17 shows the launch window for the three designs in “[Designs of Different Run Sizes](#)” on page 476.

Figure 16.17 Compare Designs Launch Window



Design Table Selection

Select up to four design tables from the list on the left.

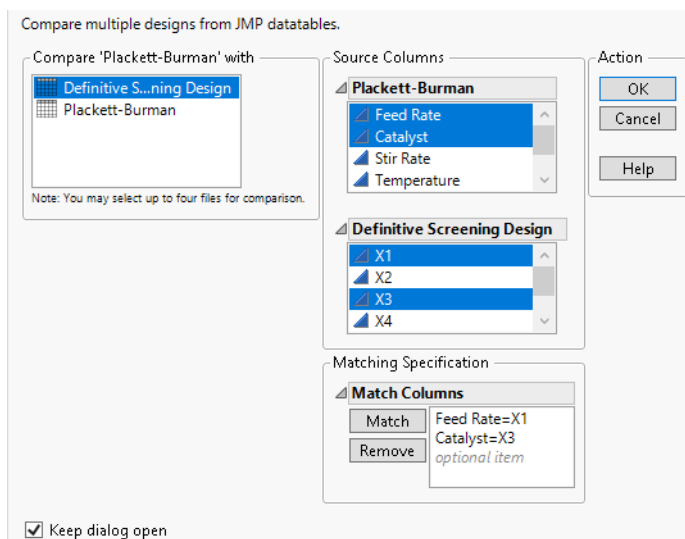
- To compare multiple designs to the reference design, you must select their design tables simultaneously from the list on the left.
- To replace a design (or designs) in the Source Columns list, select the desired table (or tables) from the list at the left. The design table (or tables) under the reference design table are replaced.

Note: The reference design table can be compared to itself, which can be useful when exploring the assignment of design columns to factors.

Match Columns

Specify which columns in each of the design tables correspond to each other in the Match Columns panel. To match columns, select the columns to match in each of the design table Source Columns lists, and then click Match.

Figure 16.18 Selection of Columns for Matching



- To match single columns in each list, select the single column in each list, and then click Match.
- To match several columns that appear in the correct matching order in each list, select them in each list. Click the Match button. They are matched in their list order (Figure 16.18). In this example, Feed Rate is matched with X1, and Catalyst is matched with X3.
- If the lists contain the same numbers of columns and your desired match order is their order of appearance in the lists, you do not have to click Match. When you click OK to run the launch window, JMP matches the columns automatically in their order of appearance. You can review the matching in the report's Matching Specification outline.

Compare Designs Window: Specify Model and Alias Terms

The Compare Designs window consists of two sets of outlines:

- Specify which effects are in the model and which effects are potentially active using the Factors, Model, and Alias Terms outlines.
- Compare the designs using the diagnostics in the Design Evaluation outlines. Changes that you make in the Model and Alias Terms outlines are updated in the Design Evaluation report.

The Compare Designs report uses the column names from the reference design.

This section describes the Reference Design, Factors, Model, and Alias Terms outlines. See [“Compare Designs Window: Design Evaluation”](#) on page 487 for a description of the Design Evaluation outlines.

Reference Design

The name of the window for the reference design appears in the outline title. The Matching Specification outline lists the specifications that you entered in the launch window.

Factors

Use the Factors outline to add effects to the Model and Alias Terms lists.

The Factors outline lists the factors, using the column names from the reference design, and coded values. Because they are not factors, whole plot and subplot columns do not appear in the Factors outline. However, they are required for the analysis.

Model

Add or remove effects to compare your designs for the effects that you believe should be in the model. The Model outline initially lists effects that are in the Model script of the reference design table and that are estimated by all designs being compared. If there is no Model script in the reference design table, the Model outline shows only the main effects that can be estimated by all designs being compared. For more information about how to add and remove effects, see [“Model”](#) on page 447 in the “Evaluate Designs” chapter.

Note: If any of the designs are supersaturated, meaning that the number of parameters to be estimated exceeds the number of runs, the Model outline lists only a set of effects that can be estimated.

Alias Terms

Add or remove effects to compare your designs for effects that might be active. The Alias Terms outline initially contains all two-factor interactions that are not in the Model outline. The effects in this outline impact the calculations in the Alias Matrix Summary and Absolute Correlations outline. See [“Alias Matrix Summary”](#) on page 496 and [“Absolute Correlations”](#) on page 498.

For more information about how to add and remove effects, see [“Alias Terms”](#) on page 448 in the “Evaluate Designs” chapter.

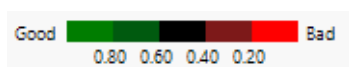
Compare Designs Window: Design Evaluation

- [“Power Analysis”](#)
- [“Prediction Variance Profile”](#)
- [“Fraction of Design Space Plot”](#)
- [“Relative Estimation Efficiency”](#)
- [“Alias Matrix Summary”](#)
- [“Absolute Correlations”](#)
- [“Design Diagnostics”](#)

Color Dashboard

Several of the Design Evaluation outlines show values colored according to a color bar. The colors are applied to diagnostic measures and they help you see which values (and designs) reflect good or bad behavior. You can edit the legend values to apply colors that reflect your definitions of good and bad behavior.

Figure 16.19 Color Dashboard



You can modify the color bar by selecting these two options in the red triangle menu for the outline or by right-clicking the color bar:

- Show Legend Values** Shows or hides the values that appear under the color bar.
- Edit Legend Values** Specify the values that define the colors.

Power Analysis

Power is the probability of detecting an active effect of a given size. The Power Analysis report helps you evaluate and compare the ability of your designs to detect effects of practical importance. For each of your designs, the Power Analysis report calculates the power of tests for the effects in the Model outline.

The Power Analysis report gives the power of tests for individual model parameters and for whole effects. It also provides a Power Plot and a Power versus Sample Size plot.

Power depends on the number of runs, the significance level, and the estimated error variation. For more information about how power is calculated, see [“Power Calculations”](#) on page 772 in the “Technical Details” appendix.

Figure 16.20 Power Analysis Outline for Three Designs

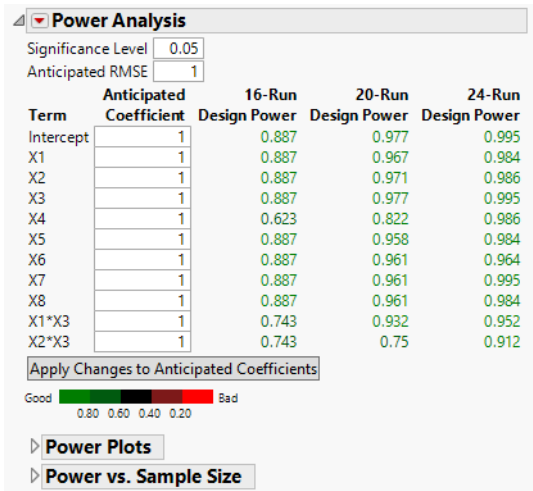


Figure 16.20 shows the Power Analysis outline for the three designs constructed in [“Designs of Different Run Sizes”](#) on page 476. Two two-way interactions have been added to the Model outline.

Power Analysis Report

When you specify values for the Significance Level and Anticipated RMSE, they are used to calculate the power of the tests for the model parameters. Enter coefficient values that reflect differences that you want to detect as Anticipated Coefficients. To update the results for all designs, click **Apply Changes to Anticipated Coefficients**.

Significance Level The probability of rejecting the hypothesis of no effect, if it is true. The power calculations update immediately when you enter a value.

Anticipated RMSE An estimate of the square root of the error variation. The power calculations update immediately when you enter a value.

The power values are colored according to a color gradient that appears under the **Apply Changes to Anticipated Coefficients** button. You can control the color legend using the options in the Power Analysis red triangle menu. See [“Color Dashboard”](#) on page 487.

For more information about the Power Plots, see [“Power Plot”](#) on page 490.

Note: If the design is supersaturated, meaning that the number of parameters to be estimated exceeds the number of runs, the Power Analysis outline lists only a set of effects that can be estimated.

Tests for Individual Parameters

The Term column contains a list of model terms. For each term, the Anticipated Coefficient column contains a value for that term. The Power value is the power of a test that the coefficient for the term is zero if the true value of the coefficient is given by the Anticipated Coefficient, given the design, and the terms in the Model outline.

Term The model term associated with the coefficient being tested.

Anticipated Coefficient A value for the coefficient associated with the model term. This value is used in the calculations for Power. When you set a new value in the Anticipated Coefficient column, click **Apply Changes to Anticipated Coefficients** to update the Power calculations.

Note: The anticipated coefficients have default values of 1 for continuous effects. They have alternating values of 1 and –1 for categorical effects.

Power The probability of rejecting the null hypothesis of no effect when the true coefficient value is given by the specified Anticipated Coefficient.

- For a coefficient associated with a numeric factor, the change in the mean response from the high to low setting (based on the model) is twice the coefficient value.

- For a coefficient associated with a categorical factor, the change in the mean response (based on the model) across the levels of the factor equals twice the absolute value of the anticipated coefficient. For more information on power for categorical factors, see [“Tests for Categorical Effects with More Than Two Levels”](#) on page 490

Calculations use the specified Significance Level and Anticipated RMSE. For more information about the power calculation, see [“Power for a Single Parameter”](#) on page 772 in the “Technical Details” appendix.

Apply Changes to Anticipated Coefficients When you set a new value in the Anticipated Coefficient column, click **Apply Changes to Anticipated Coefficients** to update the Power values.

Tests for Categorical Effects with More Than Two Levels

If your model contains a categorical effect with more than two levels, then the following columns appear below the Apply Changes to Anticipated Coefficients button:

Effect The categorical effect.

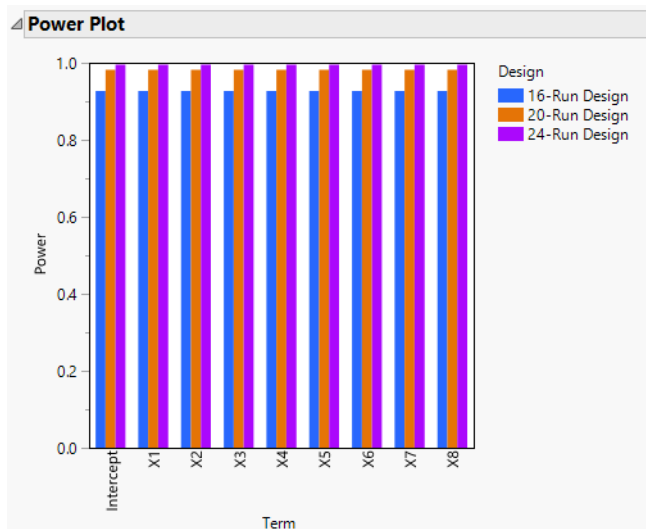
Power The power calculation for a test of no effect. The null hypothesis for the test is that all model parameters corresponding to the effect are zero. The difference to be detected is defined by the values in the Anticipated Coefficient column that correspond to the model terms for the effect. The power calculation reflects the differences in response means determined by the anticipated coefficients.

Calculations use the specified Significance Level and Anticipated RMSE. For more information about the power calculation, see [“Power for a Categorical Effect”](#) on page 774 in the “Technical Details” appendix.

Power Plot

The Power Plot shows the power values from the Power Analysis in graphical form. The plot shows the power for each effect and for each design in a side-by-side bar chart.

Figure 16.21 Power Plot for Three Designs

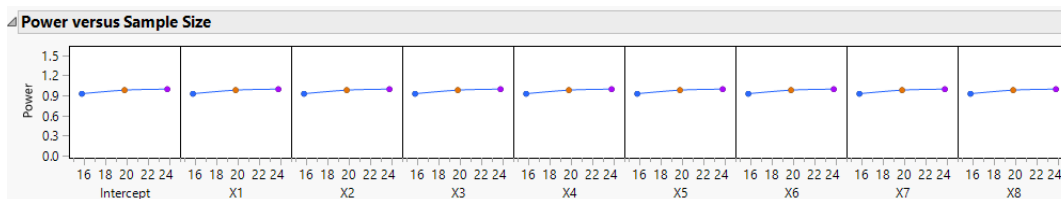


The Power Plot in Figure 16.21 is for the main effects of the three designs constructed in “Designs of Different Run Sizes” on page 476.

Power versus Sample Size

The Power versus Sample Size plots appears only when the designs that you are comparing differ in run size. The plots enables you to see how sample size affects power for each effect in the model. It conveys the same information as is in the Power Plots graph, but in a different format. The power values at integer sample sizes are connected with line segments.

Figure 16.22 Power versus Sample Size Profiler for Three Designs

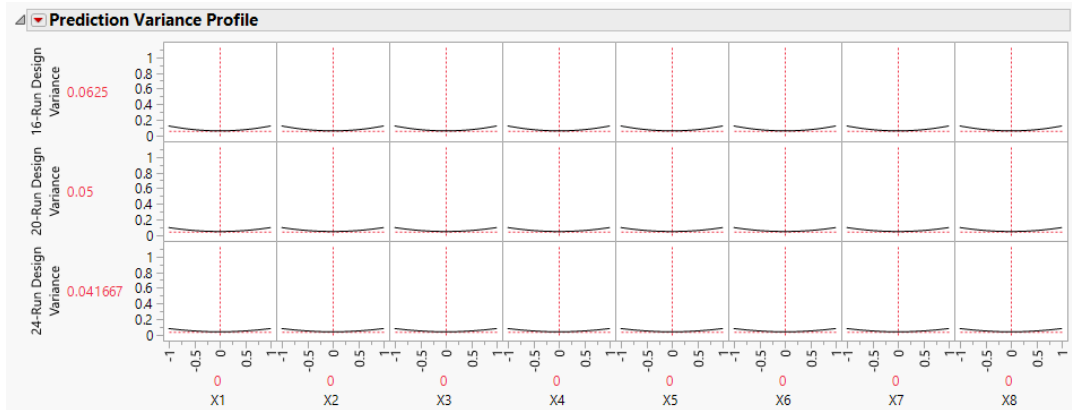


Prediction Variance Profile

The Prediction Variance Profile outline shows profilers of the relative variance of prediction for each design being compared. Each plot shows the relative variance of prediction as a function of each factor at fixed values of the other factors.

To find the maximum value of the relative prediction variance over the design space for the reference design, select the **Maximize Variance** option from the red triangle next to Prediction Variance Profile. See “[Maximize Variance](#)” on page 493.

Figure 16.23 Prediction Variance Profile for Three Designs



The Prediction Variance Profile plot in Figure 16.23 is for the three designs constructed in “[Designs of Different Run Sizes](#)” on page 476. Two two-way interactions, $X1 \times X3$ and $X2 \times X3$, have been added to the Model outline. The initial value for each continuous factor in the plot is the midpoint of its design settings. The Variance values to the left indicate that, as the number of runs increases, the variance decreases at the center point.

Relative Prediction Variance

For given settings of the factors, the prediction variance is the product of the error variance and a quantity that depends on the design and the factor settings. Before you run your experiment, the error variance is unknown, so the prediction variance is also unknown. However, the ratio of the prediction variance to the error variance is not a function of the error variance. This ratio, called the *relative prediction variance*, depends only on the design and the factor settings. Consequently, the relative variance of prediction can be calculated before acquiring the data. See “[Relative Prediction Variance](#)” on page 775 in the “Technical Details” appendix.

After you run your experiment and fit a least squares model, you can estimate the error variance using the mean squared error (MSE) of the model fit. You can estimate the actual variance of prediction at any setting by multiplying the relative variance of prediction at that setting.

Ideally, the prediction variance is small throughout the design space. Generally, the error variance drops as the sample size increases. In comparing designs, a design with lower prediction variance on average is preferable.

Maximize Variance

You can also evaluate a design or compare designs in terms of the maximum relative prediction variance. Select the **Maximize Variance** option from the red triangle next to Prediction Variance Profile. JMP uses a desirability function that maximizes the relative prediction variance. The value of the Variance in the Prediction Variance Profile is the worst (least desirable from a design point of view) value of the relative prediction variance.

Figure 16.24 Prediction Variance Profile Showing Maximum Variance for Three Designs

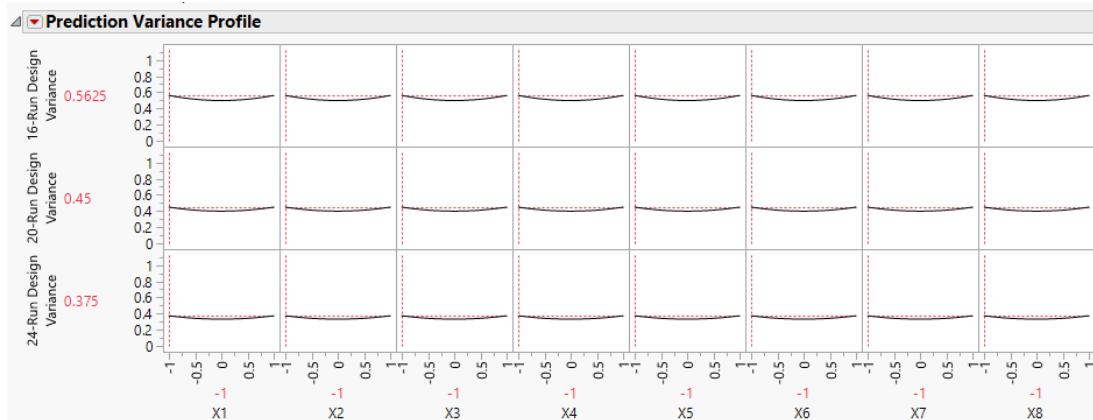


Figure 16.24 shows the Prediction Variance Profile after Maximize Variance was selected for the three designs constructed in “[Designs of Different Run Sizes](#)” on page 476. As expected, the maximum relative prediction variance decreases as the run size increases. The plot also shows values of the factors that give this worst-case relative variance. However, keep in mind that many settings can lead to this same maximum relative variance.

Fraction of Design Space Plot

The Fraction of Design Space Plot shows the proportion of the design space over which the relative prediction variance lies below a given value.

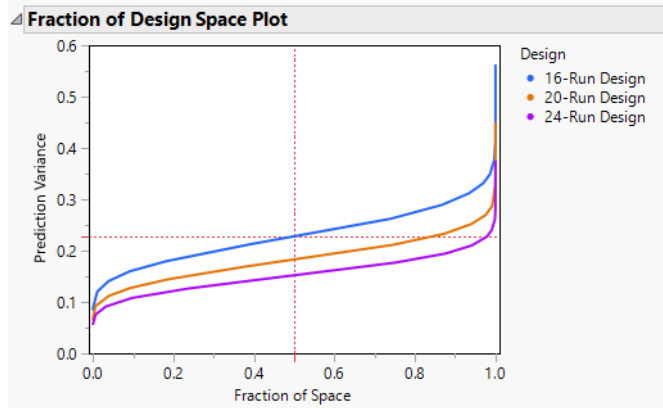
Figure 16.25 Fraction of Design Space Plot for Three Designs

Figure 16.25 shows the Fraction of Design Space plot for the three designs constructed in “[Designs of Different Run Sizes](#)” on page 476. Note the following:

- The X axis in the plot represents the proportion of the design space, ranging from 0 to 100%.
- The Y axis represents relative prediction variance values.
- For a point (x, y) that falls on a given curve, the value x is the proportion of design space with variance less than or equal to y .
- Red dotted crosshairs mark the value that bounds the relative prediction variance for 50% of design space for the reference design.

Figure 16.25 shows that the relative prediction variance for the 24-run design is uniformly smaller than for the other two designs. The 20-run design has uniformly smaller prediction variance than the 16-run design. The red dotted crosshairs indicate that the relative prediction variance for the 20-run design is less than about 0.23 over about 50% of the design space.

You can use the crosshairs tool to find the maximum relative prediction variance that corresponds to any Fraction of Space value. For example, use the crosshairs tool to see that for the 24-run design, 90% of the prediction variance values are below approximately 0.20.

Note: Plots for the same design might vary slightly, since Monte Carlo sampling of the design space is used in constructing the Fraction of Design Space Plot.

Relative Estimation Efficiency

The Relative Estimation Efficiency report compares designs in terms of the standard errors of parameter estimates for parameters in the assumed model. The standard errors control the length of confidence intervals for the parameter estimates. This report provides an efficiency ratio and the relative standard errors.

The relative estimation efficiency values are colored according to a color gradient shown under the table of relative estimation efficiency values. You can control the color legend using the options in the Relative Estimation Efficiency red triangle menu. See [“Color Dashboard”](#) on page 487.

Figure 16.26 Relative Estimation Efficiency Comparing Two Split-Plot Designs

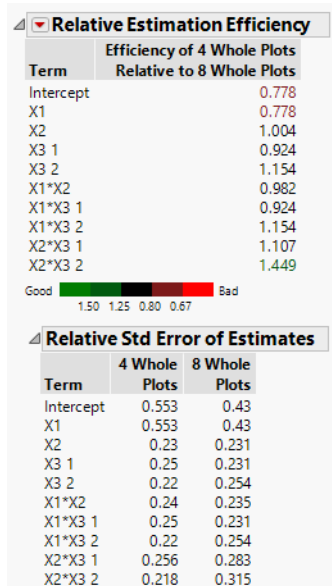


Figure 16.26 shows the Relative Estimation Efficiency outline for the split-plot designs compared in [“Split Plot Designs with Different Numbers of Whole Plots”](#) on page 480.

Relative Estimation Efficiency

For a given term, the estimation efficiency of the reference design relative to a comparison design is the relative standard error of the term for the comparison design divided by the relative standard error of the term for the reference design. A value less than one indicates that the reference design is not as efficient as the comparison design. A value greater than one indicates that it is more efficient.

Relative Standard Error of Estimates

The Relative Std Error of Estimates report gives the ratio of the standard deviation of a parameter's estimate to the error standard deviation. These values indicate how large the standard errors of the model's parameter estimates are, relative to the error standard deviation. For the i^{th} parameter estimate, the Relative Std Error of Estimate is defined as follows:

$$SE = \sqrt{(\mathbf{X}'\mathbf{X})_{ii}^{-1}}$$

where:

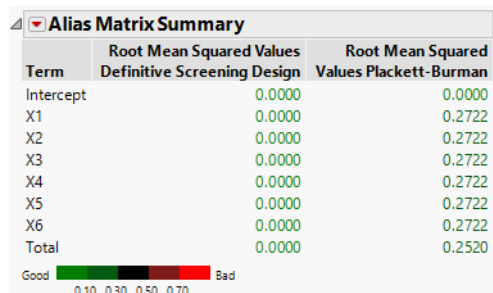
$\mathbf{X}$ is the model matrix defined in [“The Alias Matrix”](#) on page 771 in the “Technical Details” appendix, and $(\mathbf{X}'\mathbf{X})_{ii}^{-1}$ is the i^{th} diagonal entry of $(\mathbf{X}'\mathbf{X})^{-1}$.

Alias Matrix Summary

The alias matrix addresses the issue of how terms that are not included in the model affect the estimation of the model terms, if they are indeed active. In the Alias Terms outline, you list potentially active effects that are not in your assumed model but that might bias the estimates of model terms. The alias matrix entries represent the degree of bias imparted to model parameters by the Alias Terms effects. See [“Alias Terms”](#) on page 487 and [“Alias Matrix”](#) on page 497.

The Alias Matrix Summary table lists the terms in the assumed model. These are the terms that correspond to effects listed in the Model outline. Given a design, for each entry in the Term column, the square root of the sum of the squared alias matrix entries for the terms corresponding to effects in the Alias Terms outline is computed. This value is reported in the Root Mean Squared Values column for the given design. For an example, see [“Example of Calculation of Alias Matrix Summary Values”](#) on page 497.

Note: The Alias Matrix Summary report appears only if there are effects in the Alias Terms list.

Figure 16.27 Alias Matrix Summary for Two Designs

Term	Root Mean Squared Values	Root Mean Squared
	Definitive Screening Design	Values Plackett-Burman
Intercept	0.0000	0.0000
X1	0.0000	0.2722
X2	0.0000	0.2722
X3	0.0000	0.2722
X4	0.0000	0.2722
X5	0.0000	0.2722
X6	0.0000	0.2722
Total	0.0000	0.2520

Good 0.10 0.30 0.50 0.70 Bad

Figure 16.27 shows the Alias Matrix Summary report for the Plackett-Burman and Definitive Screening designs constructed in “[Designs of Same Run Size](#)” on page 471, with only main effects in the Model outline. All two-factor interactions are in the Alias Terms list. The table shows that, for the Definitive Screening Design, main effects are uncorrelated with two-factor interactions.

The Root Mean Squares Values are colored according to a color gradient shown under the Alias Matrix Summary table. You can control the color legend using the options in the Alias Matrix Summary red triangle menu. See “[Color Dashboard](#)” on page 487.

Alias Matrix

The rows of the Alias Matrix are the terms corresponding to the model effects listed in the Model outline. The columns are terms corresponding to effects listed in the Alias Terms outline. The entry in a given row and column indicates the degree to which the alias term affects the parameter estimate corresponding to the model term.

In evaluating your design, you ideally want one of two situations to occur relative to any entry in the Alias Matrix. Either the entry is small or, if it is not small, the effect of the alias term is small so that the bias is small. If you suspect that the alias term might have a substantial effect, then that term should be included in the model or you should consider an alias-optimal design. In fact, alias-optimality is driven by the squared values of the alias matrix.

For additional background on the Alias Matrix, see “[The Alias Matrix](#)” on page 771 in the “Technical Details” appendix. See also Lekivetz, R. (2014).

Example of Calculation of Alias Matrix Summary Values

This example illustrates the calculation of the values that appear in the Alias Matrix Summary outline. In this example, you compare the two designs assuming that only main effects are active.

1. Select **Help > Sample Data**, click **Open the Sample Scripts Directory**, and select **Compare Same Run Size.jsl**.

2. Right-click in the script window and select **Run Script**.

Two 13-run design tables are constructed:

- Definitive Screening Design
- Plackett-Burman

You are interested only in the Plackett-Burman design. This is the active table.

3. From the Plackett-Burman table, select **DOE > Design Diagnostics > Evaluate Design**.
4. Select X1 through X6 and click **X, Factor**.
5. Click **OK**.
6. Open the Alias Terms outline to confirm that all two-factor interactions are in the Alias Terms list.
7. Open the Alias Matrix outline.

For each model term listed in the Effect column, the entry in that row for a given column indicates the degree to which the alias term affects the parameter estimate corresponding to the model term.

For example, to obtain the Alias Matrix Summary entry in Figure 16.27 corresponding to X1, square the terms in the row for X1 in the Alias Matrix, average these, and take the square root. You obtain 0.2722.

Absolute Correlations

The Absolute Correlations report summarizes information about correlations between model terms and alias terms.

Figure 16.28 Absolute Correlations Report for Three Designs

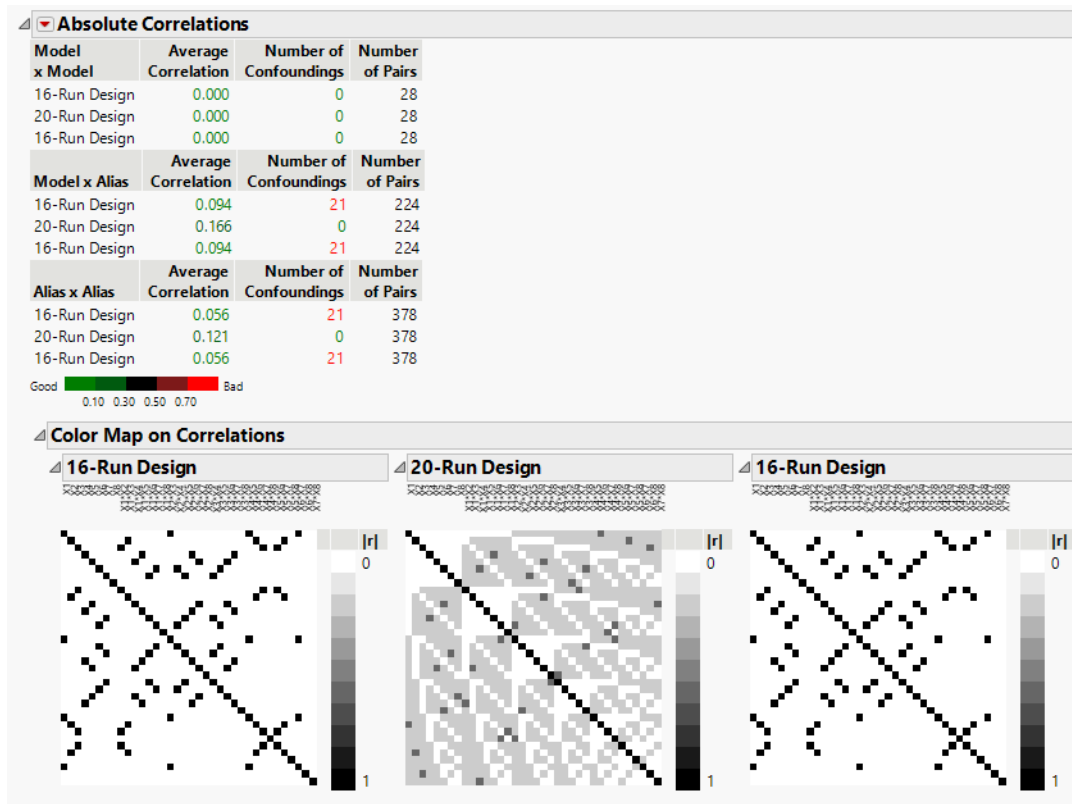


Figure 16.28 shows the Absolute Correlations report for the three designs constructed in “[Designs of Different Run Sizes](#)” on page 476, with only main effects in the Model outline.

Absolute Correlations Table

The table in the Absolute Correlations report is divided into three sections:

- Model x Model considers correlations between terms corresponding to effects in the Model list.
- Model x Alias considers correlations between terms corresponding to effects in the Model list and terms corresponding to effects in the Alias list.
- Alias x Alias considers correlations between terms corresponding to effects in the Alias list.

Note: If there are no alias terms, only the Model x Model section appears.

For each section of the report, the following are given:

Average Correlation The average of the correlations for all pairs of terms considered in this section of the report.

Number of Confoundings The number of pairs of terms consisting of confounded terms.

Number of Terms The total number of pairs of terms considered in this section of the report.

The values in the Absolute Correlations table are colored according to a color gradient shown under the table. You can control the color legend using the options in the Absolute Correlations red triangle menu. See [“Color Dashboard”](#) on page 487.

Color Map on Correlations

The Color Map on Correlations outline shows plots for each of the designs. The cells of the color map are identified above the map. There are cells for all terms that correspond to effects that appear in either the Model outline or the Alias Terms outline. Each cell is colored according to the absolute value of the correlation between the two terms.

By default, the absolute magnitudes of the correlations are represented by a blue to gray to red intensity color theme. In general terms, the color map for a good design shows a lot of blue off the diagonal, indicating orthogonality or small correlations between distinct terms. Large absolute correlations among effects inflate the standard errors of estimates.

To see the absolute value of the correlation between two effects, place your pointer over the corresponding cell. To change the color theme for the entire plot, right-click in the plot and select Color Theme.

Absolute Correlations and Color Map on Correlations Example

Figure 16.28 shows the Absolute Correlations report for the Plackett-Burman and Definitive Screening designs constructed in [“Designs of Different Run Sizes”](#) on page 476. The Model outline contains only main effects, so the Alias Terms outline contains all two-factor interactions. All main effects and two-way interactions are shown in the color maps.

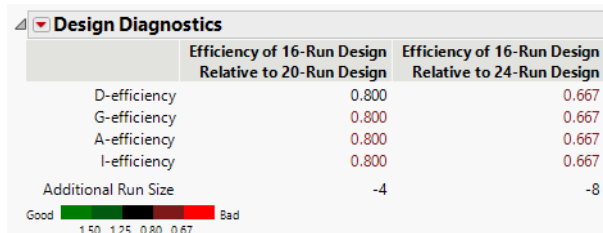
In the Color Map on Correlations for the 16-run design, the red cells off the main diagonal indicate that the corresponding terms have correlation one and therefore are completely confounded. There are nine instances where model terms (main effects) are confounded with alias terms (two factor interactions), and six instances where alias terms are confounded with each other. This is shown in the report under Pairwise Confoundings.

The color maps for the 20- and 24-run designs have no off-diagonal cells that are solid red. It follows that these designs show no instances of confounding between any pair of main or two-way interaction effects. However, it is interesting to note that the 20- and 24-run designs both have a higher Average Correlation for Model x Alias terms than does the 16-run design. Although the 16-run design shows confounding, the average amount of correlation is less than for the 20- and 24-run designs.

Design Diagnostics

The Design Diagnostics outline shows *D*-, *G*-, *A*-, and *I*-efficiencies for the reference design relative to the comparison designs. It also shows the Additional Run Size. Given two designs, the one with the higher relative efficiency measure is better.

Figure 16.29 Design Diagnostics for Three Designs



	Efficiency of 16-Run Design Relative to 20-Run Design	Efficiency of 16-Run Design Relative to 24-Run Design
D-efficiency	0.800	0.667
G-efficiency	0.800	0.667
A-efficiency	0.800	0.667
I-efficiency	0.800	0.667
Additional Run Size	-4	-8

Good 1.50 1.25 0.80 0.67 Bad

Figure 16.29 shows the Design Diagnostics report for the three designs constructed in “[Designs of Different Run Sizes](#)” on page 476, with only main effects in the Model outline.

The values in the Design Diagnostics table are colored according to a color gradient shown under the table. You can control the color legend using the options in the Design Diagnostics red triangle menu. See “[Color Dashboard](#)” on page 487.

Efficiency and Additional Run Size

Relative efficiencies for each of *D*-, *G*-, *A*-, and *I*-efficiency are shown in the Design Diagnostics report. These are obtained by computing each design’s efficiency value and then taking the appropriate ratio. The descriptions of the relative efficiency measures are given in “[Relative Efficiency Measures](#)” on page 502.

Additional Run Size is the number of runs in the reference design minus the number of runs in the comparison design. If your reference design has more runs than your comparison design, then the Additional Run Size tells you how many additional runs you need to achieve the efficiency of the reference design.

Relative Efficiency Measures

Notation

- $\mathbf{X}$ is the model matrix
- p is the number of terms, including the intercept, in the model
- $Var(\hat{y}|\mathbf{x})$ is the relative prediction variance at the point $\mathbf{x}$. See “Relative Prediction Variance” on page 775 in the “Technical Details” appendix.

Relative Efficiencies

The relative efficiency of the reference design (*Ref*) to the comparison design (*Comp*) is given by the following expressions:

D Efficiency Eff_{Ref} / Eff_{Comp} , where Eff for each design is given as follows:

$$Eff = |\mathbf{X}'\mathbf{X}|^{1/p}$$

G Efficiency Eff_{Comp} / Eff_{Ref} , where Eff for each design is given as follows:

$$Eff = Var(\hat{y}|\mathbf{x})_{max} = \underset{\mathbf{x} \text{ in } D}{\text{maximum}}[\mathbf{x}'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}]$$

Here, D denotes the design region.

Note: G-Efficiency is calculated using Monte Carlo sampling of the design space. The reported value is based on the larger of $Var(\hat{y}|\mathbf{x})_{max}$ or the prediction variance from the Monte Carlo sampling. Therefore, calculations for the same design might vary slightly.

A Efficiency Eff_{Comp} / Eff_{Ref} , where Eff for each design is given as follows:

$$Eff = \text{Trace}[(\mathbf{X}'\mathbf{X})^{-1}]$$

I Efficiency Eff_{Comp} / Eff_{Ref} , where Eff for each design is given as follows:

$$Eff = \frac{\int \mathbf{x}'(\mathbf{X}'\mathbf{X})^{-1}\mathbf{x}d\mathbf{x}}{\int d\mathbf{x}}$$

For more information about the calculation, see Section 4.3.5 in Goos and Jones (2011).

Compare Designs Options

Advanced Options > Split Plot Variance Ratio Specify the ratio of the variance of the random whole plot and the subplot variance (if present) to the error variance. Before setting this value, you must define a hard-to-change factor for your split-plot design, or hard and very-hard-to-change factors for your split-split-plot design. Then you can enter one or two positive numbers for the variance ratios, depending on whether you have specified a split-plot or a split-split-plot design.

Advanced Options > Set Delta for Power Specify the difference in the mean response that you want to detect for model effects. See [“Set Delta for Power”](#) on page 116 in the “Custom Designs” chapter.

Prospective Sample Size and Power

Use the Sample Size and Power platform to investigate the impact of sample size on your ability to answer specific questions when planning experiments or studies.

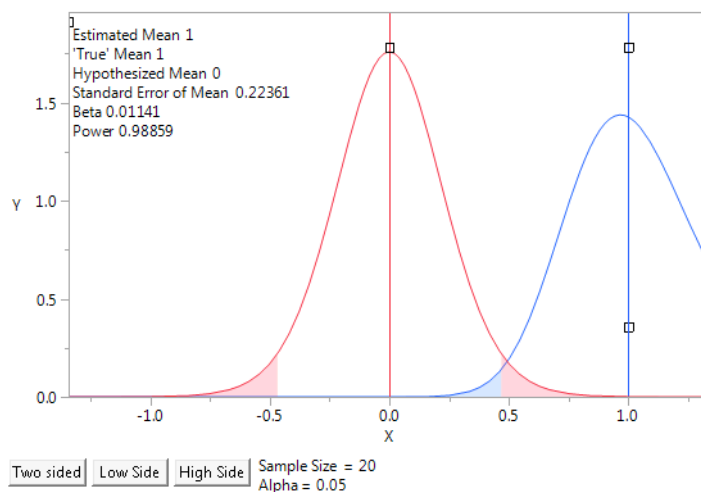
- How many units should I test?
- Will I be able to detect a difference in my treatment means?
- How many units must I test to estimate failure time?

The Sample Size and Power platform is a collection of calculators. The calculators have fields for data entry and calculated values. You can select values to specify and those to calculate. For example, you might specify power and the effect size that you want to detect and then calculate the necessary sample size. Alternatively, you might specify a sample size and power to calculate the effect size that you can detect.

Sample size calculations for hypothesis tests are based on the trade off between Type I and Type II errors. The Type I error, or α , is the probability of rejecting the null hypothesis when it is true. The Type II error, or β , is the probability of not rejecting the null hypothesis when it is false. Generally you would like α and β to be small. Power, or $1 - \beta$, is the probability of rejecting the null hypothesis when it is false.

For more information about power and sample size in JMP, see Barker (2011).

Figure 17.1 Power Animation for One Sample Mean



Contents

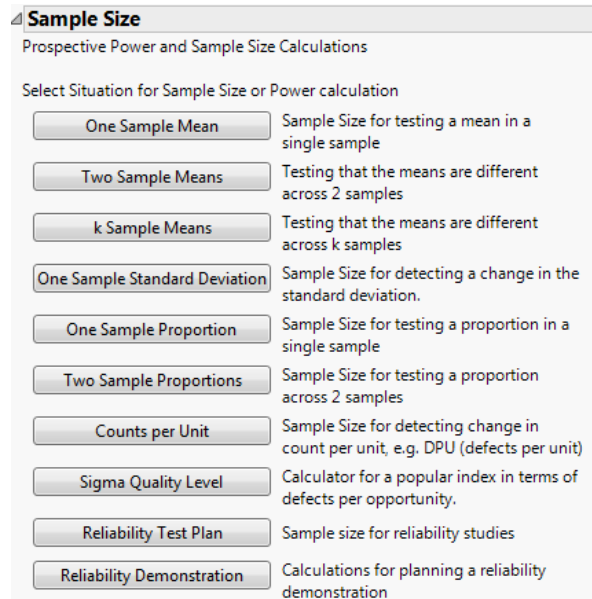
Launch the Sample Size and Power Platform	508
One Sample Mean Calculator	509
Examples of the One Sample Mean Calculator	510
One Sample Mean Calculator Fields	514
Statistical Details for the One Sample Mean Calculator	515
Two Sample Means Calculator	516
Example of the Two Sample Means Calculator	516
Two Sample Means Calculator Fields	517
Statistical Details for the Two Sample Means Calculator	518
k Sample Means Calculator	519
Example of the k Sample Means Calculator	519
k Sample Means Calculator Fields	521
Statistical Details for the k Sample Means Calculator	522
One Sample Standard Deviation Calculator	523
Example of the One Sample Standard Deviation Calculator	523
One Sample Standard Deviation Calculator Fields	524
Statistical Details for the One Sample Standard Deviation Calculator	525
One Sample Proportion Calculator	526
Example of the One Sample Proportion Calculator	527
One Sample Proportion Calculator Fields	528
Statistical Details for the One Proportion Calculator	529
Two Sample Proportions Calculator	530
Example of the Two Sample Proportions Calculator	530
Two Sample Proportions Calculator Fields	531
Statistical Details for the Two Proportions Calculator	532
Counts per Unit Calculator	533
Example of the Counts per Unit Calculator	533
Counts per Unit Calculator Fields	534
Statistical Details for the Counts per Unit Calculator	535
Sigma Quality Level Calculator	535
Example of the Sigma Quality Level Calculator	536
Example Computing the Number of Defects	536
Sigma Quality Level Calculator Fields	537
Statistical Details for the Sigma Quality Calculator	537
Reliability Test Plan Calculator	538
Example of the Reliability Test Plan Calculator	538
Reliability Test Plan Calculator Fields	539
Statistical Details for the Reliability Test Plan Calculator	541
Reliability Demonstration Calculator	542

Example of the Reliability Demonstration Calculator	542
Reliability Demonstration Calculator Fields	544
Statistical Details for the Reliability Demonstration Calculator	545

Launch the Sample Size and Power Platform

The Sample Size and Power platform is a collection of calculators for planning experiments or studies. Launch the Sample Size and Power platform by selecting **DOE > Design Diagnostics > Sample Size and Power**.

Figure 17.2 Sample Size and Power Launch Window



Choose from the following calculators:

- [“One Sample Mean Calculator”](#) on page 509
- [“Two Sample Means Calculator”](#) on page 516
- [“k Sample Means Calculator”](#) on page 519
- [“One Sample Standard Deviation Calculator”](#) on page 523
- [“One Sample Proportion Calculator”](#) on page 526
- [“Two Sample Proportions Calculator”](#) on page 530
- [“Counts per Unit Calculator”](#) on page 533
- [“Sigma Quality Level Calculator”](#) on page 535
- [“Reliability Test Plan Calculator”](#) on page 538
- [“Reliability Demonstration Calculator”](#) on page 542

Sample Size

One Mean

Testing if one mean is different from the hypothesized value.

Alpha

Std Dev

Extra Parameters

Supply two values to determine the third.
Enter one value to see a plot of the other two.

Difference to detect

Sample Size

Power

Examples of the One Sample Mean Calculator

In the following examples, suppose you are interested in demonstrating that the flammability of a new fabric being developed by your company has improved performance over current materials. Previous testing indicates that the standard deviation for time to burn of this fabric is 2 seconds.

Example of Sample Size Calculation

In this initial example, you would like to design an experiment that has 90% power to detect a difference of 1.5 seconds at a significance level of $\alpha = 0.05$. Use the One Sample Mean calculator to calculate the number of fabric samples you need to test.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click the **One Sample Mean** button.
3. Leave **Alpha** set to 0.05.
4. Enter 2 for **Std Dev**.
5. Leave **Extra Parameters** set to 0.
6. Enter 1.5 for **Difference to detect**.
7. Leave **Sample Size** blank.
8. Enter 0.9 for **Power**.
9. Click **Continue**.

Figure 17.4 One-Sample Mean Calculator

The screenshot shows a software window titled "Sample Size" with a sub-tab "One Mean". The window contains a form for calculating sample size and power. The form has two sections. The first section is for testing if one mean is different from the hypothesized value, with input fields for Alpha (0.05), Std Dev (2), and Extra Parameters (0). The second section is for supplying two values to determine the third, with input fields for Difference to detect (1.5), Sample Size (21), and Power (0.9). Below the input fields are three buttons: "Continue", "Back", and "Animation Script".

Testing if one mean is different from the hypothesized value.	
Alpha	0.05
Std Dev	2
Extra Parameters	0

Supply two values to determine the third.
Enter one value to see a plot of the other two.

Difference to detect	1.5
Sample Size	21
Power	0.9

Continue
Back
Animation Script

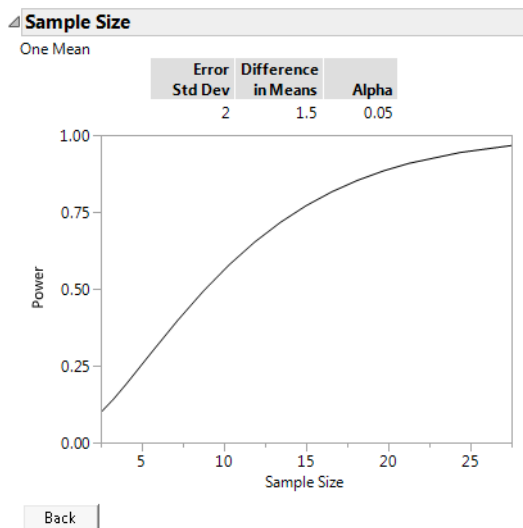
At a significance level of 0.05, 21 fabric samples are needed to have a 90% chance of detecting a significant difference of 1.5 seconds in the burn time.

Example of Power versus Sample Size Plot

To explore trade offs between sample size and power in your fabric experiment, use the plot of power versus sample size.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **One Sample Mean**.
3. Leave **Alpha** set to 0.05.
4. Enter 2 for **Std Dev**.
5. Leave **Extra Parameters** set to 0.
6. Enter 1.5 for the **Difference to detect**.
7. Leave **Sample Size** blank.
8. Leave **Power** blank.
9. Click **Continue** to launch the power by sample size plot.

Figure 17.5 Power by Sample Size



The plot shows a range of sample sizes for which the power varies from about 0.1 to about 0.95. You could reduce the number of tests in your experiment to about 15 and maintain power above 75%. However, if you ran only 10 tests, the power to detect a significant difference of 1.5 would drop to about 50%.

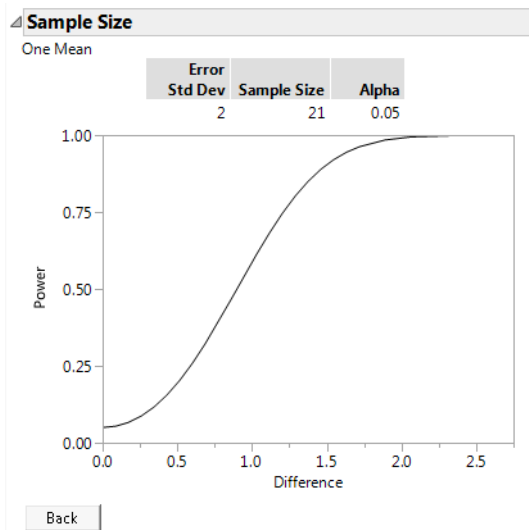
Tip: Use the crosshair tool to obtain sample size and power combinations from the plot.

Example of Power versus Difference Plot

To explore trade offs between power and the magnitude of the difference that you can detect with 21 observations in your fabric experiment, use the plot of power versus difference.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **One Sample Mean**.
3. Leave **Alpha** set to 0.05.
4. Enter 2 for **Std Dev**.
5. Leave **Extra Parameters** set to 0.
6. Leave **Difference to detect** blank.
7. Enter 21 for **Sample Size**.
8. Leave **Power** blank.
9. Click **Continue**.

Figure 17.6 Plot of Power by Difference to Detect for a Sample Size of 21



At a significance level of 0.05 and 21 observations, you can detect a difference of 1.5 seconds with 90% power. If the difference is only one second smaller, then with 21 fabric samples, you have about 50% power of detecting the difference.

Example of an Animation Script

Use an animation script to explore how changing the sample size affects power.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.

2. Click **One Sample Mean**.
3. Leave **Alpha** set to 0.05.

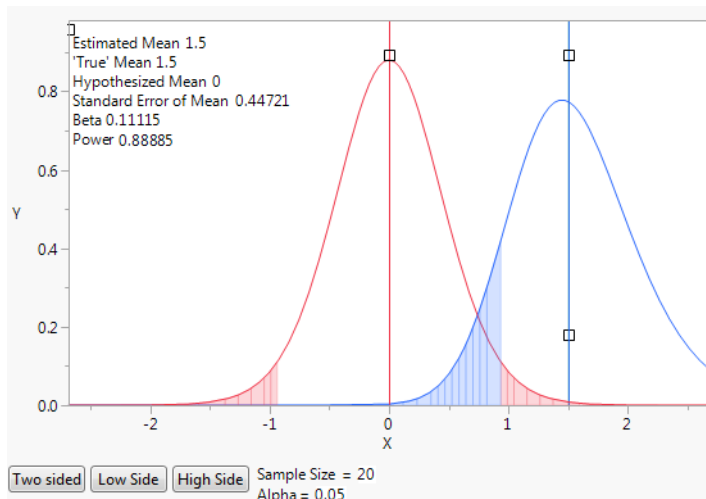
Tip: You can change the Alpha level after the animation is launched by clicking the Alpha value in the animation window.

4. Enter 2 for **Std Dev**.
5. Leave **Extra Parameters** set to 0.
6. Enter 1.5 as **Difference to detect**.
7. Leave **Sample Size** blank.

Tip: When you leave the sample size blank, the default sample size is set to 20. You will see the default of 20 after the animation script is launched. You can change the sample size before or after the animation is launched. To change it after the animation is launched, click the Sample Size value in the animation window.

8. Leave **Power** blank.
9. Click **Animation Script**.

Figure 17.7 Initial Animation Script to Illustrate Power



The initial animation plot shows two t -density curves:

- The red curve shows the t -distribution when the true mean is zero.
- The blue curve shows the t -distribution when the true mean is 1.5, which is the difference to be detected.

- The blue shading indicates the probability of committing a type II error. A type II error is the probability of not detecting a difference when there is a difference. The probability of a type II error is often denoted by β .
- The red shading indicates the probability of committing a type I error. A type I error is the probability of concluding that the difference in means is significant when there is no difference. The probability of a type I error is often denoted by α .

Select and drag the square handles to see the changes in statistics based on the positions of the curves. To change the values of Sample Size and Alpha, click their values beneath the plot.

By default, the animation shows a two sided test. Use the **Two Sided**, **Low Side**, and **High Side** buttons to toggle between not equal, less than, or greater than alternative hypotheses.

One Sample Mean Calculator Fields

Specify the following quantities:

Alpha The probability of a type I error, which is the probability of rejecting the null hypothesis when it is true. It is commonly referred to as the significance level of the test. The default alpha level is 0.05.

Tip: For a one-sided test, use $\alpha = \alpha * 2$. For a one sided test with $\alpha = 0.05$, use the two-sided calculator with $\alpha = 0.10$. The resulting values are those needed for a one-sided test with $\alpha = 0.05$.

Std Dev The assumed standard deviation. An estimate of the error standard deviation could be the root mean square error (RMSE) from a previous model fit.

Tip: Use a standard deviation of 1 to estimate the sample size needed to detect differences measured in standard deviation units.

Extra Parameters The number of parameters other than μ in the hypothesis test. This option can be used for multi-factor designs. Leave the default zero in this field for simple cases.

In a multi-factor design where effects are orthogonal, you can specify the number of additional model parameters here. For example, in a three-factor, two-level design with all three two-factor interactions, the number of additional parameters is five: two parameters for the other main effects, and three parameters for the interactions.

Specify two of the following values to calculate the third value, or specify one value to obtain a plot of the relationship between the other two:

Difference to Detect The smallest difference between the true mean and the hypothesized or reference mean you want to be able to declare statistically significant.

Sample Size The total number of observations (runs, experimental units, or samples) in your experiment.

Power The probability of rejecting the null hypothesis when it is false. With all other parameters fixed, power increases as sample size increases.

Calculator Buttons

Continue Evaluates the missing value when two parameters are specified, or launches a plot comparing two missing parameters if only one parameter is specified.

Back Returns to the previous Sample Size and Power launch window.

Animation Script Launches an interactive plot to illustrate and explore the relationship between power and the difference to detect. See [“Example of an Animation Script”](#) on page 512.

Statistical Details for the One Sample Mean Calculator

The one sample mean calculations are based on the traditional full versus reduced F test for the following hypothesis test:

$$H_0: \mu = \mu_0 \text{ vs } H_a: \mu \neq \mu_0$$

JMP calculates power as follows:

$$\Pr(\text{reject } H_0 \mid \mu = \mu_0 + \delta) = 1 - F\left(f_{1-\alpha}, 1, n-p-1, \frac{n\delta^2}{\sigma^2}\right)$$

where:

α is the significance level.

n is the sample size.

p is the number of extra parameters.

δ is the difference to detect.

$f_{1-\alpha}$ is the $(1 - \alpha)^{\text{th}}$ quantile of the $F(1, n-p-1)$ distribution.

$F(x, df_1, df_2, nc)$ is the cumulative distribution function of the non-central F distribution with degrees of freedom df_1 and df_2 and non-centrality parameter nc evaluated at x .

Because analytical solutions for δ and n do not exist, numerical solutions are used to solve for them.

For more information about calculations in JMP, see Barker (2011, Section 2.1).

Two Sample Means Calculator

Use the Two Sample Means calculator to evaluate sample size for a hypothesis test about two means. Explore questions such as how large a treatment effect is detectable for a given sample size, significance level, desired power, and assumed variability in the data. Sample size and power are associated with the following hypothesis test:

$$H_0: \mu_1 - \mu_2 = 0$$

versus the two-sided alternative:

$$H_a: \mu_1 - \mu_2 \neq 0$$

where μ_1 and μ_2 are the true means of the two populations. It is assumed that the populations of interest are normally distributed and that you want to detect a difference of δ between the means.

Example of the Two Sample Means Calculator

Suppose you are comparing two groups and want to detect a 1 standard deviation difference between the group means at a 0.05 significance level with 80% power.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **Two Sample Means**.
3. Leave **Alpha** as 0.05.
4. Enter 1 for **Std Dev**.
5. Leave **Extra Parameters** as 0.
6. Enter 1 as **Difference to detect**.
7. Leave **Sample Size** blank.
8. Enter 0.8 for **Power**.
9. Click **Continue**.

Figure 17.8 Two Sample Means Calculator Result

Sample Size

Two Means

Testing if two means are different from each other.

Alpha

Std Dev

Extra Parameters

Supply two values to determine the third.
Enter one value to see a plot of the other two.

Difference to detect

Sample Size

Power

Sample Size is the total sample size; per group would be n/2

The sample size is calculated as 34. With 17 samples in each group, you have an 80% chance of detecting a 1 standard deviation difference between the two sample means at a significance level of $\alpha = 0.05$.

Two Sample Means Calculator Fields

Specify the following quantities:

Alpha The probability of a type I error, which is the probability of rejecting the null hypothesis when it is true. It is commonly referred to as the significance level of the test. The default alpha level is 0.05.

Std Dev The assumed standard deviation where both groups are assumed to have equal standard deviation.

Extra Parameters The number of parameters other than μ_1 and μ_2 in the hypothesis test. This option can be used for multi-factor designs. Leave the default zero in this field for simple cases.

In a multi-factor balanced design, the number of additional model parameters can be specified here. For example, in a three-factor two-level design with all three two-factor interactions, the number of additional parameters is four: one parameter for the other main effects, and three parameters for the interactions.

Specify two of the following parameters to calculate the third value, or specify one value to obtain a plot of the relationship between the other two parameters:

Difference to Detect The smallest difference between the two means that you want to be able to declare statistically significant.

Sample Size The total number of observations (runs, experimental units, or samples) in your experiment. Use half of this sample size in each group.

Power The probability of rejecting the null hypothesis when it is false. With all other parameters fixed, power increases as sample size increases.

Calculator Buttons

Continue Evaluates the missing value when two parameters are specified, or launches a plot comparing two missing parameters if only one parameter is specified.

Back Returns to the previous Sample Size and Power launch window.

Statistical Details for the Two Sample Means Calculator

The two sample mean calculations are based on the traditional full versus reduced F test for the following hypothesis test:

$$H_0: \mu_1 - \mu_2 = 0 \text{ vs } H_a: \mu_1 - \mu_2 \neq 0$$

JMP calculates power as follows:

$$\text{power} = \Pr(\text{reject } H_0 \mid (\mu_1 - \mu_2 = \delta)) = 1 - F\left(f_{1-\alpha}, 1, 2n - p - 2, \frac{n\delta^2}{2\sigma^2}\right)$$

where:

α is the significance level.

n is the sample size per group.

p is the number of extra parameters.

δ is the difference to detect.

$f_{1-\alpha}$ is the $(1 - \alpha)^{\text{th}}$ quantile of the $F(1, 2n - p - 2)$ distribution.

$F(x, df_1, df_2, nc)$ is the cumulative distribution function of the non-central F distribution with degrees of freedom df_1 and df_2 and non-centrality parameter nc evaluated at x .

Because analytical solutions for δ and n do not exist, numerical solutions are used to solve for them.

For more information about calculations in JMP, see Barker (2011, Section 2.2).

k Sample Means Calculator

Use the k Sample Means calculator to determine an appropriate sample size for a hypothesis test about two to ten means. Sample size and power are associated with the following hypothesis test:

$$H_0: \mu_1 = \mu_2 = \dots = \mu_k$$

versus the two-sided alternative:

$$H_a: \text{not all means equal}$$

where:

$$X_{ij} \sim N(\mu_j, \sigma^2) \text{ for } i = 1, \dots, n, j = 1, \dots, k$$

Example of the k Sample Means Calculator

Suppose you have 4 observations from each of 4 groups and you want to test to see whether all of the group means are equal. You expect the group means to be 10, 11, 12, and 13 with a standard deviation of 0.9. Calculate the power of detecting a difference in one or more means at a 0.05 significance level.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **k Sample Means**.
3. Leave **Alpha** as 0.05.
4. Enter 0.9 for **Std Dev**.
5. Leave **Extra Parameters** as 0.
6. Enter 10, 11, 12, and 13 as the four levels of prospective means.
7. Enter 16 for **Sample Size**.
8. Leave **Power** blank.
9. Click **Continue**.

Figure 17.9 Prospective Power for 4 Means

Sample Size

k Means
Testing if there are differences among k means.

Alpha

Std Dev

Extra Parameters

Enter up to 10 Prospective Means showing separation across groups

10
11
12
13
.
.
.
.
.
.

Enter Power or Sample Size to get the other.
Enter neither to get a plot of Power vs. Sample Size

Sample Size

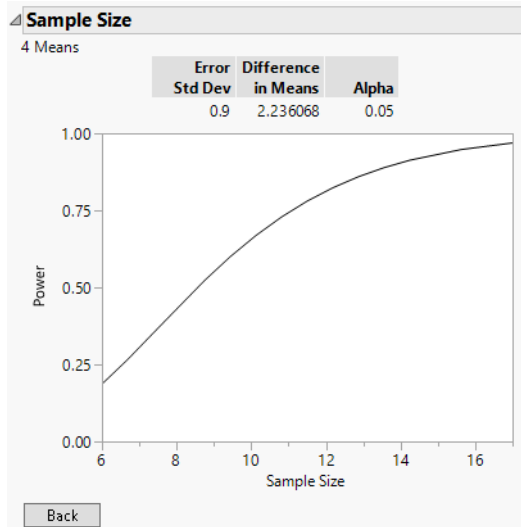
Power

Sample Size is the total sample size; per group would be n/k

The Power is calculated as approximately 0.95. This means that there is a 95% chance of detecting that at least one of the means is different at a 0.05 significance level, assuming that the population means are 10, 11, 12, and 13, with a standard deviation of 0.9 and a total sample size of 16.

10. Click the 16 for Sample Size and delete it.
11. Click the 0.95 for Power and delete it.
12. Click **Continue** to launch a plot of power versus sample size.

Figure 17.10 Prospective Power for k Means and Plot of Power by Sample Size



From the plot of power versus sample size, you can confirm that a sample size of 16 is acceptable. Place your cursor on the Sample Size axis and drag to extend the axis beyond 16. You can see that the increase in power for sample sizes above 16 is slight.

The plot also reports the difference in means, which is calculated as the square root of the sum of squared deviations from the grand mean. In this case, it is the square root of $(-1.5)^2 + (-0.5)^2 + (0.5)^2 + (1.5)^2$, which is the square root of 5. Therefore, the difference in means in this example is approximately 2.236.

k Sample Means Calculator Fields

Specify the following quantities:

Alpha The probability of a type I error, which is the probability of rejecting the null hypothesis when it is true. It is commonly referred to as the significance level of the test. The default alpha level is 0.05.

Std Dev The assumed standard deviation where all groups are assumed to have equal standard deviation.

Extra Parameters The number of parameters other than the μ in the hypothesis test. This option can be used for multi-factor designs. Leave the default 0 in this field in simple cases.

Prospective Means The assumed values for the means. At least two means must be specified.

Specify one of the following values to calculate the second, or leave both values blank to obtain a plot of the relationship between the two:

Sample Size The total number of observations (runs, experimental units, or samples) in your experiment. Use a sample size of n/k in each group, where n is the total number of observations.

Power The probability of rejecting the null hypothesis when it is false. With all other parameters fixed, power increases as sample size increases.

Calculator Buttons

Continue Evaluates the missing value when one parameter is specified, or launches a plot comparing two parameters if no parameters are specified.

Back Returns to the previous Sample Size and Power launch window.

Statistical Details for the k Sample Means Calculator

The k sample mean calculations are based on the following:

$$\text{power} = 1 - F\left(f_{1-\alpha, 1k-1, 2kn-p-2k}, \frac{n \sum_{j=1}^k (\mu_j - \bar{\mu})^2}{\sigma^2}\right)$$

where:

α is the significance level

σ is the common standard deviation among the k groups.

n is the sample size per group.

p is the number of extra parameters.

$\bar{\mu}$ is the mean of the k group prospective means

$f_{1-\alpha}$ is the $(1 - \alpha)^{\text{th}}$ quantile of the $F(k - 1, kn - p - k)$ distribution.

$F(x, df_1, df_2, nc)$ is the cumulative distribution function of the non-central F distribution with degrees of freedom df_1 and df_2 and non-centrality parameter nc evaluated at x .

Numerical solutions are used to solve for n .

Difference in Means

The difference in means is calculated as follows:

$$\sqrt{\sum_{j=1}^k (\mu_j - \bar{\mu})^2}$$

For more information about calculations in JMP, see Barker (2011, Section 2.3).

One Sample Standard Deviation Calculator

Use the One Sample Standard Deviation calculator to evaluate sample size for a hypothesis test about one standard deviation. Sample size and power are associated with the following hypothesis test:

$$H_0: \sigma = \sigma_0$$

versus the one-sided alternative:

$$H_a: \sigma > \sigma_0 \text{ or } H_a: \sigma < \sigma_0$$

where σ is the population standard deviation and σ_0 is the hypothesized value. The difference to detect is an amount, δ , away from σ_0 that one considers as important to detect based on a set of samples.

Example of the One Sample Standard Deviation Calculator

Note: This example is from the online manual of The National Institute of Standards and Technology (NIST). You can access the NIST manual examples at <https://www.itl.nist.gov/div898/handbook/prc/section2/prc232.htm>.

The variance for resistivity measurements on a lot of silicon wafers is claimed to be 100 ohm-cm squared. The buyer is unwilling to accept a lot if the variance is greater than 155 ohm-cm squared. How many wafers must you test to estimate the lot variance with precision to detect an increase in 55 ohm-cm squared from the target of 100 ohm-cm squared? In terms of standard deviation, the hypothesized standard deviation, σ_0 , is 10 (the square root of 100) and σ is 12.4499 (the square root of $100 + 55 = 155$). The difference to detect is $12.4499 - 10 = 2.4499$. The desired power is 0.99 and the significance level is 0.05.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **One Sample Standard Deviation**.

3. Leave **Alpha** as 0.05.
4. Enter 10 for **Hypothesized Standard Deviation**.
5. Select Larger for **Alternative Standard Deviation**.
6. Enter 2.4499 for **Difference to Detect**.
7. Leave **Sample Size** blank.
8. Enter 0.99 for **Power**.
9. Click **Continue**.

Figure 17.11 One-Sample Standard Deviation Calculator

Sample Size
One Sample Standard Deviation
Sample Size for detecting a difference in the standard deviation.

Alpha

Hypothesized Standard Deviation

Alternative Standard Deviation

Supply two values to determine the third.

Difference to detect

Sample Size

Power

You must test 171 wafers in order to have a 99% chance of detecting an increase in the standard deviation of 2.4499 from a standard deviation of 10, with an alpha of 0.05.

One Sample Standard Deviation Calculator Fields

Specify the following:

Alpha The probability of a type I error, which is the probability of rejecting the null hypothesis when it is true. It is commonly referred to as the significance level of the test. The default alpha level is 0.05.

Hypothesized Standard Deviation The hypothesized or baseline standard deviation to which the sample standard deviation is compared.

Alternative Standard Deviation The direction of the change that you want to detect, either Larger or Smaller.

Note: If you select Smaller from the Alternative Standard Deviation menu, enter a negative Difference to Detect.

Specify two of the following values to calculate the third value:

Difference to Detect The smallest detectable difference (how small a difference you want to be able to declare statistically significant). This is the difference between the hypothesized value and the true value.

Sample Size The total number of observations (runs, experimental units, or samples) in your experiment.

Power The probability of rejecting the null hypothesis when it is false. With all other parameters fixed, power increases as sample size increases.

Calculator Buttons

Continue Evaluates the missing value.

Back Returns to the previous Sample Size and Power launch window.

Statistical Details for the One Sample Standard Deviation Calculator

The formula that JMP uses to calculate power depends on the setting of the Alternative Standard Deviation option.

For cases where the alternative standard deviation is larger, JMP calculates power as follows:

$$\text{power} = 1 - F\left(\frac{\sigma_0^2 \cdot \chi_{1-\alpha}}{(\sigma_0 + \delta)^2}, n-1\right)$$

$$\delta = \sigma_0 \sqrt{\frac{\chi_{1-\alpha}}{\chi_\beta}} - \sigma_0$$

where:

σ_0 is the hypothesized standard deviation.

σ is the true standard deviation

X_p is the $(1 - p)^{\text{th}}$ quantile of the X_{n-1}^2 distribution.

$F(x, n-1)$ is the cumulative distribution function of the X_{n-1}^2 distribution evaluated at x .

For cases where the alternative standard deviation is smaller, JMP calculates power as follows:

$$\text{power} = 1 - F\left(\frac{\sigma_0^2 \cdot \chi_\alpha}{(\sigma_0 + \delta)^2}, n-1\right)$$

$$\delta = \sigma_0 \sqrt{\frac{\chi_\alpha}{\chi_{1-\beta}}} - \sigma_0$$

where:

σ_0 is the hypothesized standard deviation.

σ is the true standard deviation

X_p is the $(1 - p)^{\text{th}}$ quantile of the X_{n-1}^2 distribution.

$F(x, n-1)$ is the cumulative distribution function of the X_{n-1}^2 distribution evaluated at x .

Numerical solutions are used to solve for n .

For more information about calculations in JMP, see Barker (2011, Section 2.4).

One Sample Proportion Calculator

Use the One Sample Proportion calculator to evaluate sample size for a hypothesis test about one proportion. Sample size and power are associated with the following hypothesis test:

$$H_0: p = p_0$$

versus the two-sided alternative:

$$H_a: p \neq p_0$$

or versus a one-sided alternative:

$$H_a: p < p_0 \text{ or } H_a: p > p_0$$

where p is the population proportion and p_0 is the null proportion.

The One Proportion Calculator also provides the actual test size. This is the actual Type I error rate for the specified assumptions. Since the binomial distribution is discrete, the actual test size can differ significantly from the stated Alpha level for small samples or proportions near 0 or 1. To guarantee an alpha level equal to or greater than your stated level, use the Exact Clopper-Pearson method.

Example of the One Sample Proportion Calculator

Suppose that an assembly line has a historical proportion of defects equal to 0.1. Given a sample size of 100 and an alpha level of 0.05, you want to calculate the power to detect a defect rate that differs by 0.1 or more from the historical rate.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **One Sample Proportion**.
3. Leave **Alpha** as 0.05.
4. Leave 0.1 as the value for **Proportion**.
5. Leave the **Method** as Exact Agresti-Coull.
6. Leave the test type as **Two-Sided**.
7. Leave 0.2 as the value for **Null Proportion**.
8. Enter 100 as the **Sample Size**.
9. Click **Continue**.

Figure 17.12 One Sample Proportion Calculator

Sample Size

One Proportion

Testing if one proportion is different from the hypothesized value.

Alpha

Proportion H₀: P = P₀

Method:

☒ Two-Sided
☐ One-Sided

Enter one value to see a plot of the other two.

Null Proportion

Sample Size

Power

Actual Test Size =

For a sample size of 100, the power is approximately 70%. The Actual Test Size is approximately 0.0467, which is slightly less than the desired 0.05. With a sample size of 100, if the defect rate is 0.2, the probability of a Type I error of rejecting the null hypothesis is 5%. Alternatively, the probability of a Type II error of not rejecting the null hypothesis is 30% (1 - power).

One Sample Proportion Calculator Fields

Specify the following quantities and test settings:

Alpha The probability of a type I error, which is the probability of rejecting the null hypothesis when it is true. It is commonly referred to as the significance level of the test. The default alpha level is 0.05.

Proportion The true proportion, which could be known or hypothesized. The default value is 0.1.

Tip: For all other parameters fixed, the largest sample size occurs when the proportion is 0.5.

Method Select a method. Choices are Exact Agresti-Coull or Exact Clopper-Pearson. The Clopper-Pearson method tends to be more conservative (larger sample size) than the Agresti-Coull method.

One-Sided or Two-Sided Select either a one-sided or a two-sided test.

Specify two of the following parameters to calculate the third value, or specify one value to obtain a plot of the relationship between the other two parameters.

Note: The plot uses the normal approximation method rather than an exact method.

Null Proportion The proportion to test against (p_0). The default value is 0.2.

Sample Size The total number of observations (runs, experimental units, or samples) in the experiment.

Power The probability of rejecting the null hypothesis when it is false. With all other parameters fixed, power increases as sample size increases.

Calculator Buttons

Continue Evaluates the missing value when two parameters are specified, or launches a plot comparing two missing parameters if only one parameter is specified.

Back Returns to the previous Sample Size and Power launch window.

Statistical Details for the One Proportion Calculator

The one proportion sample size computations use exact methods based on the binomial distribution. Exact calculations guarantee that the stated power level is obtained.

Agresti-Coull Method

The exact Agresti-Coull method uses the adjusted Wald-based test statistic. JMP calculates power under the two-sided null hypothesis as follows:

$$\text{Power} = \sum_{y=0}^n \Pr\{Y = y | Y \sim \text{binomial}(n, p)\} I\{T(y) \geq \chi_{1-\alpha}\}$$

where:

$$T(y) = \frac{(\hat{p} - p_0)^2}{\frac{\hat{p}(1-\hat{p})}{n+4}}$$

$$\hat{p} = \frac{y+2}{n+4}$$

$I\{T(y_1, y_2) \geq \chi_{1-\alpha}\} = 1$ if $(T(y_1, y_2) \geq \chi_{1-\alpha})$ and 0 otherwise, and

$\chi_{1-\alpha}$ is the $(1 - \alpha)^{\text{th}}$ quantile of the χ_1^2 distribution.

Because there is not a closed-form expression for n or p_0 , numerical techniques are used to solve for n or p_0 .

For more information about the adjusted Wald test statistic, see Agresti and Coull (1998). For more information about calculations in JMP, see Barker (2011, Section 3.3).

Clopper-Pearson Method

The exact Clopper-Pearson method is based on the binomial distribution. This method results in an alpha level equal to or greater than the stated level. The Clopper-Pearson method is more conservative (larger sample size) than the Agresti-Coull method.

The exact Clopper-Pearson method uses the binomial distribution directly. Numerical techniques are used to solve for the unknown parameter.

For more information about the Clopper-Pearson exact method, see Clopper and Pearson (1934) or Agresti and Coull (1998, Section 1).

Two Sample Proportions Calculator

Use the Two Sample Proportions calculator to evaluate sample size for a hypothesis test about two proportions. Sample size and power are associated with the following hypothesis test:

$$H_0: p_1 - p_2 = D_0$$

versus the two-sided alternative:

$$H_a: p_1 - p_2 \neq D_0$$

or versus either of the following one-sided alternatives:

$$H_a: (p_1 - p_2) < D_0 \text{ or } H_a: (p_1 - p_2) > D_0$$

where p_1 and p_2 are the population proportions from two populations, and D_0 is the hypothesized difference in proportions.

The Two Proportions Calculator also provides the actual test size. This is the actual Type I error rate for the specified assumptions. Since the binomial distribution is discrete, the actual test size can differ significantly from the stated Alpha level for small samples or proportions near 0 or 1.

Example of the Two Sample Proportions Calculator

Suppose you are responsible for two silicon wafer assembly lines. Based on the knowledge from many runs, one of the assembly lines has a defect rate of 8% and the other line has a defect rate of 6%. You want to know the sample size necessary to have 80% power to conclude there is a difference in defect rates.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **Two Sample Proportions**.
3. Leave **Alpha** set at 0.05.
4. Enter 0.08 for **Proportion 1**.
5. Enter 0.06 for **Proportion 2**.
6. Select a **Two-Sided** test.
7. Enter 0 for **Null Difference in Proportion**.
8. Leave **Sample Size 1** and **Sample Size 2** blank.
9. Enter 0.8 for **Power**.
10. Click **Continue**.

Figure 17.13 Difference between Two Proportions for a Two-Sided Test

Sample Size

Two Proportions

Testing if two proportions are different from each other.

Alpha

Proportion 1 $H_0: P_1 - P_2 = \Delta_0$

Proportion 2

☒ Two-Sided
☐ One-Sided

Supply two of (difference, sample sizes, power) to determine the third.
When entering sample sizes, enter a value for both groups.

Null Difference in Proportion

Sample Size 1

Sample Size 2

Power

Actual Test Size = 0.0495189
Test size calculated holding P1 fixed and using P2 = P1 - Δ_0

The calculator shows that you should test a sample size of 2,554 wafers from each production line in order to detect a difference in defect rates.

Two Sample Proportions Calculator Fields

Specify the following quantities and test settings:

Alpha The probability of a type I error, which is the probability of rejecting the null hypothesis when it is true. It is commonly referred to as the significance level of the test. The default alpha level is 0.05.

Proportion 1 The proportion for population 1, which could be known or hypothesized. The default value is 0.5.

Proportion 2 The proportion for population 2, which could be known or hypothesized. The default value is 0.1.

One-Sided or Two-Sided Select either a one-sided or a two-sided test.

Specify two of the following parameters to calculate the third where the sample sizes are considered as one parameter. The sample sizes must both be entered or both be left blank for calculation.

Null Difference in Proportion The difference in proportions under the null hypothesis. The default value is 0.2.

Sample Size 1 The total number of observations (runs, experimental units, or samples) in the experiment to estimate Proportion 1.

Sample Size 2 The total number of observations (runs, experimental units, or samples) in the experiment to estimate Proportion 2.

Note: When sample size is calculated, Sample Size 1 = Sample Size 2. If you enter sample sizes, they do not have to be equal.

Power The probability of rejecting the null hypothesis when it is false. High power is desired. With all other parameters fixed, power increases with sample size.

Calculator Buttons

Continue Evaluates the missing value.

Back Returns to the previous Sample Size and Power launch window.

Statistical Details for the Two Proportions Calculator

Two proportions calculations are based on exact methods similar to those used for the one proportion calculator. For the two-sided hypothesis

$$H_0: p_1 - p_2 = D_0 \text{ vs } H_a: p_1 - p_2 \neq D_0$$

JMP calculates power under the null hypothesis as follows:

$$\text{Power} = \sum_{y_1=0}^{n_1} \sum_{y_2=0}^{n_2} \Pr(Y_1 = y_1) \Pr(Y_2 = y_2) I\{T(y_1, y_2) \geq \chi_{1-\alpha}\}$$

where the adjusted Wald statistic is defined as follows:

$$T(y_1, y_2) = \frac{(\hat{p}_1 - \hat{p}_2 - \delta_0)^2}{\frac{\hat{p}_1(1 - \hat{p}_1)}{n_1 + 2} + \frac{\hat{p}_2(1 - \hat{p}_2)}{n_2 + 2}} \text{ where } \hat{p}_j = \frac{y_j + 1}{n_j + 2} \text{ for } j = 1, 2$$

and:

$$I\{T(y_1, y_2) \geq \chi_{1-\alpha}\} = 1 \text{ if } (T(y_1, y_2) \geq \chi_{1-\alpha}) \text{ and } 0 \text{ otherwise}$$

and $\chi_{1-\alpha}$ is the $(1 - \alpha)^{\text{th}}$ quantile of the χ^2 distribution.

Because there are no closed-form expressions for the n_i or δ_0 , numerical techniques are used to solve for the n_i and δ_0 .

For more information about the adjusted Wald test statistic, see Agresti and Coull (1998). For more information about calculations in JMP, see Barker (2011, Section 3.3).

Counts per Unit Calculator

Use the Counts per Unit Calculator to evaluate sample size for a one-sided hypothesis test about the defects (counts) per unit. Sample size and power are associated with the following hypothesis test:

$$H_0: \lambda \leq \lambda_0$$

versus a one-sided alternative:

$$H_a: \lambda > \lambda_0$$

where λ is the rate of a Poisson distribution.

Example of the Counts per Unit Calculator

Consider a wafer manufacturing process with a target of 4 or fewer defects per wafer. You want to verify that a new process meets that target. How many wafers should you test to detect a difference of 1 that the count of defects per wafer is on target with 90% power at the significance level of 0.05?

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **Counts per Unit**.
1. Leave **Alpha** as 0.05 (the chance of failing the test if the new process is as good as the target).
2. Enter 4 as the **Baseline Counts per Unit**, indicating the target of 4 defects per wafer.
3. Enter 1 as the **Difference to detect**.
4. Enter a power of 0.9.

This is the chance of detecting a change larger than 1, which is equivalent to 5 or more defects per wafer. In this type of situation, the significance level (alpha) is sometimes called the *producer's risk* and the power (beta) is called the *consumer's risk*.

5. Click **Continue**.

Figure 17.14 Counts per Unit Calculator

Sample Size

Counts per Unit

Detecting change in count per unit, e.g. DPU (defects per unit)

Alpha0.05

Baseline Count per Unit4

Supply two values to determine the third.

Difference to detect1

Sample Size38

Power0.9

Using normal approximations

Continue

Back

You must test 38 wafers. The process meets the target if there are fewer than 173 defects.

Note: The 173 defects are the maximum number of defects that can occur in 38 wafers without rejecting the null hypothesis.

Counts per Unit Calculator Fields

Specify the following quantities:

Alpha The probability of a type I error, which is the probability of rejecting the null hypothesis when it is true. It is commonly referred to as the significance level of the test. The default alpha level is 0.05.

Baseline Count per Unit The number of targeted defects per unit. The default value is 0.1.

Specify two of the following quantities to calculate the third quantity:

Difference to detect The smallest detectable difference to test against specified in defects per unit.

Sample Size The number of units.

Power The probability of rejecting the null hypothesis when it is false. With all other parameters fixed, power increases as sample size increases.

Calculator Buttons

Continue Evaluates the missing value.

Back Returns to the previous Sample Size and Power launch window.

Statistical Details for the Counts per Unit Calculator

Calculations for the counts per unit calculator are based on applying a normal approximation to the Poisson distribution. The test uses the following hypothesis:

$$H_0: \lambda \leq \lambda_0 \text{ vs } H_a: (\lambda > \lambda_0)$$

The above hypothesis uses the following test statistic:

$$T = \frac{\bar{d} - \lambda_0}{\sqrt{\lambda_0/n}}$$

where n is the sample size and $\bar{d}$ is the mean number of defects per unit in the sample. We assume T to be approximately normally distributed. The power calculation is based on the distribution of T under the null and alternative hypotheses.

$$\begin{aligned} \text{power} &= \Pr(\text{reject } H_0 \mid \lambda = \lambda_0 + \delta) \\ &= \Pr[T > Z_{1-\alpha} \mid (\lambda = \lambda_0 + \delta)] = 1 - \Phi\left(\frac{Z_{1-\alpha} - \delta\sqrt{n/\lambda_0}}{\sqrt{(\lambda_0 + \delta)/\lambda_0}}\right) \end{aligned}$$

where $\Phi()$ is the standard normal cumulative distribution function and Z_{1-p} is the $(1 - p)^{\text{th}}$ quantile of the standard normal distribution.

Using $1 - \beta$ to denote the desired power to reject the null hypothesis, the sample size is calculated as follows:

$$n = \frac{\lambda_0}{\delta^2} \left(Z_{1-\alpha} - Z_\beta \sqrt{\frac{\lambda_0 + \delta}{\lambda_0}} \right)^2$$

Because an analytical solution for δ does not exist, numerical methods are used to solve for δ given power and n .

For more information about calculations in JMP, see Barker ([2011](#), Section 2.5).

Sigma Quality Level Calculator

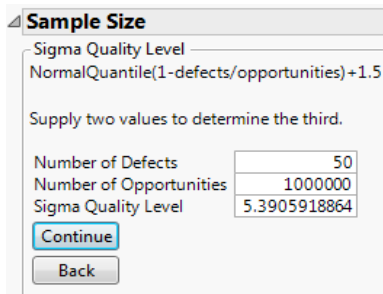
Use the Sigma Quality Level Calculator to explore relationships between the number of defects, number of opportunities for defects, and the Sigma Quality Level. The Sigma Quality Level is a statistic that relates a given defect rate to a six-sigma scale. For example, a defect rate of 3.4 per one million opportunities results in a six-sigma process.

Example of the Sigma Quality Level Calculator

Suppose you want to estimate the Sigma Quality Level for 50 defects in 1,000,000 opportunities at your plant. An opportunity is the unit for the defect count. It could be a single item or you might have multiple defects on a single item where each potential defect is an opportunity.

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click the **Sigma Quality Level**.
3. Enter 50 for the **Number of Defects**.
4. Enter 1000000 as the **Number of Opportunities**.
5. Click **Continue**.

Figure 17.15 Sigma Quality Level Calculator



The screenshot shows a dialog box titled "Sample Size" with a tab labeled "Sigma Quality Level". The formula $\text{NormalQuantile}(1 - \text{defects}/\text{opportunities}) + 1.5$ is displayed. Below the formula, it says "Supply two values to determine the third." There are three input fields: "Number of Defects" with the value 50, "Number of Opportunities" with the value 1000000, and "Sigma Quality Level" with the calculated value 5.3905918864. At the bottom, there are "Continue" and "Back" buttons.

Field	Value
Number of Defects	50
Number of Opportunities	1000000
Sigma Quality Level	5.3905918864

The Sigma Quality level is approximately 5.4.

Example Computing the Number of Defects

To calculate the maximum number of defects you can have out of a million opportunities and still have a six-sigma quality level process, follow these steps:

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click the **Sigma Quality Level** button.
3. Leave **Number of Defects** blank.
4. Enter 1000000 as the **Number of Opportunities**.
5. Enter 6 as **Sigma Quality Level**.
6. Click **Continue**.

Figure 17.16 Six-Sigma Quality Level for 1,000,000 Opportunities

The screenshot shows a window titled "Sample Size". Inside, it displays the formula for Sigma Quality Level: $\text{NormalQuantile}(1 - \text{defects}/\text{opportunities}) + 1.5$. Below this, it says "Supply two values to determine the third." There are three input fields: "Number of Defects" with the value 3.3976731247, "Number of Opportunities" with the value 1000000, and "Sigma Quality Level" with the value 6. At the bottom, there are two buttons: "Continue" and "Back".

The computation shows that the **Number of Defects** cannot be more than 3 in 1,000,000 opportunities to achieve a six-sigma quality level process.

Sigma Quality Level Calculator Fields

Specify two of the following quantities to calculate the third quantity:

Number of Defects The number of defects.

Number of Opportunities The number of opportunities for a defect.

Sigma Quality Level The defect rate in standard deviation units.

Calculator Buttons

Continue Evaluates the missing value.

Back Returns to the previous Sample Size and Power launch window.

Statistical Details for the Sigma Quality Calculator

Calculations for the sigma quality calculator are based on the definition of sigma quality level σ_q for d defects in n opportunities:

$$\sigma_q = Z_{1 - \frac{d}{n}} + 1.5$$

From this definition, d and n can be calculated as follows:

$$d = n[1 - \Phi(\sigma_q - 1.5)]$$

$$n = d[1 - \Phi(\sigma_q - 1.5)]^{-1}$$

where $\Phi()$ is the standard normal cumulative distribution function and $Z_{1-d/n}$ is the $(1 - d/n)^{\text{th}}$ quantile from the standard normal distribution.

Reliability Test Plan Calculator

Use the Reliability Test Plan Calculator to determine the sample size, study length, or precision needed for a reliability study. Reliability studies are used to estimate failure times and failure probabilities of a product. The test plan calculator uses expected confidence intervals to define the precision of the estimates. These estimates are based on a fitted failure distribution. To fit a failure distribution, a minimal number of failures must be observed.

Example of the Reliability Test Plan Calculator

A company has developed a new product and wants to calculate how many units to test to estimate the time until 20% of units fail, using a 95% confidence interval with two-sided absolute precision of 200 hours. In other words, when a confidence interval is created for the estimated time, the difference between the upper and lower limits should be approximately 200 hours. The company can run the test for 2,500 hours. In addition, from studies of similar products, they believe the approximate failure distribution to be a Weibull distribution with location parameter $\alpha = 2000$ and scale parameter $\beta = 3$.

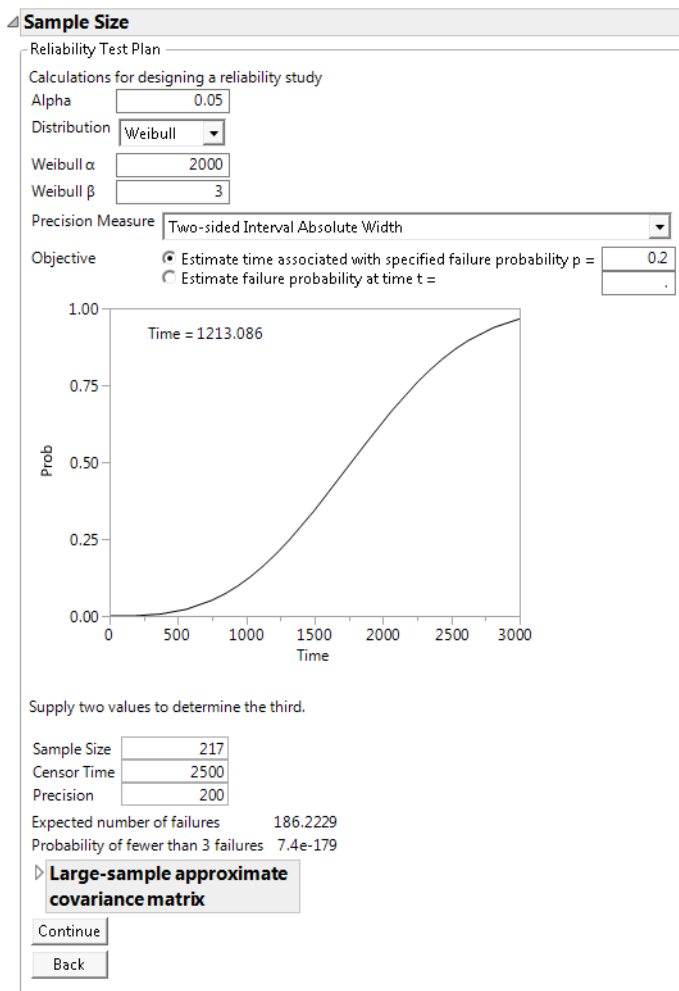
To compute the required sample size:

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **Reliability Test Plan**.
3. Leave Alpha set at 0.05.
4. Select **Weibull** from the Distribution list.
5. Enter 2000 for the **Weibull α** parameter.
6. Enter 3 for the **Weibull β** parameter.
7. Select **Two-sided Interval Absolute Width** from the Precision Measure list.
8. Select **Estimate time associated with specified failure probability** and enter 0.2 for **p**.

Note: The cumulative distribution function plot is labeled according to the study objective. Here, Time = 1213 is the time estimate for 20% of unit failures.

9. Leave **Sample Size** blank.
10. Enter 2500 for **Censor Time**.
11. Enter 200 for **Precision**.
12. Click **Continue**.

Figure 17.17 Reliability Test Plan Calculator



To estimate the time until 20% of units fail with a precision of 200 hours requires 217 units on test for 2,500 hours. The expected number of failures in this test is approximately 186, much larger than the minimum of 3 needed to estimate the failure distribution.

Reliability Test Plan Calculator Fields

Specify the following quantities and test settings:

Alpha The significance level of the confidence interval used to define the precision measure.

Distribution The assumed failure distribution. Distributions available are: Weibull, Lognormal, Frechet, Loglogistic, SEV, Normal, LEV, and logistic. For more information

about these distributions, see the Life Distribution chapter in *Reliability and Survival Methods*.

Location The location parameter for the failure distribution.

Note: The Location field is denoted Weibull α when the Distribution is set to Weibull.

Scale The scale parameter for the failure distribution.

Note: The Scale field is denoted Weibull β when the Distribution is set to Weibull.

Precision Measure The definition of the precision measure. Definitions are based on the expected confidence interval of the quantity being estimated (failure time or probability). The choices for the precision measure are as follows:

Interval Ratio Defines precision as the square root of the ratio of the upper limit to the lower limit. This ratio is always greater than one since the upper limit is greater than or equal to the lower limit. The interval ratio decreases as the precision in the estimate increases.

Two-sided Interval Absolute Width Defines precision as the width of the confidence interval or the difference between the upper and lower limits.

Lower One-sided Interval Absolute Width Defines precision as the width of the lower side of the interval or the difference between the estimate and the lower limit of the confidence interval for the estimate.

Two-sided Interval Relative Width Defines precision as the width of the confidence interval relative to the estimate. This is the difference between the upper and lower limits divided by the estimate.

Lower One-sided Interval Relative Width Defines precision as the width of the lower side of the interval relative to the estimate. This is the difference between the estimate and the lower limit divided by the estimate.

Objective The objective of the study. Select one of the following objectives and enter the corresponding value:

- Estimate time associated with specified failure probability p .
- Estimate failure probability at time t .

Note: The plot is the cumulative distribution function of the failure distribution. The plot is labeled with the estimate of time or probability based on the study objective.

Specify two of the following quantities to calculate the third quantity:

Sample Size The number of units to include in the reliability test.

Censor Time The amount of time available to run the reliability test.

Precision The level of precision. The definition of the units on this value corresponds to the chosen Precision Measure.

Calculator Buttons

Continue Evaluates the missing value.

Back Returns to the previous Sample Size and Power launch window.

Additional Reliability Test Plan Calculations

In addition to calculating the sample size, censor time, or precision, the following quantities are also calculated:

Expected number of failures The expected number of failures for the specified reliability test.

Probability of fewer than 3 failures The probability that the specified reliability test results in fewer than three failures. This is important because a minimum of three failures is required to obtain stable estimates for the location and scale parameters of the failure distribution. With only one or two failures, the estimates are unstable. If this probability is large, you risk not observing enough failures to reliably estimate the distribution parameters. Increasing the sample size or censor time are both ways to lower the probability of fewer than three failures.

Large-sample approximate covariance matrix Provides the approximate variances and covariance for the location and scale parameters of the failure distribution.

Statistical Details for the Reliability Test Plan Calculator

The reliability test plan is designed to estimate either a quantile:

$$q(p) = \exp[\sigma\Phi^{-1}(p) + \mu]$$

or a failure probability:

$$p(t) = \Phi\left[\frac{\log(t) - \mu}{\sigma}\right]$$

at a specified level of significance with an assumed failure time distribution and precision measure. $\Phi()$ is the standard cumulative distribution function of the assumed failure time distribution with location parameter μ and scale parameter σ . Wald confidence intervals and the precision measure define precision as a function of t (time) and n (sample size). Numerical methods are used to solve for the desired quantity.

For more information about calculations in JMP, see Barker (2011, Section 6).

Reliability Demonstration Calculator

Use the Reliability Demonstration calculator to test a specified number of units for a specified period of time. If fewer than k units fail, the demonstration passes, and you can conclude that the product reliability meets or exceeds a reliability standard. We pose the reliability demonstration as a hypothesis test:

$$H_0: p < p^* \text{ vs } H_a: p \geq p^*$$

where p is the probability of survival time at t^* for the new product and p^* is the standard probability of survival at time t^* .

Example of the Reliability Demonstration Calculator

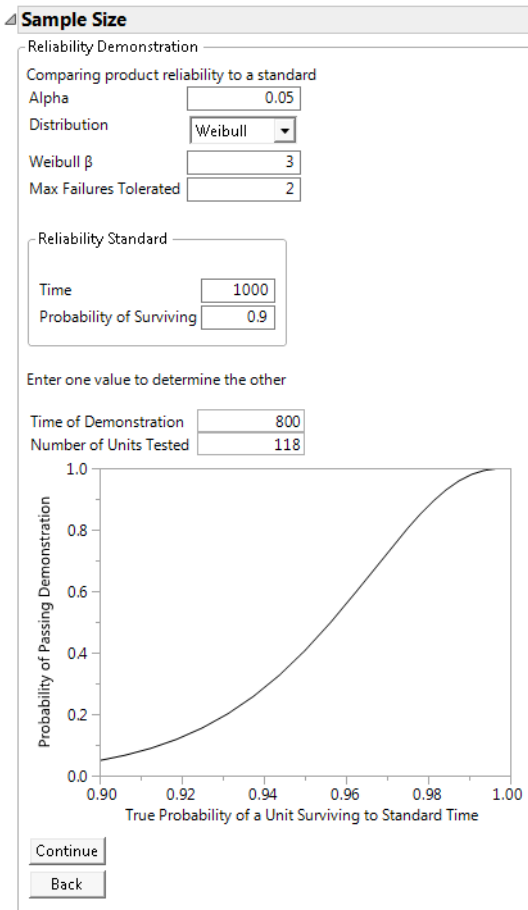
A company wants to estimate the sample size needed for assessing the reliability of a new product against a historical reliability standard of 90% survival after 1,000 hours. From prior studies on similar products, it is believed that the failure time distribution is Weibull, with a β parameter of 3. The company can afford to run the demonstration for 800 hours, and wants the experiment to result in no more than 2 failures.

To compute the required sample size:

1. Select **DOE > Design Diagnostics > Sample Size and Power**.
2. Click **Reliability Demonstration**.
3. Leave Alpha set at 0.05.
4. Select **Weibull** from the Distribution list.
5. Enter 3 for the **Weibull β** .
6. Enter 2 for **Max Failures Tolerated**.
7. Enter 1000 for **Time**.
8. Enter 0.9 for **Probability of Surviving**.
9. Enter 800 for **Time of Demonstration**.
10. Leave **Number of Units Tested** blank.

11. Click **Continue**.

Figure 17.18 Reliability Demonstration Calculator



The company needs to run 118 units in the demonstration. Furthermore, if they observe 2 or fewer failures by 800 hours, they can conclude that the new product reliability is at least as reliable as the standard.

The plot shows that as the true probability of a unit surviving to the time in the reliability standard increases, the probability of passing the demonstration increases. In this example, if a unit has only a 92% chance of surviving to the standard time, then the chance of the new product passing the demonstration is only about 12.5%.

Reliability Demonstration Calculator Fields

Specify the following quantities and test settings:

Alpha The probability of a type I error, which is the probability of rejecting the null hypothesis when it is true. It is commonly referred to as the significance level of the test. The default alpha level is 0.05.

Distribution The assumed failure time distribution. Distributions available are: Weibull, Lognormal, Frechet, Loglogistic, SEV, Normal, LEV, and logistic. For more information about these distributions, see the Life Distribution chapter in *Reliability and Survival Methods*.

Scale The scale parameter for the failure time distribution.

Note: The Scale field is denoted Weibull β when the specified Distribution is Weibull.

Max Failures Tolerated The maximum number of failures allowed for a successful test demonstration.

Reliability Standard The definition of the standard that you are testing. This standard is based on a length of time and the probability of an items survival to that time.

Time The minimum length of time that an item under test should survive.

Probability of Surviving The probability that the item under test survives until the defined Time of the reliability standard

Enter one of the following two values to calculate the other:

Time of Demonstration The length of the demonstration in time.

Number of Units Tested The number of units needed for the demonstration.

Calculator Buttons

Continue Evaluates the missing value.

Back Returns to the previous Sample Size and Power launch window.

Statistical Details for the Reliability Demonstration Calculator

The reliability demonstration depends on the assumed failure time distribution with scale parameter σ . The reliability standard, or probability of survival at time t and location μ is stated as follows:

$$p = 1 - \Phi\left[\frac{\log t - \mu}{\sigma}\right]$$

where μ is solved for using the following:

$$\mu = \log(t) - \sigma\Phi^{-1}(1 - p).$$

To calculate sample size and the size of the test, the probability of survival at time t is posed as a hypothesis test:

$$H_0: p < p^* \text{ vs } H_a: p \geq p^*$$

where p^* is the standard probability of survival at time t^* .

We want to test the hypothesis at the α level or as follows:

$$\alpha = \Pr(k \text{ or few failures} \mid H_0 \text{ true}).$$

Since the test is of n independent units, the number of failures has a binomial (n, p) distribution where p is the probability of a unit failing before time t . Therefore, we can express α as a function of t and n :

$$\alpha = \Pr\left(x \leq k \mid x \sim \text{Bin}\left\{n, \Phi\left[\frac{\log t - \mu^*}{\sigma^*}\right]\right\}\right)$$

where μ^* and σ^* are from the assumed reliability standard.

Properties of the binomial and beta distributions result in being able to solve for t using:

$$t = \frac{t^* \exp\{\sigma^* \Phi^{-1}[1 - B^{-1}(\alpha; n - k, k + 1)]\}}{\exp[\sigma^* \Phi^{-1}(1 - p^*)]}$$

For n , Brent's method is used to find the root of:

$$B^{-1}(\alpha; n - k, k + 1) - 1 + \Phi\left[\frac{\log t - (\log t^*) + \sigma^* \Phi^{-1}(1 - p^*)}{\sigma^*}\right] = 0$$

where:

$B^{-1}(\alpha; n - k, k + 1)$ is the α quantile of the Beta($n - k; k + 1$) distribution
and $\Phi()$ is the cumulative distribution function of the assumed failure time distribution.

For more information about calculations in JMP, see Barker ([2011](#), Section 5).

Chapter 18

Choice Designs

Create a Design for Selecting Preferred Product Profiles

A choice (or *discrete choice*) experiment provides data for modeling discrete preferences. Study participants are presented with sets of potential products (or product profiles) with varying attributes. From each set of profiles, a participant selects a preferred profile. For example, in designing a high-end laptop, a computer company might be interested in the relative importance of key features such as: processor speed, hard disk size, screen size, battery life, and price. A choice experiment addresses the relative values of these features to a customer and indicates an optimal set of trade-offs among product features.

The results of a choice experiment are analyzed using conjoint analysis methods. See the Choice Models chapter in *Consumer Research*.

Figure 18.1 A Survey with Eight Choice Sets

	Choice Set	Choice ID	Disk Size	Speed	Battery Life	Price
1	1	1	40 GB	2.0 GHz	6 Hrs	\$1000
2	1	2	80 GB	1.5 GHz	4 Hrs	\$1500
3	2	1	80 GB	1.5 GHz	4 Hrs	\$1200
4	2	2	40 GB	1.5 GHz	6 Hrs	\$1500
5	3	1	40 GB	1.5 GHz	4 Hrs	\$1200
6	3	2	80 GB	2.0 GHz	6 Hrs	\$1500
7	4	1	40 GB	2.0 GHz	4 Hrs	\$1000
8	4	2	80 GB	1.5 GHz	6 Hrs	\$1500
9	5	1	80 GB	2.0 GHz	6 Hrs	\$1200
10	5	2	40 GB	1.5 GHz	6 Hrs	\$1000
11	6	1	40 GB	1.5 GHz	4 Hrs	\$1500
12	6	2	80 GB	1.5 GHz	6 Hrs	\$1000
13	7	1	80 GB	1.5 GHz	4 Hrs	\$1000
14	7	2	40 GB	2.0 GHz	6 Hrs	\$1200
15	8	1	80 GB	1.5 GHz	6 Hrs	\$1200
16	8	2	80 GB	2.0 GHz	4 Hrs	\$1500

Contents

Overview of Choice Designs	549
Choice Design Terminology	549
Bayesian D-Optimality	550
Example of a Choice Design	550
Example of a Choice Design with Analysis	553
Create a Choice Design for a Pilot Study	553
Analyze the Pilot Study Data	555
Design the Final Choice Experiment Using Prior Information	556
Run the Design and Analyze the Results	558
Choice Design Window	561
Attributes	561
Model	562
Design Generation	563
Design	564
Make Table	566
Choice Design Options	566
Technical Details	567
Bayesian D-Optimality and Design Construction	567
Utility-Neutral and Local D-Optimal Designs	568

Overview of Choice Designs

Discrete choice experiments support the process of designing a product. They help prioritize product features for a company's market so that the company can design a product that people want to buy. The Choice Design platform creates experiments using factors that are attributes of a product. Selecting the attributes to be studied and their values is of critical importance. You must include all attributes that are likely to influence a consumer's decision to buy the product. For more information and guidelines for designing a choice experiment, see Sall ([2008](#)).

Choice Design Terminology

The following terminology is associated with choice designs:

- An *attribute* is a feature of a product.
- A *profile* is a specification of product attributes.
- A *choice set* is a collection of profiles.
- A *survey* is a collection of choice sets.
- A *partial profile* is a profile in a choice design where only a specified number of attributes are varied within each choice set. The remaining attributes are not varied.

In a discrete choice experiment, respondents are presented with a survey containing several choice sets. Choice sets usually contain only a small number of profiles to facilitate the decision process. Within each choice set, each respondent specifies which of the profiles he or she prefers. For example, attributes for a laptop experiment might include speed, storage, screen size, battery life, and price. Different combinations of these attributes comprise product profiles. A choice set might consist of two profiles. From each choice set, a respondent chooses the profile that he or she prefers.

In cases where many attributes are involved, you can construct surveys where each choice set contains *partial profiles*. In a choice set with partial profiles, only a specified number of attributes are varied and the remaining attributes are held constant. This reduces the complexity of the choice task.

Bayesian D-Optimality

Because discrete choice models are nonlinear in their parameters, the efficiency of a choice design depends on the unknown parameters. The Choice Design platform uses a Bayesian approach, optimizing the design over a prior distribution of likely parameter values that you specify. The Bayesian D-optimality criterion is the expected logarithm of the determinant of the information matrix, taken with respect to the prior distribution. The Choice Design platform maximizes this expectation with respect to the prior probability distribution. See [“Bayesian D-Optimality and Design Construction”](#) on page 567 and Kessels et al. (2011).

You can also generate the following types of designs:

- Utility-neutral designs - In a utility-neutral design, all choices within a choice set are equally probable. The prior mean is set to 0.
- Local D-optimal designs - A local D-optimal design takes into account the prior on the mean, but does not include any information from a prior covariance matrix.

For more information about utility neutral and local D-optimal designs, see [“Utility-Neutral and Local D-Optimal Designs”](#) on page 568.

Example of a Choice Design

About the Experiment

In this example, a coffee shop is interested in making an ideal cup of coffee to satisfy the majority of its customers. The manager has asked you to determine which factors affect customer preferences. Specifically, you need to determine which settings of the following factors (attributes) result in an ideal cup of coffee:

- Grind size (medium or coarse)
- Temperature (195°, 200°, 205°)
- Brewing time (3 minutes, 3.5 minutes, or 4 minutes)
- Charge (1.6 grams/ounce, 2 grams/ounce, or 2.4 grams/ounce)

Each combination of factor levels is a *profile*. Trying to obtain information about preferences by having every respondent sample every possible profile is not practical. However, you can ask a respondent to select a preferred profile from a *choice set* consisting of a small number of profiles.

In this example, you design an experiment where each respondent indicates his or her preference in several choice sets. Your design will have the following structure:

- ten respondents

- twelve choice sets
- two profiles per choice set
- one survey per respondent containing all 12 choice sets

The experiment results in 12 responses per respondent. Analysis of these preferences can be used to draw conclusions about how to make a cup of coffee that pleases most customers.

Create the Design

You can enter factors either manually or automatically using a preexisting table that contains the factors and settings. In this example, for convenience, you use a preexisting table. But, if you are designing a new experiment, you must first enter the factors manually. For more information about entering factors manually, see [“Attributes”](#) on page 561.

1. Select **DOE > Consumer Studies > Choice Design**.
2. Select **Help > Sample Data Library** and open Design Experiment/Coffee Choice Factors.jmp.
3. Click the Choice Design red triangle and select **Load Factors**.

Figure 18.2 Choice Design Window with Attributes Defined

The screenshot shows the 'Choice Design' window with the 'Attributes' tab selected. It displays a table of attributes and their levels. Below the table is a 'Specify Attributes' dialog box with a 'Continue' button.

Name	Role	Attribute Levels		
Grind	Categorical	Medium		
Temperature	Categorical	195	200	205
Time	Categorical	3	3.5	4
Charge	Categorical	1.6	2	2.4

Specify Attributes
Add an attribute by clicking the Add Factor button. Double-click an attribute name or level to edit it.
Continue

4. Click **Continue**.
5. (Optional) Open the DOE Model Controls outline.
Notice that there only main effects. In this example, you are only interested in a model that contains the main effects of your four factors. However, if you wanted your design to be capable of estimating additional effects, you add them in this outline.
6. In the Design Generation panel:
 - Keep the **Number of attributes that can change within a choice set** at 4.
 - Keep the **Number of profiles per choice set** at 2.
 - Type 12 for the **Number of choice sets per survey**.

In this example, the respondents evaluate 12 choice sets.

- Keep the **Number of surveys** at 1.
- Type 10 for the **Expected number of respondents per survey**.

In this example, there are ten respondents.

Figure 18.3 Completed Design Generation Panel

The screenshot shows a 'Design Generation' panel with the following settings:

- Number of attributes that can change within a choice set: 4
- Number of profiles per choice set: 2
- Number of choice sets per survey: 12
- Number of surveys: 1
- Expected number of respondents per survey: 10

At the bottom, there are two buttons: 'Make Design' and 'Back'.

Note: Setting the Random Seed in step 7 reproduces the exact results shown in this example. In constructing a design on your own, this step is not necessary.

- (Optional) Click the Choice Design red triangle and select **Set Random Seed**. Type 12345 and click **OK**.
- Click **Make Design**.

There are 12 choice sets, each consisting of two coffee profiles.

- Select **Output separate tables for profiles and responses**.

This places the descriptions of the choice sets in a data table called Choice Profiles. A second data table (called Choice Runs) is constructed to facilitate entry of the response information.

- Click **Make Table**.

The Choice Profiles table shows the 12 choice sets, each consisting of two profiles. The Choice Runs table enables you to record the preferred profile in the column **Response** using the profile ID. Enter 1 if Choice 1 is the preferred profile or 2 if Choice 2 is the preferred profile. Alternatively, if the respondent has no preference, leave the response missing.

The **Choice** script in the Choice Profiles table facilitates analysis of experimental results. It opens a completed launch window for a Choice Model. For information about Choice Models, see the Choice Models chapter in *Consumer Research*.

The **DOE Dialog** script in the Choice Profiles table relaunches the design dialog.

Example of a Choice Design with Analysis

In this example, a computer manufacturer is interested in manufacturing a new laptop and wants information about customer preferences before beginning an expensive development process. The manufacturer decides to construct a design consisting of two sets of profiles that will be administered to ten respondents. The goal of the choice design is to understand how potential laptop purchasers view the advantages of a collection of four attributes:

- size of hard drive disk (40 GB or 80 GB)
- speed of processor (1.5 GHz or 2.0 GHz)
- battery life (4 Hrs or 6 Hrs)
- cost of computer (\$1000, \$1200 or \$1500)

To construct the design for the ten respondents, you first conduct a small pilot study using only one respondent. Then you analyze the results and use the parameter estimates as prior information in designing the final study for the ten respondents.

Create a Choice Design for a Pilot Study

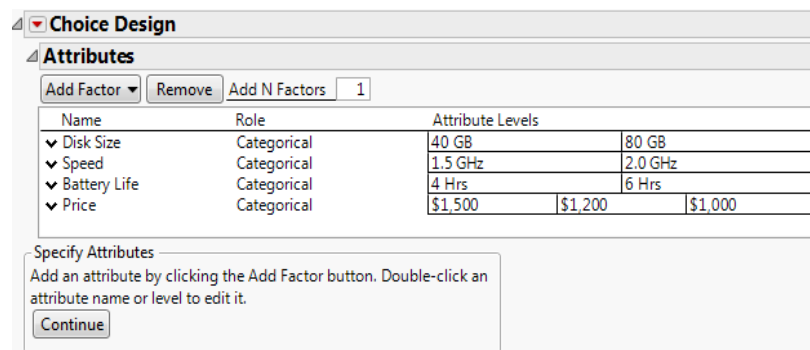
In this section, you construct a choice design for a one-respondent study.

Define Factors and Levels

In this example, you load the factors from an existing table. When designing a new experiment on your own, enter the factors manually.

1. Select **DOE > Consumer Studies > Choice Design**.
2. Select **Help > Sample Data Library** and open Design Experiment/Laptop Factors.jmp.
3. Click the Choice Design red triangle and select **Load Factors**.

Figure 18.4 Choice Design Window with Attributes Defined

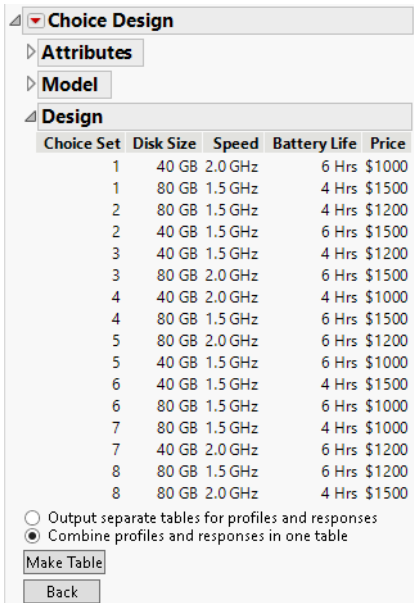


Create the Design

- 1. Click **Continue**.
- 2. This pilot survey will be given to a single respondent. The default values in the DOE Model Controls, Prior Specification, and Design Generation panels are appropriate as is.
- 3. (Optional) Click the Choice Design red triangle and select **Set Random Seed**. Type 12345 and click **OK**.
- 4. Click **Make Design**.

Note: Setting the Random Seed in step 3 reproduces the exact results shown in this example. In constructing a design on your own, this step is not necessary.

Figure 18.5 Pilot Design



The single survey contains eight choice sets, each consisting of two laptop profiles.

- 5. Verify that the **Combine profiles and responses in one table** option is selected.
This places the choice sets and the survey results in the same table.
- 6. Click **Make Table**.

This survey was designed assuming no prior information. For this reason, some choice sets might not elicit useful information. The plan is to obtain survey results from the single respondent, analyze the results, and then use the results from the pilot survey as prior information in designing the final survey.

Analyze the Pilot Study Data

Now that the pilot survey design is complete, it is administered to single respondent. The respondent chooses one profile from each set, entering 1 for the chosen profile and 0 for the rejected profile. You will analyze the results using the Choice platform.

Note: For more information about the Choice platform, see the Choice Models chapter in *Consumer Research*.

1. Select **Help > Sample Data Library** and open Design Experiment/Laptop Design.jmp.
2. Run the **Choice** script.

Figure 18.6 Choice Model Launch Window

Data Format: One Table, Stacked

Select Data Table: Laptop Design

Select Columns: 6 Columns

- Choice Set
- Response Indicator
- Disk Size
- Speed
- Battery Life
- Price

Pick Role Variables

Response Indicator: Response Indicator (required)

Choice Set ID: Choice Set (optional)

By: optional

Construct Profile Effects

Add: Disk Size, Speed, Battery Life, Price

Cross:

Nest:

Macros: Degree 2

Transform: none

Construct Subject Effects (Optional)

Add:

Cross:

Nest:

Macros: Degree 2

Transform: none

Run Model

Help

Remove

☒ Firth Bias-Adjusted Estimates

☐ Hierarchical Bayes

Number of Bayesian Iterations: 5000

☐ Respondent is allowed to select "None" or "No Choice"

The only grouping variable is Choice Set because there is a single survey and a single respondent.

3. Click **Run Model**.

Figure 18.7 Parameter Estimates for Pilot Survey

Parameter Estimates		
Term	Estimate	Std Error
Disk Size[40 GB]	-0.007125177	0.4991763970
Speed[1.5 GHz]	0.214880936	0.5265848733
Battery Life[4 Hrs]	0.052925888	0.4895109018
Price[\$1500]	-0.978517706	0.7062009731
Price[\$1200]	0.076206148	0.7442825255
AICc	43.487115	
BIC	13.884322	
-2*LogLikelihood	3.4871145	
-2*Firth LogLikelihood	-2.203446	
Converged in Gradient		
Firth Bias-Adjusted Estimates		

To construct the final choice design that you will give to ten respondents, you need prior means and variances for the parameter estimates. The analysis in Figure 18.7 gives estimates of the parameter means (Estimate) and estimates of their standard errors (Std Error). You will treat the standard errors as prior estimates of the standard deviations. Next, to calculate estimates of the variances of the attributes, construct a JMP table and square the standard errors.

4. Right-click in the Parameter Estimates report and select **Make into Data Table**.
5. In the new data table, right-click in the Std Error column header and select **New Formula Column > Transform > Square**.
- A column called Std Error^2 is added to the data table. Its values will serve as your estimates of the prior variance for the choice model parameters.

Figure 18.8 Untitled Data Table with Variance Estimates in Last Column

	Term	Estimate	Std Error	Std Error^2
1	Disk Size[40 GB]	-0.007125177	0.4991763970	0.2491770753
2	Speed[1.5 GHz]	0.214880936	0.5265848733	0.2772916288
3	Battery Life[4 Hrs]	0.052925888	0.4895109018	0.2396209229
4	Price[\$1500]	-0.978517706	0.7062009731	0.4987198144
5	Price[\$1200]	0.076206148	0.7442825255	0.5539564778

Note: Do not close the Untitled data table at this point.

Design the Final Choice Experiment Using Prior Information

In this section, you use the prior information obtained from the pilot laptop study to construct a final design. The final design will be administered to a set of ten participants.

1. Select **Help > Sample Data Library** and open Design Experiment/Laptop Factors.jmp.
2. Select **DOE > Consumer Studies > Choice Design**.

3. Click the Choice Design red triangle and select **Load Factors**.
4. Click **Continue**.
5. From the Untitled table, enter the values in the Estimate column into the Prior Mean outline in the Choice Design window (Figure 18.9).

You can copy and paste the entire column from the Untitled table, then click in the Disk Size text box under **Prior Mean** in the Prior Mean outline, right-click, and select Paste.

6. From the Untitled table, enter the values in the Std Error^2 column into the diagonal entries in the Prior Variance outline in the Choice Design window (Figure 18.9). Enter these one-by-one, rounded to three decimal places.

Figure 18.9 Prior Mean and Variance Information from Pilot Study

▲ Prior Specification

☐ Ignore prior specifications. Generate the Utility Neutral design.

▲ Prior Mean

Effect	Prior Mean
Disk Size	0.078
Speed	-0.23
Battery Life	-0.02
Price 1	1.028
Price 2	0.004

☐ Ignore prior variance. Generate the local design for the prior mean.

▲ Prior Variance Matrix

Effect	Disk Size	Speed	Battery Life	Price 1	Price 2
Disk Size	0.230	0.000	0.000	0.000	0.000
Speed		0.282	0.000	0.000	0.000
Battery Life			0.279	0.000	0.000
Price 1				0.583	0.000
Price 2					0.467

7. In the Design Generation panel, enter 2 for the **Number of surveys** and five for the **Expected number of respondents per survey**.

This gives instruments for a total of 10 respondents and allows for two different sets of profiles.

8. Click **Make Design**.
9. Click **Make Table**.

The design table has 160 rows. There are 16 rows for each of the ten study respondents. Each respondent has 8 choice sets, each with 2 profiles. There are two surveys, each given to 5 respondents. The 160 rows result from the following calculation: 2 profiles * 8 choice sets * 2 surveys * 5 respondents = 160 rows.

The final design is now ready to be administered to the 10 respondents.

Run the Design and Analyze the Results

In this section, you analyze the results obtained when you obtain results from the final design. In particular, you want to know how changing the price or other characteristics of a laptop affects its desirability as perceived by potential buyers. This desirability is called the *utility value* of the laptop attributes.

Determine Significant Attributes

1. Select **Help > Sample Data Library** and open **Design Experiment/Laptop Results.jmp**.
2. Run the **Choice** script.

Figure 18.10 Choice Model Launch Window

Data Format: One Table, Stacked

Select Data Table: Laptop Results

Select Columns: 8 Columns

- Respondent
- Survey
- Choice Set
- Response Indicator
- Disk Size
- Speed
- Battery Life
- Price

Pick Role Variables:

Response Indicator: Response Indicator

Subject ID: required

Choice Set ID: required

Grouping: Respondent, Survey, Choice Set

By: optional

Construct Profile Effects:

Add, Cross, Nest, Macros

Degree: 2

Transform: [dropdown]

Construct Subject Effects (Optional):

Add, Cross, Nest, Macros

Degree: 2

Transform: [dropdown]

Run Model, Help, Remove

☒ Firth Bias-Adjusted Estimates

☐ Hierarchical Bayes

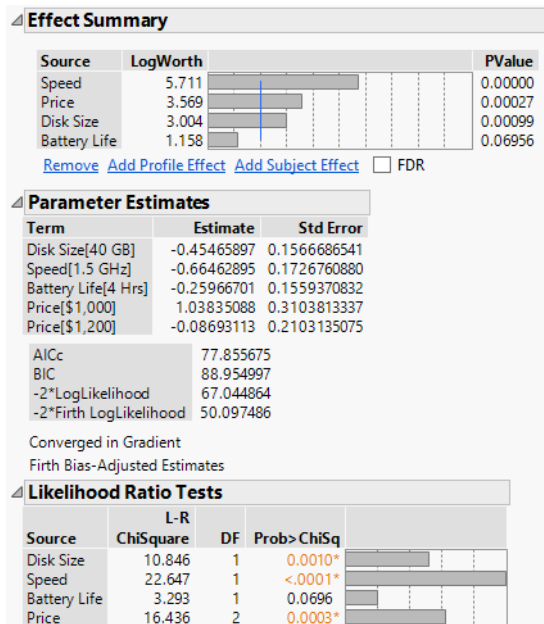
Number of Bayesian Iterations: 5000

☐ Respondent is allowed to select "None" or "No Choice"

There are three grouping variables, Respondent, Survey, and Choice Set, because there are multiple surveys and respondents.

3. Click **Run Model**.

Figure 18.11 Initial Analysis of the Final Laptop Design



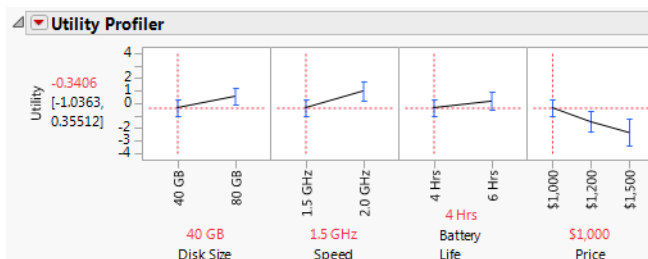
The Effect Summary and Likelihood Ratio Tests outlines indicate that Disk Size, Speed, and Price are significant at the 0.05 level, and that Battery Life is marginally significant.

Find Unit Cost and Trade Off Costs

Next, use the Profiler to see the utility value and how it changes as the laptop attributes change.

1. Click the Choice Model red triangle and select **Utility Profiler**.

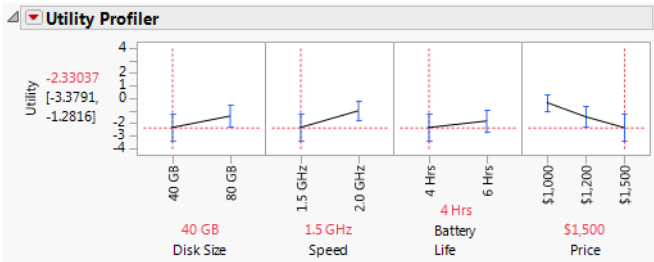
Figure 18.12 Utility Profiler at Price = \$1000



When each attribute value is set to its lowest value, the Utility value is -0.3406. The first thing that you want to do is determine the unit utility cost.

2. Move the slider for Price to \$1,500.

Figure 18.13 Utility Profiler at Price = \$1500

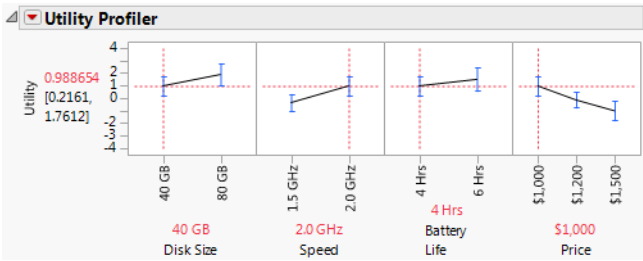


When Price changes from \$1,000 to \$1,500, the Utility changes from -0.3406 to -2.3303. That is, raising the price of a laptop by \$500.00 lowers the utility (or desirability) approximately 2 units. Therefore, you can estimate the unit utility cost to be approximately \$250.00.

With this unit utility cost estimate, you can now vary the other attributes, note the change in utility, and find an approximate monetary value associated with attribute changes. For example, the most significant attribute is Speed (Figure 18.11).

3. In the Utility Profiler, set Price back to \$1,000, its lowest value, and change Speed to 2.0 GHz, its higher value.

Figure 18.14 Utility Value of Higher Speed



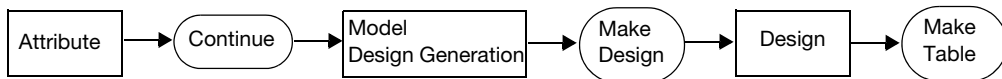
The Utility value changes from the original value shown in Figure 18.12 of -0.3406 to 0.9886, for a total change of 1.3292 units. Using the utility cost estimate or \$250.00, the increase in price for a 2.0 GHz laptop over a 1.5 GHz laptop can be computed to be $1.3292 \times \$250.00 = \$332.30$. This is the dollar value that the Choice study indicates that the manufacturer can use as a basis for pricing this laptop attribute. You can make similar calculations for the other attributes.

Choice Design Window

The Choice Design window walks you through the steps to construct a choice design for modeling attribute preferences. You can specify an assumed model and prior information.

The Choice Design window updates as you work through the design steps. The outlines, separated by buttons that update the outlines, follow the flow in Figure 18.15.

Figure 18.15 Choice Design Flow



Attributes

Attributes in a choice design can only be categorical.

Tip: When you have completed the Attributes outline, consider selecting **Save Factors** from the red triangle menu. This saves the attribute names, roles, and levels in a data table that you can later reload and reuse.

Figure 18.16 Attributes Outline

Name	Role	Attribute Levels
▼ X1	Categorical	L1 L2
▼ X2	Categorical	L1 L2 L3

Specify Attributes

Add an attribute by clicking the Add Factor button. Double-click an attribute name or level to edit it.

[Continue](#)

The Attributes outline contains the following buttons:

Add Factor Adds an attribute with the selected number of levels.

Remove Removes the selected attributes.

Add N Factors Adds multiple attributes. Enter the number of attributes to add, click **Add Factor**, and then select the number of levels. Repeat **Add N Factors** to add multiple attributes with different numbers of levels.

The Attributes outline contains the following columns:

Name The name of the attribute. Attributes are given default names of X1, X2, and so on. To change a name, double-click it and type the desired name.

Role Specifies the Design Role of the attribute as Categorical.

Attribute Levels The attribute name or description. To insert Attribute Levels, click the default levels and type the desired names.

Editing the Attributes Outline

- To edit the Name of an attribute, double-click the attribute name.
- To edit an Attribute Level, click the level.

Attribute Column Properties

For each attribute, JMP saves the Value Order column property to the data tables constructed by the Choice platform. The Value Order column property specifies that levels appear in reports using the ordering specified in the Attributes outline. See [“Value Order”](#) on page 759 in the “Column Properties” appendix.

Model

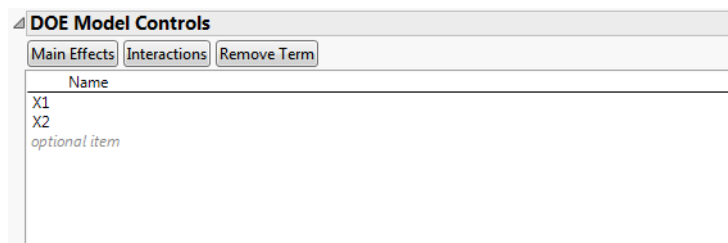
The model outline consists of two parts:

- You can specify your assumed model, which contains all the effects that you want to estimate. See [“DOE Model Controls”](#) on page 562.
- You can specify prior knowledge about the attribute levels, which can result in a better design. See [“Prior Specification”](#) on page 563.

DOE Model Controls

Specify your assumed model in the DOE Model Controls outline. All main effects are included by default. Click the Interactions button to add all two-way interactions.

Figure 18.17 DOE Model Controls Outline



When you construct your design table, JMP saves a Choice script to the data table. The Choice script launches the Choice Platform with the model selected in the DOE Model Controls outline.

The DOE Model Controls outline contains the following buttons:

Main Effects Adds main effects for all attributes in the model.

Interactions Adds all second-order interactions. If you do not want to include all of the interactions, select the interactions that you want to remove and click **Remove Term**.

Remove Term Removes selected effects.

Prior Specification

Enter specifications for a multivariate normal prior distribution on the model parameters. Enter the prior distribution's mean in the Prior Mean outline and its covariance matrix in the Prior Variance outline.

Figure 18.18 Prior Specification Outline

☐ Ignore prior specifications. Generate the Utility Neutral design.

Prior Mean

Effect	Prior Mean
X1	0.000
X2 1	0.000
X2 2	0.000

☐ Ignore prior variance. Generate the local design for the prior mean.

Prior Variance Matrix

Effect	X1	X2 1	X2 2
X1	1.000	0.000	0.000
X2 1		1.000	0.000
X2 2			1.000

You can ignore the prior specifications using these options:

Ignore prior specifications. Generate the Utility Neutral design. Sets the prior means to 0 and generates a locally D-optimal design. This design is called a *utility-neutral* design. See Huber and Zwerina (1996).

Ignore prior variance. Generate the local design for the prior mean. Generates a locally D-optimal design. The local design takes into account the prior on the mean but ignores the covariance matrix. See Huber and Zwerina (1996).

Design Generation

Enter specifications that define the structure for your design in the Design Generation panel.

Figure 18.19 Design Generation Panel for Coffee Example

Design Generation

4 Number of attributes that can change within a choice set

2 Number of profiles per choice set

12 Number of choice sets per survey

1 Number of surveys

10 Expected number of respondents per survey

Make Design

Back

Note: Figure 18.19 is taken from the coffee example. See [“Example of a Choice Design”](#) on page 550.

Enter a number for each of the following items:

Number of attributes that can change within a choice set Enter a number less than or equal to the total number of attributes. This is often set to the total number of attributes. However, if you are comparing many attributes and want to simplify the selection process for respondents, set this to a number that is smaller than the total number of attributes.

Number of profiles per choice set Enter the number of profiles that a respondent must choose from in stating a preference.

Number of choice sets per survey Enter the number of preferences that you want to obtain from each respondent.

Number of surveys Enter the number of distinct collections of choice sets. This is useful if you want to administer surveys to multiple respondents.

Expected number of respondents per survey Enter the total number of respondents divided by the number of surveys.

Make Design

Once you have completed the Design Generation outline, click **Make Design** to generate the design. The design appears in the Design outline.

Design

The Design outline shows the runs for a design that is optimal, given your selections. Review the design to ensure that it meets your needs.

Figure 18.20 Design Outline for Coffee Example

Design				
Choice Set	Grind	Temperature	Time	Charge
1	Coarse	195	3	2
1	Medium	200	4	1.6
2	Medium	205	3	2
2	Coarse	200	3.5	2.4
3	Coarse	195	3	1.6
3	Medium	200	3.5	2.4
4	Coarse	200	3	2.4
4	Medium	205	3.5	2
5	Coarse	195	3.5	1.6
5	Coarse	200	3	2
6	Medium	195	3.5	2
6	Medium	205	4	2.4
7	Coarse	195	3.5	2.4
7	Medium	200	3	1.6
8	Medium	195	4	2
8	Coarse	205	3	1.6
9	Coarse	200	3.5	1.6
9	Medium	195	3	2.4
10	Medium	200	3	1.6
10	Coarse	205	4	2
11	Coarse	200	3.5	2
11	Medium	195	4	1.6
12	Coarse	200	4	2
12	Medium	205	3.5	2.4

☐ Output separate tables for profiles and responses
☒ Combine profiles and responses in one table

Note: Figure 18.20 is taken from the coffee example. See [“Example of a Choice Design”](#) on page 550.

Note: The algorithm for finding an optimal design is based on a random starting design. Because of this, the design you obtain is not unique. The design algorithm will generate different designs when you click the **Back** and **Make Design** buttons repeatedly.

Output Options

Select one of the following output options:

Output separate tables for profiles and responses Displays two data tables:

- The Choice Profiles table lists the profiles in each row, identified by Survey and Choice Set columns. Within a choice set, the profiles are identified by Choice ID. This table is useful for constructing the survey instruments.
- The Choice Runs table provides an empty Response column where you can enter respondent preferences. Each row corresponds to a single choice set. The rows are sorted by Respondent, Survey, and Choice Set. The choice set IDs are given in the next columns, followed by the Response column. Enter the choice set ID for the respondent’s preference in the Response column.

Combine profiles and responses in one table Provides a single Choice Profiles table with an empty Response Indicator column where you can enter respondent preferences. Each

row corresponds to a single profile. The table is sorted by Respondent, Survey, and Choice Set. Enter the value 1 (or another nonzero numerical indicator) for the respondent's preferred profile and a 0 indicator for the other profiles in that choice set.

Note: The values you enter in the Response Indicator column must be numerical.

Make Table

Click **Make Table** to construct the table or tables that you selected in “[Output Options](#)” on page 565. In the table panel of the Choice Profiles table, there is a **Choice** script. Run the script and then click **Run Model** to analyze your experimental results.

For information about the Choice Model report, see the Choice Models chapter in *Consumer Research*.

Choice Design Options

Save Factors Saves the information in the Factors panel to a new data table. Each factor's column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Note: It is possible to create a factors table by entering data into an empty table, but remember to assign each column an appropriate Design Role. Do this by right-clicking on the column name in the data grid and selecting **Column Properties > Design Role**. In the Design Role area, select the appropriate role.

Load Factors Loads factors that you saved using the Save Factors option.

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction
- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Simulate Responses Adds response probabilities and a column containing a choice simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A new column that contains a simulated choice response formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of the marginal utilities for model effects.
- A script called **Choice Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the marginal utilities.

When you click Apply, the formula for the simulated choice response values is updated in the Choice Simulated column. If you click Apply again, the formula and values in this column are updated.

See [“Simulate Responses”](#) on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Number of Starts Enables you to specify the number of random starts used in constructing the design. See [“Bayesian D-Optimality and Design Construction”](#) on page 567.

Advanced Options Not available for Choice Designs.

Save Script to Script Window Creates the script for the design that you specified in the Choice Design window and saves it in an open script window.

Technical Details

Bayesian D-Optimality and Design Construction

The Bayesian D-optimality criterion is the expected logarithm of the determinant of the information matrix of the maximum likelihood of the parameter estimators in the multinomial logit model, taken with respect to the prior distribution. The Choice Design platform maximizes this expectation with respect to a sample of parameter vectors that represents the prior probability distribution. See Kessels et al. (2011).

For partial profile designs, JMP uses a two-stage design algorithm:

1. The constant attributes in each choice set are determined using an *attribute balance* approach.
2. The levels of the non-constant attributes are determined using Bayesian D-optimality.

Attribute balance means that the algorithm attempts to balance the number of times each attribute is held constant in the entire design. If two or more attributes are held constant, the algorithm attempts to balance the occurrence of pairs of attributes held constant in the design.

The levels of the non-constant attributes are determined to optimize the Bayesian D-optimal criterion. A random starting design is found. Then levels of the non-constant attributes are generated using a coordinate-exchange algorithm and evaluated until the Bayesian D-optimality criterion is optimized. The calculations, which involve integration with respect to a multivariate normal prior, use the quadrature method described in Gotwalt et al. (2009).

Note: The Bayesian D-optimality criterion can result in choice sets where some non-constant attributes have identical levels. This situation occurs when varying the non-constant levels within a profile would result in uninformative choice sets where all profiles have very high or very low probabilities.

Utility-Neutral and Local D-Optimal Designs

You can use the Choice Design platform to generate a utility-neutral design by setting prior means to 0. In a utility-neutral design, all choices within a choice set are equally probable. See Huber and Zwerina (1996).

You can also generate a local D-optimal design. The local design takes into account the prior of the mean, but does not include any information from a prior covariance matrix. See Huber and Zwerina (1996).

Chapter 19

MaxDiff Design

Create a Design for Selecting Best and Worst Items

MaxDiff (*maximum difference scaling*) studies are an alternative to studies that use standard preference scales to determine the relative importance of items being rated. In a MaxDiff study, a respondent reports only the most and least preferred options from among a small set of choices. This forces respondents to rank options in terms of preference, which often results in rankings that are more definitive than rankings obtained using standard preference scales.

Use the MaxDiff platform when you need to construct a design consisting of choice sets for one factor that can be presented to respondents as part of a MaxDiff study. Conduct your study and then analyze your data use the MaxDiff analysis platform.

Figure 19.1 A MaxDiff Design Table

	Subject	Choice Set	Candy	Choice
1	1	1	Reese's Cups	•
2	1	1	Hershey Bar	•
3	1	1	Snickers	•
4	1	1	Butterfinger	•
5	1	2	Butterfinger	•
6	1	2	Heath Bars	•
7	1	2	Plain M&Ms	•
8	1	2	Snickers	•
9	1	3	Plain M&Ms	•
10	1	3	Snickers	•
11	1	3	Peanut M&Ms	•
12	1	3	Hershey Bar	•
13	1	4	Snickers	•
14	1	4	Peanut M&Ms	•
15	1	4	Reese's Cups	•
16	1	4	Heath Bars	•

Contents

- Overview of the MaxDiff Design Platform 571
- Example of a MaxDiff Design 571
- MaxDiff Design Launch Window 575
- MaxDiff Window..... 575
 - Design Options Outline 575
 - Design Outline..... 576
 - Pairwise Incidence Matrix 576
 - Make Table 576
- MaxDiff Options 576

Overview of the MaxDiff Design Platform

A *choice set* is a collection of items from which a respondent must select an item that is most preferred (*best* item) and one that is least preferred (*worst* item). Use the MaxDiff platform to specify the number of items that appear in a choice set and the number of choice sets to be presented to a respondent.

The MaxDiff platform provides you with a single survey. If you have several respondents, you can administer the same survey to all respondents or construct a survey for each respondent.

The MaxDiff design is constructed to match the incidence matrix of a balanced incomplete block design as closely as possible. This implies that each pair of items occurs equally often in the design. If the number of choice sets allows a balanced incomplete block to be constructed for the specified number of items in a choice set, the design that is constructed is a balanced incomplete block design.

The items or profiles are considered to be the levels of the treatment factor. The choice sets are considered to be blocks. In a MaxDiff study, the block size is smaller than the number of treatments.

Example of a MaxDiff Design

You are the purchaser for your company's office supplies and you need to buy candy for a holiday party. First, you want to determine which candy types people prefer. To figure this out, you conduct a MaxDiff study.

You ask five randomly chosen associates to rate seven types of candies. Based on your experience with previous studies, you realize that it is difficult for raters to rank seven types of items in order of preference. Instead, you create a design that consists of choice sets of size four. To keep the study manageable, you structure the survey as a MaxDiff study: you ask each associate to specify his or her most preferred and least preferred candy in each of the seven choice sets. (These selections result in a balanced incomplete block design. See Cochran and Cox (1957).) You administer the same survey to each associate.

Create the Design

Construct a table that lists the items or profiles for your choice sets. In this example, your table of items, `Candy Profiles.jmp`, is already constructed.

1. Select **Help > Sample Data Library** and open `Design Experiment/Candy Profiles.jmp`.

The table lists the seven candy types of interest.

2. Select **DOE > Consumer Studies > MaxDiff Design**.

- 3. From the Select Columns list, select Candy and click **X, Factor**.
- 4. Click **OK**.
- 5. In the Design Options outline, do the following:
 - Set the **Number of Profiles per Choice Set** to 4.
 - Set the **Number of Choice Sets** to 7.

Note: Setting the Random Seed in step 6 reproduces the exact results shown in this example. In constructing a design on your own, this step is not necessary.

- 6. (Optional) Click the MaxDiff Study red triangle and select **Set Random Seed**. Type 12345 and click **OK**.
- 7. Click **Make Design**.

The Design outline shows 7 choice sets, each consisting of 4 candy types. The Pairwise Incidence Matrix provides a matrix of the number of choice sets in which each pair of candy types appear together. The diagonal of the matrix is the number of choice sets each candy type appears in. For this example, each candy appears in 4 choice sets and appears with each other candy types in two of the four choice sets.
- 8. Click **Make Table**.

Figure 19.2 Partial Design Table for Candy Preference Survey

	Subject	Choice Set	Candy	Choice
1	1	1	Reese's Cups	•
2	1	1	Hershey Bar	•
3	1	1	Snickers	•
4	1	1	Butterfinger	•
5	1	2	Butterfinger	•
6	1	2	Heath Bars	•
7	1	2	Plain M&Ms	•
8	1	2	Snickers	•
9	1	3	Plain M&Ms	•
10	1	3	Snickers	•
11	1	3	Peanut M&Ms	•
12	1	3	Hershey Bar	•
13	1	4	Snickers	•
14	1	4	Peanut M&Ms	•
15	1	4	Reese's Cups	•
16	1	4	Heath Bars	•

The design table contains a Choice column for recording preferences. For each choice set, record a 1 for the most preferred candy, a -1 for the least preferred, and a 0 for the other two candies.

Analyze the Study Results

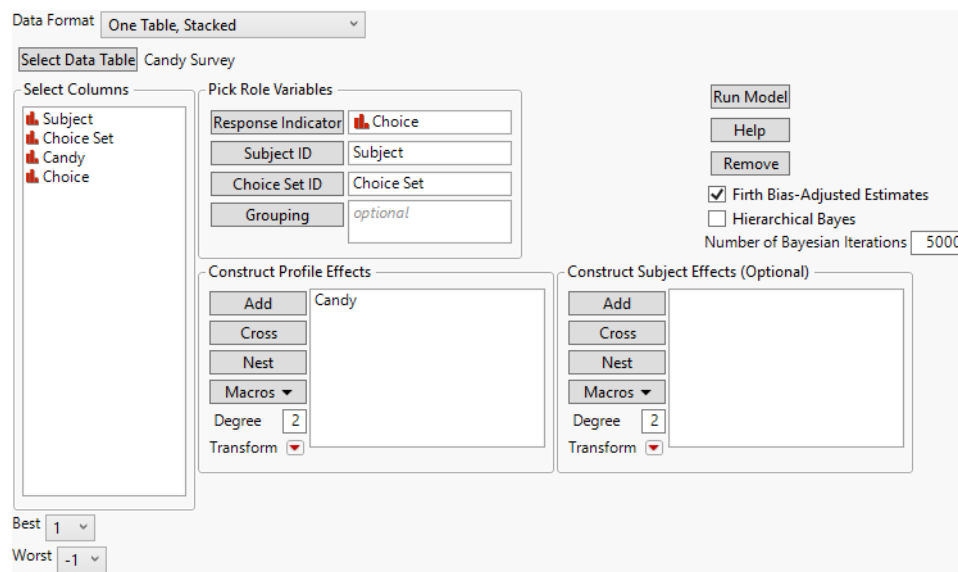
You conduct the study and record your data in Candy Survey.jmp.

1. Select **Help > Sample Data Library** and open Design Experiment/Candy Survey.jmp.

The table shows the results of presenting the survey to each of five respondents, listed in the Subject column.

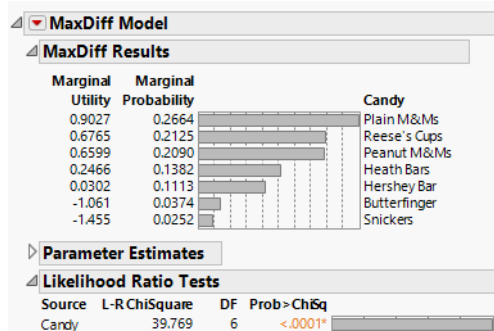
2. Select **Analyze > Consumer Research > MaxDiff**.
3. Click **Select Data Table**, select **Candy Survey**, and click **OK**.
4. Assign roles to columns as follows:
 - Select **Choice** and click **Response Indicator**.
 - Select **Subject** and click **Subject ID**.
 - Select **Choice Set** and click **Choice Set ID**.
 - Select **Candy** and click **Add** in the **Construct Profile Effects** panel.

Figure 19.3 Completed MaxDiff Analysis Launch Window



Because you designated the Best choice as 1 and the Worst choice as -1, you make no change to the Best and Worst choice indicators at the bottom left of the launch window.

5. Click **Run Model**.

Figure 19.4 MaxDiff Report


The report indicates that Candy is significant. The three candy types with the highest utilities are Plain M&Ms, Reese's Cups, and Peanut M&Ms.

- Click the MaxDiff Model red triangle and select **All Levels Comparison Report**.

Figure 19.5 All Levels Comparison Report for Candy Types

All Levels Comparison Report							
Difference (Row-Column) Standard Error of Difference Wald p-Value	Butterfinger	Heath Bars	Hershey Bar	Peanut M&Ms	Plain M&Ms	Reese's Cups	Snickers
Butterfinger	0	-1.3079	-1.0915	-1.7212	-1.964	-1.7378	0.39343
	0	0.52954	0.52038	0.54692	0.56495	0.54938	0.5059
		0.01964	0.04478	0.0038	0.00162	0.00364	0.44306
Heath Bars	1.30792	0	0.21642	-0.4133	-0.656	-0.4299	1.70135
	0.52954	0	0.47619	0.48741	0.49094	0.46959	0.54758
	0.01964		0.65286	0.40343	0.19185	0.36752	0.0042
Hershey Bar	1.0915	-0.2164	0	-0.6297	-0.8725	-0.6463	1.48493
	0.52038	0.47619	0	0.47406	0.49984	0.49096	0.53808
	0.04478	0.65286		0.19443	0.09149	0.19835	0.00992
Peanut M&Ms	1.72121	0.41329	0.62971	0	-0.2427	-0.0166	2.11463
	0.54692	0.48741	0.47406	0	0.49706	0.49054	0.56901
	0.0038	0.40343	0.19443		0.62896	0.97325	0.00086
Plain M&Ms	1.96396	0.65604	0.87246	0.24275	0	0.22616	2.35738
	0.56495	0.49094	0.49984	0.49706	0	0.49327	0.58562
	0.00162	0.19185	0.09149	0.62896		0.65002	0.00037
Reese's Cups	1.7378	0.42988	0.6463	0.01659	-0.2262	0	2.13123
	0.54938	0.46959	0.49096	0.49054	0.49327	0	0.5714
	0.00364	0.36752	0.19835	0.97325	0.65002		0.00083
Snickers	-0.3934	-1.7013	-1.4849	-2.1146	-2.3574	-2.1312	0
	0.5059	0.54758	0.53808	0.56901	0.58562	0.5714	0
	0.44306	0.0042	0.00992	0.00086	0.00037	0.00083	

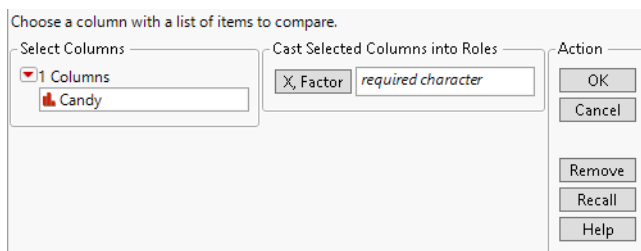
The comparison report indicates which pairs of candy types differ significantly in terms of utility. The third entry in each cell is the p -value for the difference defined by the row item's utility minus the column item's utility. The intensity of the color for the p -value indicates how significant a difference is. The shading, blue or red, indicates whether the difference (row - column) is negative or positive. The p -values are not adjusted to control the multiple comparison error rate and should be used only as a guide. For more information about the All Level Comparisons Report, see the MaxDiff chapter in *Consumer Research*.

MaxDiff Design Launch Window

To use the MaxDiff Design platform, you need a starting data table. Your starting data table must contain a column of character data that lists the items to be presented to respondents for rating.

With your starting data table active in JMP, select **DOE > Consumer Studies > MaxDiff Design**. If you have no active data tables, you are prompted to navigate to your starting data table.

Figure 19.6 MaxDiff Launch Window using Candies.jmp

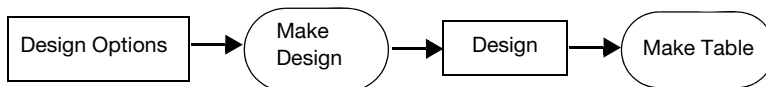


X, Factor The character column that contains the items that respondents will rate.

MaxDiff Window

The MaxDiff Study window updates as you work through the design steps. The outlines, separated by buttons that update the outlines, follow the flow in Figure 19.7

Figure 19.7 MaxDiff Design Flow



The MaxDiff Study window opens showing the Design Options outline. Once you click Make Design, the Design outline appears. To construct the design table, click Make Table.

Design Options Outline

Specify the following:

Number of Profiles per Choice Set The number of items to be included in each choice set.

Number of Choice Sets The total number of sets of items to be presented to and rated by respondents.

Design Outline

The Design outline identifies the choice sets using consecutive positive integers. The items that comprise each choice set are listed.

Pairwise Incidence Matrix

The Pairwise Incidence Matrix outline provides a matrix of the number of choice sets in which each pair of factor levels appear together. The diagonal of the matrix is the number of choice sets each factor level appears in.

Make Table

The Make Table button creates the design table. The design table consists of four columns:

Subject Initially populated by ones. Replace with appropriate identifiers for respondents.

Tip: To easily add respondent identifiers, see the Enter and Edit Data chapter in *Using JMP*. Use Columns > Recode to change the identifiers to names.

Choice Set A designator for each choice set.

Factor The levels of the factor that you specified. These are the items in the choice set.

Choice A column where you can enter results. Use a numeric value to indicate the best choice, the worst choice, and choices in between. The values 1, -1, and 0 are typically used and are required for analysis by the MaxDiff analysis platform (located at Analyze > Consumer Research > MaxDiff).

MaxDiff Options

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction

- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table

Simulate Responses Select this option to simulate response values. When you click Make Table, a Simulate Choice window opens along with the design table, and Probability and Choice Simulated columns are added to the design table. The Choice Simulated formula column contains random responses and the Probability formula column contains their probabilities.

To change the model used in simulating the responses, enter values in the Marginal Utility column in the Simulate Choice window for all factor levels but the last. Because the marginal utilities must sum to zero, you are not permitted to edit the Marginal Utility for the last level. Once you have specified the remaining Marginal Utility values, the last level of the factor is adjusted accordingly.

Number of Starts The number of random starts used in constructing the design. This value is set to 10 by default.

Advanced Options Not applicable for MaxDiff Designs.

Chapter 20

JMP[®] PRO Covering Arrays

Detecting Component Interaction Failures

Covering arrays are used in testing deterministic systems where failures occur as a result of interactions among components or subsystems. The design goal is to reveal if any interaction induces a failure in the system. Application areas include software, circuit, and network design.

Since the tests are deterministic, the emphasis driving the design is the need to cover all required interactions. The Covering Arrays platform constructs highly efficient covering arrays. You can also exclude factor level combinations that are not feasible for your testing protocol.

Figure 20.1 Strength 3 Covering Array

Design					
Run	X1	X2	X3	X4	X5
1	L1	L1	L1	L1	L1
2	L1	L1	L1	L2	L2
3	L2	L2	L2	L2	L2
4	L2	L1	L2	L2	L1
5	L2	L2	L2	L1	L2
6	L2	L2	L1	L1	L1
7	L2	L1	L1	L1	L2
8	L1	L2	L2	L1	L1
9	L1	L2	L1	L2	L2
10	L1	L2	L2	L2	L1
11	L1	L1	L2	L1	L2
12	L2	L2	L1	L2	L1

Contents

Overview of Covering Arrays	581
Example of a Covering Array with No Factor Level Restrictions	583
Create the Design	583
Analyze the Experimental Data	585
Example of a Covering Array with Factor Level Restrictions	587
Create the Design	588
Analyze the Experimental Data	592
Covering Array Window	594
Factors	594
Restrict Factor Level Combinations	595
Design	598
Metrics	600
Output Options	601
The Covering Array Data Table	601
Covering Array Options	603
Technical Details for Covering Arrays	604
Algorithm for Optimize	604
Formulas for Metrics	604

Overview of Covering Arrays

You can use covering arrays to test systems where failures occur as a result of interactions among components or subsystems. Covering arrays are often used in areas such as software, circuit, and network design, where the following conditions are likely to be true:

- The cost of testing is usually high.
- Testing focuses on revealing interactions for which failures occur.
- A test run is typically deterministic and results in either success or failure.
- Replicate runs are wasteful because they yield identical results.
- The efficiency of a design is based on how many of the possible conditions are covered without including redundant runs.

Because systems testing is expensive, reducing the amount of testing is critical. Testing all possible interactions is usually prohibitive and often unnecessary. Experience shows that most failures result from the interaction of a small number of components. The size of the largest combination of components likely to drive a failure, called the *strength*, drives the size of the design.

In the Covering Array platform, you specify the required strength of your design. If appropriate, you define factor level combinations that are not permitted. The Covering Array platform constructs a highly efficient design that meets your requirements. It provides metrics that you can use to assess the quality of the design in terms of its coverage. It also provides a script in the data table for the design that enables you to analyze your results.

Covering arrays are often used in situations where certain combinations of factor level settings are not feasible. The Covering Array platform is able to find very efficient covering arrays even when restrictions are placed on factor level combinations.

For background on the structure of covering arrays and algorithms for computing them, see Colbourn (2004), Colbourn et al. (2011), Hartman and Raskin (2004), and Martirosyan (2003). For more information about covering arrays with restrictions on factor levels, see Cohen et al. (2007) and Morgan (2009).

Covering Arrays and Strength

A *covering array of strength t* is a design that tests all combinations of t factor level settings. Consider an interaction defined by specific settings for k factors. If failures occur for all tests involving that interaction, then that interaction *detects* a failure. Using this terminology, a strength t design enables you to detect failures associated with any interaction of up to t factors.

In the literature, covering arrays are also referred to as *factor covering designs*. For background and more information, see Yilmaz et al. (2014), Cohen et al. (2003), and Dalal and Mallows (1998).

To illustrate the nature of covering arrays, consider a situation involving seven categorical factors each with two levels. You want to test all pairwise combinations of factor levels.

A design that might be used in this situation is the 8-run resolution III main effects design given as follows:

Figure 20.2 A Resolution III Design with Strength 2

Design								
Run	X1	X2	X3	X4	X5	X6	X7	
1	L1	L1	L2	L1	L1	L2	L2	
2	L2	L2	L2	L2	L1	L1	L2	
3	L1	L1	L2	L2	L2	L1	L1	
4	L2	L1	L1	L1	L2	L1	L2	
5	L1	L2	L1	L1	L1	L1	L1	
6	L2	L2	L2	L1	L2	L2	L1	
7	L1	L2	L1	L2	L2	L2	L2	
8	L2	L1	L1	L2	L1	L2	L1	

Note that this factorial design is a strength 2 covering array because all pairwise combinations of levels of any two factors appear. For example, for X1 and X2, the following combinations each appear twice:

- L2 and L1
- L2 and L2
- L1 and L1
- L1 and L2

However, a strength 2 covering array needs only 6 runs :

Figure 20.3 Strength 2 Covering Array

Design								
Run	X1	X2	X3	X4	X5	X6	X7	
1	2	1	1	2	2	1	2	
2	1	2	1	2	1	2	2	
3	1	1	2	1	2	2	1	
4	2	2	2	2	2	2	1	
5	2	2	2	1	1	1	2	
6	1	1	1	1	1	1	1	

All pairwise combinations of levels of any two factors appear at least once in the six runs. The Covering Array design is more efficient than the Resolution III design because it achieves strength 2 coverage with fewer runs.

The efficiency of a covering array is measured by the number of runs required to achieve the required coverage. The smaller the number of runs, the more efficient the design.

Example of a Covering Array with No Factor Level Restrictions

The data in this example pertain to interoperability in the area of software testing. There are four factors of interest:

- Web Browser (Safari, IE, Firefox, Chrome, Other)
- Operating System (Windows or macOS)
- RAM (4, 8, or 16 MB)
- Connection Speed at three settings (0-1, 1-5, or greater than 5 Mbps)

You are interested in finding out which combinations of these factors are likely to cause failures.

The response is whether the system functions properly for each combination of factor settings.

Testing each combination of settings would require 90 (5x2x3x3) trials. To keep the run size manageable, you decide to require Strength 3 coverage, indicating that all combinations of any three factors are tested.

Create the Design

Create the Strength 3 covering array by following these steps.

1. Select **Help > Sample Data Library** and open Design Experiment/Software Factors.jmp.
The Software Factors.jmp data table contains the factors and their settings.
2. Select **DOE > Special Purpose > Covering Array**.
3. From the menu next to **Strength: t =**, select 3.
4. Click the Covering Array red triangle, select **Load Factors** and click **Continue**.
The Factors outline is populated with the four factors and their levels.

Figure 20.4 Factors Outline for Software Factors

Factors

Add Factor

Remove

Add N Factors

1

Strength: t =

3

Name	Role	Values
Web Browser	Categorical	Safari IE Firefox Chrome Other
Operating System	Categorical	Macintosh Windows
RAM	Categorical	4 MB 8 MB 16 MB
Connection Speed	Categorical	>5 Mbps 0-1 Mbps 1-5 Mbps

Specify Factors

Add a factor by clicking the Add Factor button. Double click on a factor name or level to edit it.

Continue

5. Click **Continue**.
- The Restrict Factor Level Combinations outline opens, where you can enter restrictions on the design settings. Because there are no restrictions for this design, do not change the default selection of **None**.
6. Click **Make Design**.
- The Design outline opens to show a 45-run design.

Figure 20.5 Design and Metrics Outlines for Software Factors

Design

Run	Web Browser	Operating System	RAM	Connection Speed
1	Safari	Macintosh	4 MB	>5 Mbps
2	Safari	Windows	4 MB	0-1 Mbps
3	Safari	Macintosh	4 MB	1-5 Mbps
4	Safari	Windows	8 MB	>5 Mbps
5	Safari	Macintosh	8 MB	0-1 Mbps
6	Safari	Macintosh	8 MB	1-5 Mbps
7	Safari	Macintosh	16 MB	>5 Mbps
8	Safari	Macintosh	16 MB	0-1 Mbps
9	Safari	Windows	16 MB	1-5 Mbps
10	IE	Windows	4 MB	>5 Mbps
11	IE	Macintosh	4 MB	0-1 Mbps

Optimize

Maximum iterations:

250

Metrics

Number of Runs:

45

t	Coverage	Diversity
3	100.00	68.33
4	50.00	100.00

In the Metrics outline, consider the row that corresponds to $t = 3$. The Coverage is 100%, indicating that the design covers 100% of the three-factor interactions. This is what you want, because you requested a Strength 3 design. For $t = 3$, the Diversity column indicates that 68.33% of the three-factor interactions that appear are distinct. There is some minor repetition of three-factor combinations.

For $t = 4$, the Coverage is 50%, indicating that the design covers half of the four-factor interactions. There are 90 possible distinct combinations of the four factor settings. The 45 runs in the design comprise one-half of these distinct combinations. The Diversity value of 100% reinforces the fact that none of the four-way interactions are repeated.

7. Click **Make Table**.

Figure 20.6 Partial Design Table for Software Factors

	Response	Web Browser	Operating System	RAM	Connection Speed
1		• Safari	Macintosh	4 MB	>5 Mbps
2		• Safari	Windows	4 MB	0-1 Mbps
3		• Safari	Macintosh	4 MB	1-5 Mbps
4		• Safari	Windows	8 MB	>5 Mbps
5		• Safari	Macintosh	8 MB	0-1 Mbps
6		• Safari	Macintosh	8 MB	1-5 Mbps
7		• Safari	Macintosh	16 MB	>5 Mbps
8		• Safari	Macintosh	16 MB	0-1 Mbps
9		• Safari	Windows	16 MB	1-5 Mbps
10		• IE	Windows	4 MB	>5 Mbps
11		• IE	Macintosh	4 MB	0-1 Mbps
12		• IE	Macintosh	4 MB	1-5 Mbps
13		• IE	Macintosh	8 MB	>5 Mbps
14		• IE	Macintosh	8 MB	0-1 Mbps
15		• IE	Windows	8 MB	1-5 Mbps
16		• IE	Macintosh	16 MB	>5 Mbps

The design is presented in a data table. Notice the following in the Table panel at the top left:

- The Design note indicates that this is a strength 3 covering array.
- The DOE Dialog script reproduces the Covering Array window settings.
- The Analysis script analyzes the experimental data.

JMP PRO Analyze the Experimental Data

Now that you have your design table, you can conduct your experiment and record your data in the Response column of the design table (Figure 20.6). Your experimental results are in the Software Data.jmp sample data table.

1. Select **Help > Sample Data Library** and open Design Experiment/Software Data.jmp.
2. In the Table panel, click the green triangle next to the **Analysis** script.

Figure 20.7 Analysis of Software Experimental Data

Summary		
Success Runs	38	
Failure Runs	3	
Missing	4	
Failure Analysis Details		
3 Factor Interactions		
Factors	Failure Levels	Failure Count
Web Browser, Operating System, RAM	Firefox, Macintosh, 8 MB	2
Web Browser, RAM, Connection Speed	Firefox, 4 MB, 0-1 Mbps	1
	Firefox, 8 MB, 1-5 Mbps	1
	Firefox, 8 MB, >5 Mbps	1

The Summary outline indicates that three tests failed and four tests did not result in a pass or fail outcome.

The Failure Analysis Details outline gives a breakdown of failures in terms of the associated three-way interactions. The outline lists only combinations of factor levels where all tests resulted in failure. If any test that involves a given three-way combination of settings results in success, then that three-way combination of settings cannot be responsible for system failure.

Two failures were associated with Web Browser set to Firefox, Operating System set to macOS, and RAM set to 8 MB. Notice that this combination led to failure regardless of the setting of Connection Speed.

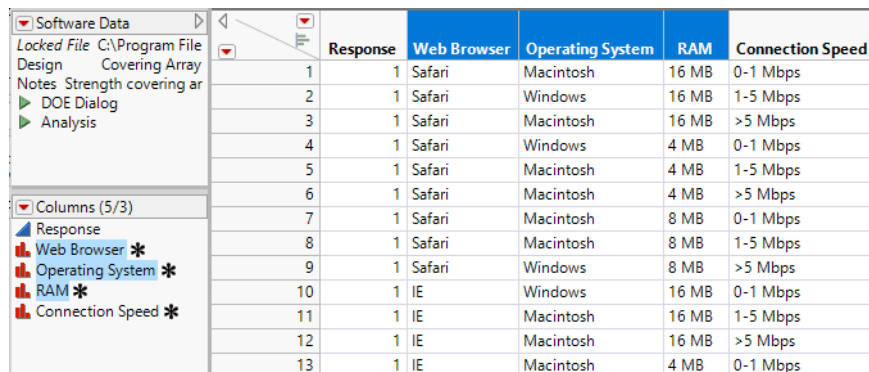
- 3. Select the first line in the 3 Factor Interactions report.

Figure 20.8 Selection of an Interaction in the Analysis Report

Failure Analysis Details		
3 Factor Interactions		
Factors	Failure Levels	Failure Count
Web Browser, Operating System, RAM	Firefox, Macintosh, 8 MB	2
Web Browser, RAM, Connection Speed	Firefox, 4 MB, 0-1 Mbps	1
	Firefox, 8 MB, 1-5 Mbps	1
	Firefox, 8 MB, >5 Mbps	1

This action selects the corresponding rows and columns in the data table.

Figure 20.9 Corresponding Selection of Rows and Columns in the Data Table



	Response	Web Browser	Operating System	RAM	Connection Speed
1	1	Safari	Macintosh	16 MB	0-1 Mbps
2	1	Safari	Windows	16 MB	1-5 Mbps
3	1	Safari	Macintosh	16 MB	>5 Mbps
4	1	Safari	Windows	4 MB	0-1 Mbps
5	1	Safari	Macintosh	4 MB	1-5 Mbps
6	1	Safari	Macintosh	4 MB	>5 Mbps
7	1	Safari	Macintosh	8 MB	0-1 Mbps
8	1	Safari	Macintosh	8 MB	1-5 Mbps
9	1	Safari	Windows	8 MB	>5 Mbps
10	1	IE	Windows	16 MB	0-1 Mbps
11	1	IE	Macintosh	16 MB	1-5 Mbps
12	1	IE	Macintosh	16 MB	>5 Mbps
13	1	IE	Macintosh	4 MB	0-1 Mbps

Three failures were associated with combinations of Web Browser, RAM, and Connection Speed. Note that two of these failures, Firefox, 8 MB, 1-5 Mbps and Firefox, 8 MB, >5 Mbps, are among the two failures for the Web Browser, Operating System, and RAM interaction. Selecting any of these rows in the report selects the corresponding rows and columns in the data table.

JMP PRO Example of a Covering Array with Factor Level Restrictions

The following example is patterned after an example described in Dalal and Mallows (1998). An originating phone (Near Phone) calls a receiving phone (Far Phone). Each phone call goes through an interface of type A or B. Five factors are of interest:

- Market: USA, UK, Canada, France, Mexico
- Near Phone: ISDN, Bus (Business), Coin, Res (Residential)
- Near Interface: A or B
- Far Phone: ISDN, Bus (Business), Coin, Res (Residential)
- Far Interface: A or B

You are interested in which combinations of pairs of these factors are likely to cause failures. However, certain combinations are not possible:

- An ISDN line on either phone (Near or Far) cannot use interface A.
- Business and Residential lines on the originating phone (Near) cannot use interface B.

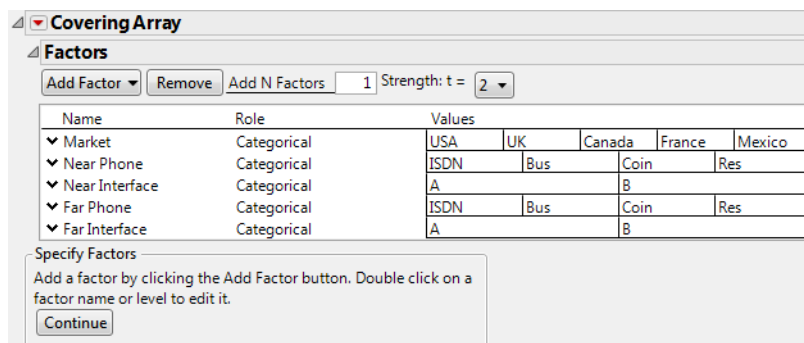
JMP PRO Create the Design

The factors and their settings are given in the data table Phone Factors.jmp. Create a Strength 2 covering array by following these steps.

JMP PRO Load Factors

1. Select **Help > Sample Data Library** and open Design Experiment/Phone Factors.jmp.
The Phone Factors.jmp data table contains the factors and their settings.
2. Select **DOE > Special Purpose > Covering Array**.
Notice that the menu next to **Strength: t =** , is set to 2 by default.
3. Click the Covering Array red triangle and select **Load Factors**.
The Factors outline is populated with the five factors and their levels.

Figure 20.10 Factors Outline for Phone Factors



4. Click **Continue**.
The Restrict Factor Level Combinations outline opens.

JMP PRO Restrict Factor Level Combinations

You can specify disallowed combinations in two ways:

- Use Disallowed Combinations Filter
- Use Disallowed Combinations Script

The filter gives an intuitive way to specify disallowed combinations. The script provides a quick and easy way to specify disallowed combinations, but requires that you have written or saved a script. In this example, if you do not want to specify combinations using the filter, skip to [“Specify Disallowed Combinations Using a Script”](#) on page 590.

Recall that the restrictions are the following:

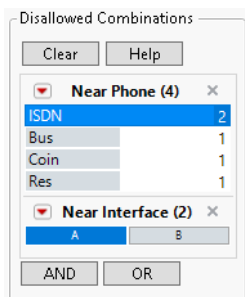
- An ISDN line on either phone (Near or Far) cannot use interface A.
- Business and Residential phones on the originating phone (Near) cannot use interface B.

JMP PRO Specify Disallowed Combinations Using the Filter

Use this approach to enter disallowed combinations using the filter interface. Alternatively, you can paste a script as shown in [“Specify Disallowed Combinations Using a Script”](#) on page 590.

1. Select **Use Disallowed Combinations Filter**.
2. From the Add Filter Factors list, select Near Phone and Near Interface and click **Add**.
3. Hold down the Ctrl key and click ISDN under Near Phone and A under Near Interface.

Figure 20.11 Disallowed Combinations Panel Showing First Constraints



Both blocks should turn dark. You have added the constraint that an ISDN line on the originating phone (Near) cannot use interface A.

4. Click **OR**.
5. From the Add Filter Factors list, select Far Phone and Far Interface and click **Add**.
6. Hold down the Ctrl key and click ISDN under Far Phone and A under Far Interface.
You have added the constraint that an ISDN line on the receiving phone (Far) cannot use interface A.
7. Click **OR**.
8. From the Add Filter Factors list, select Near Phone and Near Interface and click **Add**.
9. Hold down the Ctrl key and click Bus and Res under Near Phone and B under Near Interface.

You have added the restriction that Business and Residential lines on the originating phone (Near) cannot use interface B.

Figure 20.12 Completed Disallowed Combinations Filter

Disallowed Combinations

Clear Help

☒ Near Phone (4) ×

ISDN	2
Bus	1
Coin	1
Res	1

☒ Near Interface (2) ×

A	B

OR

☒ Far Phone (4) ×

ISDN	2
Bus	1
Coin	1
Res	1

☒ Far Interface (2) ×

A	B

OR

☒ Near Phone (4) ×

ISDN	2
Bus	1
Coin	1
Res	1

☒ Near Interface (2) ×

A	B

AND OR

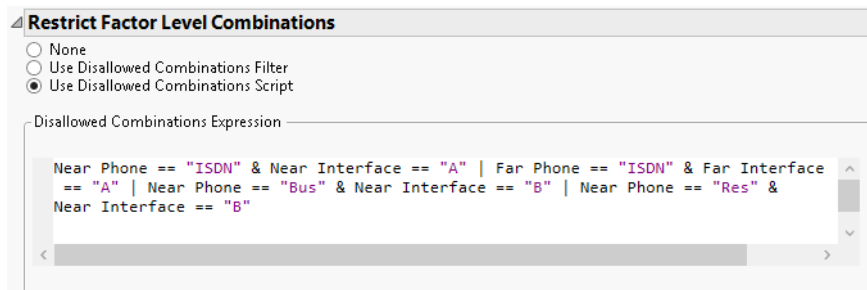
JMP^{PRO} Specify Disallowed Combinations Using a Script

Alternatively, you can specify disallowed combinations by constructing a script. After loading your factors ("[Load Factors](#)" on page 588), do the following:

1. Click **Continue**.
2. Select **Use Disallowed Combinations Script**.
3. Copy the following script and paste it in the Disallowed Combinations Expression script box:

```
(Near Phone == "ISDN" & Near Interface == "A") |
(Far Phone == "ISDN" & Far Interface == "A") |
(Near Phone == "Bus" & Near Interface == "B") |
(Near Phone == "Res" & Near Interface == "B")
```

Figure 20.13 Completed Disallowed Combinations Script Window



JMP PRO Construct the Design Table

Note: Setting the Random Seed in the next two steps reproduces the exact results shown in this example. When constructing a design on your own, these steps are not necessary.

1. Click the Covering Array red triangle and select **Set Random Seed**.
2. Enter 632 and click **OK**.
3. Click **Make Design**.

The Design outline opens to show a 20-run design. A Metrics outline is added to the window.

Figure 20.14 Metrics Outline for Phone Design

Metrics		
Number of Runs:		20
t	Coverage	Diversity
2	100.00	54.00
3	63.14	83.00
4	30.29	97.00

The Metrics outline indicates that Strength 2 coverage is 100%. This means that all permissible two-factor combinations are represented in the design. The design also covers 65% of all three-factor combinations.

4. Click **Make Table**.

The design is placed in a design table. A column for the response is provided, as well as various scripts.

Figure 20.15 Covering Array Design Table

	Response	Market	Near Phone	Near Interface	Far Phone	Far Interface
1		• Mexico	Res	A	Res	A
2		• Mexico	Bus	A	Bus	A
3		• France	Coin	B	Bus	B
4		• UK	Res	A	Coin	A
5		• UK	ISDN	B	Bus	B
6		• France	ISDN	B	Res	A
7		• Canada	Bus	A	Res	A
8		• Mexico	ISDN	B	ISDN	B
9		• Canada	ISDN	B	Coin	B
10		• UK	Bus	A	ISDN	B
11		• Canada	Res	A	Bus	B
12		• USA	ISDN	B	ISDN	B
13		• UK	Coin	B	Res	B
14		• Mexico	Coin	B	Coin	B
15		• USA	Bus	A	Bus	A
16		• USA	Coin	A	Coin	A
17		• USA	Res	A	Res	A
18		• France	Res	A	ISDN	B
19		• France	Bus	A	Coin	B
20		• Canada	Coin	B	ISDN	B

JMP® PRO Analyze the Experimental Data

Now that you have your design table, you can conduct your experiment and record your data in the Response column of the design table. Your experimental results are in the Phone Data.jmp sample data table.

1. Select **Help > Sample Data Library** and open Design Experiment/Phone Data.jmp.
2. In the Table panel, click the green triangle next to the **Analysis** script.

Figure 20.16 Analysis of Phone Experimental Data

Summary		
Success Runs	17	
Failure Runs	3	
Missing	0	
Failure Analysis Details		
2 Factor Interactions		
Factors	Failure Levels	Failure Count
Near Interface, Far Phone	A, Coin	3
Far Phone, Far Interface	Coin, A	2
Market, Near Phone	USA, Coin	1
Market, Near Phone	UK, Res	1
Market, Near Phone	France, Bus	1
Market, Far Phone	USA, Coin	1
Market, Far Phone	UK, Coin	1
Market, Far Phone	France, Coin	1
Market, Far Interface	UK, A	1
Near Phone, Near Interface	Coin, A	1
Near Phone, Far Phone	Bus, Coin	1
Near Phone, Far Phone	Res, Coin	1
Near Phone, Far Interface	Coin, A	1

The Summary outline indicates that three tests failed.

The Failure Analysis Details outline contains a 2 Factor Interactions report, because a two-way interaction is the lowest level interaction that detects a failure.

The 2 Factor Interactions report shows the combinations that might have caused the three failures. It is possible that one combination, Near Interface set to A and Far Phone set to Coin, is responsible for all three failures. Or it is possible that two or three other combinations caused the three failures.

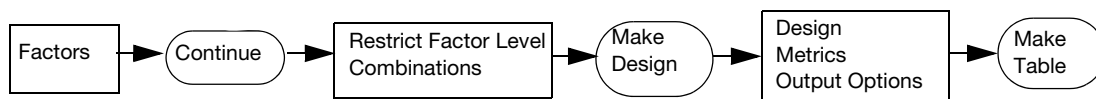
3. Select the first line in the 2 Factor Interactions report.

In the data table, rows 4, 16, and 19 are selected. Failures occur for these combinations, regardless of the settings for Market, Near Phone, and Far Interface. But note that other combinations of factor settings could account for these failures as well.

JMP PRO Covering Array Window

The covering array window updates as you work through the design steps. For more information about the general flow of DOE windows, see [“The DOE Workflow: Describe, Specify, Design”](#) on page 57. The outlines, separated by buttons that update the outlines, follow the flow in Figure 20.17.

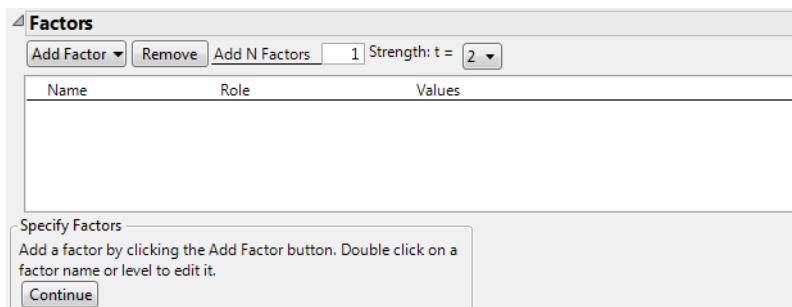
Figure 20.17 Covering Array Flow



JMP PRO Factors

Add factors in the Factors outline.

Figure 20.18 Factors Outline



Add Factor Enters the number of factors specified in **Add N Factors**. All factors are categorical. Select or specify the number of levels.

Remove Removes the selected factors.

Add N Factors Adds multiple factors with a specific number of levels. Enter the number of factors to add, click **Add Factor**, and select or specify the number of levels. Repeat **Add N Factors** to add multiple factors with different numbers of levels.

Strength t = Select a value to specify the strength of the array.

Tip: When you have completed your Factors panel, select **Save Factors** from the red triangle menu. This saves the factor names and values in a data table that you can later reload. See [“Covering Array Options”](#) on page 603.

JMP PRO Factors Table

The Factors table contains the following columns:

- Name** The name of the factor. When a factor is added, it is given a default name of X1, X2, and so on. To change this name, double-click it and enter the desired name.
- Role** Specifies the Design Role of the factor. The Design Role for all covering array factors is Categorical. The Design Role column property for the factor is saved to the data table. This property ensures that the factor is modeled appropriately.
- Values** The settings for the factors. To insert Values, click the default values and enter the desired values. The value ordering in the design table is the order of the values as entered from left to right.

JMP PRO Editing the Factors Table

In the Factors outline, notice the following:

- To edit a factor name, double-click the factor name.
- Categorical factors have a down arrow to the left of the factor name. Click the arrow to add a level.
- To remove a factor level, click the value, click **Delete**, and click outside the text box.
- To edit a value, click the value in the Values column.

JMP PRO Factor Column Properties

For each factor, the Value Labels column property is saved to the design table. The Value Labels column property represents values in a column with specified labels. These labels are shown in the data table and are used in plots and reports. See [“Value Labels”](#) on page 763 in the “Column Properties” appendix.

JMP PRO Restrict Factor Level Combinations

When you complete the Factors outline and click Continue, the Restrict Factor Level Combinations outline appears. This outline enables you to specify factor level combinations that are prohibited. Unless you have loaded a constraint or included one as part of a script, the **None** option is selected. To specify constraints, select one of the other options:

- Use Disallowed Combinations Filter** Defines sets of constraints based on restricting values of individual factors. You can define both AND and OR constraints. See [“Use Disallowed Combinations Filter”](#) on page 596.

Use Disallowed Combinations Script Defines disallowed combinations and other constraints as Boolean JSL expressions in a script editor box. See [“Use Disallowed Combinations Script”](#) on page 597.

Use Disallowed Combinations Filter

This option uses an adaptation of the Data Filter to facilitate specifying disallowed combinations. For detailed information about using the Data Filter, see the JMP Reports chapter in *Using JMP*.

To add disallowed combinations:

1. Select factors from the Add Filter Factors list and click **Add**.
2. Specify the disallowed combinations by selecting levels.

Note: The red triangle options in the Add Filter Factors menu are the same as those found in the Select Columns panel of many platform launch windows. See the Get Started chapter in *Using JMP*.

When you click Add, the initial panel is updated. The Disallowed Combinations control panel shows the selected factors and provides options for further control.

The Covering Array platform allows only categorical factors. For categorical factors, the possible levels are shown either as labeled blocks or, when the number of levels is large, as list entries. Select a level to disallow it. To select multiple levels, hold down the Ctrl key. The block or list entries are highlighted to indicate the levels that have been disallowed. When you add a factor to the Disallowed Combinations panel, the number of levels of the categorical factor is given in parentheses following the factor name.

Disallowed Combinations Options

Clear Clears all disallowed factor level settings that you have specified. This does not clear the selected factors.

AND Opens the Add Filter Factors list. Selected factors become an AND group. Any combination of factor levels specified within an AND group is disallowed.

To add a factor to an AND group later on, click the group’s outline to see a highlighted rectangle. Select AND and add the factor.

To remove a single factor, select **Delete** from its red triangle menu.

OR Opens the Add Filter Factors list. Selected factors become a separate AND group. For AND groups separated by OR, a combination is disallowed if it is specified in at least one AND group.

Red Triangle Options for Factors

A factor can appear in several OR groups. An occurrence of the factor in a specific OR group is referred to as an *instance* of the factor.

Delete Removes the selected instance of the factor from the Disallowed Combinations panel.

Clear Selection Clears any selection for that instance of the factor.

Invert Selection Deselects the selected values and selects the values not previously selected for that instance of the factor.

Display Options Changes the appearance of the display. Options include the following:

Blocks Display Shows each level as a block.

List Display Shows each level as a member of a list.

Single Category Display Shows each level.

Check Box Display Adds a check box next to each value.

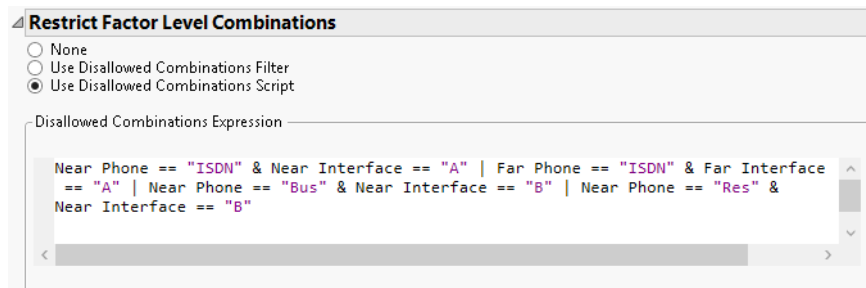
Radio Box Display Adds a radio selection next to each value for a single selection.

Find Available only for categorical factors. Provides a text box beneath the factor name where you can enter a search string for levels of the factor. Press the Enter key or click outside the text box to perform the search. Once **Find** is selected, Find options appear in the red triangle menu, such as Does not contain, Match Case, Starts with or Ends with, or Clear Find.

Use Disallowed Combinations Script

This option opens a script window where you insert a script that identifies the combinations that you want to disallow. The script must evaluate as a Boolean expression. When the expression evaluates as true, the specified combination is disallowed.

When creating the expression, use the name of the level in quotation marks. Do not use the ordinal value of the level. For example, Figure 20.19 shows the script that you entered in the phone interface example, [“Specify Disallowed Combinations Using a Script”](#) on page 590.

Figure 20.19 Script Window Showing Names of Levels in Quotes

JMP PRO Design

When you click Make Design, the Design and Metrics outlines appear. For designs that require extensive computation, a progress bar appears.

The Design outline shows the design that you have constructed. The first column lists a Run order. You might need to use the scroll bar to view all the runs. The remaining columns show factor settings for each run.

Optimize

Select Optimize to reduce the size of a design that was constructed by the Covering Array platform or that you have loaded using the Load Design red triangle option. Optimize is not available for designs constructed by the Covering Array platform that are known to be optimal. In particular, all unconstrained strength 2 designs for two-level factors constructed by the platform are optimal. Also, any unconstrained strength t design for $t+1$ factors is optimal for any t .

For more information about the algorithm, see [“Algorithm for Optimize”](#) on page 604.

Note: Optimize is time intensive, but can be run repeatedly to yield incrementally better designs.

Use the **Maximum iterations** option to specify a maximum number of iterations to be used in optimizing the design.

JMP PRO Unsatisfiable Constraints

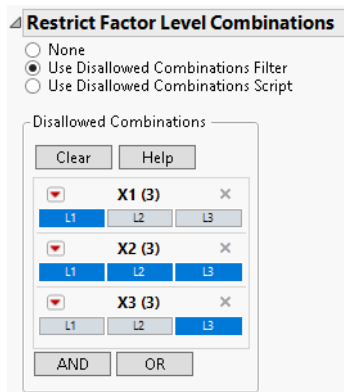
If a set of constraints prohibits the construction of a covering array where all required factor levels are represented, it is said to be *unsatisfiable*.

Example of Unsatisfiable Constraints

Consider a strength 2 design for three factors, each at three levels.

1. Select **DOE > Special Purpose > Covering Array**.
2. Next to **Add N Factors**, type 3.
3. From the **Add Factor** menu, select **3 Level**.
4. Click **Continue**.
5. Select **Use Disallowed Combinations Filter**.
6. From the **Add Filter Factors** list, select all three factors and click **Add**.
7. Hold down the Ctrl key and select the following levels:
 - For X1, select L1.
 - For X2, select L1, L2, and L3.
 - For X3, select L3.

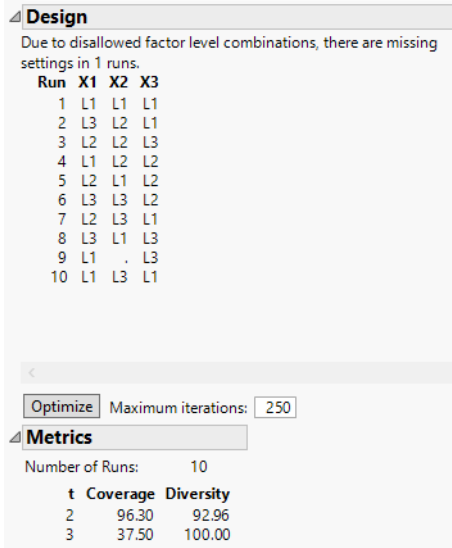
Figure 20.20 Completed Restrict Factor Level Combinations Panel



Note: Setting the Random Seed in the next two steps reproduces the exact results shown in this example. When constructing a design on your own, these steps are not necessary.

8. Click the Covering Array red triangle and select **Set Random Seed**.
9. Enter 12345 and click **OK**.
10. Click **Make Design**.

Figure 20.21 Design and Metrics Outlines



A note beneath the Design outline indicates that one run has a missing setting due to the constraints. That run is run 9. To ensure that the covering array has strength 2, the combination of X1 set to L1 and X3 set to L3 is required. But for these settings, the constraints prohibit all settings for X2.

JMP PRO Metrics

The Metrics outline gives you information about how well the design meets the strength requirements. See Dalal and Mallows (1998) for background on these metrics for unconstrained designs.

t The number of factors.

Coverage The ratio of the number of distinct *t*-factor settings that appear in the design to the total possible number of *t*-factor settings, expressed as a percentage. A *t*-coverage of 100% indicates that all possible *t*-factor settings are covered by the design. Note that each *t*-factor setting can appear multiple times.

For constrained and unsatisfiable designs, the definition of Coverage is adjusted for the number of *t*-factor settings that are possible once the constraints have been applied to all *t*-factor combinations. See “Formulas for Metrics” on page 604.

Diversity The ratio of the number of distinct *t*-factor settings in the design to the total number of occurrences of *t*-factor settings in the design, expressed as a percentage. The *t*-diversity measures how well the design avoids replication. A *t*-diversity of 100%

indicates that no t -factor settings are repeated. A t -diversity of 50% indicates that the average number of times that distinct t -factor settings appear is two.

For constrained and unsatisfiable designs, the definition of Diversity is adjusted for the number of runs with missing settings. See [“Formulas for Metrics”](#) on page 604.

JMP PRO Output Options

Make Table Constructs the Covering Array data table.

Back Takes you back to the Factors outline. You can make changes to the previous outlines and regenerate the design.

Note: If you have defined Disallowed Combinations in the Restrict Factor Level Combinations outline, these are retained as a script. The script is shown in the Use Disallowed Combinations Script panel when you click Continue.

JMP PRO The Covering Array Data Table

The Covering Array data table contains a first column where you can enter the response. The remaining columns give the factor settings.

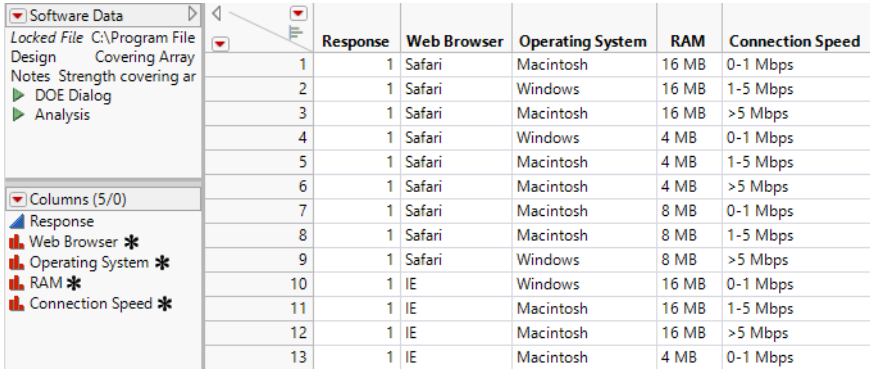
The Table panel in the upper left contains a Design note indicating that the design is a Covering Array and giving the Strength of the design. The Table panel also contains the following scripts.

DOE Dialog Re-creates the Covering Array window that you used to generate the design table.

Disallowed Combinations Shows factor level restrictions that you entered in the Restrict Factor Level Combinations outline.

Analysis Provides an analysis of your experimental data. For more information about the script, see [“Analysis Script”](#) on page 602. For background, see Zhang and Zhang (2011).

Tip: To run a script, click the green triangle next to the script name.

Figure 20.22 Partial View of Covering Array Table for Software Data.jmp Showing Scripts


	Response	Web Browser	Operating System	RAM	Connection Speed
1	1	Safari	Macintosh	16 MB	0-1 Mbps
2	1	Safari	Windows	16 MB	1-5 Mbps
3	1	Safari	Macintosh	16 MB	>5 Mbps
4	1	Safari	Windows	4 MB	0-1 Mbps
5	1	Safari	Macintosh	4 MB	1-5 Mbps
6	1	Safari	Macintosh	4 MB	>5 Mbps
7	1	Safari	Macintosh	8 MB	0-1 Mbps
8	1	Safari	Macintosh	8 MB	1-5 Mbps
9	1	Safari	Windows	8 MB	>5 Mbps
10	1	IE	Windows	16 MB	0-1 Mbps
11	1	IE	Macintosh	16 MB	1-5 Mbps
12	1	IE	Macintosh	16 MB	>5 Mbps
13	1	IE	Macintosh	4 MB	0-1 Mbps

JMP PRO Analysis Script

The Analysis script assumes the following about the Response column:

- The responses are recorded as 0 for failure and 1 for success.
- Missing values are permitted.
- The Response column is continuous.
- You can rename or move the Response column.

The Analysis script produces a report with two outlines:

- The Summary outline gives the number of runs resulting in Success, Failure, and the number of runs for which the response is Missing.
- The Failure Analysis Details report contains a <k> Factor Interactions report. The value of k is the smallest number of interactions that detect a failure. (For a definition of *detect*, see [“Covering Arrays and Strength”](#) on page 581.) The three columns contain the following:
 - The **Factors** column lists all k -factor combinations that detect failures.
 - The **Failure Levels** column lists the values of the k factors in the **Factors** column that detect failures.
 - The **Failure Count** column gives the number of failures corresponding to the k -factor combination of **Failure Levels**.

Note: A failure observation can appear in more than one of the k -factor combinations listed in the **Failure Levels** column.

The rows in the <k> Factor Interactions report are dynamically linked to the data table. If you select one or more rows in the report, the corresponding rows are selected in the data table.

Covering Array Options

The red triangle menu in the Covering Array platform contains these options:

Save Factors Creates a data table containing a column for each factor that contains its factor levels. Each factor's column contains these column properties: Design Role, Value Order, and Factor Changes. Saving factors enables you to quickly load them into a DOE window.

Note: You can create a factors table for a Covering Array by entering data into an empty table, but remember to assign each column an appropriate Design Role of Categorical. Right-click the column name in the data grid and select **Column Properties > Design Role**. In the Design Role area, select Categorical.

Load Factors Loads factors that you have saved using the Save Factors option into the Factors outline.

Load Design Loads a design from the active data table. If no data table is active, you are prompted to open one. When you select Load Design, a menu appears that enables you to select the columns that you want to specify as factors in the design. All columns are imported as categorical. Columns and their values are listed in the Factors outline. The Design outline shows a Run for each row in the data table and gives the values of the factors for each run.

The Load Design options enables you to obtain metrics, modify, or construct an Analysis script for an existing design:

- The Metrics outline shows *t*-Coverage and *t*-Diversity for the specified design.
- You can click **Back** to impose factor level restrictions and then construct a new design.
- Clicking **Make Table** constructs a design table where you can enter responses. The table contains an Analysis script for the design.

Set Random Seed Sets the random seed that JMP uses to control certain actions that have a random component. For Covering Arrays, the seed selects a starting design and an iteration count. For most designs, the random seed guarantees reproducibility of the design, but not of the run order.

Note: Upper limits on time as well as iteration count are used to limit design construction time. For some large and high strength designs, depending on the machine, the time limit might override the iteration limit. For such designs, the random seed does not guarantee reproducibility.

To reproduce a design, enter the random seed used to generate it before clicking Make Design.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Advanced Options Not available for Covering Arrays.

Save Script to Script Window Creates the script for the design that you specified in the Covering Array window and places it in an open script window.

JMP PRO Technical Details for Covering Arrays

- [“Algorithm for Optimize”](#)
- [“Formulas for Metrics”](#)

JMP PRO Algorithm for Optimize

The Optimize button invokes an algorithm that is conceptually similar to a class of covering array optimizers sometimes referred to as *post-construction randomized optimizers* (Nayeri et al., 2013). However, the algorithm in JMP differs from most in that it also addresses designs with constraints. In particular, it optimizes constrained covering arrays as well as unsatisfiable, constrained covering arrays.

The algorithm assumes that the design to be optimized is a covering array of the specified strength. For a K factor design of strength t , the algorithm iteratively examines all ${}_K C_t$ factor projections to determine whether runs can be eliminated or merged. Consequently, as K or t increases, the run time of the algorithm quickly escalates. To improve performance, the JMP implementation is threaded to use as many CPU cores as are available on your workstation.

JMP PRO Formulas for Metrics

The formulas for Coverage and Diversity depend on whether there are constraints. The following notation is used:

${}_u C_v$ is the number of combinations of u things taken v at a time

t is the strength of the design

K is the number of factors

$M = {}_K C_t$

$i = 1, 2, \dots, M$ is an index that orders all combinations, or *projections*, of t factors

v_{ik} is the number of levels for the k^{th} factor

n_i is the number of distinct t tuples in the design for the i^{th} projection

p_i is the product of the v_{ik} for the factors in the i^{th} projection

r is the number of runs in the design

JMP PRO Unconstrained Design

Coverage and Diversity are given by the following:

$$\text{Coverage} = \frac{1}{M} \sum_{i=1}^M n_i / p_i$$

$$\text{Diversity} = \frac{1}{M} \sum_{i=1}^M n_i / r$$

JMP PRO Constrained Design

In a constrained design, certain t tuples are not allowed. This can result in missing values for some t tuples. For some combinations of t factors, there might be no valid t tuples whatsoever. Coverage and diversity must be defined in terms of the possible valid combinations. For this reason, the formulas for constrained designs require additional notation:

a_i is the number of invalid t tuples arising from factors in the i^{th} projection

m is the number of projections where there are no valid t tuples

q_i is the number of runs in the design with missing values for any factor in the i^{th} projection

$$r_i = r - q_i$$

$$M' = M - m$$

Coverage and Diversity are given by the following:

$$\text{Coverage} = \frac{1}{M'} \sum_{i=1}^{M'} n_i / (p_i - a_i)$$

$$\text{Diversity} = \frac{1}{M'} \sum_{i=1}^{M'} n_i / r_i$$

If there are no invalid t tuples ($M' = M$) and if there are no missing values ($r_i = r$, for all i), then the definitions for coverage and diversity for constrained designs reduce to the definitions for unconstrained designs. See Morgan (2014).

Chapter 21

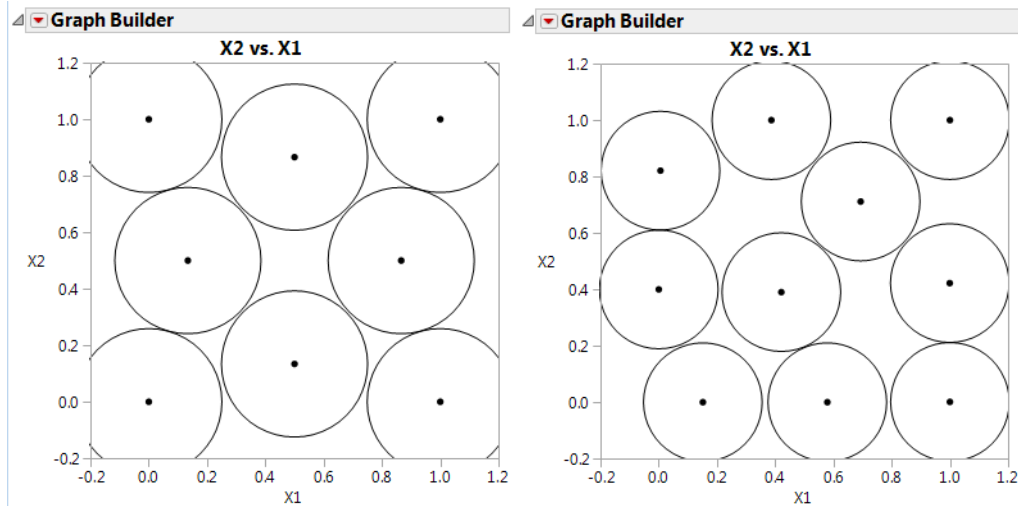
Space-Filling Designs

Space-filling designs are useful in situations where run-to-run variability is of far less concern than the form of the model. Consider a sensitivity study of a computer simulation model. In this situation, and for any mechanistic or deterministic modeling problem, any variability is small enough to be ignored. For systems with no variability, replication, randomization, and blocking are irrelevant.

The Space Filling platform provides designs for situations with both continuous and categorical factors. For continuous factors, space-filling designs have two objectives:

- maximize the distance between any two design points
- space the points uniformly

Figure 21.1 Space-Filling Design



Contents

Overview of Space-Filling Designs.....	609
Space Filling Design Window.....	610
Responses.....	610
Factors.....	612
Define Factor Constraints.....	614
Space Filling Design Methods.....	617
Design.....	618
Design Diagnostics.....	618
Design Table.....	618
Space Filling Design Options.....	618
Sphere-Packing Designs.....	621
Creating a Sphere-Packing Design.....	621
Visualizing the Sphere-Packing Design.....	622
Latin Hypercube Designs.....	624
Creating a Latin Hypercube Design.....	624
Visualizing the Latin Hypercube Design.....	625
Uniform Designs.....	627
Comparing Sphere-Packing, Latin Hypercube, and Uniform Methods.....	629
Minimum Potential Designs.....	630
Maximum Entropy Designs.....	633
Gaussian Process IMSE Optimal Designs.....	634
Fast Flexible Filling Designs.....	636
FFF Optimality Criterion.....	636
Set Average Cluster Size.....	637
Constraints.....	637
Creating and Viewing a Constrained Fast Flexible Filling Design.....	638
Creating a Space-Filling Design for a Map Shape.....	640
Example of a Sphere-Packing Design.....	645
Create the Sphere-Packing Design for the Borehole Data.....	645
Guidelines for the Analysis of Deterministic Data.....	646

Overview of Space-Filling Designs

Space-filling designs are useful for modeling systems that are deterministic or near-deterministic. One example of a deterministic system is a computer simulation. Such simulations can be very complex involving many variables with complicated interrelationships. A goal of designed experiments on these systems is to find a simpler empirical model that adequately predicts the behavior of the system over limited ranges of the factors.

In experiments on systems where there is substantial random noise, the goal is to minimize the variance of prediction. In experiments on deterministic systems, there is no variance but there is *bias*. Bias is the difference between the approximation model and the true mathematical function. The goal of space-filling designs is to bound the bias.

One approach to bound the bias is to spread the design points out as far from each other as possible while staying inside the experimental boundaries. The other approach is to space the points out evenly over the region of interest.

The **Space Filling** designer supports the following design methods:

Sphere Packing Maximizes the minimum distance between pairs of design points. See [“Sphere-Packing Designs”](#) on page 621 and [“Create the Sphere-Packing Design for the Borehole Data”](#) on page 645.

Latin Hypercube Maximizes the minimum distance between design points but requires even spacing of the levels of each factor. This method produces designs that mimic the uniform distribution. The Latin Hypercube method is a compromise between the Sphere-Packing method and the Uniform design method. See [“Latin Hypercube Designs”](#) on page 624.

Uniform Minimizes the discrepancy between the design points (which have an empirical uniform distribution) and a theoretical uniform distribution. See [“Uniform Designs”](#) on page 627.

Minimum Potential Spreads points out inside a sphere around the center. See [“Minimum Potential Designs”](#) on page 630.

Maximum Entropy Measures the amount of information contained in the distribution of a set of data. See [“Maximum Entropy Designs”](#) on page 633.

Gaussian Process IMSE Optimal Creates a design that minimizes the integrated mean squared error of the Gaussian process over the experimental region. See [“Gaussian Process IMSE Optimal Designs”](#) on page 634.

Fast Flexible Filling The Fast Flexible Filling method forms clusters from random points in the design space. These clusters are used to choose design points according to an optimization criterion. This is the only method that can accommodate categorical factors

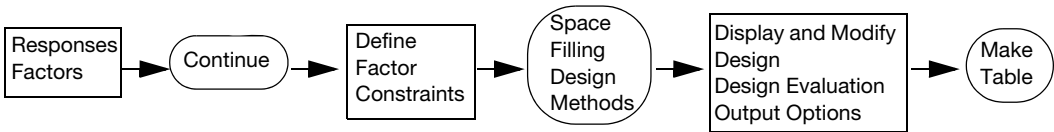
and constraints on the design space. You can specify linear constraints and disallowed combinations. See [“Fast Flexible Filling Designs”](#) on page 636 and [“Creating and Viewing a Constrained Fast Flexible Filling Design”](#) on page 638.

Note: If the number of runs is 500 or less, a Gaussian Process model is saved to the data table. If the number of runs exceeds 500, a Neural model is saved to the data table.

Space Filling Design Window

The Space Filling Design window updates as you work through the design steps. The outlines that appear, separated by buttons that update the window, follow the flow in Figure 21.2.

Figure 21.2 Space Filling Design Flow



Responses

Use the Responses outline to specify one or more responses.

Tip: When you have completed the Responses outline, consider selecting **Save Responses** from the red triangle menu. This option saves the response names, goals, limits, and importance values in a data table that you can later reload in DOE platforms.

Figure 21.3 Responses Outline

The screenshot shows the 'Responses' window with a table containing one response named 'Y' with a goal of 'Maximize'. The table has columns for Response Name, Goal, Lower Limit, Upper Limit, and Importance. There are buttons for 'Add Response', 'Remove', and 'Number of Responses...' at the top.

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.

Add Response Enters a single response with a goal type of Maximize, Match Target, Minimize, or None. If you select Match Target, enter limits for your target value. If you select Maximize or Minimize, entering limits is not required but can be useful if you intend to use desirability functions.

Remove Removes the selected responses.

Number of Responses Enters additional responses so that the number that you enter is the total number of responses. If you have entered a response other than the default Y, the Goal for each of the additional responses is the Goal associated with the last response entered. Otherwise, the Goal defaults to Match Target. Click the Goal type in the table to change it.

The Responses outline contains the following columns:

Response Name The name of the response. When added, a response is given a default name of Y, Y2, and so on. To change this name, double-click it and enter the desired name.

Goal, Lower Limit, Upper Limit The Goal tells JMP whether you want to maximize your response, minimize your response, match a target, or that you have no response goal. JMP assigns a Response Limits column property, based on these specifications, to each response column in the design table. It uses this information to define a desirability function for each response. The Profiler and Contour Profiler use these desirability functions to find optimal factor settings. For further details, see the Profiler chapter in *Profiler* and [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

- A Goal of Maximize indicates that the best value is the largest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Minimize indicates that the best value is the smallest possible. If there are natural lower or upper bounds, you can specify these as the Lower Limit or Upper Limit.
- A Goal of Match Target indicates that the best value is a specific target value. The default target value is assumed to be midway between the Lower Limit and Upper Limit.
- A Goal of None indicates that there is no goal in terms of optimization. No desirability function is constructed.

Note: If your target response is not midway between the Lower Limit and the Upper Limit, you can change the target after you generate your design table. In the data table, open the Column Info window for the response column (**Cols > Column Info**) and enter the desired target value.

Importance When you have several responses, the Importance values that you specify are used to compute an overall desirability function. These values are treated as weights for the responses. If there is only one response, then specifying the Importance is unnecessary because it is set to 1 by default.

Editing the Responses Outline

In the Responses outline, note the following:

- Double-click a response to edit the response name.
- Click the goal to change it.
- Click on a limit or importance weight to change it.
- For multiple responses, you might want to enter values for the importance weights.

Response Limits Column Property

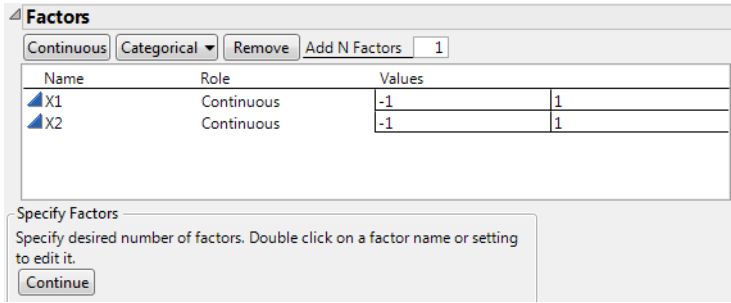
The Goal, Lower Limit, Upper Limit, and Importance that you specify when you enter a response are used in finding optimal factor settings. For each response, the information is saved in the generated design data table as a Response Limits column property. JMP uses this information to define the desirability function. The desirability function is used in the Prediction Profiler to find optimal factor settings. For further details about the Response Limits column property and examples of its use, see [“Response Limits”](#) on page 732 in the “Column Properties” appendix.

If you do not specify a Lower Limit and Upper Limit, JMP uses the range of the observed data for the response to define the limits for the desirability function. Specifying the Lower Limit and Upper Limit gives you control over the specification of the desirability function. For more details about the construction of the desirability function, see the Profiler chapter in *Profilers*.

Factors

Add factors in the Factors outline.

Figure 21.4 Factors Outline



The Factors outline contains these options:

- Continuous** Enters the number of continuous factors specified in **Add N Factors**.
- Categorical** Enters the number of nominal factors specified in **Add N Factors**.

Remove Removes the selected factors.

Add N Factors Adds multiple factors of a given type. Enter the number of factors to add and click Continuous or Categorical. Repeat **Add N Factors** to add multiple factors of different types.

Tip: When you have completed your Factors panel, select **Save Factors** from the red triangle menu. This saves the factor names and values in a data table that you can later reload. See [“Space Filling Design Options”](#) on page 618.

Factors Outline

The Factors outline contains the following columns:

Name The name of the factor. When added, a factor is given a default name of X1, X2, and so on. To change this name, double-click it and enter the desired name.

Role Specifies the Design Role of the factor. The Design Role column property for the factor is saved to the data table. This property ensures that the factor type is modeled appropriately.

Values The experimental settings for the factors. To insert Values, click the default values and enter the desired values.

Editing the Factors Outline

In the Factors outline, note the following:

- To edit a factor name, double-click the factor name.
- Categorical factors have a down arrow to the left of the factor name. Click the arrow to add a level.
- To remove a factor level, click the value, click **Delete**, and click outside the text box.
- To edit a value, click the value in the Values column.

Factor Types

Continuous Numeric data types only. A continuous factor is a factor that you can conceptually set to any value between the lower and upper limits you supply, given the limitations of your process and measurement system.

Categorical Either numeric or character data types. For a categorical factor, the value ordering is the order of the values as entered from left to right. This ordering is saved in a Value Order column property after the design data table is created.

Factor Column Properties

For each factor, various column properties are saved to the data table.

Design Role Each factor is assigned the Design Role column property. The Role that you specify in defining the factor determines the value of its Design Role column property. The Design Role property reflects how the factor is intended to be used in modeling the experimental data. Design Role values are used in the Augment Design platform.

Factor Changes Each factor is assigned the Factor Changes column property with a setting of Easy. In space-filling designs, it is assumed that factor levels can be changed for each experimental run. Factor Changes values are used in the Evaluate Design and Augment Design platforms.

Coding If the Role is Continuous, the Coding column property for the factor is saved. This property transforms the factor values so that the low and high values correspond to -1 and +1, respectively. The estimates and tests in the Fit Least Squares report are based on the transformed values.

Value Order If the Role is Categorical or Blocking, the Value Order column property for the factor is saved. This property determines the order in which levels of the factor appear.

Define Factor Constraints

Note: Constraints can be specified only for designs constructed using the Fast Flexible Filling method.

Use Define Factor Constraints to restrict the design space. Unless you have loaded a constraint or included one as part of a script, the **None** option is selected. To specify constraints, select one of the other options:

Specify Linear Constraints Specifies inequality constraints on linear combinations of factors. Available only for factors with a Role of Continuous or Mixture. See [“Specify Linear Constraints”](#).

Note: When you save a script for a design that involves a linear constraint, the script expresses the linear constraint as a *less than or equal to* inequality ($\leq$).

Use Disallowed Combinations Filter Defines sets of constraints based on restricting values of individual factors. You can define both AND and OR constraints. See [“Use Disallowed Combinations Filter”](#).

Use Disallowed Combinations Script Defines disallowed combinations and other constraints as Boolean JSL expressions in a script editor box. See [“Use Disallowed Combinations Script”](#).

Specify Linear Constraints

In cases where it is impossible to vary continuous factors independently over the design space, you can specify linear inequality constraints. Linear inequalities describe factor level settings that are allowed.

Click **Add** to enter one or more linear inequality constraints.

Add Adds a template for a linear expression involving all the continuous factors in your design. Enter coefficient values for the factors and select the direction of the inequality to reflect your linear constraint. Specify the constraining value in the box to the right of the inequality. To add more constraints, click **Add** again.

Note: The Add option is disabled if you have already constrained the design region by specifying a Sphere Radius.

Remove Last Constraint Removes the last constraint.

Check Constraints Checks the constraints for consistency. This option removes redundant constraints and conducts feasibility checks. A JMP alert appears if there is a problem. If constraints are equivalent to bounds on the factors, a JMP alert indicates that the bounds in the Factors outline have been updated.

Use Disallowed Combinations Filter

This option uses an adaptation of the Data Filter to facilitate specifying disallowed combinations. For detailed information, see the JMP Reports chapter in *Using JMP*.

Select factors from the Add Filter Factors list and click **Add**. Then specify the disallowed combinations by using the slider (for continuous factors) or by selecting levels (for categorical factors).

The red triangle options for the Add Filter Factors menu are those found in the Select Columns panel of many platform launch windows. See the Get Started chapter in *Using JMP*.

When you click Add, the Disallowed Combinations control panel shows the selected factors and provides options for further control. Factors are represented as follows, based on their modeling types:

Continuous Factors For a continuous factor, a double-arrow slider that spans the range of factor settings appears. An expression that describes the range using an inequality appears above the slider. You can specify disallowed settings by dragging the slider arrows or by clicking on the inequality bounds in the expression and entering your desired constraints. In the slider, a solid blue highlight represents the disallowed values.

Categorical Factor For a categorical factor, the possible levels are displayed either as labeled blocks or, when the number of levels is large, as list entries. Select a level to disallow it. To

select multiple levels, hold the Control key. The block or list entries are highlighted to indicate the levels that have been disallowed. When you add a categorical factor to the Disallowed Combinations panel, the number of levels of the categorical factor is given in parentheses following the factor name.

Disallowed Combinations Options

The control panel has the following controls:

Clear Clears all disallowed factor level settings that you have specified. This does not clear the selected factors.

Start Over Removes all selected factors and returns you to the initial list of factors.

AND Opens the Add Filter Factors list. Selected factors become an AND group. Any combination of factor levels specified within an AND group is disallowed.

To add a factor to an AND group later on, click the group's outline to see a highlighted rectangle. Select AND and add the factor.

To remove a single factor, select **Delete** from its red triangle menu.

OR Opens the Add Filter Factors list. Selected factors become a separate AND group. For AND groups separated by OR, a combination is disallowed if it is specified in at least one AND group.

Red Triangle Options for Factors

A factor can appear in several OR groups. An occurrence of the factor in a specific OR group is referred to as an *instance* of the factor.

Delete Removes the selected instance of the factor from the Disallowed Combinations panel.

Clear Selection Clears any selection for that instance of the factor.

Invert Selection Deselects the selected values and selects the values not previously selected for that instance of the factor.

Display Options Available only for categorical factors. Changes the appearance of the display. Options include showing each level as a block, list, or single category, or adding a check box next to each value.

Find Available only for categorical factors. Provides a text box beneath the factor name where you can enter a search string for levels of the factor. Press the Enter key or click outside the text box to perform the search. Once **Find** is selected, Find options appear in the red triangle menu, such as Does not contain, Match Case, Starts with or Ends with, or Clear Find.

Use Disallowed Combinations Script

Use this option to disallow particular combinations of factor levels using a JSL script. This option can be used with continuous factors or mixed continuous and categorical factors.

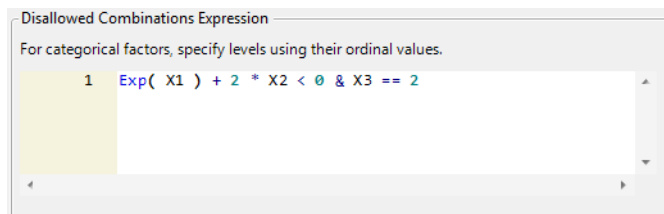
This option opens a script window where you insert a script that identifies the combinations that you want to disallow. The script must evaluate as a Boolean expression. When the expression evaluates as true, the specified combination is disallowed.

When forming the expression for a categorical factor, use the ordinal value of the level or the name of the level. If a factor's levels are high, medium, and low, specified in that order in the Factors outline, their associated ordinal values are 1, 2, and 3. For example, suppose that you have two continuous factors, X1 and X2, and a categorical factor X3 with three levels: L1, L2, and L3, in order. You want to disallow levels where the following is true:

$$e^{X_1} + 2X_2 < 0 \text{ and } X_3 = L2$$

Enter the expression `(Exp(X1) + 2*X2 < 0) & (X3 == 2)` into the script window.

Figure 21.5 Expression in Script Editor



(In the figure, unnecessary parentheses were removed by parsing.) Notice that functions can be entered as part of the Boolean expression. The expression `(Exp(X1) + 2*X2 < 0) & (X3 == "L2")` is also valid.

Space Filling Design Methods

The following methods for constructing space-filling designs are available:

- [“Sphere-Packing Designs”](#) on page 621
- [“Latin Hypercube Designs”](#) on page 624
- [“Uniform Designs”](#) on page 627
- [“Minimum Potential Designs”](#) on page 630
- [“Maximum Entropy Designs”](#) on page 633
- [“Gaussian Process IMSE Optimal Designs”](#) on page 634

- [“Fast Flexible Filling Designs”](#) on page 636

Design

The Design outline shows the runs for the space-filling design.

Design Diagnostics

The Design Diagnostics outline shows the values for the factors scaled from zero to one. The Minimum Distance is based on these scaled values and is the minimum distance from each point to its nearest neighbor. The row number for the nearest neighbor is given in the Nearest Point column. The discrepancy value shown below the table is the integrated difference between the design points based and a uniform distribution.

The MaxPro (maximum projection) criterion is provided for the full design as well as for each level of categorical variables when used. Smaller values are better. For more information about the MaxPro criteria see [“MaxPro”](#) on page 636.

Note: If for any dimension, two points can have the same value (that is, $x_{ik} = x_{jk}$), the MaxPro criterion is undefined. In this case, the Design Diagnostics shows a missing value for MaxPro.

Design Table

Make Table Constructs the Space Filling Design data table.

Back Takes you back to where you were before clicking **Make Design**. You can make changes to the previous outlines and regenerate the design.

Space Filling Design Options

The red triangle menu in the Space Filling Design platform contains these options:

Save Responses Saves the information in the Responses panel to a new data table. You can then quickly load the responses and their associated information into most DOE windows. This option is helpful if you anticipate reusing the responses.

Load Responses Loads responses that you saved using the Save Responses option.

Save Factors Saves the information in the Factors panel to a new data table. Each factor's column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Note: It is possible to create a factors table by entering data into an empty table, but remember to assign each column an appropriate Design Role. Do this by right-clicking on the column name in the data grid and selecting **Column Properties > Design Role**. In the Design Role area, select the appropriate role.

Load Factors Loads factors that you saved using the Save Factors option.

Save Constraints (Unavailable for some platforms) Saves factor constraints that you defined in the Define Factor Constraints or Linear Constraints outline into a data table, with a column for each constraint. You can then quickly load the constraints into most DOE windows.

In the constraint table, the first rows contain the coefficients for each factor. The last row contains the inequality bound. Each constraint's column contains a column property called `ConstraintState` that identifies the constraint as a "less than" or a "greater than" constraint. See "[ConstraintState](#)" on page 766 in the "Column Properties" appendix.

Load Constraints (Unavailable for some platforms) Loads factor constraints that you saved using the Save Constraints option.

Set Random Seed Sets the random seed that JMP uses to control actions that have a random component. These actions include one or more of the following:

- initializing search algorithms for design generation
- randomizing Run Order for design construction
- selecting a starting design for designs based on random starts

To reproduce a design, enter the random seed that generated the design before clicking **Make Design**.

Note: The random seed associated with a design is included in the DOE Dialog script that is saved to the design data table.

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.

- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click **Apply**, a formula for the simulated response values is saved in a new column called <Y> Simulated, where Y is the name of the response. Clicking **Apply** again updates the formula and values in <Y> Simulated.

For additional details, see [“Simulate Responses”](#) on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

FFF Optimality Criterion For the Fast Flexible Filling design method, enables you to select between the MaxPro criterion (the default) and the Centroid criterion. See [“FFF Optimality Criterion”](#) on page 636.

Number of Starts Specifies the number of times that the algorithm for the chosen design type initiates to construct a new design. The best design, based on the criterion for the given design type, is returned. Set to 1 by default for all design types. Not used for Fast Flexible Filling Designs.

Advanced Options > Set Average Cluster Size For the Fast Flexible Filling design method, enables you to specify the average number of randomly generated points used to define each cluster or, equivalently, each design point.

Advanced Options > MaxPro Categorical Weight For the Fast Flexible Filling design method, enables you to specify the weight used for categorical factors in the MaxPro calculation. Points that have the same categorical level are multiplied by the weight. This provides better space-filling properties at the sub-design for each level of a categorical factor.

Save Script to Script Window Creates the script for the design that you specified in the Custom Design platform and saves it in an open script window.

Sphere-Packing Designs

The Sphere-Packing design method maximizes the minimum distance between pairs of design points. The effect of this maximization is to spread the points out as much as possible inside the design region.

- “Creating a Sphere-Packing Design”
- “Visualizing the Sphere-Packing Design”

Creating a Sphere-Packing Design

1. Select **DOE > Special Purpose > Space Filling Design**.
2. Enter responses and factors.
See “Responses” on page 610.
3. Alter the factor level values, if necessary. For example, Figure 21.6 shows the two existing factors, X1 and X2, with values that range from 0 to 1 (instead of the default –1 to 1).

Figure 21.6 Space-Filling Dialog for Two Factors

Space Filling Design

Responses

Add Response Remove Number of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y <i>optional item</i>	Maximize	.	.	.

Factors

Continuous Categorical Remove Add N Factors 1

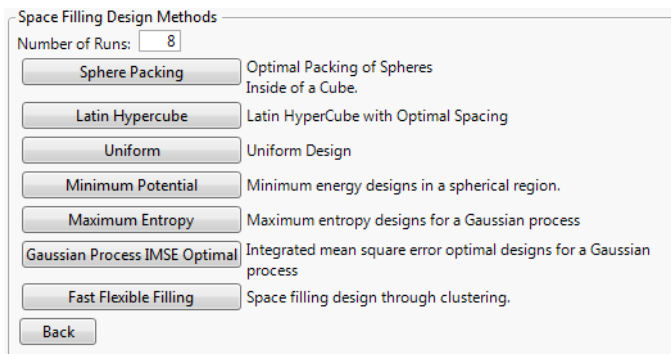
Name	Role	Values
X1	Continuous	0 1
X2	Continuous	0 1

Specify Factors

Specify desired number of factors. Double click on a factor name or setting to edit it.

Continue

4. Click **Continue**.
5. In the design specification dialog, specify a sample size (**Number of Runs**). Figure 21.7 shows a sample size of eight.

Figure 21.7 Space-Filling Design Dialog

6. Click **Sphere Packing**.

JMP creates the design and displays the design runs and the design diagnostics.

Figure 21.8 shows the Design Diagnostics panel with 0.518 as the **Minimum Distance**. Your results might differ slightly from the ones below, but the minimum distance is the same. The MaxPro statistic is undefined for Sphere Packing as points can have identical values for one dimension.

Figure 21.8 Sphere-Packing Design Diagnostics

Design Diagnostics

Run	ScaledX1	ScaledX2	Minimum Distance	Nearest Point
1	0.50000	0.86602	0.518	3
2	0.86603	0.50000	0.518	6
3	0.13397	0.50000	0.518	1
4	0.50000	0.13398	0.518	7
5	0.00000	1.00000	0.518	1
6	1.00000	1.00000	0.518	2
7	1.00000	0.00000	0.518	4
8	0.00000	0.00000	0.518	3

discrepancy = 0.046969

MaxPro

Full Design

Design Table

Make Table Back

7. Click **Make Table**. Use this table to complete the visualization example, described next.

Visualizing the Sphere-Packing Design

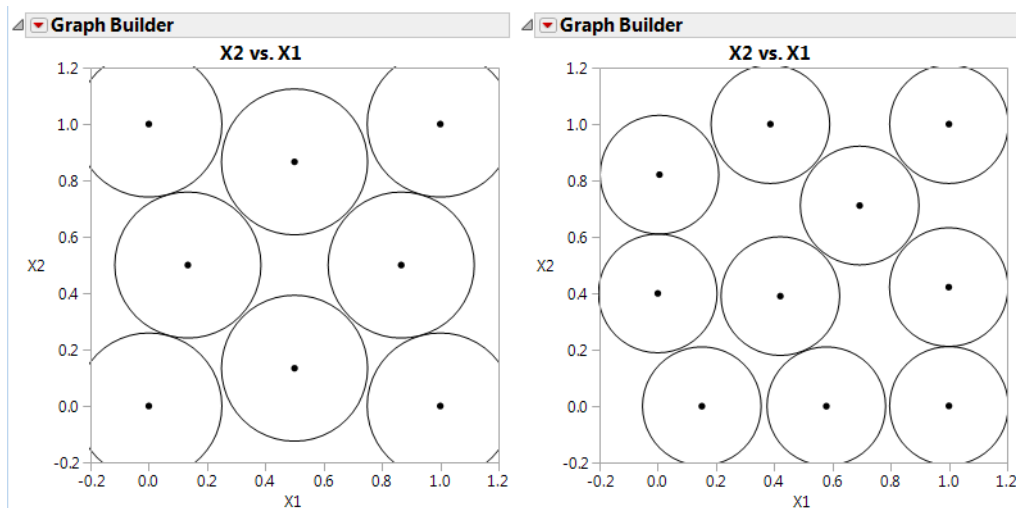
To visualize the nature of the Sphere-Packing technique, create a plot with Graph Builder. Using the table you just created, proceed as follows:

1. Select **Graph > Graph Builder**.
2. Specify X1 as **X** and X2 as **Y**.

3. Adjust the frame size so that the frame is square by right-clicking the plot and selecting **Graph > Size/Scale > Size to Isometric**.
4. Right-click the plot and select **Customize**. When the Customize panel appears, click the plus sign to see a text edit area and enter the following script:
`For Each Row(Circle({:X1, :X2}, 0.518/2))`
 where 0.518 is the minimum distance number that you noted in the Design Diagnostics panel. This script draws a circle centered at each design point with radius 0.259 (half the diameter, 0.518), as shown on the left in Figure 21.9. This plot shows the efficient way JMP packs the design points.
5. Now repeat the procedure exactly as described in the previous section, but with a sample size of 10 instead of eight.

Remember to change 0.518 in the graphics script to the minimum distance produced by 10 runs. When the plot appears, again set the frame size and create a graphics script using the minimum distance from the diagnostic report as the diameter for the circle. You should see a graph similar to the one on the right in Figure 21.9. Note the irregular nature of the sphere packing. In fact, you can repeat the process a third time to get a slightly different picture because the arrangement is dependent on the random starting point.

Figure 21.9 Sphere-Packing Example with Eight Runs (left) and 10 Runs (right)



Latin Hypercube Designs

In a Latin Hypercube, each factor has as many levels as there are runs in the design. The levels are spaced evenly from the lower bound to the upper bound of the factor. Like the sphere-packing method, the Latin Hypercube method chooses points to maximize the minimum distance between design points, but with a constraint. The constraint maintains the even spacing between factor levels.

- [“Creating a Latin Hypercube Design”](#)
- [“Visualizing the Latin Hypercube Design”](#)

Creating a Latin Hypercube Design

To use the Latin Hypercube method:

1. Select **DOE > Special Purpose > Space Filling Design**.
2. Enter responses, if necessary, and factors.
See [“Responses”](#) on page 610.
3. Alter the factor level values, if necessary. Figure 21.10 shows adding two factors to the two existing factors and changing their values to 1 and 8 instead of the default –1 and 1.

Figure 21.10 Space-Filling Dialog for Four Factors

Space Filling Design

Responses

Add ResponseRemoveNumber of Responses...

Response Name	Goal	Lower Limit	Upper Limit	Importance
Y	Maximize	.	.	.
optional item				

Factors

ContinuousCategoricalRemoveAdd N Factors1

Name	Role	Values
X1	Continuous	18
X2	Continuous	18
X3	Continuous	18
X4	Continuous	18

Specify Factors

Specify desired number of factors. Double click on a factor name or setting to edit it.

Continue

4. Click **Continue**.

5. In the design specification dialog, specify a sample size (**Number of Runs**). This example uses a sample size of eight.
6. Click **Latin Hypercube** (Figure 21.7). Factor settings and design diagnostics results appear similar to those in Figure 21.11, which shows the Latin Hypercube design with four factors and eight runs.

Note: The purpose of this example is to show that each column (factor) is assigned each level only once, and each column is a different permutation of the levels.

Figure 21.11 Latin Hypercube Design for Four Factors and Eight Runs with Eight Levels

4 Factors

Space Filling Latin Hypercube

Design

Run	X1	X2	X3	X4
1	4.00000	7.00000	1.00000	4.00000
2	2.00000	6.00000	5.00000	7.00000
3	1.00000	4.00000	6.00000	1.00000
4	7.00000	5.00000	3.00000	8.00000
5	5.00000	2.00000	8.00000	6.00000
6	6.00000	8.00000	7.00000	3.00000
7	8.00000	3.00000	4.00000	2.00000
8	3.00000	1.00000	2.00000	5.00000

Design Diagnostics

Run	ScaledX1	ScaledX2	ScaledX3	ScaledX4	Minimum Distance	Nearest Point
1	0.42857	0.85714	0.00000	0.42857	0.782	2
2	0.14286	0.71429	0.57143	0.85714	0.782	1
3	0.00000	0.42857	0.71429	0.00000	0.926	2
4	0.85714	0.57143	0.28571	1.00000	0.795	2
5	0.57143	0.14286	1.00000	0.71429	0.845	2
6	0.71429	1.00000	0.85714	0.28571	0.892	7
7	1.00000	0.28571	0.42857	0.14286	0.892	6
8	0.28571	0.00000	0.14286	0.57143	0.892	1

discrepancy = 0.039317

MaxPro

Full Design 11.40992

Visualizing the Latin Hypercube Design

To visualize the nature of the Latin Hypercube technique, create a plot with Graph Builder:

1. Create another Latin Hypercube design using the default X1 and X2 factors.
2. Be sure to change the factor values so that they are 0 and 1 instead of the default -1 and 1.
3. Click **Continue**.
4. Specify a sample size of eight (**Number of Runs**).
5. Click **Latin Hypercube**.

Figure 21.12 Latin Hypercube Design with Two Factors and Eight Runs

2 Factors
Space Filling Latin Hypercube

Design		
Run	X1	X2
1	0.28571	0.00000
2	0.00000	0.85714
3	0.57143	0.28571
4	0.42857	0.71429
5	0.14286	0.42857
6	1.00000	0.14286
7	0.85714	0.57143
8	0.71429	1.00000

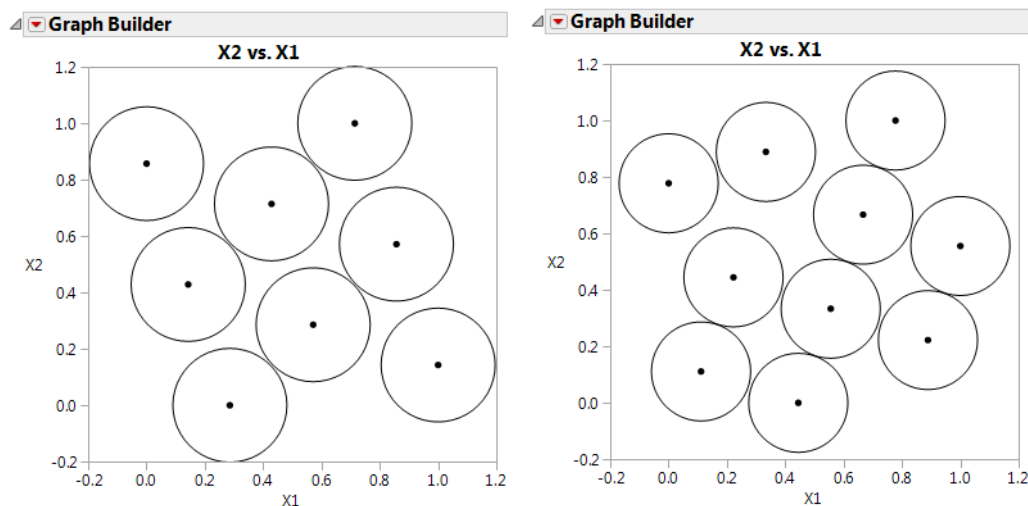
Design Diagnostics				
Run	ScaledX1	ScaledX2	Minimum Distance	Nearest Point
1	0.28571	0.00000	0.404	3
2	0.00000	0.85714	0.452	4
3	0.57143	0.28571	0.404	1
4	0.42857	0.71429	0.404	5
5	0.14286	0.42857	0.404	4
6	1.00000	0.14286	0.452	7
7	0.85714	0.57143	0.404	3
8	0.71429	1.00000	0.404	4

discrepancy = 0.009279
MaxPro
Full Design 11.32801

6. Click **Make Table**.
7. Select **Graph > Graph Builder**.
8. Specify X1 as **X** and X2 as **Y**.
9. Right-click the plot and select **Graph > Size/Scale > Size to Isometric** to adjust the frame size so that the frame is square.
10. Right-click the plot, select **Customize** from the menu. In the Customize panel, click the large plus sign to see a text edit area, and enter the following script:
 For Each Row(Circle({:X1, :X2}, 0.404/2))
 where 0.404 is the minimum distance number that you noted in the Design Diagnostics panel (Figure 21.12). This script draws a circle centered at each design point with radius 0.202 (half the diameter, 0.404), as shown on the left in Figure 21.13. This plot shows the efficient way JMP packs the design points.
11. Repeat the above procedure exactly, but with 10 runs instead of eight (step 5). Remember to change 0.404 in the graphics script to the minimum distance produced by 10 runs.

You should see a graph similar to the one on the right in Figure 21.13. Note the irregular nature of the sphere packing. In fact, you can repeat the process to get a slightly different picture because the arrangement is dependent on the random starting point.

Figure 21.13 Comparison of Latin Hypercube Designs with Eight Runs (left) and 10 Runs (right)



Note that the minimum distance between each pair of points in the Latin Hypercube design is smaller than that for the Sphere-Packing design. This is because the Latin Hypercube design constrains the levels of each factor to be evenly spaced. The Sphere-Packing design maximizes the minimum distance without any constraints.

Uniform Designs

The Uniform design minimizes the discrepancy between the design points (empirical uniform distribution) and a theoretical uniform distribution.

Note: These designs are most useful for getting a simple and precise estimate of the integral of an unknown function. The estimate is the average of the observed responses from the experiment.

1. Select **DOE > Special Purpose > Space Filling Design**.
2. Enter responses, if necessary, and factors.
See [“Responses”](#) on page 610.
3. Alter the factor level values to 0 and 1.
4. Click **Continue**.
5. In the design specification dialog, specify a sample size. This example uses a sample size of eight (**Number of Runs**).

- Click the **Uniform** button. JMP creates this design and displays the design runs and the design diagnostics as shown in Figure 21.14.

Note: The emphasis of the Uniform design method is not to spread out the points. The minimum distances in Figure 21.14 vary substantially.

Figure 21.14 Factor Settings and Diagnostics for Uniform Space-Filling Designs with Eight Runs

2 Factors
Space Filling Uniform Design

Design		
Run	X1	X2
1	0.93092	0.69041
2	0.69276	0.18514
3	0.06860	0.32218
4	0.18696	0.81304
5	0.56496	0.93540
6	0.81410	0.43443
7	0.32218	0.06860
8	0.43643	0.56357

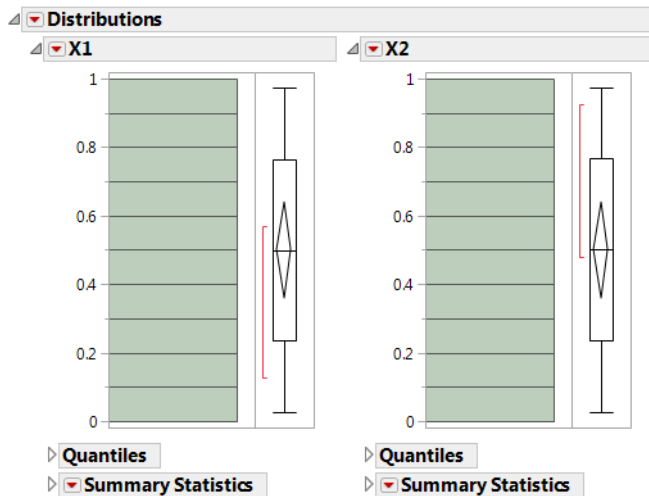
Design Diagnostics				
Run	ScaledX1	ScaledX2	Minimum Distance	Nearest Point
1	0.93092	0.69041	0.281	6
2	0.69276	0.18514	0.277	6
3	0.06860	0.32218	0.359	7
4	0.18696	0.81304	0.353	8
5	0.56496	0.93540	0.393	8
6	0.81410	0.43443	0.277	2
7	0.32218	0.06860	0.359	3
8	0.43643	0.56357	0.353	4

discrepancy = 0.004663
MaxPro
 Full Design 15.40271

- Click **Make Table**.

A Uniform design does not guarantee even spacing of the factor levels. However, increasing the number of runs and running a distribution on each factor (use **Analyze > Distribution**) shows flat histograms.

Figure 21.15 Histograms Are Flat for Each Factor When Number of Runs Is Increased to 20



Comparing Sphere-Packing, Latin Hypercube, and Uniform Methods

To compare space-filling design methods, create the **Sphere Packing**, **Latin Hypercube**, and **Uniform** designs, as shown in the previous examples. The Design Diagnostics tables show the values for the factors scaled from zero to one. The minimum distance is based on these scaled values and is the minimum distance from each point to its closest neighbor. The discrepancy value is the integrated difference between the design points and the uniform distribution.

Figure 21.16 shows a comparison of the design diagnostics for three eight-run space-filling designs. Note that the discrepancy for the Uniform design is the smallest (best). The discrepancy for the Sphere-Packing design is the largest (worst). The discrepancy for the Latin Hypercube takes an intermediate value that is closer to the optimal value. The MaxPro criteria is smaller for the Latin Hypercube design as compared to the Uniform Design.

Also note that the minimum distance between pairs of points is largest (best) for the Sphere-Packing method. The Uniform design has pairs of points that are only about half as far apart. The Latin Hypercube design behaves more like the Sphere-Packing design in spreading the points out.

For both spread and discrepancy, the Latin Hypercube design represents a healthy compromise solution.

Figure 21.16 Comparison of Diagnostics for Three Eight-Run Space-Filling Methods

Space Filling Sphere Packing

▷ Design

▢ Design Diagnostics

Run	ScaledX1	ScaledX2	Minimum Distance	Nearest Point
1	0.50000	0.86602	0.518	3
2	0.86603	0.50000	0.518	6
3	0.13397	0.50000	0.518	1
4	0.50000	0.13398	0.518	7
5	0.00000	1.00000	0.518	1
6	1.00000	1.00000	0.518	2
7	1.00000	0.00000	0.518	4
8	0.00000	0.00000	0.518	3

discrepancy = 0.046969

MaxPro

Full Design

Space Filling Latin Hypercube

▷ Design

▢ Design Diagnostics

Run	ScaledX1	ScaledX2	Minimum Distance	Nearest Point
1	0.28571	0.00000	0.404	3
2	0.00000	0.85714	0.452	4
3	0.57143	0.28571	0.404	1
4	0.42857	0.71429	0.404	5
5	0.14286	0.42857	0.404	4
6	1.00000	0.14286	0.452	7
7	0.85714	0.57143	0.404	3
8	0.71429	1.00000	0.404	4

discrepancy = 0.009279

MaxPro

Full Design 11.32801

Space Filling Uniform Design

▷ Design

▢ Design Diagnostics

Run	ScaledX1	ScaledX2	Minimum Distance	Nearest Point
1	0.93092	0.69041	0.281	6
2	0.69276	0.18514	0.277	6
3	0.06860	0.32218	0.359	7
4	0.18696	0.81304	0.353	8
5	0.56496	0.93540	0.393	8
6	0.81410	0.43443	0.277	2
7	0.32218	0.06860	0.359	3
8	0.43643	0.56357	0.353	4

discrepancy = 0.004663

MaxPro

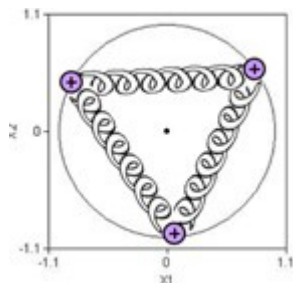
Full Design 15.40271

Another point of comparison is the time it takes to compute a design. The Uniform design method requires the most time to compute. Also, the time to compute the design increases rapidly with the number of runs. For comparable problems, all the space-filling design methods take longer to compute than the D -optimal designs in the Custom Designer.

Minimum Potential Designs

The Minimum Potential design spreads points out inside a sphere. To understand how this design is created, imagine the points as electrons with springs attached to every other point, as illustrated in Figure 21.17. The coulomb force pushes the points apart, but the springs pull them together. The design is the spacing of points that minimizes the potential energy of the system.

Figure 21.17 Minimum Potential Design



Minimum Potential designs:

- have spherical symmetry
- are nearly orthogonal
- have uniform spacing

To construct a Minimum Potential example design:

1. Select **DOE > Special Purpose > Space Filling Design**.
2. Add 3 continuous factor.
See “Factors” on page 612.
3. Alter the factor level values to 0 and 1.
4. Click **Continue**.
5. In the design specification dialog (shown on the left in Figure 21.18), enter a sample size (**Number of Runs**). This example uses a sample size of 12.
6. Click the **Minimum Potential** button. JMP creates this design and displays the design runs and the design diagnostics (shown on the right in Figure 21.18).

Figure 21.18 Space-Filling Methods and Design Diagnostics for Minimum Potential Design

Space Filling Design Methods

Number of Runs:

Sphere Packing Optimal Packing of Spheres
Inside of a Cube.

Latin Hypercube Latin HyperCube with Optimal Spacing

Uniform Uniform Design

Minimum Potential Minimum energy designs in a spherical region.

Maximum Entropy Maximum entropy designs for a Gaussian process

Gaussian Process IMSE Optimal Integrated mean square error optimal designs for a Gaussian process

Fast Flexible Filling Space filling design through clustering.

Back

Design Diagnostics

Run	ScaledX1	ScaledX2	ScaledX3	Minimum Distance	Nearest Point
1	0.49845	0.50584	0.50273	0.488	12
2	0.39334	0.06387	0.70753	0.498	1
3	0.20787	0.68269	0.14320	0.485	9
4	0.61776	0.92918	0.27211	0.49	12
5	0.84983	0.15239	0.47495	0.499	1
6	0.33657	0.14577	0.19726	0.499	1
7	0.73480	0.45365	0.06714	0.498	1
8	0.26209	0.52317	0.93916	0.49	12
9	0.00665	0.45300	0.52058	0.485	3
10	0.75354	0.46263	0.92467	0.485	11
11	0.95475	0.69233	0.54729	0.485	10
12	0.38435	0.93546	0.70339	0.488	1

discrepancy = 0.0165

Design Table

Make Table **Back**

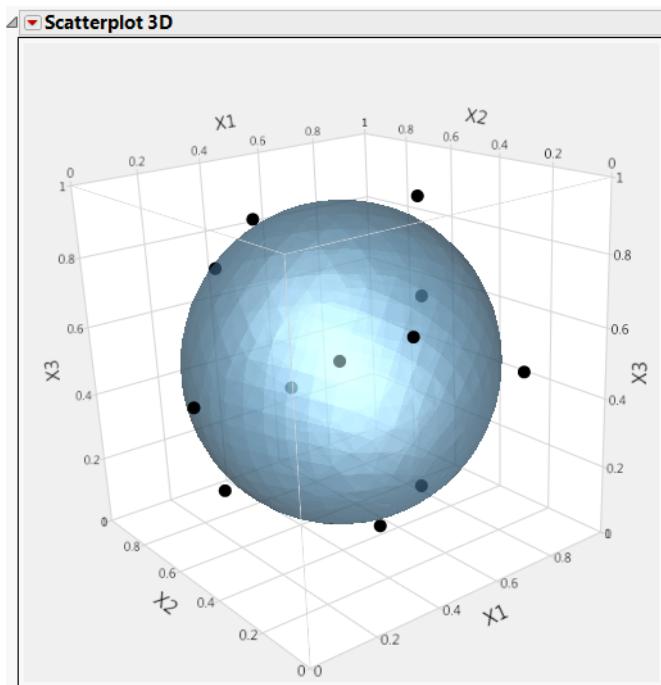
7. Click **Make Table**.

You can see the spherical symmetry of the Minimum Potential design using the Scatterplot 3D graphics platform.

1. After you make the JMP design table, choose the **Graph > Scatterplot 3D** command.
2. In the Scatterplot 3D launch dialog, select X1, X2, and X3 as **Y, Columns** and click **OK** to see the initial three-dimensional scatterplot of the design points.
3. To see the results similar to those in Figure 21.19:
 - Click the Scatterplot 3D red triangle and select the **Normal Contour Ellipsoids**.
 - Right-click the plot and select **Settings**, and then increase the marker size using the **Marker Size** slider.

Now the even spread of the points on the surface of the ellipsoid can be visualized.

Figure 21.19 Minimum Potential Design Points on Sphere

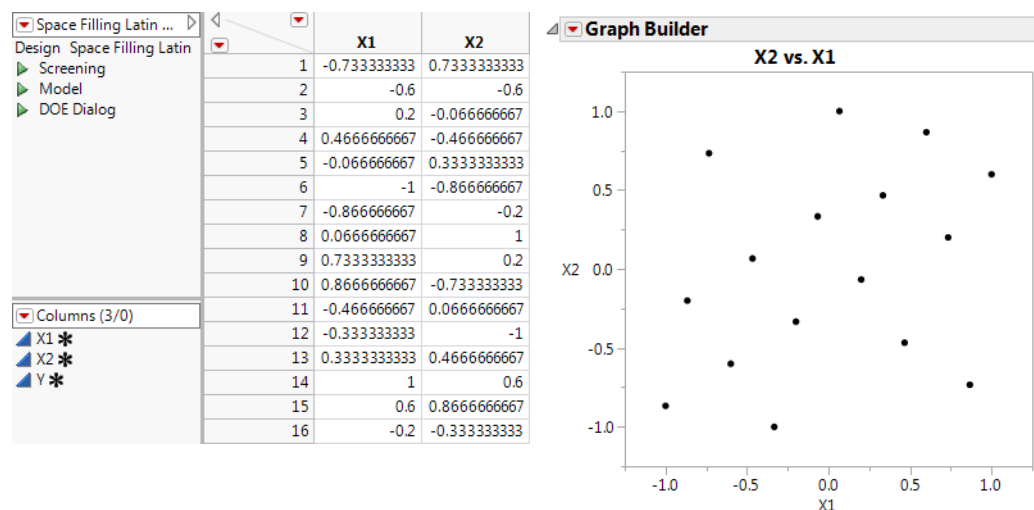


Maximum Entropy Designs

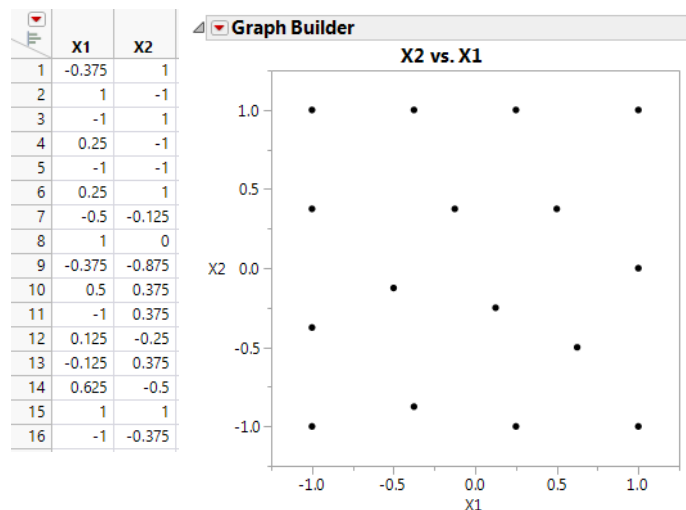
The Latin Hypercube design is a popular design to use along with a Gaussian-Process model. Computer simulation experts like to use the Latin Hypercube design because all projections onto the coordinate axes are uniform.

However, as the example in Figure 21.20 shows, the Latin Hypercube design might not provide you with optimal space filling properties. This example is a two-factor 16 run Latin Hypercube with factor level settings set between -1 and 1. The plot of the two factors shows that this design has regions that are missing coverage. In particular, there is poor coverage for X_1 near 0 and X_2 near -1.

Figure 21.20 Two-factor Latin Hypercube Design



The Maximum Entropy design is an alternative to the Latin Hypercube design for computer experiments. The Maximum Entropy design optimizes a measure of the amount of information contained in an experiment. See the technical note below. With the factor levels set between -1 and 1, the two-factor Maximum Entropy design shown in Figure 21.21 covers the region better than the Latin hypercube design in Figure 21.20. The space-filling property of the design improves as the number of runs increases.

Figure 21.21 Two-Factor Maximum Entropy Design


Technical Maximum Entropy designs maximize the Shannon information (Shewry and Wynn (1987)) of an experiment, assuming that the data come from a normal $(m, s^2 R)$ distribution, where

$$R_{ij} = \exp\left(-\sum_k \theta_k (x_{ik} - x_{jk})^2\right)$$

is the correlation of response values at two different design points, x_i and x_j . Computationally, these designs maximize $|\mathbf{R}|$, the determinant of the correlation matrix of the sample. If x_i and x_j are far apart, then $\mathbf{R}_{ij}$ approaches zero. If x_i and x_j are close together, then $\mathbf{R}_{ij}$ is near one.

Gaussian Process IMSE Optimal Designs

The Gaussian process IMSE optimal design method constructs designs that are suitable for Gaussian process models. Gaussian process models fit a wide variety of surfaces. Gaussian process IMSE optimal designs minimize the integrated mean squared error of the Gaussian process model over the experimental region. The Gaussian process IMSE optimal design method uses a correlation structure similar to that of the kriging model. See Jones and Johnson (2009).

Covariance Parameter Vector

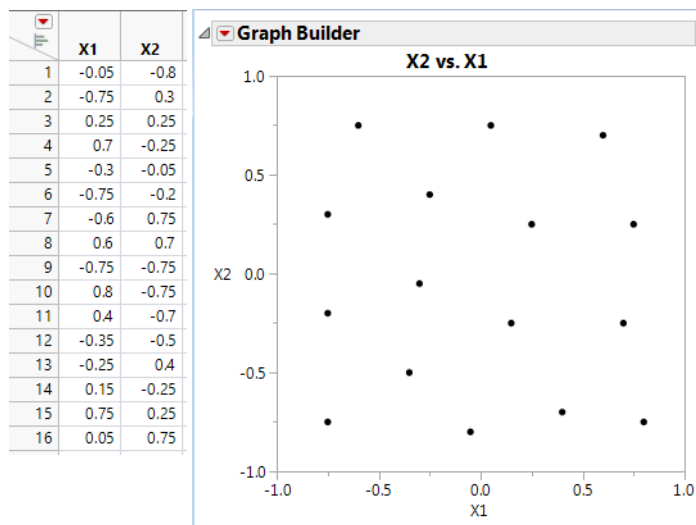
In a Gaussian Process IMSE Optimal Design formulation of the Gaussian process model, the covariance parameter vector determines the correlation structure. There is a Theta for each factor. A theta equal to 0 corresponds to a correlation of 1, causing the fitted surface to be flat in the corresponding factor's direction. As theta increases, the correlation decreases, allowing the surface to be flexible in the factor's direction.

In the Covariance Parameter Vector outline, in the list of values under Thetas, you can enter values that reflect your prior knowledge of the surface.

Comparison of Gaussian Process IMSE Optional Design with Latin Hypercube Design

Gaussian process IMSE optimal designs are alternatives to the Latin Hypercube design. You can compare the IMSE optimal design to the Latin Hypercube (shown previously in Figure 21.20). The table and overlay plot in Figure 21.22 show a Gaussian IMSE optimal design. You can see that the design provides uniform coverage of the factor region.

Figure 21.22 Comparison of Two-factor Latin Hypercube and Gaussian IMSE Optimal Designs



Note: Both the Maximum Entropy design and the Gaussian Process IMSE Optimal design were created using 100 random starts.

Fast Flexible Filling Designs

Note: If you have Categorical factors or factor constraints, then Fast Flexible Filling is the only Method available.

FFF Optimality Criterion

The algorithms for Fast Flexible Filling designs begin by generating a large number of random points within the specified design region. These points are then clustered using a Fast Ward algorithm into a number of clusters that equals the Number of Runs that you specified.

The final design points can be obtained by using the default MaxPro (*maximum projection*) optimality criterion or by selecting the Centroid criterion. You can find these options under FFF Optimality Criterion in the report's red triangle menu.

MaxPro For p factors and n equal to the specified Number of Runs, the MaxPro criterion strives to find points in the clusters that minimize the following criterion:

$$C_{MaxPro} = \sum_i^{n-1} \sum_{j=i+1}^n \left[1 / \prod_{k=1}^p (x_{ik} - x_{jk})^2 \right]$$

The MaxPro criterion maximizes the product of the distances between potential design points in a way that involves all factors. This supports the goal of providing good space-filling properties on projections of factors. See Joseph et al. (2015). The Max Pro option is the default.

Centroid This method places a design point at the centroid of each cluster. It has the property that the average distance from an arbitrary point in the design space to its closest neighboring design point is smaller than for other designs.

Note: You can set a preference to always use a given optimality criterion. Select **File > Preferences > Platforms > DOE**. Select **FFF Optimality Criterion** and select your preferred criterion.

Categorical Factors

When you have categorical factors, the algorithm proceeds as follows:

- The total number of design points is balanced across the total number of combinations of levels of the categorical factors. Suppose that there are m combinations of levels and that k design points are allocated to each of these.

- A large number of points within the design space defined by the continuous variables is generated. These are grouped into k primary clusters.
- Each of the k primary clusters of points is further clustered into m sub-clusters.
- Within each primary cluster, a design point is calculated for each of the m sub-clusters using the specified FFF optimality criterion.
- For each of the k primary clusters, one of the m combinations of levels is randomly assigned to each of the m sub-cluster design points. This yields a total of km design points.
- For each of the k primary clusters, a design point is chosen for each of the m combinations of levels according to the MaxPro criterion, where points having the same level of a categorical factor are multiplied by the MaxPro Categorical Weight. (For a description of MaxPro Categorical Weight, see [“Space Filling Design Options”](#) on page 618.) This process continues through all k primary clusters 10 times, or until no improvement is found by changing the given design points.

Set Average Cluster Size

The Set Average Cluster Size option is found under Advanced Options in the Space Filling Design red triangle menu. This option enables you to specify the average number of randomly- generated points used to define each cluster or, equivalently, each design point.

By default, if the Number of Runs is set to 200 or less, a total of 10,000 randomly generated points are used as the basis for the clustering algorithm. When the number of Runs exceeds 200, a default value of 50 is used. Increasing this value can be particularly useful in designs with a large number of factors or where disallowed combinations restrict the distribution of points used in the clustering algorithm.

Note: Depending on the number of factors and the specified value for Number of Runs, you might want to increase the average number of initial points per design point by selecting **Advanced Options > Set Average Cluster Size**.

Constraints

Once you complete the Factors outline, click **Continue**. The Define Factor Constraints outline appears. Use this outline to restrict the design region. For more information about the outline, see [“Define Factor Constraints”](#) on page 614.

You can use the Use Disallowed Combinations Filter and Use Disallowed Combinations Script options to specify disallowed factor level combinations. Or, you can use the Specify Linear Constraints option to specify bounds in terms of linear inequalities. However, the design is generated differently for these two methods.

Use Disallowed Combinations Filter and Use Disallowed Combinations Script

When disallowed combinations are specified, the random points that form the basis for the clustering algorithm are randomly distributed within the unconstrained design region. Then disallowed points are removed and clustering proceeds with the remaining points.

Note: Depending on the nature of the constraints and the specified Number of Runs, the default coverage of the unconstrained design space by the initial randomly generated points might not be sufficient to produce the required Number of Runs. In this case, you might obtain a JMP Alert indicating that the algorithm “Could not find sufficient number of points.” To increase the initial number of points that form the basis for the clustering algorithm, specify a larger average number of initial points per design point by selecting **Advanced Options > Set Average Cluster Size**. (See [“Set Average Cluster Size”](#) on page 637).

Specify Linear Constraints

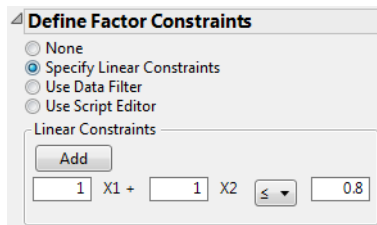
When you use the **Specify Linear Constraints** option, the random points that form the basis for the clustering algorithm are randomly distributed within the constrained design region. The clustering algorithm uses these points.

Creating and Viewing a Constrained Fast Flexible Filling Design

Constructing the Design

1. Select **DOE > Special Purpose > Space Filling Design**.
2. Add 3 continuous factors.
3. Alter the factor level values to 0 and 1 and click **Continue**.
4. In the Define Linear Constraints outline, select **Specify Linear Constraints**.
Notice that Fast Flexible Filling is the only available Space Filling Design Method.
5. Select **Add**.
6. Enter the following coefficients and bound:
 - 1 for X1
 - 1 for X2
 - 0.8 for the bound

Figure 21.23 Linear Constraint



7. Type 200 next to **Number of Runs**.

8. Select **Fast Flexible Filling**.

JMP creates a design that satisfies the constraints. Open the Design outline to view the design.

9. Select **Make Table** to construct the data table.

Constructing the Plot

1. With the data table active, select **Graph > Graph Builder**.

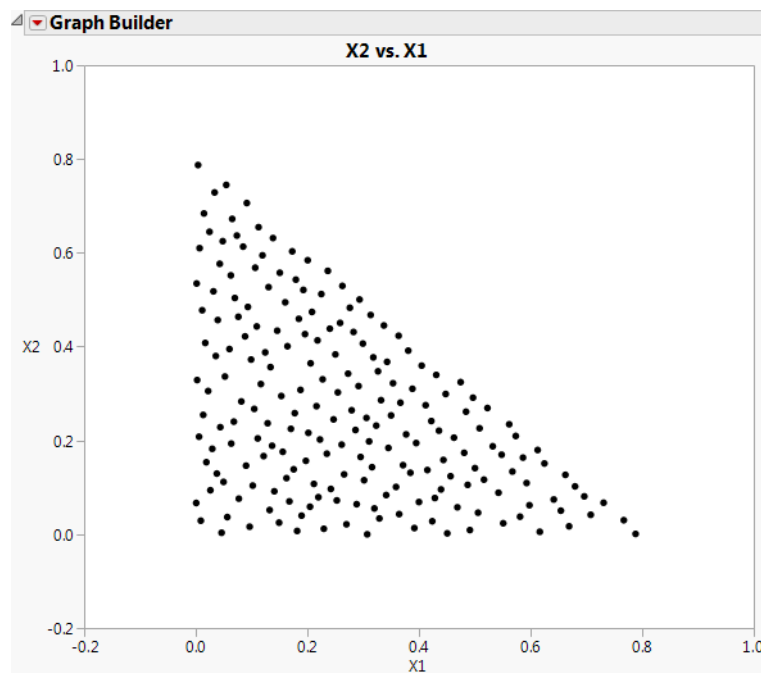
2. Drag X1 to the drop zone labeled **X**.

3. Drag X2 to the drop zone labeled **Y**.

4. Remove the **Smoother** by clicking the smoother icon.

5. Click the Graph Builder red triangle and click **Show Control Panel** to deselect it.

You should see a graph similar to the one in Figure 21.24. Note that the points satisfy the linear constraint $X1 + X2 \leq 0.8$.

Figure 21.24 Fast Flexible Filling Design with One Linear Constraint

Creating a Space-Filling Design for a Map Shape

A fast flexible filling (FFF) design can be used to design an experiment to be executed across a region such as a map shape. For example, if you had two sample collection methods (perhaps for air or soil samples) that you wanted to compare at locations spread throughout a US state, consider a FFF design.

Supposed the US state of interest is Georgia. You want 40 samples, 20 for each sampling method, spread throughout the state. To create a FFF design for a map shape, you need a file with the latitude and longitude of the map boundaries (see [“Background Maps”](#) on page 264 in the “Maps” chapter).

In the following example, you determine the longitude and latitude for the US state of Georgia, define the shape of the state, then create the design. The design will provide the location and the sampling method for each of your 40 samples.

Find Longitude and Latitude for Georgia

1. Navigate to the Maps directory, usually found here:
Windows: C:/Program Files/SAS/JMP/<Version>/Maps
macOS: /Library/Application Support/JMP/<Version>/Maps

2. Open US-State-XY.jmp and US-State-Name.jmp.
Notice that in US-State-Name.jmp, Shape ID 11 corresponds to the US state of Georgia. In this example, that is the state of interest.
3. In US-State-XY.jmp, select **Rows > Row Selection > Select Where...**
4. Select Shape ID and enter 11 in the text box, then click **OK**.
This selects all rows with a Shape ID of 11, representing Georgia.
5. Select **Tables > Subset** and click **OK** to obtain a subset of the US-State-XY.jmp table of the latitudes and longitudes data for the state of Georgia (the selected rows).
6. Close US-State-XY.jmp and US-State-Name.jmp.

Get the Range of Values

To define the space for the space filling design, you need to know the range of values for the latitude and longitudes in Georgia. Then, you can create design factors with these values.

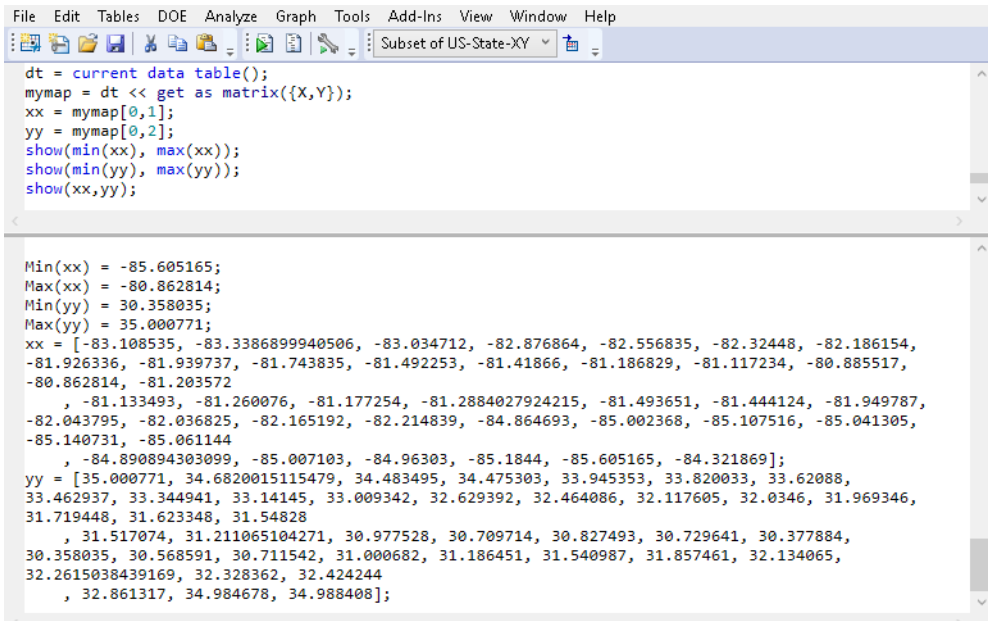
1. Select **File > New > Script**. In the script window, right-click and select **Show Embedded Log**.
2. Copy the following and paste it into the new script:

```
dt = current data table();  
mymap = dt << get as matrix({X,Y});  
xx = mymap[0,1];  
yy = mymap[0,2];  
show(min(xx), max(xx));  
show(min(yy), max(yy));  
show(xx,yy);
```

3. With the subset of the US-State-XY.jmp table active, run the script.

The range of values for the latitudes and longitudes in Georgia and the **xx** and **yy** matrices appear in the log.

Figure 21.25 Range Finding Script and Results in Log



Add Factors

1. Select **DOE > Special Purpose > Space Filling Design**.
2. Add two continuous factors to the design. Enter 2 in the **Add N Factors** text box and click **Continuous**.
3. Rename the factors longitude and latitude. Enter the min and max values, as follows:
 - For longitude, enter -85.6 and -80.9 as the values. These are the min and max longitude values from the log for *x*.
 - For latitude, enter 30.4 and 35.0 as the values. These are the min and max latitude values from the log for *y*.
4. Add a categorical factor. Click **Categorical > 2 Level**. Rename the factor Method.

Figure 21.26 Factors

Factors			
Name	Role	Values	
Longitude	Continuous	-85.6	-80.7
Latitude	Continuous	30.3	35
Method	Categorical	L1	L2

5. Click **Continue**.

Define the Shape

To further define the space for the space filling design, you need to define the shape of the state using a polygon.

1. Select **Use Disallowed Combinations Script** to open the Disallowed Combinations Expression text box.

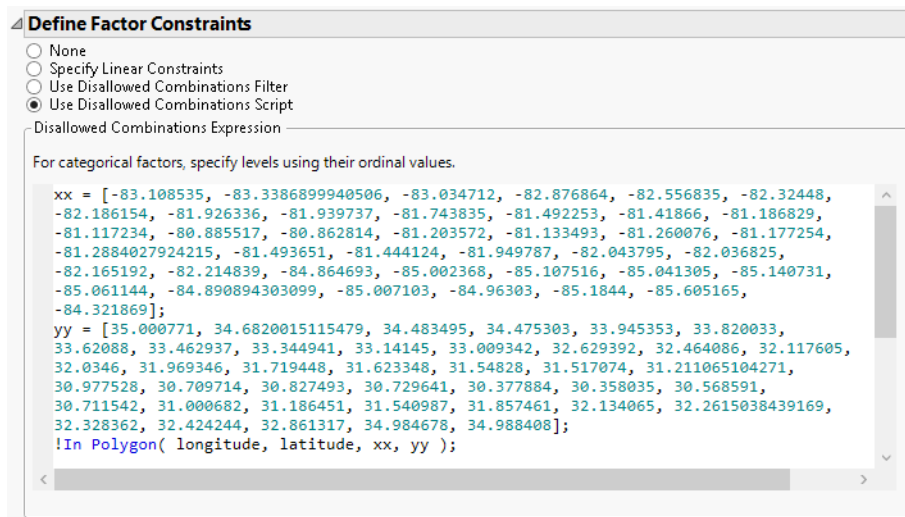
The ranges of longitude and latitude define a square. The state of Georgia is a polygon, not a square. You want to disallow points that fall outside of the state of Georgia polygon.

2. Go to your script window and copy the **xx** and **yy** matrices from your script log. Paste this into the Disallowed Combinations Expression text box.
3. At the end, add the following line:

```
!In Polygon(longitude, latitude , xx, yy);
```

This excludes points that are not in the polygon defined by the **xx** and **yy** coordinates.

Figure 21.27 Disallowed Combinations Expression

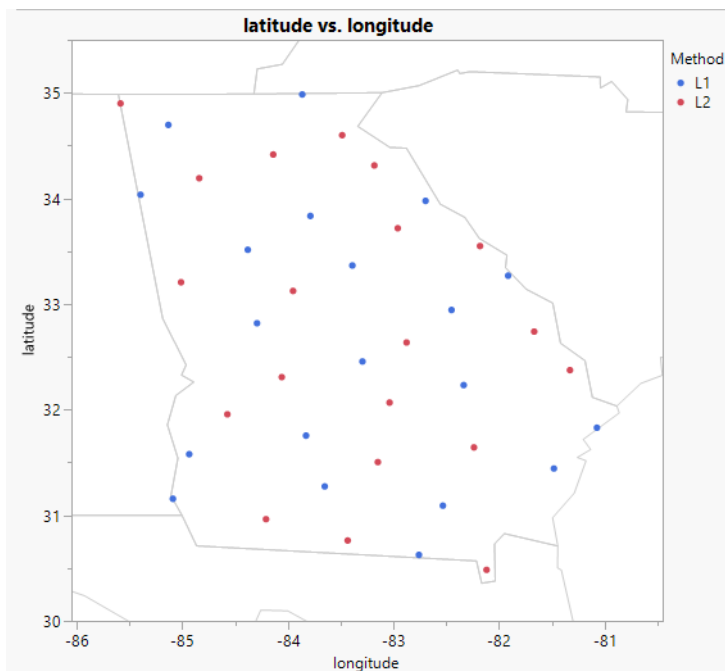


Make the Design

1. Change the **Number of Runs** to 40 and click **Fast Flexible Filling**.
2. Click **Make Table**.
3. From your design table, select **Graph > Graph Builder**.
4. Drag these columns into these zones:
 - longitude to the **X** zone
 - latitude to the **Y** zone

- Method to the **Color** zone
- 5. Above the graph, click the Smoother element icon to turn it off.
- 6. To add the state shape, right-click the graph and select **Graph > Background Map**.
- 7. Under **Boundaries**, select **US States** and click **OK**.

Figure 21.28 Example of FFF Design for the US State of Georgia



Note: The exact location of the points on your map may differ from the one shown due to the random starting point for the FFF algorithm.

The map identifies the 40 locations where you should collect your samples. The sampling method (L1 or L2) is indicated by the color of the points.

The Graph > Background Map window has these additional options:

- To determine the counties for your testing, under **Boundaries**, select **US Counties**.
- To map your testing locations at the street level, under **Images**, select **Street Map Service**.

Example of a Sphere-Packing Design

Worley (1987) presented a model of the flow of water through a borehole that is drilled from the ground surface through two aquifers. The response variable y is the flow rate through the borehole in m^3/year and is determined by the following equation:

$$y = \frac{2\pi T_u (H_u - H_l)}{\ln(r/r_w) \left[1 + \frac{2LT_u}{\ln(r/r_w)r_w^2 K_w} + \frac{T_u}{T_l} \right]}$$

There are eight factors in this model:

r_w = radius of borehole, 0.05 to 0.15 m

r = radius of influence, 100 to 50,000 m

T_u = transmissivity of upper aquifer, 63,070 to 115,600 m^2/year

H_u = potentiometric head of upper aquifer, 990 to 1100 m

T_l = transmissivity of lower aquifer, 63.1 to 116 m^2/year

H_l = potentiometric head of lower aquifer, 700 to 820 m

L = length of borehole, 1120 to 1680 m

K_w = hydraulic conductivity of borehole, 9855 to 12,045 m/year

You can use a sphere-packing design to obtain a set of conditions for which to calculate the response y . Then you can build a model to estimate the true model over the range of inputs used in your design. Evaluation of your estimated model can help you understand the impact of each of the eight factors on the response.

Create the Sphere-Packing Design for the Borehole Data

To create a sphere-packing design for the borehole model, you can use a data table of saved factor settings.

1. Select **Help > Sample Data Library** and open Design Experiment/Borehole Factors.jmp
2. Select **DOE > Special Purpose > Space Filling Design**
3. Click the Space Filling Design red triangle and select **Load Factors**.

Figure 21.29 Factors Panel with Factor Values Loaded for Borehole Example

Factors			
Continuous	Categorical	Remove	Add N Factors 1
Name	Role	Values	
log10 R _w	Continuous	-1.3	-0.82
log10 R	Continuous	2	4.7
Tu	Continuous	63070	115600
Tl	Continuous	63.1	116
Hu	Continuous	990	1110
Hl	Continuous	700	820
L	Continuous	1120	1680
Kw	Continuous	9855	12045

Note: The logarithm of r and r_w are used as factors.

4. Click **Continue**.
5. Set the Number of Runs to 32 and press Enter.
6. Click **Sphere Packing** to produce the design.
7. Click **Make Table** to make a table showing the design settings for the experiment.

Note: The design table includes a Model table script. This script runs a Gaussian Process model for the response y .

8. (Optional) To see a completed data table for this example, select **Help > Sample Data Library** and open Design Experiment/ Borehole Sphere Packing.jmp.

Because the designs are generated from a random seed, the settings that you obtain can differ from those shown in the completed table.

Guidelines for the Analysis of Deterministic Data

It is important to remember that deterministic data have no random component. The same input values generate the same output. As a result, p -values from fitted statistical models do not have their usual meanings. A large F statistic (low p -value) is an indication of an effect due to a model term. However, you cannot construct valid confidence intervals for effects or model predictions.

Residuals from a model fit to deterministic data are not a measure of noise. Instead, residuals are a measure of the model bias. Bias is the difference between the true value and the predicted value. Distinct patterns in the residuals indicate that additional terms should be considered for the model in order to reduce bias.

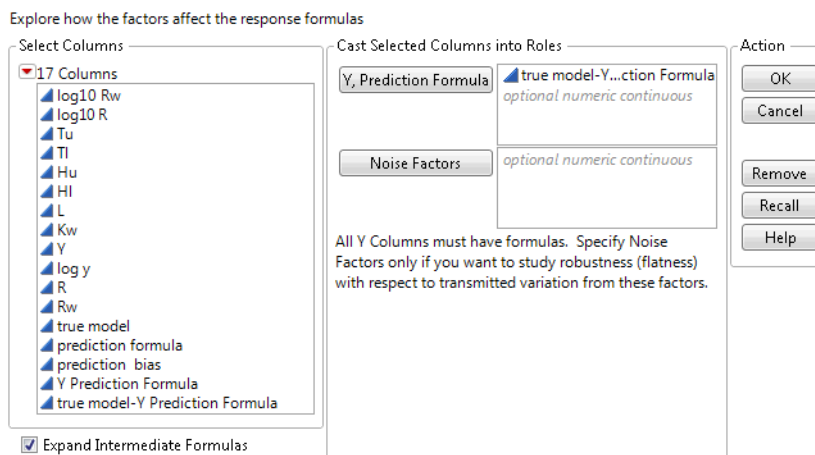
Analysis of the Borehole Sphere-Packing Design

Often, the true model is not available in a simple analytical form. As a result, the prediction bias is known only at observed data points. However, in this example, the functional form of the true model is known. In the Borehole Sphere Packing.jmp data table, the true model column contains the formula of the known function. This formula enables you to profile the prediction bias over the factor input region.

1. Select **Help > Sample Data Library** and open Design Experiment/Borehole Sphere Packing.jmp.
2. Click the green triangle next to the Model (GP from DOE) script.
Use the Gaussian Process Model report to explore the relationships between the factors and the outcome Y.
3. Click the red triangle next to Gaussian Process Model of Y and select **Save Prediction Formula**.
4. Go back to the Borehole Sphere Packing.jmp data table.
5. In the data grid, select the column headings for true model and Y Prediction Formula.
6. Right-click and select **New Formula Column > Combine > Difference**.
This creates a new column containing the bias.
7. From the Borehole Sphere Packing.jmp data table, select **Graph > Profiler**.
8. Select true model-Y Prediction Formula and click **Y, Prediction Formula**
9. Select **Expand Intermediate Formulas**.

This option shows the bias as a function of the eight design factors.

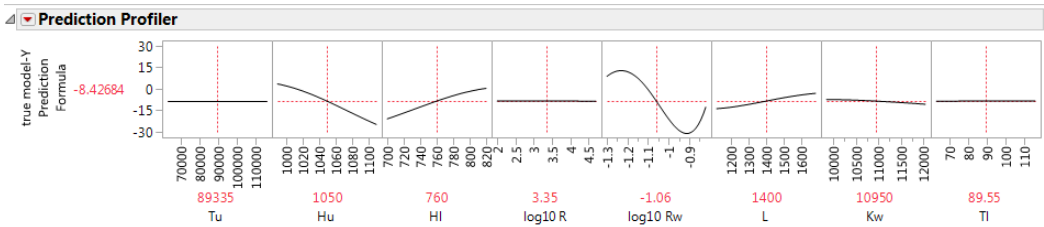
Figure 21.30 Profiler Dialog for Borehole Sphere-Packing Data



10. Click **OK**.

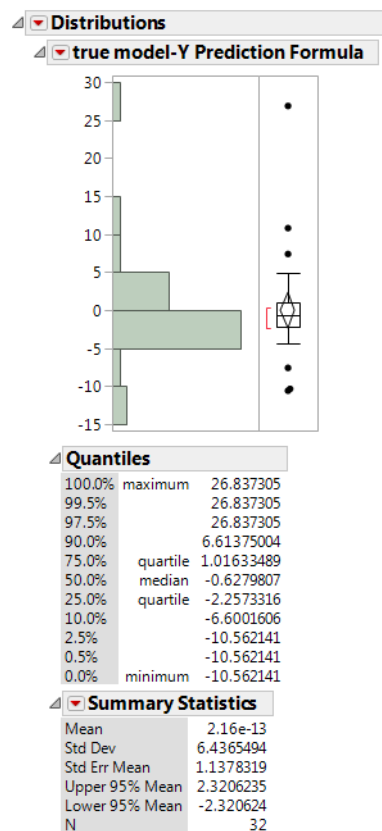
The profiler defaults to the center of the design region. If there were no bias, all profile traces would be constant between the value ranges of each factor. In this example, the variables logRw, Hu, and HI show the largest effects on the bias.

Figure 21.31 Profiler for Bias of the Borehole GP Model with Y Axis Set at -30 to 30



You can use the profiler to explore the range of the prediction bias over the entire domain. To find points of minimum and maximum bias, select **Optimization and Desirability > Desirability Functions** from the Prediction Profiler red triangle menu. See the Profiler chapter in *Profilers*. To evaluate the prediction bias over the design points, select **Analyze > Distribution** to see a distribution analysis.

Figure 21.32 Distribution of the Prediction Bias



Keep in mind that, in this example, the true model is known. In many applications, the response at any factor setting is unknown. The prediction bias over the experimental data can underestimate the bias throughout the design domain.

Chapter 22

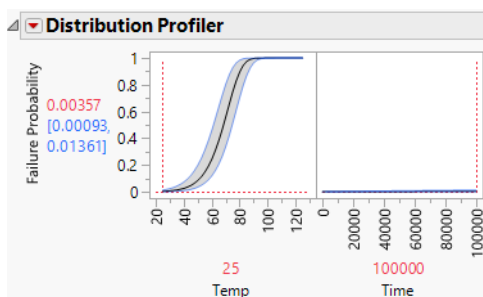
Accelerated Life Test Designs

Designing Experiments for Accelerated Life Tests

Product reliability at normal use conditions is often so high that the time required to test the product until it fails is prohibitive. Rather than test the product at normal use conditions, you can test the product under conditions that are more severe than normal use conditions. These severe conditions cause the product to degrade faster and fail more quickly. You can then use this accelerated failure data to predict product reliability at normal use conditions.

Use the Accelerated Life Test (ALT) Design platform to design plans for accelerated life testing experiments. You can design initial experiments or augment existing experiments.

Figure 22.1 Profiler Showing Failure Probabilities for ALT Experiment



Contents

Overview of Accelerated Life Test Designs.....	653
Example of an Accelerated Life Test Design.....	653
Obtain Prior Estimates.....	654
Enter Basic Specifications.....	655
Enter Prior Information and Remaining Specifications.....	656
Make the Design.....	658
Run Simulations.....	660
Save Design.....	661
Example of Augmenting an Accelerated Life Test Design.....	661
Accelerated Life Test Plan Window.....	666
Specify the Design Structure.....	667
Specify Acceleration Factors.....	667
Specify Design Details.....	668
Review Balanced Design Diagnostics and Update Specifications.....	671
Create and Assess the Optimal Design.....	672
Create Design Tables.....	674
Platform Options.....	674
Statistical Details for the ALT Design Platform.....	675

Overview of Accelerated Life Test Designs

When product reliability at normal use conditions is high, you can test the product under conditions that are more severe than normal to reduce the time until failures occur. You can use this accelerated failure data to predict product reliability at normal use conditions.

Such a test is called an *accelerated life test* (ALT). The factors that are set at higher than normal use levels are called *acceleration factors*. The models for accelerated life tests are typically nonlinear models. For more information about nonlinear models, see [“Nonlinear Models”](#) on page 699 in the “Nonlinear Designs” chapter.

The ALT Design platform creates and evaluates designs for situations involving one or two acceleration factors. You can optimize your design using D-optimality, quantile estimation, or failure probabilities.

Creating an ALT design requires initial estimates of the acceleration model parameters. Use subject matter knowledge to define starting estimates. If you do not have information to inform your starting estimates, you can specify a multivariate normal prior distribution to describe their uncertainty.

You can also use the Accelerated Life Test (ALT) Design platform to augment a current experiment. You might want to augment a current design to decrease the variance of estimates.

Example of an Accelerated Life Test Design

In this example, suppose that you need to design an accelerated life test for a mechanical component. The single acceleration factor is torque, and the normal use stress is 35 Nm (newton meters). You want to estimate the time at which 10% of the units fail at the normal use stress. You construct a failure probability optimal ALT test.

Your test plan has the following characteristics:

- A total of 100 units are available for testing.
- The life distribution is assumed to be Weibull.
- The life-stress relationship is given by the logarithmic transformation.
- You plan to test torque at three stress levels: 50, 75, and 100 Nm.
- You have prior knowledge to predict initial failure times at the test levels. See [“Obtain Prior Estimates”](#) on page 654.
- The test will continue for 5,000 cycles.
- You will monitor the process for failures on a continuous basis.

Obtain Prior Estimates

To create an accelerated life test design, you must provide prior estimates of the parameters. Here is an approach to obtaining prior estimates:

1. Use your process knowledge to create a table of hypothetical, but likely, failure times at a small number of stress levels.
2. Use the Fit Life by X platform to fit a model and obtain estimates of the model parameters.
3. Use these estimates as your prior values for creating a design using the ALT Design platform.

Following the approach outlined above, you create a data table containing estimates of the number of failure cycles for a balanced design. Your table consists of five units at each of the three stress levels to be used in your design. Your data table is Torque Prior.jmp, found in the Sample Data Library.

1. Select **Help > Sample Data Library** and open Design Experiment/Torque Prior.jmp.
2. Select **Analyze > Reliability and Survival > Fit Life by X**.
3. Select Cycles and click **Y, Time to Event**.
4. Select Torque and click **X**.
5. From the **Relationship** list, select **Log**.
6. Enter 35 for the **Use Condition**.
7. From the **Distribution** list, select **Weibull**.

Figure 22.2 Fit Life by X Launch Window

Models life distributions parameterized by a single regression factor.

Select Columns
☒ 2 Columns
☒ Torque
☒ Cycles

Censor Code:

Relationship
 Log
☒ Nested Model Tests
 Use Condition

Distribution
 Weibull

Select Confidence Interval Method
 Wald

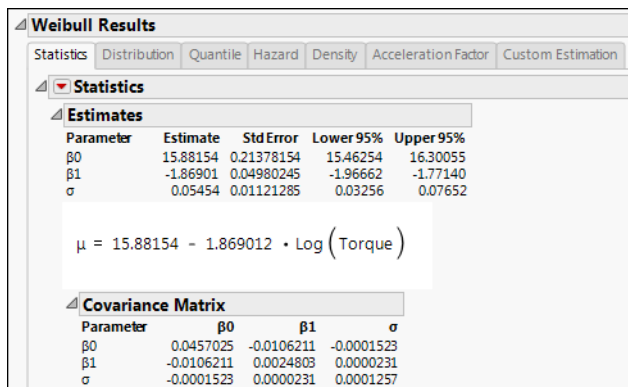
Cast Selected Columns into Roles

Y, Time to Event	☒ Cycles <i>optional numeric</i>
X	☒ Torque <i>optional</i>
Censor	<i>optional</i>
Freq	<i>optional numeric</i>
By	<i>optional</i>

Action
 OK
 Cancel
 Remove
 Recall
 Help

8. Click **OK**.
9. Scroll down to the Weibull Results outline and open the Covariance Matrix outline.

Figure 22.3 Fit Life by X Model for Prior Data



The model for the mean is given in the Estimates outline. The Estimate column contains the parameter estimates for the intercept (β_0), the linear coefficient (β_1), and the scale (σ). The estimated variances and covariances for the parameter estimates are given in the Covariance Matrix outline. Use these parameter estimates and their covariances as your prior values when you construct your ALT design.

Enter Basic Specifications

1. Select **DOE > Special Purpose > Accelerated Life Test Design**.
Notice that **Design for one accelerating factor** and **Continuous Monitoring** are selected by default.
2. Click **Continue**.
3. Under **Factor Name**, click X1 and type Torque.
Notice that the **Number of Levels** is set to 3 by default.
4. Select **Log** under Factor Transformation.
5. Enter 35 for both the **Low Usage Condition** and the **High Usage Condition**.
Setting both the low and high usage conditions at the common value of 35 indicates that 35 represents the normal usage condition.
6. Enter 50 for the **Low Test Condition** and 100 for the **High Test Condition**.

Figure 22.4 Completed ALT Specification Window

The figure shows the 'Accelerated Life Test Plan' dialog box with the following specifications:

Factor Name	Number of Levels	Factor Transformation	Low Usage Condition	High Usage Condition	Low Test Condition	High Test Condition
Torque	3	Log	35	35	50	100

Below the table is a 'Continue' button.

7. Click **Continue**.
8. Notice the **Torque Level Values** are set to 50, 75, and 100.
JMP sets the values of the levels evenly spaced between the low and high test conditions.
9. Ensure that **Weibull** is selected as the **Distribution Choice**.

Enter Prior Information and Remaining Specifications

1. Under **Prior Mean**, enter the prior acceleration model parameter estimates that you obtained using the Fit Life by X platform with the hypothetical data. Round the values shown in the Estimates outline in Figure 22.3 to two decimal places:
 - Enter 15.88 for **Intercept**.
 - Enter -1.87 for **Power (Torque)**.
 - Enter 0.05 for **scale**.
2. Open the Prior Variance Matrix outline.
3. De-select the option **Ignore prior variance. Generate the local design for the prior mean**.
4. Enter the estimated covariances from the prior acceleration model. Round the values shown in the Covariance Matrix outline in Figure 22.3 to five decimal places:
 - In the row and column for Intercept, click 0.10000 and type 0.04570.
 - In the row for Intercept and column for Torque, click 0.00000 and type -0.01062.
 - In the row for Intercept and column for scale, click 0.00000 and type -0.00015.
 - In the row and column for Torque, click 0.10000 and type 0.00248.
 - In the row for Torque and column for scale, click 0.00000 and type 0.00002.
 - In the row and column for scale, click 0.10000 and type 0.00013.

Figure 22.5 Completed Prior Specification Outline

Accelerated Life Test Plan

Factors

Prior Specification

Distribution Choice

☐ LogNormal
☒ Weibull

Prior Mean

☒ Specify Intercept
☐ Specify Quantile

Effect	Prior Mean
Intercept	15.8800
Power (Torque)	-1.8700
scale	0.05000

Prior Variance Matrix

Effect	Intercept	Torque	scale
Intercept	0.04570	-0.0106	-0.0002
Torque		0.00248	0.00002
scale			0.00013

☐ Ignore prior variance. Generate the local design for the prior mean.

Design Choices

Click the Accelerated Life Test Plan red triangle and select **ALT Optimality Criterion > Make Failure Probability Optimal**.

The test will be conducted over 5,000 cycles. However, you are interested in predictions for as many as 10,000 cycles.

- In the Diagnostic Choices outline, enter 10,000 for both boxes for **Time range of interest**.
- In the Design Choices outline. Enter 5,000 for **Length of test** and 100 for **Number of units under test**.

Figure 22.6 Completed Design Details Window

Accelerated Life Test Plan

Factors

Torque

Level Values

50

75

100

Prior Specification

Distribution Choice

☐ LogNormal

☒ Weibull

Prior Mean

☒ Specify Intercept

☐ Specify Quantile

Effect	Prior Mean
Intercept	15.8800
Power (Torque)	-1.8700
scale	0.05000

Prior Variance Matrix

Effect	Intercept	Torque	scale
Intercept	0.04570	-0.0106	-0.0002
Torque		0.00248	0.00002
scale			0.00013

☐ Ignore prior variance. Generate the local design for the prior mean.

Diagnostic Choices

Time range of interest

10000

10000

Design Choices

Length of test

5000

Number of units under test

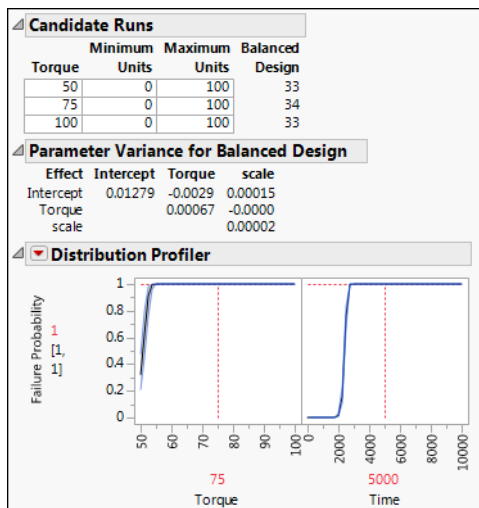
100

Continue

Make the Design

- 1. Click Continue.

Figure 22.7 Balanced Design Diagnostics



The number of runs in a balanced design appears in the Candidate Runs outline. The Parameter Variance for Balanced Design outline shows the covariance matrix for the parameters for this design. The Distribution Profiler also appears. When you obtain the optimal design, you can compare the Parameter Variance and Distribution Profiler results to those for the balanced design to see the reduction in uncertainty.

2. Click **Make Design**.

The optimal experimental design appears, along with other results.

Figure 22.8 Optimal Design

Design			
Torque	N Units	Expected Failures	All Censored Probability
50	72	23.2	0.0
75	8	8.0	0.0
100	20	20.0	0.0

The optimal design is computed based on the levels of the test runs, the total number of units to be tested, and the prior information that you specified. The optimal design consists of testing the following number of units at each torque level:

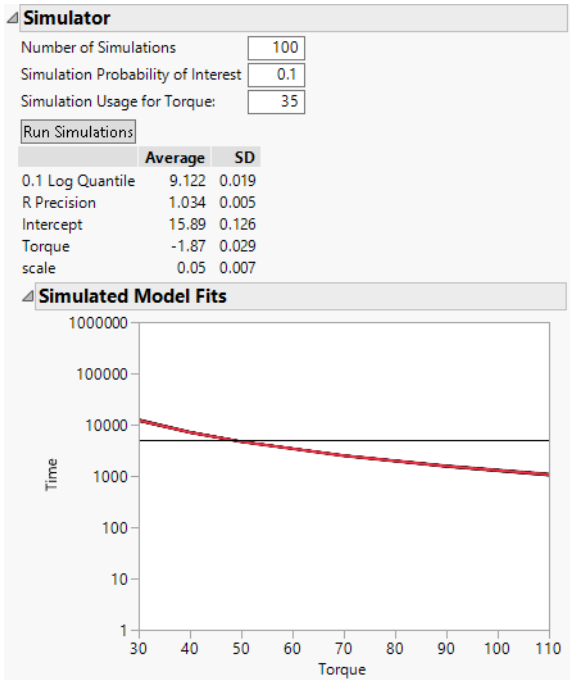
- 72 units at 50 Nm
- 8 units at 75 Nm
- 20 units at 100 Nm

Run Simulations

Use the simulator to simulate the results of your test plan based on the assumptions used to build the design. This is a way to evaluate if the test plan is sufficient before running the test plan.

1. Enter 100 for the **Number of Simulations**.
2. Enter 35 for the **Simulation Usage for Torque**.
3. Click **Run Simulations** to run 100 simulations of the experiment. This is to check to see whether the 100 units that you plan to test will be sufficient given your prior assumptions for the design.

Figure 22.9 Simulation Results



The simulated model fits shows that the 100 simulations of the model are all similar as they all fall close to the average shown by the red curve. The Y axis is on a log scale.

4. Double-click the Y axis and select **Linear** from the Type menu and click **OK**.

The variability in the simulated fits around the 35 normal usage condition is now easier to view. You expect variability in your estimated of the normal usage condition. The simulation helps you evaluate if your ALT plan is acceptable. If there is too much variability, you might need to increase the number of units on test or the test conditions.

Save Design

1. Click **Make Test Plan** to save the design to a data table. This table is a summary of the design.
2. Click **Make Table** to save the design to a data table. This table has 100 rows, one for each test run.

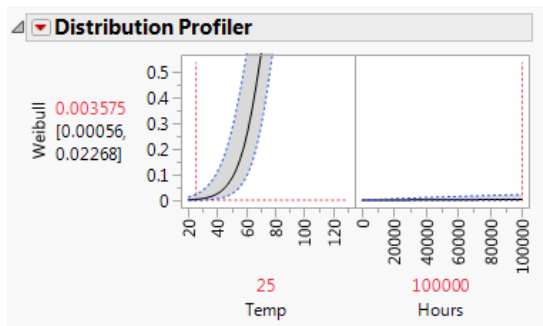
Example of Augmenting an Accelerated Life Test Design

This example shows how to use the Accelerated Life Test Design platform to augment an existing design. In this example, 150 capacitor units are tested across three temperatures (85°, 105°, and 125° Celsius) for 1500 hours. The results are recorded in the Capacitor ALT.jmp sample data table. The resulting model is used to predict the fraction of the population that is failing at 100,000 hours at a normal usage temperature of 25° Celsius.

Review Current Predictions

1. Select **Help > Sample Data Library** and open Design Experiment/Capacitor ALT.jmp.
2. Click the green triangle to run the **Fit Life by X** table script.
3. In the Distribution Profiler, found in the Comparisons outline on the Distribution tab, do the following:
 - Click 105 above Temp and change it to 25.
 - Click 750.5 above Hours and change it to 100,000.

Figure 22.10 Distribution Profiler for Capacitor Model



Based on your current study, the predicted fraction of the population that fails at 25° at 100,000 hours is 0.00358, with a 95% confidence interval of 0.00056 to 0.02268. You want to estimate the failure fraction more precisely. To decrease the width of the confidence interval, you can augment your study with additional tests.

Augment the Design

You want to augment your design to obtain more precise estimates of the predicted failure fractions. Your original design used temperature settings of 85, 105, and 125. In your augmented design, you want to test at temperature values of 90, 110, and 125. Note that two of these settings are new. To augment the design with optimally selected runs, follow these steps:

- 1. Select **DOE > Special Purpose > Accelerated Life Test Design**.
- 2. Select **Design for one accelerating factor**, select **Continuous Monitoring**, and click **Continue**.
- 3. Enter Temp for **Factor Name**.
- 4. Enter 5 for **Number of Levels**.

Even though your augmented runs span only three levels (90, 110, and 125), you must also specify the levels used in the original experiment for a total of five levels. The Factor Transformation is set to Arrhenius Celsius by default.

- 5. Enter 25 for both **Low Usage Condition** and **High Usage Condition**.
- 6. Enter 85 for the **Low Test Condition** and 125 for the **High Test Condition**.
- 7. Click **Continue**.
- 8. Enter 85, 90, 105, 110, and 125 for the **Temp Level Values**.

There are three levels from the original experiment (85, 105, and 125). The augmented design has two new levels (90 and 110) and one of the levels from the first experiment (125).

Note: All levels must be listed.

- 9. Ensure that **Weibull** is selected as the **Distribution Choice**.
- 10. Under **Prior Mean**, select **Specify Intercept**. Enter the acceleration model parameter estimates from the Fit Life by X Estimates outline, found in the Weibull Results outline on the Statistics tab.

Figure 22.11 Parameter Estimates and Fitted Model from Weibull Results Outline

Estimates				
Parameter	Estimate	Std Error	Lower 95%	Upper 95%
β0	-35.19979	4.6912686	-44.39451	-26.00508
β1	1.38891	0.1566297	1.08192	1.69589
σ	1.30471	0.1119128	1.08536	1.52405

$$\mu = -35.19979 + \frac{(1.388906 \cdot 11605)}{(Temp + 273.15)}$$

- Enter -35.200 for **Intercept**.
- Enter 1.389 for **Activation Energy (Temp)**.

This is an estimate of the activation energy and is the coefficient of the inverse temperature, measured in degrees Kelvin, multiplied by Boltzmann's constant.

- Enter 1.305 for **scale**.

For the Weibull distribution, JMP uses a parameterization that depends on a location parameter μ and scale parameter σ . In terms of the usual α and β parameterization, the scale parameter is $\sigma = 1/\beta$. See "[Weibull](#)" on page 676.

In the Accelerated Life Test Plan window, you can specify uncertainty about your prior means in the Prior Variance Matrix outline. In this example, do not make any changes to the Prior Variance Matrix outline. Your design is created assuming that the values specified in the Prior Means outline are the true parameter values.

11. Under Design Choices, enter 1500 for **Length of test**.

The test is conducted over 1500 hours, which was the length of the original design.

12. Enter 250 for **Number of units under test**.

The previous experiment tested 150 units, and the augmented experiment tests an additional 100 units, for a total of 250 units.

Figure 22.12 Completed Design Details Window

Accelerated Life Test Plan

Factors

Temp Level Values
85
90
105
110
125

Prior Specification

Distribution Choice

☐ LogNormal
☒ Weibull

Prior Mean

☒ Specify Intercept
☐ Specify Quantile

Effect	Prior Mean
Intercept	-35.2
Activation Energy (Temp)	1.389
scale	1.305

Prior Variance Matrix

Effect	Intercept	Temp	scale
Intercept	0.10000	0.00000	0.00000
Temp		0.10000	0.00000
scale			0.10000

☒ Ignore prior variance. Generate the local design for the prior mean.

Diagnostic Choices

Probability of interest

Design Choices

Length of test
 Number of units under test

13. Click **Continue**.

14. To account for the units at each setting of Temp from the previous experiment, enter the following under **Candidate Runs**.

- For the Temperatures of 85, 105, and 125, enter Minimum Units of 50 for each.

15. Click the Accelerated Life Test Plan red triangle and select **ALT Optimality Criterion > Make Failure Probability Optimal**.

For more information about this criterion, see [“Make Failure Probability Optimal”](#) on page 675.

16. Click **Make Design**.

The optimal experimental design appears along with other results.

Figure 22.13 Optimal Design

Design			
Temp	N Units	Expected Failures	All Censored Probability
85	50	6.9	0.0
90	89	18.8	0.0
105	50	30.0	0.0
110	0	0.0	0.0
125	61	60.5	0.0

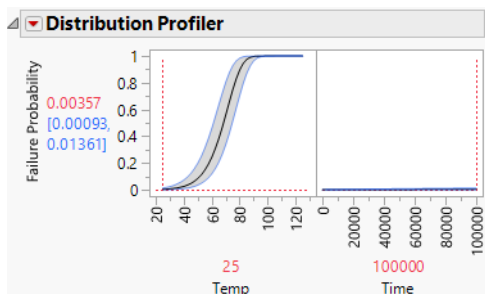
The optimal design is computed based on the levels of the test runs, the minimum number of units under test, the total number of units to be tested (this is the information in the Candidate Runs outline), and other information that you specified. The optimal design consists of testing the following number of units at each temperature level:

- 50 units at 85°C. Since the previous experiment already tested 50 units at 85°C, no additional units are needed.
- 89 units at 90°C.
- 50 units at 105°C. Since the previous experiment test 50 units at 105°C, no additional units are needed.
- 0 units at 110°C. This level is not needed.
- 61 units at 125°C. Since the previous experiment test 50 units at 125°C, 11 additional units are needed.

Compare the Augmented Design to the Original Study

1. In the Distribution Profiler, enter 25 for Temperature and 100,000 for Time. The estimate of the fraction of the population that is failing is 0.00357, with a 95% confidence interval of 0.00093 to 0.01361. This interval is narrower than the one from the initial experiment, which was from 0.00056 to 0.02268 (Figure 22.10).

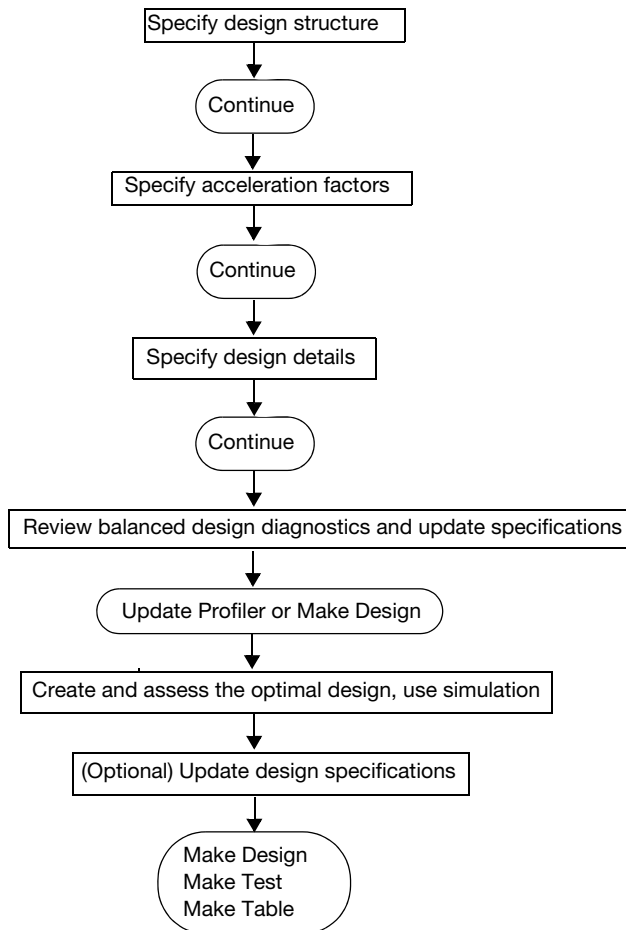
Figure 22.14 Distribution Profiler for Temp = 25 and Time = 100000



Accelerated Life Test Plan Window

The Accelerated Life Test Plan window is updated as you work through the design steps. The outlines that appear, separated by buttons that update the window, follow the flow in Figure 22.15.

Figure 22.15 Accelerated Life Test Plan Flow

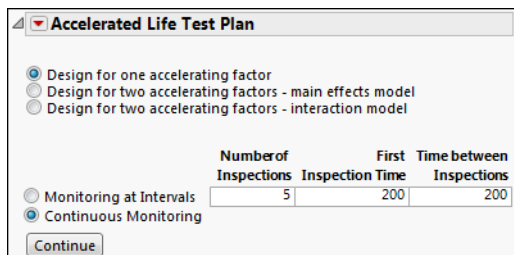


This section describes the outlines in the Accelerated Life Test Plan flow.

Specify the Design Structure

Use the Accelerated Life Test Plan outline to define the model structure and the type of inspection for your design.

Figure 22.16 Initial ALT Design Window



The screenshot shows the 'Accelerated Life Test Plan' window. It has three radio buttons for design types: 'Design for one accelerating factor' (selected), 'Design for two accelerating factors - main effects model', and 'Design for two accelerating factors - interaction model'. Below these are two more radio buttons: 'Monitoring at Intervals' and 'Continuous Monitoring' (selected). A table with four columns is visible: 'Number of Inspections', 'First Inspection Time', and 'Time between Inspections'. The 'Number of Inspections' column has a value of 5. The 'First Inspection Time' column has a value of 200. The 'Time between Inspections' column has a value of 200. A 'Continue' button is at the bottom.

	Number of Inspections	First Inspection Time	Time between Inspections
Monitoring at Intervals	5	200	200

Design for one accelerating factor Select for a design with one factor.

Design for two accelerating factors

main effects model Select for a design with two factor design for a main effects model

interaction model Select for a design with two factor design for a model that contains main effects and an interaction term for the two factors.

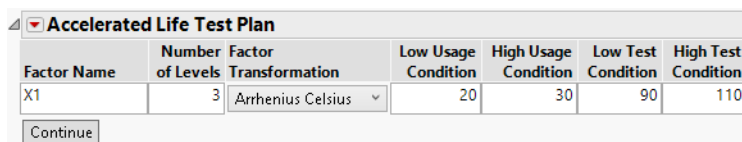
Monitoring at Intervals Assumes that units are inspected for failures at intervals. Failure times are interval censored. Enter the number of inspections, the time of the first inspection, and a time between inspections. For inspection intervals that are irregular, you can change the inspection times later in the Design Choices outline.

Continuous Monitoring Assumes that exact failure times are recorded. Failure times beyond the length of the test are right censored.

Specify Acceleration Factors

Specify details about the acceleration factor or factors.

Figure 22.17 ALT Specification Window



The screenshot shows the 'Accelerated Life Test Plan' window with a table for specifying acceleration factors. The table has seven columns: 'Factor Name', 'Number of Levels', 'Factor Transformation', 'Low Usage Condition', 'High Usage Condition', 'Low Test Condition', and 'High Test Condition'. There is one row with the following values: 'X1', 3, 'Arrhenius Celsius' (with a dropdown arrow), 20, 30, 90, and 110. A 'Continue' button is at the bottom.

Factor Name	Number of Levels	Factor Transformation	Low Usage Condition	High Usage Condition	Low Test Condition	High Test Condition
X1	3	Arrhenius Celsius	20	30	90	110

Factor Name Enter a name for each acceleration factor.

Number of Levels For each acceleration factor, enter the number of proposed levels that you want to include in the experiment.

Factor Transformation For each acceleration factor, select a transformation function. This transformation describes the life-stress relationship, which is the manner in which the life distribution changes across stress levels. The transformations available are Arrhenius Celsius, Reciprocal, Log, Square Root, and Linear.

Low Usage Condition For each acceleration factor, enter a lower bound for its value in typical usage conditions.

High Usage Condition For each acceleration factor, enter an upper bound for its value in typical usage conditions.

Note: The Low Usage Condition and High Usage Condition values can be equal. Use equal values when the usage condition is a single value.

Low Test Condition For each acceleration factor, enter the low test condition.

High Test Condition For each acceleration factor, enter the high test condition.

Note: The low and high test conditions define the endpoints of the acceleration factors in the design. The initial candidate will use equally spaced levels between the low and high test conditions. You can adjust the levels before constructing the design.

Specify Design Details

Specify the factor levels, details of the prior distribution, the time range and probability of interest, the length of the test, and the number of units to be tested.

Figure 22.18 Distribution Details

Accelerated Life Test Plan

Factors

X1 Level Values

90
100
110

Prior Specification

Distribution Choice

☐ LogNormal
☒ Weibull

Prior Mean

☒ Specify Intercept
☐ Specify Quantile

Effect	Prior Mean
Intercept	-40.000
Activation Energy (X1)	1.50000
scale	2.00000

Prior Variance Matrix

Effect	Intercept	X1	scale
Intercept	0.10000	0.00000	0.00000
X1		0.10000	0.00000
scale			0.10000

☒ Ignore prior variance. Generate the local design for the prior mean.

Design Choices

Length of test: 1000

Number of units under test: 150

Continue

Factors Enter the levels for the acceleration factors. By default, the levels are evenly spaced between the low and high test conditions.

Distribution Choice Select a life distribution (Weibull or LogNormal) for each acceleration factor. See [“Statistical Details for the ALT Design Platform”](#) on page 675.

Prior Mean Enter prior estimates of the acceleration model parameters. The prior estimates are hyperparameters in a Bayesian prior distribution. The Prior Mean values can be a best guess based on subject matter knowledge or they can be based on a previous study.

When there is only one acceleration factor, there is a choice to specify the intercept or the quantile.

Specify Intercept Enter the priors for the model parameters including the intercept.

Specify Quantile Enter the priors for the acceleration factor, the expected failure proportion at a given time and value for the acceleration factor. This information is used to calculate the intercept.

Prior Variance Matrix Enter values for the variances and covariances of the prior distribution for the acceleration model parameters. These variances and covariances reflect uncertainty relative to the prior estimates of the acceleration model parameters. Large variances indicate greater uncertainty.

Ignore prior variance Select this option to ignore the prior variances and covariances for the prior distribution. When these variances and covariances are ignored, the design is created by treating the values entered under Prior Mean as the true parameter values. This design is close to optimal if the prior mean parameters are close to the true values. However, this design is not robust to misspecification of the parameter estimates. If you are unsure about your prior estimates, use the Prior Variance Matrix to reflect your uncertainty.

Diagnostic Choices The choices available depend on the design optimality criteria. For a D-optimal design there are no diagnostic choices.

Time range of interest (Available only for a failure probability optimal design.) Specify the time interval over which you want to estimate the fraction of the population that is failing. Enter the lower value in the left box and the upper value in the right box. If you are interested in a specific time point, enter that value in both boxes.

Probability of interest (Available only for a quantile estimate optimal design.) Specify the failure fraction for which you want an estimate of time. For example, if you want to estimate the time at which 10% of the units fail, then enter 0.10.

Design Choices Specify values relating to the length of the test, inspection intervals, and the number of units being tested. Enter values for the following:

Length of test (Available only for Continuous Monitoring.) Specify the length of time during which units are on test. When you make the design table, record each unit's failure time or whether it was right censored.

Inspection Times (Available only for Monitoring at Intervals.) Specify the times at which inspections are conducted. When you make the design table, these times are used to construct Start Time and End Time columns. The number of units failing in each interval is recorded.

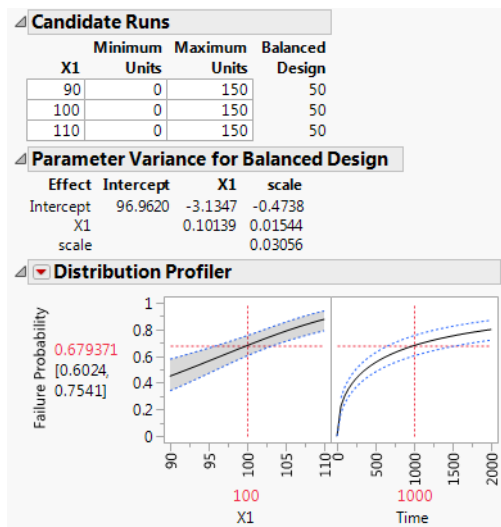
Number of units under test The number of units in the experiment.

- If you are designing an initial experiment, enter the number of units that you plan to test.
- If you are augmenting a previous experiment, enter the number of units tested in the previous experiment plus the number of units for the next experiment.

Review Balanced Design Diagnostics and Update Specifications

After you specify the design details and click **Continue**, three new outlines are added to the window.

Figure 22.19 Additional Outline Nodes



Candidate Runs Enter the minimum and maximum number of units allowed at each level of the acceleration factors. If you are augmenting a previous experiment, for each setting, include the number of units already run at each level in the Minimum Units column.

Parameter Variance for Balanced Design Gives a matrix proportional to the covariance matrix for the estimates of the acceleration model parameters for the balanced design.

Denote the matrix of first partial derivatives of the model with respect to the parameters, θ , by $\mathbf{X}$. Denote the error variance by σ^2 . Under general conditions, the least squares estimator of θ is asymptotically unbiased with asymptotic covariance matrix as follows:

$$\text{Cov}(\hat{\theta}) = \sigma^2 (\mathbf{X}'\mathbf{X})^{-1}$$

The Parameter Variance for Balanced Design outline gives $(\mathbf{X}'\mathbf{X})^{-1}$, where derivatives are calculated numerically. The calculation assumes that the values specified in the Prior Mean outline are the true parameter values.

See “[Nonlinear Models](#)” on page 699 in the “Nonlinear Designs” chapter.

Distribution Profiler Enables you to explore failure probabilities, based on the balanced design, as you vary the acceleration factors and time. The probabilities are based on the assumption that the values specified in the Prior Mean are the true parameter values.

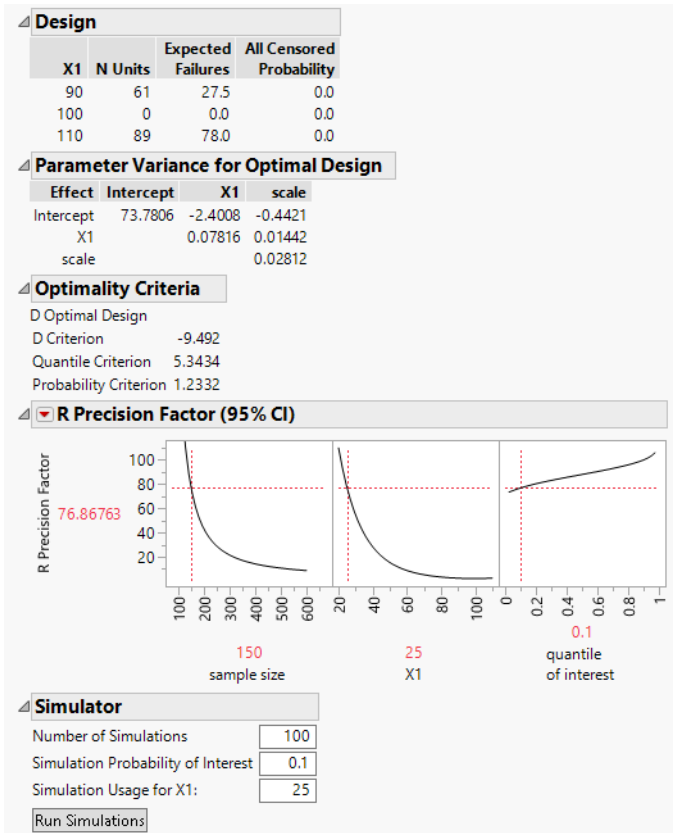
Tip: Click **Update Profiler** to update the profiler if changes are made to Distribution Choice, Prior Mean, or Design Choices.

Create and Assess the Optimal Design

Click **Make Design** to create the optimal design and see results that address the quality of the design. By default, the optimal design is D-optimal. You can change the optimality criterion by selecting the red triangle option ALT Optimality Criterion.

When you click **Make Design**, the following report appears:

Figure 22.20 Design Outlines



Design Shows the number of units to be tested at each combination of acceleration levels. For a single factor, the first column gives the levels of the factor. For two factors, the first two columns give the levels of the factors.

N Units The number of units to be tested at each combination of acceleration levels.

Expected Failures Expected number of failures for the design setting. The expected number is computed using the prior model specification.

All Censored Probability Probability that none of the units tested at the combination of acceleration levels fail. The probability is computed using the prior model specification.

Note: The expected failures and censoring probabilities enable you to judge whether your prior specifications are reasonable.

Parameter Variance for Optimal Design Shows a matrix proportional to the covariance matrix for the estimates of the acceleration model parameters for the optimal design. The calculation assumes that the values specified in the Prior Mean outline are the true parameter values. See [“Parameter Variance for Balanced Design”](#) on page 671.

Note: Compare the values in the Parameter Variance for Optimal Design matrix to those in the Parameter Variance for Balanced Design matrix to determine the extent to which the optimal design reduces the variance of estimates.

Optimality Criteria Values for three optimality criteria appear:

D Criterion D-optimality of the design. See [“Make D-Optimal Design”](#) on page 674.

Quantile Criterion Time I-optimality of the design. See [“Make Quantile Estimate Optimal”](#) on page 674.

Probability Criterion Probability I-optimality of the design. See [“Make Failure Probability Optimal”](#) on page 675.

R Precision Factor (95% CI) A profiler to evaluate the impact of sample size, level of the factor(s), and the quantile of interest on the R Precision Factor. The R Precision Factor (95% CI) reflects the confidence interval for the quantile of interest.

Simulator Use the simulator to explore your accelerated model before collecting data.

Number of Simulations The number of simulation trials to run.

Simulation Probability of Interest The failure probability of interest.

Simulation Usage Values for the usage conditions of the acceleration factors.

After clicking Run Simulations, a table of summary statistics and a plot of model fits appear.

Create Design Tables

Make Design Updates the optimal design if any changes are made to the Distribution Choice, Prior Mean, Prior Variance Matrix, Design Choices, or Candidate Runs.

Make Test Plan Creates a data table where each row corresponds to a distinct design setting. The table shows the acceleration factor design settings and the number of units to test at those design settings.

Make Table Creates a table that you can use for recording your failure time data.

- For Continuous Monitoring, the table contains a row for each unit to be tested and the design settings for that unit. If a unit fails during the experiment, record the unit's failure time in the Failure Time column and a 0 in the Censored column. If a unit does not fail during the experiment, record the final inspection time in the Failure Time column and a 1 in the Censored column.
- For Monitoring at Intervals, the table contains a row for each design setting and time interval combination. The time intervals are defined by the Start Time and End Time columns, which are based on the Inspection Times entered in the Design Choices outline. For each setting and time interval, record the number of units that failed in the Number Failing column.

Platform Options

The Accelerated Life Test Plan red triangle menu contains the following options:

Simulate Responses Adds simulated responses to the table when you click **Make Table**. The simulated responses are created by taking random draws from the chosen distribution at the parameter values specified in the Prior Mean outline. If a simulated response exceeds the specified test length, the observation is censored at the test length value.

ALT Optimality Criterion Contains the following choices for design optimality:

Make D-Optimal Design Creates a Bayesian D-optimal design if the number of Monte Carlo spheres is greater than 0. The optimality criterion is the expectation of the logarithm of the determinant of the information matrix with respect to the prior distribution. If the number of Monte Carlo spheres is 0, then the design is a locally D-optimal design. It follows that D-optimality focuses on precise estimates of the coefficients.

Make Quantile Estimate Optimal Default optimality criteria. Creates a design that minimizes the variance of a specified percentile at the specified usage conditions.

Make Failure Probability Optimal Creates a design that minimizes the variance of the estimate of failure probability over the time range of interest at the specified usage conditions.

Advanced Options Contains the following advanced option:

N Monte Carlo Spheres Enables you to set the number of nonzero radius values used in the integration. To find a nonlinear design that optimizes a given optimality criterion, JMP must minimize the integral of the log of the determinant of the Fisher information matrix with respect to the prior distribution of the parameters. Such an integral must be calculated numerically. For more information about how the integration is performed, see [“Nonlinear Design Options”](#) on page 698 in the “Nonlinear Designs” chapter.

Tip: By default N Monte Carlo Spheres is set to four. Higher values result in better numerical accuracy but with more computation time.

Save Script to Script Window Creates the script for the design that you specified in the Accelerated Life Test Plan window and places it in an open script window.

Statistical Details for the ALT Design Platform

In the ALT design platform, you can select either a lognormal or Weibull failure distribution.

This section contains the parameterizations for the probability density function (pdf) and cumulative distribution function (cdf) for each distribution. For more information about the Weibull distribution, see the Life Distribution chapter in *Reliability and Survival Methods*.

Lognormal

Lognormal distributions are used commonly for failure times when the range of the data is several powers of 10. This distribution is often conceptualized as the multiplicative product of many small positive independently and identically distributed random variables. This distribution is appropriate when the logarithms of the data values appear normally distributed. The probability distribution function is usually characterized by strong right-skewness.

The lognormal family is parameterized by a location parameter, μ , and a shape parameter, σ . The lognormal pdf and cdf are given as follows, where the logarithm is to the base e :

$$f(x; \mu, \sigma) = \frac{1}{x\sigma} \phi_{\text{nor}}\left[\frac{\log(x) - \mu}{\sigma}\right], \quad x > 0$$

$$F(x;\mu,\sigma) = \Phi_{\text{nor}}\left[\frac{\log(x) - \mu}{\sigma}\right],$$

The functions

$$\phi_{\text{nor}}(z) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{z^2}{2}\right)$$

and

$$\Phi_{\text{nor}}(z) = \int_{-\infty}^z \phi_{\text{nor}}(w) dw$$

are the pdf and cdf, respectively, for the standard normal distribution (N(0,1)).

Weibull

The Weibull distribution can be used to model failure time data with either an increasing or a decreasing hazard rate. It is used frequently in reliability analysis because of its tremendous flexibility in modeling many different types of data, based on the values of the shape parameter.

The Weibull pdf and cdf are commonly represented as follows:

$$f(x;\alpha,\beta) = \frac{\beta}{\alpha} x^{(\beta-1)} \exp\left[-\left(\frac{x}{\alpha}\right)^\beta\right]; \quad x > 0, \alpha > 0, \beta > 0$$

$$F(x;\alpha,\beta) = 1 - \exp\left[-\left(\frac{x}{\alpha}\right)^\beta\right]$$

where α is a scale parameter, and β is a shape parameter. The Weibull distribution reduces to an exponential distribution when $\beta = 1$.

Weibull Parameterization

In JMP, the Weibull parameterization defines σ as the scale parameter and μ as the location parameter. These are related to the α and β Weibull parameterization as follows:

$$\alpha = \exp(\mu)$$

and

$$\beta = \frac{1}{\sigma}$$

With these parameters, the pdf and the cdf of the Weibull distribution are expressed as a log-transformed smallest extreme value distribution (SEV) using a location-scale parameterization with $\mu = \log(\alpha)$ and $\sigma = 1/\beta$:

$$f(x; \mu, \sigma) = \frac{1}{x\sigma} \phi_{\text{sev}} \left[\frac{\log(x) - \mu}{\sigma} \right], \quad x > 0, \sigma > 0$$

$$F(x; \mu, \sigma) = \Phi_{\text{sev}} \left[\frac{\log(x) - \mu}{\sigma} \right]$$

where

$$\phi_{\text{sev}}(z) = \exp[z - \exp(z)]$$

and

$$\Phi_{\text{sev}}(z) = 1 - \exp[-\exp(z)]$$

are the pdf and cdf, respectively, for the standardized smallest extreme value ($\mu = 0, \sigma = 1$) distribution.

R Precision Factor (95% CI)

The R precision factor is a Wald confidence interval for a positive quantity t:

Chapter 23

Nonlinear Designs

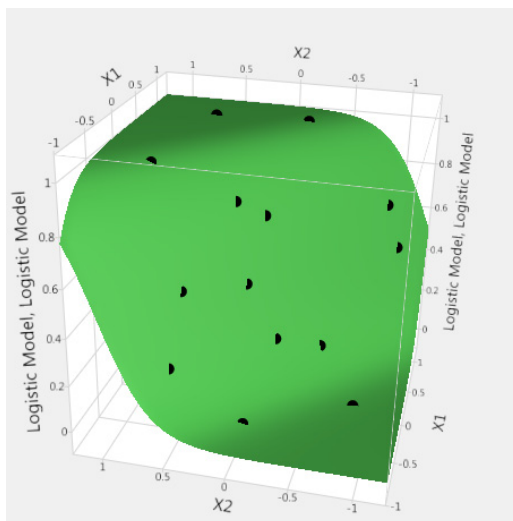
When the goal of your experiment is to fit a model that is nonlinear in the unknown parameters, use a nonlinear design to place design points in areas that are key to fitting the nonlinear model. Although you could use an orthogonal design that is optimal for a linear model, such designs, in general, do not place design points in locations that minimize the uncertainty (or maximize the precision) of the estimates of the fitted parameters.

The efficiency of a nonlinear design depends on the values of the unknown parameters. This creates a circular problem in that to find the best design, you need to know the parameters in advance. JMP uses a Bayesian approach to construct a nonlinear design that maximizes the average efficiency over specified ranges of the values of the parameter. To properly specify these ranges, you must have some insight about the system of interest.

Nonlinear designs offer these advantages compared to designs for linear models:

- Predictions using a well-chosen model are likely to be good over a wider range of factor settings.
- It is possible to model response surfaces with complex curvature and with asymptotic behavior.

Figure 23.1 Design Points for a Nonlinear Model



Contents

Overview of Nonlinear Designs	681
Examples of Nonlinear Designs	682
Create a Nonlinear Design with No Prior Data	682
Augment a Design Using Prior Data	687
Create a Design for a Binomial Response	691
Nonlinear Design Launch Window	695
Nonlinear Design Window	696
Factors	696
Parameters	697
Design Generation	697
Design	698
Make Table or Augment Table	698
Nonlinear Design Options	698
Statistical Details for Nonlinear Designs	699
Nonlinear Models	699
Radial-Spherical Integration of the Optimality Criterion	700
Finding the Optimal Design	700

Overview of Nonlinear Designs

Construct designs to fit models that are nonlinear in their parameters using the Nonlinear Design platform. You can construct optimal designs or optimally augment existing data for nonlinear models. Nonlinear designs based on information that is descriptive of the underlying process can yield more accurate estimates of model parameters and prediction of process behavior than is possible with standard designs for polynomial models. For background on nonlinear models, see [“Nonlinear Models”](#) on page 699.

The efficiency of a design for a nonlinear model depends on the unknown values of the parameters that the design is intended to estimate. For this reason, JMP uses a Bayesian approach to construct designs that are efficient over a wide range of likely parameter values. You can specify a range of values for the unknown parameters and a distribution for the prior. The prior distribution choices include Uniform, Normal, Lognormal, and Exponential.

The Nonlinear Design platform uses a Bayesian approach, optimizing the design over a prior distribution of likely parameter values that you specify. The Bayesian D-optimality criterion is the expectation of the logarithm of the determinant of the information matrix with respect to a sample of parameter vectors that represents this prior probability distribution. The information matrix entries depend on the prediction variances at the design points. Little information is contributed by observations with low variance, where the response is almost certain. It follows that an optimal design places some design settings at high-variance points. See Gotwalt et al. (2009).

A principal of optimal design is that, over the feasible region of experimentation, the optimal design places points in locations with the highest variance of prediction. Though this may seem counter-intuitive, if an alternative design put points at other locations, the prediction variance at the design points of the optimal design would be even higher. For models that are linear in the parameters, the high-variance points tend to be at the vertices of the experimental region. But this is not necessarily true for models that are nonlinear in the parameters.

Note: Nonlinear designs are computed using a random starting design. For this reason, nonlinear designs that you obtain for identical specifications usually differ.

To use the Nonlinear Design platform, you must have an existing data table. That data table must contain the following:

- A column for the response.
- A column for each factor.
- A column that contains a formula showing the relationship between the factors and the response. This formula must include the unknown parameters.

Note: This is the same format as is required for a data table used in the Nonlinear platform for modeling.

Your table can come in one of two forms:

- It might be a template, containing only column information and no rows. See [“Create a Nonlinear Design with No Prior Data”](#) on page 682.
- It might contain rows with predictor information. In this case, the predictor values are included in the nonlinear design. See [“Augment a Design Using Prior Data”](#) on page 687.

Examples of Nonlinear Designs

- [“Create a Nonlinear Design with No Prior Data”](#)
- [“Augment a Design Using Prior Data”](#)
- [“Create a Design for a Binomial Response”](#)

Create a Nonlinear Design with No Prior Data

This example shows how to create a design when you have not yet collected data, but have a guess for the unknown parameters. In this example, you model the fractional yield (Observed Yield) of an intermediate product in a chemical reaction. The fractional yield is a function of reaction time and temperature. See Box and Draper (1987).

Create the Design

1. Select **Help > Sample Data Library** and open Design Experiment/Reaction Kinetics Start.jmp. Notice the following:
 - The data table contains no rows because no data have been collected.
 - The columns for the predictors, Reaction Temperature and Reaction Time, have the Coding, Design Role, and Factor Changes properties. To see these properties, click  in the Columns panel. They tell JMP how to treat these predictors when constructing a design. For information about how to save these column properties, see the [“Column Properties”](#) chapter on page 729.
 - The Observed Yield column will contain response data obtained by running the experiment.
 - The Yield Model column contains the formula that relates the predictors to the response, Observed Yield. Click  in the Columns panel to see the formula. The formula is nonlinear in the parameters t1 and t3.

2. Select **DOE > Special Purpose > Nonlinear Design**.
3. Select Observed Yield and click **Y, Response**.
4. Select Yield Model and click **X, Predictor Formula**.
5. Click **OK**.

In this example, the values 510 and 540 for Reaction Temperature and 0.1 and 0.3 for Reaction Time were specified using the Coding column property. Alternatively, you can specify a reasonable range of values directly in the Factors outline.

6. Change the values of the parameter t1 to 25 and 50, and t3 to 30 and 35.

These new values represent a reasonable range of parameter values for the experimental situation. The default values were constructed based on the initial parameter values that were specified in the definition of the prediction formula. For information about constructing formulas, see the Formula Editor chapter in *Using JMP*.

Notice that the prior distribution, shown under Distribution, for each of t1 and t3 is set to Normal by default.

7. Change the number of runs to 12 in the Design Generation panel.

Figure 23.2 Completed Outlines for Reaction Kinetics Experiment

Nonlinear Design

Factors

Name	Role	Values
Reaction Temperature	Continuous	510 540
Reaction Time	Continuous	0.1 0.3

optional item

Parameters

Name	Distribution	Values
t1	Normal	25 50
t3	Normal	30 35

optional item

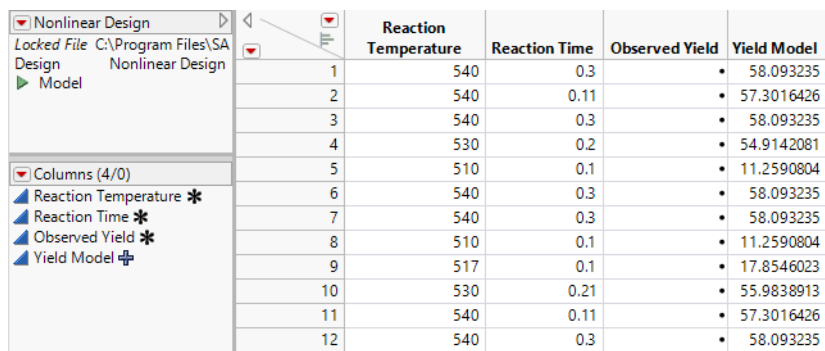
Design Generation

Number of Runs: 12

Make Design

8. Click **Make Design**.
9. Click **Make Table**.

Figure 23.3 Design Table



	Reaction Temperature	Reaction Time	Observed Yield	Yield Model
1	540	0.3		• 58.093235
2	540	0.11		• 57.3016426
3	540	0.3		• 58.093235
4	530	0.2		• 54.9142081
5	510	0.1		• 11.2590804
6	540	0.3		• 58.093235
7	540	0.3		• 58.093235
8	510	0.1		• 11.2590804
9	517	0.1		• 17.8546023
10	530	0.21		• 55.9838913
11	540	0.11		• 57.3016426
12	540	0.3		• 58.093235

Your design should be similar to the one shown in Figure 23.3. The runs might be in a different order, and the values for Reaction Temperature and Reaction Time, and consequently those computed for Yield Model, can be slightly different. Notice that values appear in the Yield Model column because the column contains the formula for the model. Also notice that the table contains a **Model** script that you can use to fit a nonlinear model to your observations.

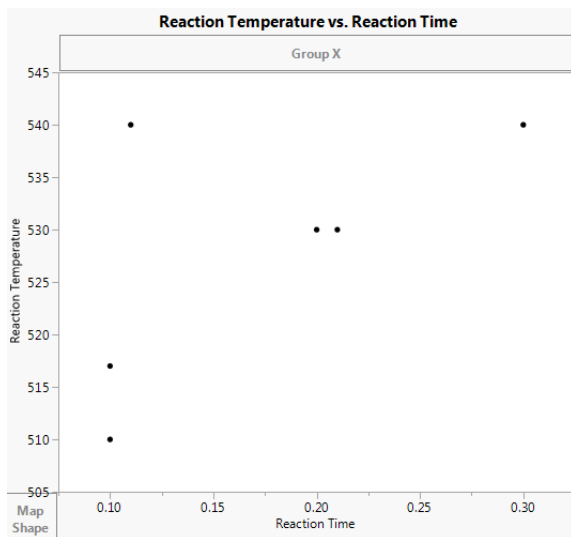
Now that you have created your design table, run your experiment, and record the responses in the Observed Yield column. The data table Reaction Kinetics.jmp, found in the Design Experiment folder contains observed results for the design.

Explore the Design

Before analyzing your results, construct a plot to see the design settings.

1. Select **Help > Sample Data Library** and open Design Experiment/Reaction Kinetics.jmp.
1. Select **Graph > Graph Builder**.
2. Drag and drop Reaction Temperature into the **Y** zone.
3. Drag and drop Reaction Time into the **X** zone.
4. Click the second icon above the graph to deselect the Smoother.

Figure 23.4 Design Settings

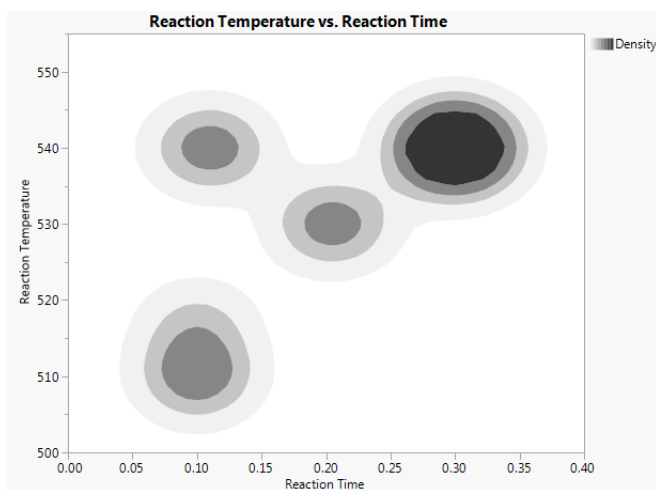


Notice that the points are located in three areas. There are no points at low temperature and high time (the lower right corner of the graph). Unlike orthogonal designs, nonlinear designs do not necessarily place design points at the corners of the design region. In this example, design points at low temperature and high time would be inefficient.

To see the density of design points in the remaining three corners, use the Contour tool.

5. Click  to turn on the Contour tool.
6. Click **Done**.

Figure 23.5 Design Settings with Density Contours



Notice that there are comparatively few points at low time and high temperature. From the design table, you can see that there are only two such points. Because of the model and the parameter specifications, the optimal design places more design points at high time and high temperature.

Analyze the Results

Now that you visually explored your design, analyze your results.

Note: Rather than conduct step 1 through step 4, you can run the **Model** script.

1. Select **Analyze > Specialized Modeling > Nonlinear**.
2. Select Observed Yield and click **Y, Response**.
3. Select Yield Model and click **X, Predictor Formula**.

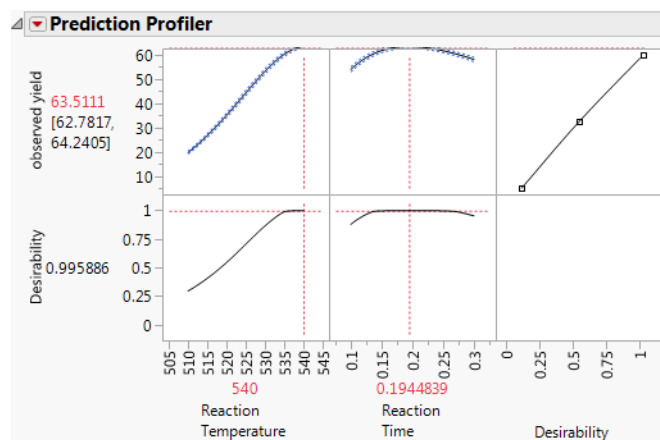
Notice that the model appears in the Options for fitting custom formulas panel.

4. Click **OK**.
5. Click **Go** in the Control Panel.

The iterative search for a solution proceeds until one of the Stop Limit values is reached. Then, the Solution and Correlation of Estimates reports appear.

6. Click the Nonlinear Fit red triangle and select **Profilers > Profiler**.
7. To maximize the yield, click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

Figure 23.6 Time and Temperature Settings for Maximum Yield



The estimated maximum yield is approximately 63.5% at a reaction temperature of 540 degrees Kelvin and a reaction time of 0.1945 minutes.

Augment a Design Using Prior Data

This example shows how to create a nonlinear design when you have prior data. In this example, the data pertain to a chemical reaction. You want to model the rate of uptake (velocity) of available organic substrate as a function of the concentration of that substrate. See Meyers (1986). You have already run an experiment, but you want to leverage your results to obtain more precise estimates of the parameters.

Obtain Prior Parameter Estimates

Use your existing experimental data to obtain better estimates of the parameter values.

1. Select **Help > Sample Data Library** and open Nonlinear Examples/Chemical Kinetics.jmp.
2. Click the plus sign next to **Model (x)** in the Columns panel. The formula editor opens.
3. The **Parameters** outline in the middle bottom of the formula editor shows the current values of the model parameters. The values ($b_1 = 1$ and $b_2 = 1$) are your initial estimates. They are used to compute the **Model (x)** values in the data table. For your next experiment, you want to replace these with better estimates.
4. Click **Cancel** to close the formula editor window.
5. Select **Analyze > Specialized Modeling > Nonlinear**.
6. Select **Velocity (y)** and click **Y, Response**.
7. Select **Model (x)** and click **X, Predictor Formula**.

Notice that the formula given by **Model (x)** appears in the Options for fitting custom formulas panel.

Figure 23.7 Nonlinear Analysis Launch Window

Fits nonlinear models using least squares or a custom loss function.

Select Columns
☒ 3 Columns
☒ Velocity (y)
☒ Concentration
☒ Model (x)

 Model Library

Cast Selected Columns into Roles

Y, Response	Velocity (y)
X, Predictor Formula	Model (x)
Group	optional
Weight	optional numeric
Freq	optional numeric
Loss	optional numeric
By	optional

The X Predictor column either has a formula with parameters, or is an independent variable to use with a builtin model.

Options for fitting custom formulas
 Predictor

Parameter({b1 = 1, b2 = 1}, b1 / (b2 + :Concentration))

 Loss

☐ Numeric Derivatives Only
☐ Expand Intermediate Formulas

Action

8. Click **OK**.

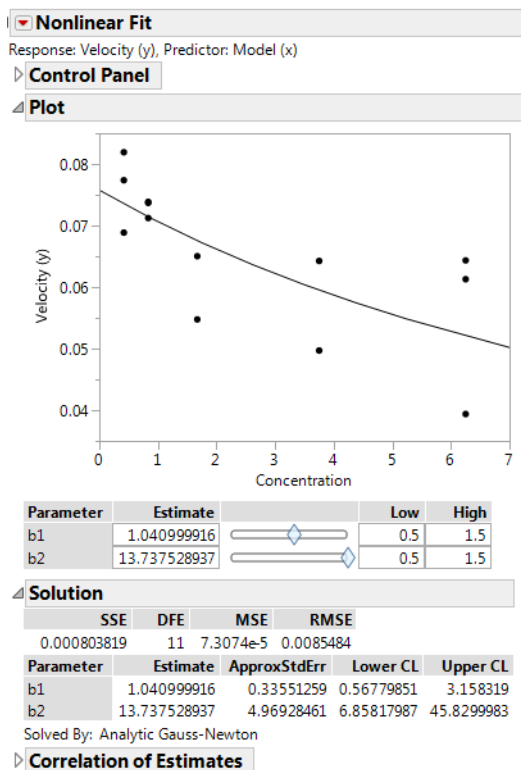
9. In the Control Panel, click **Go**.

The iterative search for a solution proceeds until one of the Stop Limit values is reached. Then, the Solution and Correlation of Estimates reports appear. Also, an option appears in the Control Panel enabling you to add confidence limits to the Solution report.

10. In the Control Panel, click **Confidence Limits**.

Confidence intervals for the parameters b1 and b2 appear in the Solution report.

Figure 23.8 Nonlinear Fit Results



The Lower CL and Upper CL values for b1 and b2 define ranges of values for these parameters. Next, use these intervals to define a range for the prior values in your augmented nonlinear design.

Note: Do not close the Nonlinear Fit report because these results are needed in the next steps.

Augment the Design

Now, create a design to estimate the nonlinear parameters more precisely.

1. With the Chemical Kinetics.jmp data table active, select **DOE > Special Purpose > Nonlinear Design**.
2. Select Velocity (y) and click **Y, Response**.
3. Select Model (x) and click **X, Predictor Formula**.
4. Click **OK**.

Figure 23.9 Nonlinear Design Outlines for Factors and Parameters

Nonlinear Design

Factors

Name	Role	Values	
▲ Concentration <i>optional item</i>	Continuous	0.417	6.25

Parameters

Name	Distribution	Values	
b1	Normal	0.5	1.5
b2	Normal	0.5	1.5

optional item

Design Generation

Enter Number of Runs (counting 13 included runs):

In the Chemical Kinetics.jmp data, the values for Concentration range from 0.417 to 6.25. Therefore, these values initially appear as the low and high values in the Factors outline. You want to change these values to encompass a broader interval.

- Click 0.417 and type 0.1. Click 6.25 and type 7.
- Leave the prior Distribution for each parameter set to Normal.

The range of Values for the parameters reflects the uncertainty of your knowledge about them. You could specify a range that you think covers 95% of possible parameter values. The confidence limits from the Nonlinear Fit report shown in Figure 23.8 provide such a range. Replace the Values for the parameters in the Parameters outline with the confidence limits, rounding to three decimal places.

7. In the DOE Nonlinear Design window, enter these values into the Parameters for b1 and b2:
 - b1: 0.568 and 3.158
 - bb2: 6.858 and 45.830

Figure 23.10 Updated Values for Factor and Parameters

Nonlinear Design

Factors

Name	Role	Values
Concentration <i>optional item</i>	Continuous	0.417 6.25

Parameters

Name	Distribution	Values
b1	Normal	0.568 3.158
b2	Normal	6.858 45.83

Design Generation

Enter Number of Runs (counting 13 included runs): 17

Make Design

8. Enter 40 for the **Number of Runs** in the Design Generation panel.
9. Click **Make Design**.

The Design outline opens, showing the Concentration and Velocity (y) values for the original 13 runs and new Concentration settings for the additional 27 runs.

10. Click **Make Table**.

This creates a new JMP design table that contains the settings and results for the original 13-run design and settings for 27 new runs. Instead of creating a new data table, you can add the new runs to your existing data table by clicking **Augment Table** instead of **Make Table**.

The new runs reflect the broader interval of Concentration values and the range of values for b1 and b2 obtained from the original experiment, which are used to define the prior distribution. Both should lead to more precise estimates of b1 and b2.

Create a Design for a Binomial Response

In some applications, the measurement of interest is a pass/fail (binomial) measurement. In this example, you will construct a nonlinear design for a binomial response with one design factor. For more information on designs in nonlinear settings and the binomial case in particular, see Gotwalt et al. (2009).

You plan to model the probability of success for your binomial response as a function of a single factor x , using a generalized linear model with a logistic link function

$$Y(x) = \frac{1}{1 + e^{-(\beta_0 + \beta_1 x)}}$$

and the variance of Y as the weight.

This model is nonlinear in the unknown parameters β_0 , and β_1 . Use the nonlinear designer to plan an experimental design. Your goal is to model the effect of x on your binomial response Y .

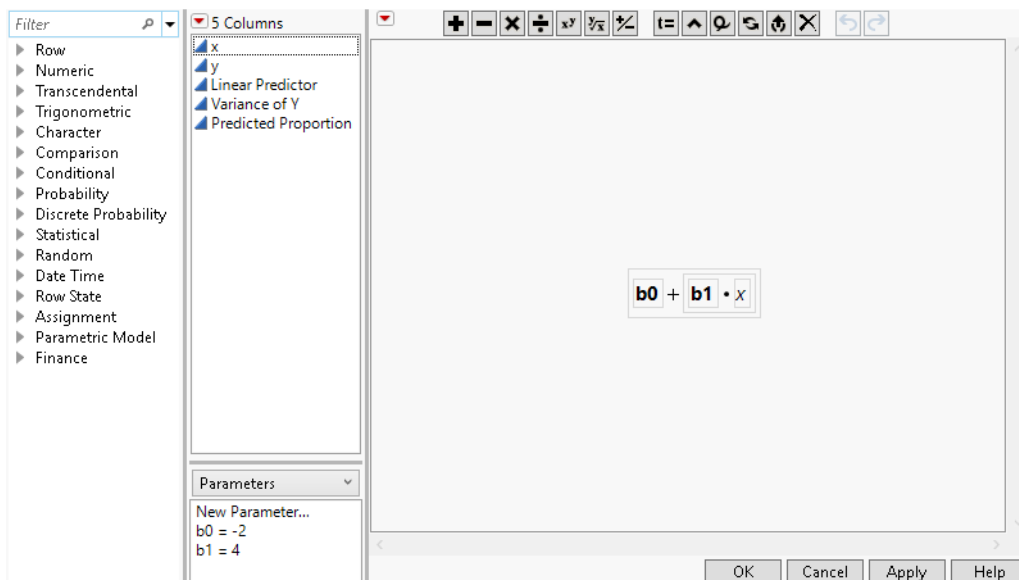
To generate the desired non-linear design you must have a data table containing columns for the predictor, a column containing a formula for the nonlinear response model that you are fitting, and a column with a formula for your weights.

Data Table for Design Construction

The One Factor Logistic Design.jmp data table, found in the Design Experiment folder, contains the following:

- Column x for the predictor. The Coding property defined for this column sets the low value to 0 and the high value to 1.
- Column Y , for the response. This is the column for collected responses (0/1) from your testing.
- Column Linear Predictor that contains a formula for the linear portion of the link function. To view the formula, click the plus sign to the right of Linear Predictor in the Columns panel. The model formula includes the initial estimates for the parameters **b0** and **b1**. Initial parameter values are set when you define a formula. These values are shown in the formula element panel in the lower center of the formula editor window.

Figure 23.11 Linear Predictor Formula with Initial Parameter Estimates



- Column Variance of Y that contains the formula for the variance of the response based on the assumed logistic model. This is $p(1-p)$ where p is the logistic function. This column is used as the weight column.
- Column Predicted Proportion that contains the logistic link function.

Create the Design

1. Select **Help > Sample Data Library** and open Design Experiment/One Factor Logistic Design.jmp.
2. Select **DOE > Special Purpose > Nonlinear Design**.
3. Select Y and click **Y, Response**.
4. Select Linear Predictor and click **X, Linear Predictor**.
5. Select Variance of Y and click **Weight**.
6. Click **OK**.

Figure 23.12 Nonlinear Design Window

Nonlinear Design

Factors

Name	Role	Values
x <i>optional item</i>	Continuous	0 1

Parameters

Name	Distribution	Values
b0	Normal	-1 -3
b1 <i>optional item</i>	Normal	2 6

Design Generation

Number of Runs: 10

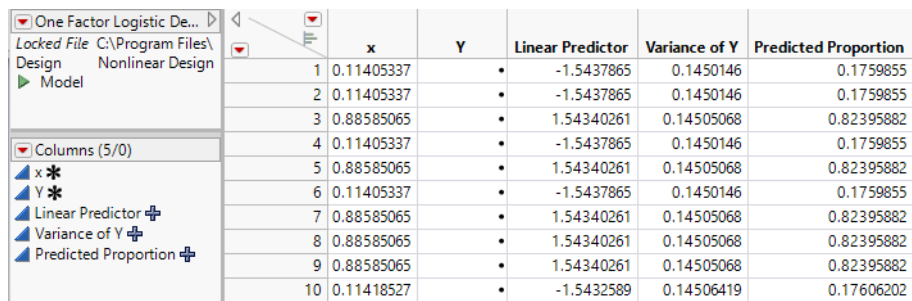
Make Design

The Factors outline shows the factor x with the specified values of 0 and 1. The Parameters outline shows the two model parameters with each prior distribution set to a normal distribution. JMP computes default Values based on your initial settings for the parameter values. You can change the assumed distribution and ranges for the parameters in this outline. Leave all settings as they appear. The default number of runs is 10.

7. Click the Nonlinear Design red triangle and select **Advanced Options > N Monte Carlo Spheres**. Set the number of spheres to 0 to obtain a locally optimal design. Click **OK**.
8. Click **Make Design**.
9. Click **Augment Table**.

This adds the 10 run design to One Factor Logistic Design.jmp (Figure 23.13). Your design table will be different because the optimization algorithm has a random component.

Figure 23.13 Augmentation of One Factor Logistic Design.jmp



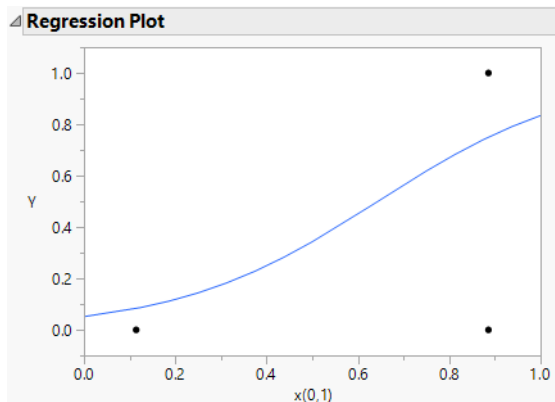
	x	Y	Linear Predictor	Variance of Y	Predicted Proportion
1	0.11405337	•	-1.5437865	0.1450146	0.1759855
2	0.11405337	•	-1.5437865	0.1450146	0.1759855
3	0.88585065	•	1.54340261	0.14505068	0.82395882
4	0.11405337	•	-1.5437865	0.1450146	0.1759855
5	0.88585065	•	1.54340261	0.14505068	0.82395882
6	0.11405337	•	-1.5437865	0.1450146	0.1759855
7	0.88585065	•	1.54340261	0.14505068	0.82395882
8	0.88585065	•	1.54340261	0.14505068	0.82395882
9	0.88585065	•	1.54340261	0.14505068	0.82395882
10	0.11418527	•	-1.5432589	0.14506419	0.17606202

The model saved to the data table is a nonlinear model to fit the “Linear Predictor” used in the creation of the design. To model the effect of x on Y use a generalized linear model (GLM) with a logit link function. Simulate values for Y to explore the use of a GLM to evaluate your design.

Simulate Responses and Analyze

1. Right-click the Y column heading and select formula.
2. From the Random function list select Random Binomial.
3. Enter 1 for n and Predicted Proportion for p . Click **OK**
4. Select **Analyze > Fit Model**
5. Select Y and click **Y**.
6. Select x and click **Add**.
7. Select Generalized Linear Model as the personality.
8. Select Binomial as the Distribution.
9. Select Logit as the Link Function.
10. Check Firth Bias-Adjusted Estimates.
11. Click **Run**.

Figure 23.14 Regression Plot for Binomial DOE



Nonlinear Design Launch Window

To use the Nonlinear Design platform, you must have an existing data table that contains the following:

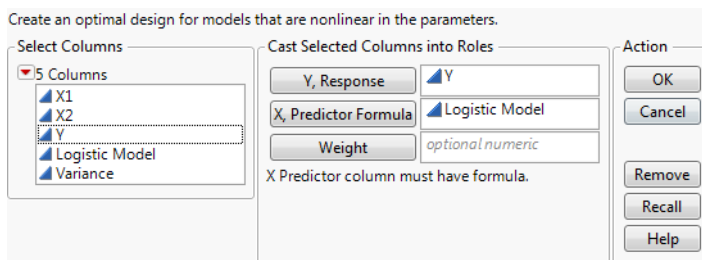
- A column for the response.
- A column for each factor.
- A column that contains a formula showing the relationship between the factors and the response. This formula must include the unknown parameters.

For information about formulas, see the Formula Editor chapter in *Using JMP*.

The table can contain values for the predictors and response. If it does, the design that you construct augments the design that is implicit in the table. There can be no row containing missing predictor values.

With your starting data table active, select **DOE > Special Purpose > Nonlinear Design**.

Figure 23.15 Nonlinear Launch Window

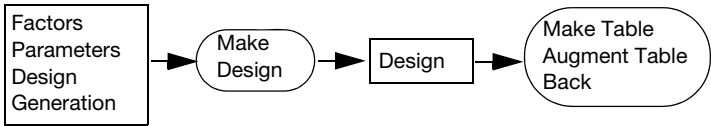


- Y, Response** The numeric column for response values.
- X, Predictor Formula** The numeric column that contains the formula for the nonlinear model. This formula must contain parameters.
- Weight** (Optional) A numeric column that assigns weights to the observations.

Nonlinear Design Window

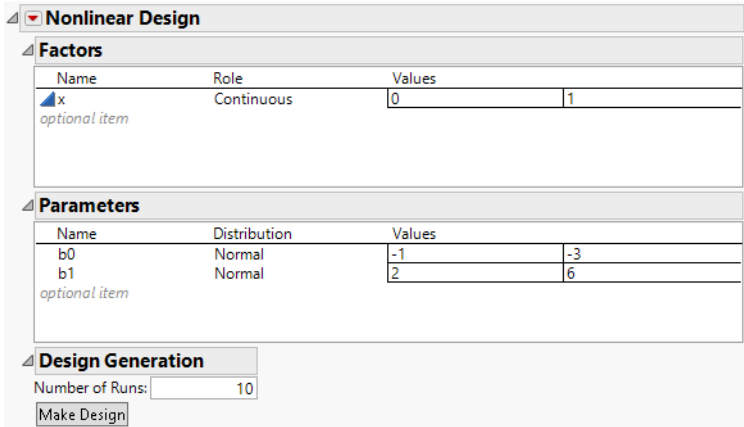
The Nonlinear Design window updates as you work through the design steps. The outlines, separated by buttons that update the outlines, follow the flow in Figure 23.16.

Figure 23.16 Nonlinear Design Flow



The initial design window shows the Factors, Parameters, and Design Generation outlines.

Figure 23.17 Initial Design Window



Factors

The column names used in the model formula are automatically inserted in the Name column of the Factors outline. Each factor’s role is set to Continuous.

For each factor, the Values are initially set to -1 and 1. Or, if you have defined a value using the Coding column property, those values are used instead. You can change these values in the factors outline.

Parameters

The parameter names used in the model formula are automatically inserted in the Name column of the Parameters outline.

For each parameter, the Values are initially set to a symmetric interval around the initial value specified in the parameter definition. This interval is obtained by taking the initial value's distance to 0, and constructing an interval of this width around the initial value. These values are used in defining prior distributions for the model parameters.

Note: Adjust the Values for the parameters in conjunction with your choice of distribution to reflect your uncertainty about the model parameters.

Four families of prior distributions are listed under Distribution. The Values that you specify for the parameters determine which member of the family of prior distributions that is used. Denote the low value by *low* and the high value by *high*. Then the distributions are determined as follows:

- Uniform: The distribution is uniform on the interval (*low*, *high*).
- Normal, Lognormal, Exponential: The distribution is the one where *low* is the 0.025 quantile and where *high* is the 0.975 quantile.

Design Generation

JMP provides a suggested number of runs, determined as follows:

- If you are not augmenting a data table, the number of runs is four times the number of parameters plus two.
- If you are augmenting a data table, the number of runs is the number of runs in the data table plus two times the number of parameters.

Note: If you are augmenting a design, the number of runs that JMP suggests or that you specify includes those runs corresponding to observations in your data table. Adjust the number of runs appropriately.

Design

When you click Make Design, JMP constructs the design and adds a Design outline to the Nonlinear Design window. In the Design outline, you can review the factor level settings.

Make Table or Augment Table

The Make Table button creates a new design table. If your original table included existing runs, the new table also includes the existing runs. The **Augment Table** button adds the new runs to your existing table.

Nonlinear Design Options

The red triangle menu in the Nonlinear Design platform contains these options:

Save Responses Saves the information in the Responses panel to a new data table. You can then quickly load the responses and their associated information into most DOE windows. This option is helpful if you anticipate reusing the responses.

Load Responses Not available.

Save Factors Saves the information in the Factors panel to a new data table. Each factor's column contains its levels. Other information is stored as column properties. You can then quickly load the factors and their associated information into most DOE windows.

Load Factors Not available.

Save Constraints Not available.

Load Constraints Not available.

Simulate Responses Adds response values to the design table. Select this option before you click Make Table. Then, in the resulting design table, the response columns contain simulated values.

Note: To set a preference to always simulate responses, select **File > Preferences > Platforms > DOE** and select **Simulate Responses**.

Number of Starts Sets the number of times that a nonlinear design is created using the quadrature method. Among the designs created, the platform selects the design that maximizes the optimality criterion.

Advanced Options > Number of Monte Carlo Samples Sets the number of octahedra per sphere used in computing the optimality criterion. The default value is one octahedron. See [“Radial-Spherical Integration of the Optimality Criterion”](#) on page 700.

Advanced Options > N Monte Carlo Spheres Sets the number of nonzero radius values used in computing the optimality criterion. The default is two. See [“Radial-Spherical Integration of the Optimality Criterion”](#) on page 700.

Note: If N Monte Carlo Spheres (the number of radii) is set to zero, then only the center point is used in the calculations. This gives a local design that is optimal for the initial values of the parameters. For some situations, this is adequate.

Statistical Details for Nonlinear Designs

- [“Nonlinear Models”](#)
- [“Radial-Spherical Integration of the Optimality Criterion”](#)
- [“Finding the Optimal Design”](#)

Nonlinear Models

Denote the vector of n responses by $\mathbf{Y} = (Y_1, Y_2, \dots, Y_n)'$. A nonlinear model is defined by the following properties:

- The Y_i are independent and identically distributed with an exponential family distribution.
- The expected value of each Y_i given a vector of predictor values $\mathbf{x}_i$ is a nonlinear function of parameters, θ . Denote this function as follows:

$$E(Y_i(\mathbf{x})) = f(\theta, \mathbf{x}_i)$$

- Each Y_i is expressed as follows:

$$Y_i = f(\theta, \mathbf{x}_i) + \varepsilon_i$$

- The vector of errors, $\varepsilon = (\varepsilon_1, \varepsilon_2, \dots, \varepsilon_n)'$ has mean $\mathbf{0}$ and covariance matrix $\sigma^2 \mathbf{I}$, where $\mathbf{I}$ is the $n \times n$ identity matrix.

Denote the matrix of first partial derivatives of the function f with respect to the parameters θ by X . Under general conditions, the least squares estimator of θ is asymptotically unbiased, with asymptotic covariance matrix given as follows:

$$\text{Cov}(\hat{\theta}) = \sigma^2 (X'X)^{-1}$$

For the proof of this result, see Wu (1981) and Jennrich (1969).

Radial-Spherical Integration of the Optimality Criterion

The optimality criterion is the expectation of the logarithm of the determinant of the information matrix with respect to the prior distribution. Consequently, finding an optimal nonlinear design requires maximizing the integral of the log of the determinant of the Fisher information matrix with respect to the prior distribution of the parameters. This integral must be calculated numerically. The approach used in the Nonlinear Design platform is based on Gotwalt et al. (2009).

For normal distribution priors, the integral is reparameterized into a radial direction and a number of angular directions equal to the number of parameters minus one. The radial part of the integral is computed using Radau-Gauss-Laguerre quadrature with an evaluation at radius = 0. This is done by constructing a certain number of hyperoctahedra and randomly rotating each of them.

If the prior distribution is not normal, then the integral is reparameterized so that the new parameters have a normal distribution. Then the radial-spherical integration method is applied.

Note: If the prior distribution for the parameters does not lend itself to a solution and the process fails, a message is added to the window that the Fisher information matrix is singular in a region of the parameter space. When this occurs, consider changing the prior distribution or the ranges of the parameters.

Finding the Optimal Design

The method used to find an optimal design is similar to the coordinate exchange algorithm described in Meyer and Nachtsheim (1995). For more information about how the nonlinear optimal design is obtained, see Gotwalt et al. (2009). The general approach proceeds as follows:

- Random designs are tested until a nonsingular starting design is found.
- Iterations are conducted, where each iteration consists of a pass through all the runs.
- For each run, factors are optimized one at a time.

- The objective function is the Bayesian D-optimality criterion. This is the expectation of the logarithm of the determinant of the information matrix with respect to the prior distribution.
- Iterations terminate once the change in the objective function is small.

Chapter 24

Balanced Incomplete Block Designs

A Balanced Incomplete Block Design (BIBD) can be used when it is not possible to include all treatments or factor combinations in every block. Use BIBDs for testing a treatments in b blocks when only k treatments can be run in any one block (and $k < a$). For example, you might have a situation where there are limitations on the number of treatments that can be tested at the same time. The BIBD design is constructed such that each treatment and each pair of treatments occurs together in blocks an equal number of times. This makes the design *balanced*. Not including all treatments in every block makes the design *incomplete*.

Use the Balanced Incomplete Block Platform to construct BIBDs. In addition, you can use the platform to construct complete block designs where each treatment occurs in every block.

Figure 24.1 BIBD for 4 Treatments in 4 Blocks of Size 3

Design			
Block			
1	L2	L3	L4
2	L3	L4	L1
3	L4	L1	L2
4	L1	L2	L3

Contents

Example of Balanced Incomplete Block Designs	705
Balanced Incomplete Block Design Launch Window	707
Balanced Incomplete Block Design Window	708
Specify Treatments, Blocks, and Make Design	708
Verify Design and Make Table	709
Balanced Incomplete Block Design Options	709
Statistical Details for Balanced Incomplete Block Designs	710

Example of Balanced Incomplete Block Designs

In this example, you want to test three new product formulations in your plant, and you want to block by day. There is a limit of two formulations that can be run per day. In addition, you need to finish your experiment in one work week. You need a design with at most five blocks.

1. Select **DOE > Special Purpose > Balanced Incomplete Block Design**.
2. Set the Treatment Size to 3.

Tip: Treatment size can refer to the number of levels of a single factor or combinations of levels of two or more factors.

3. Select 2 for the **Block Size**.
4. Select 3 for **Allowable Blocks**.

You can select 3 or 6 blocks. In this example, each block is a day and you can have no more than 5 days or 5 blocks. The number of runs is 6 because you have 3 blocks of size 2.

Note: You can enter a Treatment Name, Block Name, and Treatment Labels to customize your design table.

5. Click **Make Design**.

Figure 24.2 BIBD for 3 Treatments in 3 Blocks of Size 2

Balanced Incomplete Block Design

Design Options

Treatments

Treatment Name:

Treatment Size:

Treatment Labels:

Blocks

Block Name:

Block Size:

Allowable Blocks

☐ Include Block Multiples

Make Design

Block Design

Run	Block	Treatment
1	1	L2
2	1	L3
3	2	L3
4	2	L1
5	3	L1
6	3	L2

Make Table ☐ Include order column

Design

Incidence Matrix

The design has six runs in three blocks with two treatments in each block as seen in the Block Design outline. For additional design details open the Design and Incidence Matrix outlines. See [“Verify Design and Make Table”](#) on page 709.

Initially, you had to select a design with 3 or 6 blocks. You selected 3 blocks because of time constraints in your plant. However, suppose you were allowed to run your experiment for 6 days. You could then update the design to use 6 blocks.

6. Select 6 in the **Allowable Blocks** menu.
7. Click **Make Design** to update the design.
8. Open the **Incidence Matrix** outline.

Figure 24.3 BIBD Design Details for 3 Treatments in 6 Blocks of Size 2

Incidence Matrix			
Block	Treatment		
	L1	L2	L3
1	1	0	1
2	1	0	1
3	0	1	1
4	0	1	1
5	1	1	0
6	1	1	0

Pairwise Treatment Frequencies			
Treatment	L1	L2	L3
L1	4	2	2
L2		4	2
L3			4

Positional Frequencies		
Treatment	Position	
	1	2
L1	4	0
L2	2	2
L3	0	4

Now the design has 6 blocks. The incidence matrix shows the treatments that are in each block. The Pairwise Treatment Frequencies shows that each treatment appears four times (diagonal entries) and that each treatment pair occurs twice in the design (off diagonal entries). The blocks each have two positions. The Positional Frequencies show that treatment L1 always occurs in position 1, L2 is split between the two positions, and L3 occurs in position 2.

Balanced Incomplete Block Design Launch Window

Launch the Balanced Incomplete Block Design platform by selecting **DOE > Special Purpose > Balanced Incomplete Block Design**.

Figure 24.4 Balanced Incomplete Block Design Launch Window

Balanced Incomplete Block Design

Design Options

Treatments

Treatment Name

Treatment

Treatment Size

2

Treatment Labels:

L1

L2

Blocks

Block Name

Block

Block Size

2

Allowable Blocks

2

☐ Include Block Multiples

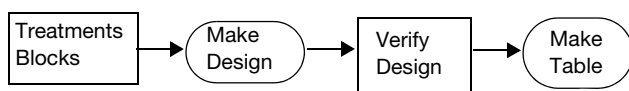
Make Design

For more information about this window, see [“Specify Treatments, Blocks, and Make Design”](#) on page 708.

Balanced Incomplete Block Design Window

The BIBD window updates as you work through the design steps. The outlines that appear, separated by buttons that update the outlines, follow the flow in Figure 24.5.

Figure 24.5 Balanced Incomplete Block Design Flow



The initial design window shows the Design Options outline.

Specify Treatments, Blocks, and Make Design

The following options are in the Design Options outline:

Treatment Name The name of the treatment factor.

Treatment Size The number of treatments.

Tip: Treatments can be factor-level combinations.

Treatment Labels Labels for the treatments. The default labeling is L1 to L a where a is the total number of treatments.

Block Name The name of the blocking factor.

Block Size The number of treatments that can be run in any one block.

Note: Values for Block Size are determined by the number of treatments. If a value is not in the list, it is because the BIBD platform cannot create a design with that block size for the given number of treatments.

Allowable Blocks Allowable number of blocks in the design based on the number of treatments and the block size.

Include Block Multiples Allows for the selection of a design with the number of blocks equal to a multiple (that you specify) of the initial allowable block sizes. The additional (multiple) blocks are cyclic permutations of the original design.

Specify Multiplier (Available when Include Block Multiples is selected.) Multiplier that you want to apply to the initial allowable block sizes.

Make Design Generates the design.

Verify Design and Make Table

Block Design Shows the BIBD in a traditional design layout.

Make Table Creates the BIBD data table.

Include order column Includes a column in the design table for the order of runs within each block.

Design Shows the BIBD in a block layout where each row represents a block and the treatments in that block.

Incidence Matrix Shows the treatments that occur in each block. Each row of the matrix corresponds to a block in the design with a 1 if the treatment is in the block and 0 otherwise.

Pairwise Treatment Frequencies Shows the number of times each treatment occurs in the same block with every other treatment on the off-diagonal entries. The diagonal entries are the total number of times each treatment occurs.

Positional Frequencies Shows the number of times each treatment occurs in each block position.

Balanced Incomplete Block Design Options

The red triangle menu in the Balanced Incomplete Block Design platform contains these options:

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.

- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click **Apply**, a formula for the simulated response values is saved in a new column called **<Y> Simulated**, where Y is the name of the response. Clicking **Apply** again updates the formula and values in **<Y> Simulated**.

For additional details, see [“Simulate Responses”](#) on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Save Script to Script Window Creates the script for the design that you specified in the Balanced Incomplete Block window and places it in an open script window.

Statistical Details for Balanced Incomplete Block Designs

The total number of observations in a BIBD is determined by the following equation:

$$N = ar = bk$$

where

a is the number of treatments

r is the number of times each treatment occurs in the design

b is the number of blocks

k is the number of treatments in each block

The number of blocks in which each pair of treatments occurs is given by the following:

$$\lambda = \frac{r(k-1)}{a-1}$$

where λ must be an integer.

These conditions are necessary for a BIBD to exist, but not sufficient. The BIBD platform presents values for which it can generate a BIBD.

When the Include Block Multiples is selected, the original BIBD is copied the number of times specified. The additional (multiple) blocks are cyclic permutations of the original design. The positions of treatments in the copies differs from the original.

For more information about BIBDs, see SAS Institute Inc. ([2013](#)).

Group Orthogonal Supersaturated Designs

Create Group Orthogonal Supersaturated Designs

Group orthogonal supersaturated designs are a special case of screening designs. They are appropriate in early stage work when the number of factors to be investigated is larger than the number of feasible runs. A supersaturated design is a design with fewer observations than model parameters. That is, the design cannot estimate all main effects simultaneously. This makes detection of significant effects difficult. A group orthogonal supersaturated design is a special class of two-level supersaturated designs with properties that are desirable for model selection.

Contents

[Overview of Group Orthogonal Supersaturated Designs](#) 715

[Example of Group Orthogonal Supersaturated Designs](#) 715

[Launch the Group Orthogonal Supersaturated Designs Platform](#) 717

[Group Orthogonal Supersaturated Design Window](#)..... 717

[Specify Factors, Structure, and Make Design](#) 718

[Verify Design and Make Table](#) 718

[Group Orthogonal Supersaturated Design Options](#) 719

[Statistical Details for Group Orthogonal Supersaturated Designs](#) 720

Overview of Group Orthogonal Supersaturated Designs

Group orthogonal supersaturated designs (GOSSD) are constructed to have a specific structure to facilitate model selection. The design factors are partitioned into groups such that factors within a group are correlated with other factors within the same group yet factors are orthogonal to factors in other groups. One group consists of three or more “fake” factors. These fake factors are not assigned to design factors. They are used only in the analysis of the design. The fake factors provide an unbiased estimate of the variance assuming that second-order and higher-order effects are negligible.

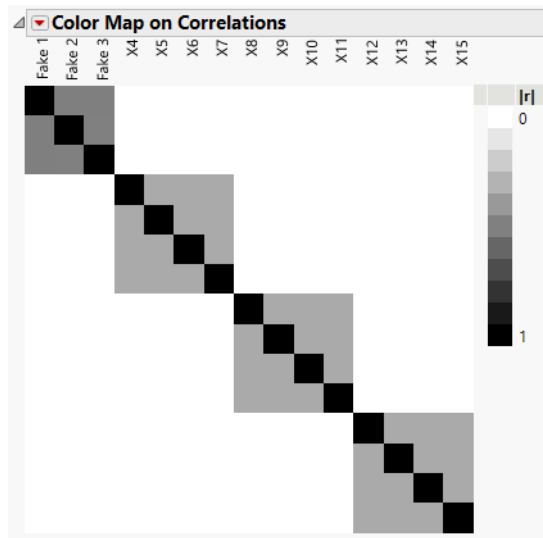
Practical considerations for setting up a GOSSD include:

- More groups are better than fewer groups.
- Factors that are thought to be active should be placed in separate groups.
- Choose the levels of factors so that anticipated coefficients are positive.

Example of Group Orthogonal Supersaturated Designs

In this example, you want to design an experiment to study 12 factors in at most 12 runs. You want a design with at least 16 parameters. One parameter is needed for the intercept and at least three fake factors are desired.

1. Select **DOE > Special Purpose > Group Orthogonal Supersaturated > Group Orthogonal Supersaturated Design**.
2. Select **Number of Runs**.
3. Enter 12 for the **Number of Runs** and click away from the text box.
4. Select the second option in the Structure window with 16 parameters in 4 groups of size 4.
5. Click on the X1 factor name and change the name to Fake 1.
6. Repeat for X2 and X3 to name them as Fake 2 and Fake 3, respectively.
7. Click **Make Design**.

Figure 25.1 GOSSD Color Map on Correlations

The structure of the design is evident in the color map. Factors within a group are correlated. Factors in different groups are not correlated, which indicates that they are orthogonal.

Launch the Group Orthogonal Supersaturated Designs Platform

Launch the Group Orthogonal Supersaturated Design platform by selecting **DOE > Special Purpose > Group Orthogonal Supersaturated > Group Orthogonal Supersaturated Design**.

Figure 25.2 Group Orthogonal Supersaturated Design Launch Window

Group Orthogonal Supersaturated Design

☒ Number of Runs
Must be a multiple of 2 (or preferably 4).

☐ Number of Factors
Add 1 for the Intercept.

Structure

Number of Groups	Group Size	Number of Parameters
2	4	8

Note: Select an option above.

Factors

Factor Name	Role	Lower	Upper
Intercept	Continuous	1	1
X1	Continuous	-1	1
X2	Continuous	-1	1
X3	Continuous	-1	1
X4	Continuous	-1	1
X5	Continuous	-1	1
X6	Continuous	-1	1

Group Structure

Make Design

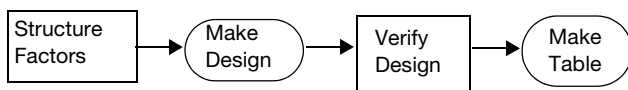
For more information about this window, see [“Specify Factors, Structure, and Make Design”](#) on page 718

For more information about the options in the Select Columns red triangle menu, see the Get Started chapter in *Using JMP*.

Group Orthogonal Supersaturated Design Window

The GOSSD launch window updates as you work through the design steps. The outlines that appear, separated by buttons that update the outlines, follow the flow in Figure 25.3.

Figure 25.3 Group Orthogonal Supersaturated Design Flow



Specify Factors, Structure, and Make Design

The following options are available in the Design Options outline.

Tip: Provide either the number of runs or the number of factors to define possible design structures.

Number of Runs Defines the number of design runs.

Number of Factors Defines the number of factors. Include a factor for the intercept and three to five fake factors.

Structure Provides a list of possible structures for your design. The columns in this list are dependent on whether you set the number of runs or the number of factors.

Number of Groups The number of groups in the design. The minimum is 2.

Group Size The number of factors in each group.

Number of Runs (Available if you set the number of factors) The number of runs in the design.

Number of Parameters (If you set the number of runs.) The number of parameters in the design. The parameters include one intercept, your design factors, and the remaining parameters are assigned to fake factors.

Factors

Factor Name The name for the intercept is fixed. The remaining factors can be renamed.

Role Enables you to define a factor as continuous or categorical.

Lower Enables you to set the lower limit of each factor.

Upper Enables you to set the upper limit of each factor.

Group Structure Shows factors that appear in each group.

Swap Enables you to swap factors between groups.

Make Design Generates the design.

Verify Design and Make Table

Design Displays the design matrix.

Make Table Creates the GOSSD data table. The design table includes the following scripts:

Model (Executes only if the Response column is populated.) Runs the **DOE > Special Purpose > Group Orthogonal Supersaturated > Fit Group Orthogonal Supersaturated** platform.

Doe Dialog Re-creates the GOSSD Design window that you used to generate the design table.

GOSSD Structure Shows the structure of the design (number of runs, number of parameters, number of groups, group size).

Multivariate Runs the **Analyze > Multivariate** platform for the design factors.

Model Matrix Contains the design matrix.

DOE Simulate (Available when you simulate responses.) Launches the simulate window.

Design Evaluation

Correlation Matrix Shows the matrix of correlations between effects.

Color Map on Correlations Shows the absolute correlations between effects on a plot using an intensity scale. Includes a red triangle menu that contains an option to save the correlations to a data table.

Group Orthogonal Supersaturated Design Options

The red triangle menu in the Group Orthogonal Supersaturated Design platform contains the following options:

Simulate Responses Adds response values and a column containing a simulation formula to the design table. Select this option before you click Make Table.

When you click Make Table, the following occur:

- A set of simulated response values is added to each response column.
- For each response, a new column that contains a simulation model formula is added to the design table. The formula and values are based on the model that is specified in the design window.
- A Model window appears where you can set the values of coefficients for model effects and specify one of three distributions: Normal, Binomial, or Poisson.

- A script called **DOE Simulate** is saved to the design table. This script reopens the Model window, enabling you to re-simulate values or to make changes to the simulated response distribution.

Make selections in the Model window to control the distribution of simulated response values. When you click **Apply**, a formula for the simulated response values is saved in a new column called <Y> Simulated, where Y is the name of the response. Clicking **Apply** again updates the formula and values in <Y> Simulated.

For additional details, see “[Simulate Responses](#)” on page 110 in the “Custom Designs” chapter.

Note:  You can use Simulate Responses to conduct simulation analyses using the JMP Pro Simulate feature. For more information and DOE examples, see the Simulate chapter in *Basic Analysis*.

Save Script to Script Window Creates the script for the design that you specified in the Group Orthogonal Supersaturated Design window and places it in an open script window.

Statistical Details for Group Orthogonal Supersaturated Designs

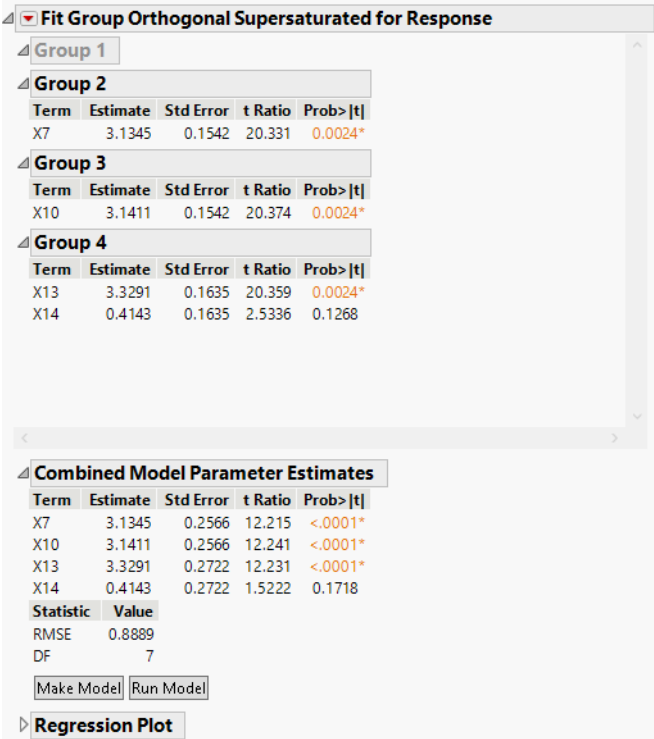
For statistical details of design construction see Jones, et.al. ([2019](#)).

The Fit Group Orthogonal Supersaturated Platform

Analyze Data from Group Orthogonal Supersaturated Designs

Use the Fit Group Orthogonal Supersaturated platform to analyze group orthogonal supersaturated designs. This analysis technique takes advantage of the group orthogonal structure of group orthogonal supersaturated designs.

Figure 26.1 Fit Group Orthogonal Supersaturated Results



Contents

[Overview of the Fit Group Orthogonal Supersaturated Platform](#)..... 723

[Example of Fit Group Orthogonal Supersaturated](#) 723

[Launch the Fit Group Orthogonal Supersaturated Platform](#) 726

[The Fit Group Orthogonal Supersaturated Report](#) 727

[Fit Group Orthogonal Supersaturated Platform Options](#)..... 728

Overview of the Fit Group Orthogonal Supersaturated Platform

A group orthogonal supersaturated design (GOSSD) is constructed to facilitate the identification of active effects. These designs group factors such that within a group, factors are correlated but between groups, factors are orthogonal. This structure leads to two analysis features:

- Fake, or dummy, factors assigned to the first group are used to estimate variance, assuming that the second- or higher-order effects are negligible.
- Between-group orthogonality enables a two-step model selection procedure.

The Fit Group Orthogonal Supersaturated Platform performs a two-step model selection procedure, enables you to view a regression plot of all factors, and then provides a model for the active effects with the option to examine this model with the Fit Model platform.

Example of Fit Group Orthogonal Supersaturated

In this example, you want to design an experiment to study 12 factors in at most 12 runs. For a GOSSD, you need a design with at least 16 parameters. One parameter is needed for the intercept, at least three parameters are needed for the fake factors, and 12 parameters are needed for the design factors. Use simulated responses to fit a model.

Tip: You are not limited to three fake factors. More can be used.

Generate the Design

1. Select **DOE > Special Purpose > Group Orthogonal Supersaturated > Group Orthogonal Supersaturated Design**.
2. Select **Number of Runs**.
3. Enter 12 for the **Number of Runs** and click away from the text box.
4. Select the second option in the Structure section with 16 parameters in 4 groups of size 4.
5. Click on the X1 factor name and change the name to Fake 1.
6. Repeat for X2 and X3 to name them as Fake 2 and Fake 3, respectively.
7. Click the Group Orthogonal Supersaturated Design red triangle and select **Simulate Responses**.
8. Click **Make Design**.
9. Click **Make Table**.

Figure 26.2 Simulate Responses Window

Simulate Responses

▸ **Group 1**

▾ **Group 2**

Effects	Response
X4	0
X5	0
X6	0
X7	3.09

▾ **Group 3**

Effects	Response
X8	0
X9	0
X10	3.18
X11	0

▾ **Group 4**

Effects	Response
X12	0
X13	3.22
X14	0
X15	0

Reset Coefficients

▾ **Distribution**

☒ Normal Error σ : 1

☐ Binomial

☐ Poisson

Apply

The Simulate Responses window shows the settings that were used to generate the design table. Note that the outline box for Group 1 is closed. Group 1 contains the “fake” factors. These are all set to 0 as they are assumed to have no impact on the model. Your values might differ from those shown in Figure 26.2. This simulation has one active effect per group.

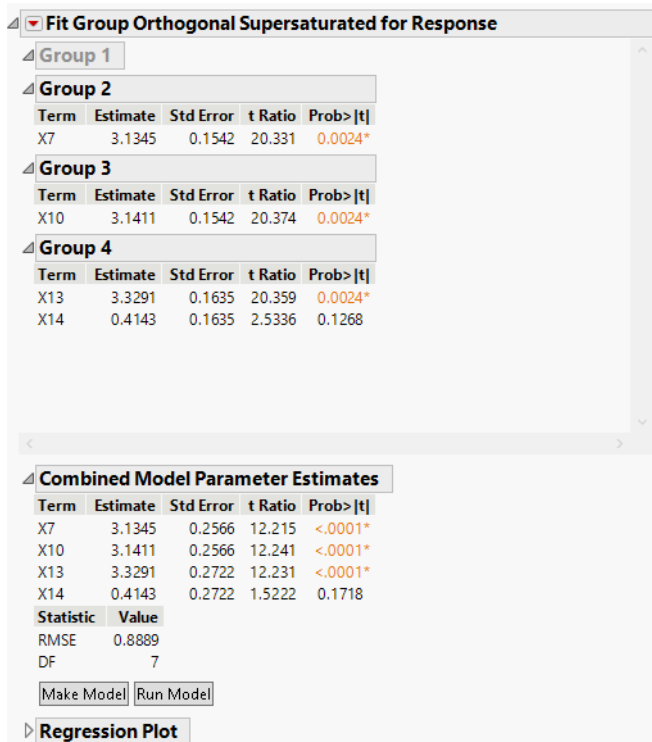
Tip: Use the Simulate Responses Window to enter new values for a different simulation.

Analyze the Simulated Response

1. Select **DOE > Special Purpose > Group Orthogonal Supersaturated > Fit Group Orthogonal Supersaturated**.
2. Select Response and click **Y**.
3. Select Fake 1 through X15 and click **X**.
4. Click **OK**.

Tip: The design table generated using the Group Orthogonal Supersaturated platform contains a script called Model. Run this script to run this analysis directly.

Figure 26.3 Fit Group Orthogonal Supersaturated Results



Your simulation set X7, X10, and X13 as factors with effects. The Fit GOSSD procedure identified those three factors as significant. In addition, case X14 was also identified. In practice, you would need to determine whether the size and significance of the effect warrants further attention.

Note: Your simulation results might differ from those shown in Figure 26.3 due to the generation of random results. You might see an alert that a group has one or more additional effects that might be active. This is a limitation of a supersaturated design; your findings might not be conclusive.

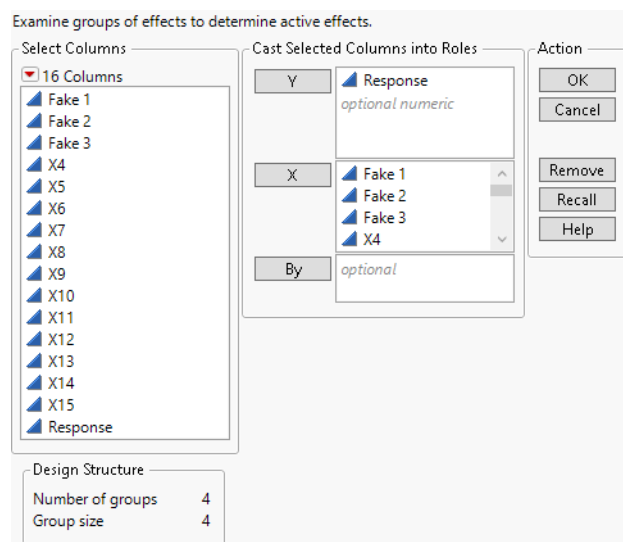
5. Click **Run Model** to run the identified model.

Launch the Fit Group Orthogonal Supersaturated Platform

Launch the Fit Group Orthogonal Supersaturated Design platform by selecting **DOE > Special Purpose > Group Orthogonal Supersaturated > Fit Group Orthogonal Supersaturated**.

Note: If you created your design in JMP using the Group Orthogonal Supersaturated platform, the design table contains a script called Model. Run this script to run the analysis directly.

Figure 26.4 Fit Group Orthogonal Supersaturated Launch Window



- Y** One or more numeric response variables.
- X** Two-level continuous or categorical factors. Because the platform uses the unique structure of a GOSSD in performing the analysis, these factors must define a GOSSD.

For more information about the options in the Select Columns red triangle menu, see the Get Started chapter in *Using JMP*.

The Fit Group Orthogonal Supersaturated Report

The Fit Group Orthogonal Supersaturated Report has two sections. The first section reports the estimates of active effects within each group, fitting one group at a time. Only active groups are shown. This is Stage 1 of the analysis. In Stage 2 a combined model is constructed

Figure 26.5 Stage 1 Findings

Fit Group Orthogonal Supersaturated for Response

Group 1

Group 2

Term	Estimate	Std Error	t Ratio	Prob> t
X7	3.1345	0.1542	20.331	0.0024*

Group 3

Term	Estimate	Std Error	t Ratio	Prob> t
X10	3.1411	0.1542	20.374	0.0024*

Group 4

Term	Estimate	Std Error	t Ratio	Prob> t
X13	3.3291	0.1635	20.359	0.0024*
X14	0.4143	0.1635	2.5336	0.1268

Term Main effects that are identified as active.

Estimate Parameter estimate for a regression fit of Y on the main effect.

Std Error The standard error of the estimate, computed using the fake factors in Group 1.

t Ratio The Estimate divided by its Std Error.

Prob>|t| The *p*-value computed using the t Ratio and the degrees of freedom for error (DF).

The second section of the report contains the results of Stage 2 of the analysis. Stage 2 constructs the model used to evaluate the design results.

Figure 26.6 Stage 2 Results

Combined Model Parameter Estimates

Term	Estimate	Std Error	t Ratio	Prob> t
X7	3.1345	0.2566	12.215	<.0001*
X10	3.1411	0.2566	12.241	<.0001*
X13	3.3291	0.2722	12.231	<.0001*
X14	0.4143	0.2722	1.5222	0.1718

Statistic	Value
RMSE	0.8889
DF	7

Make Model Run Model

Regression Plot

Term Main effects that are identified as active.

Estimate Parameter estimate for a regression fit of Y on the main effect.

Std Error The standard error of the estimates obtained by fitting the specified model and intercept using the full data set.

t Ratio The Estimate divided by its Std Error.

Prob>|t| The p -value computed using the t Ratio and the degrees of freedom for error (DF).

RMSE The square root of the mean square error.

DF The degrees of freedom associated with the error estimate used to construct the RMSE.

Fit Group Orthogonal Supersaturated Platform Options

The red triangle menu in the Fit Group Orthogonal Supersaturated Design platform contains the following options:

See the JMP Reports chapter in *Using JMP* for more information about the following options:

Redo Contains options that enable you to repeat or relaunch the analysis. In platforms that support the feature, the Automatic Recalc option immediately reflects the changes that you make to the data table in the corresponding report window.

Save Script Contains options that enable you to save a script that reproduces the report to several destinations.

Statistical Details for the Fit Group Orthogonal Supersaturated Platform

A two-stage strategy is used to fit group orthogonal supersaturated designs. The first stage identifies which groups contain active effects.

- Obtain an estimate of σ from the fake factors and then discard the fake factors from further analyses.
- Test for significance of each group of factors.

The second stage considers significance of individual factors in active groups. This stage uses a guided subsets procedure. See Jones and Nachtsheim (2016). For complete statistical details, see Jones, et al. (2019).

Appendix A

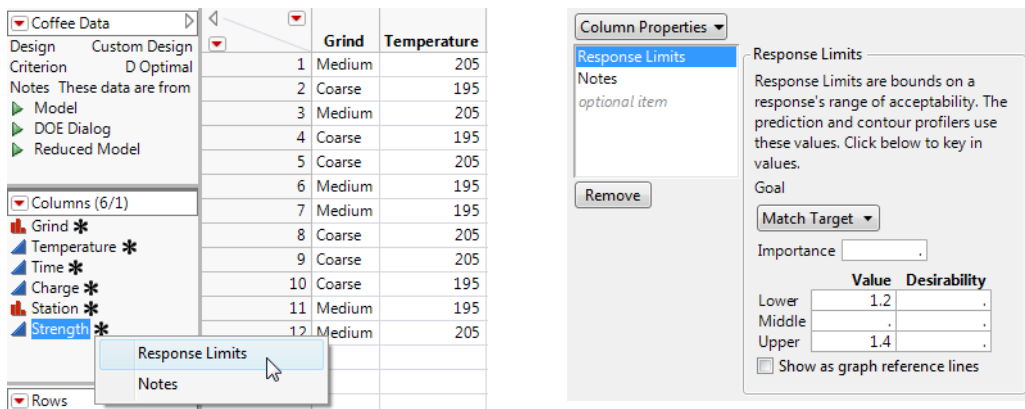
Column Properties

Understanding Column Properties Assigned by DOE

When you construct a design using the DOE platforms, column properties are saved to the data table that contains the resulting design. This appendix provides detail about only those column properties that are saved to designs that the DOE platform constructs. Examples illustrate how each column property is assigned and used by the DOE platforms. Descriptions of column properties not assigned by the DOE platforms are provided in The Column Info Window chapter in *Using JMP*.

Some of the column properties described in this appendix are useful in general modeling situations. To use the properties more generally, you can specify them yourself. This ability is particularly useful when your design has not been created by a DOE platform. Some of the examples in this appendix illustrate situations where you add a column property on your own.

Figure A.1 Column Property Asterisks and Column Info Window



Contents

Adding and Viewing Column Properties	731
Response Limits.....	732
Response Limits Example	734
Editing Response Limits	737
Design Role	740
Design Role Example	740
Coding.....	742
Low and High Values	743
Coding Column Property and Center Polynomials.....	744
Coding Example	744
Assigning Coding	747
Mixture	750
Mixture Example.....	752
Factor Changes	755
Factor Changes Example	757
Value Order	759
Value Order Example	761
Assigning Value Order.....	762
Value Labels	763
Value Labels Example.....	763
RunsPerBlock.....	765
RunsPerBlock Example	766
ConstraintState	766
ConstraintState Example	767

Adding and Viewing Column Properties

The DOE platforms automatically save certain column properties to the design tables that they construct. However, some of the column properties associated with designed experiments are useful in general modeling situations. To use these column properties with data tables that have not been created using DOE platforms, you can add them yourself.

Adding a Column Property

To assign a column property to one or more columns, do the following:

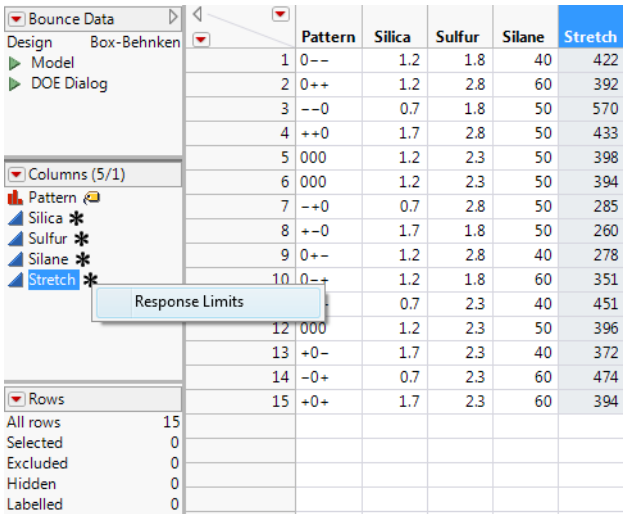
1. Select the column or columns to which you want to assign a property.
2. Do one of the following:
 - Right-click the header area, select **Column Properties**, and select the property.
 - Right-click the header area, select **Column Info**, and select the property from the Column Properties menu.
 - Select **Cols > Column Info** and select the property from the Column Properties menu.
3. In the column property panel that appears, specify values and select options as appropriate.
4. Click **Apply** to add the column property or click **OK** to add the column property and close the column properties window.

Tip: A column might already contain a property that you want to apply to other columns. Use the Standardize Attributes command to apply that property to other columns. See The Column Info Window chapter in *Using JMP*.

Viewing a Column Property

To view the properties assigned to a specific column, in the columns panel, click the column property asterisk icon *****. Click a property to see its settings or to edit it. Figure A.2 shows the column properties assigned to Stretch in the Bounce Data.jmp sample data table, located in the Design Experiment folder.

Figure A.2 Asterisk Icon for Stretch Revealing Two Column Properties



You can also view column properties by accessing the Column Properties list in the Column Info menu. Select the column or columns whose column properties you want to view and do one of the following:

- Right-click the header area, select **Column Info**, and select the property from the Column Properties list.
- Select **Cols > Column Info** and select the property from the Column Properties list.

Response Limits

Using the Response Limits column property, you can specify the following:

- bounds on the range of variation for a response
- a desirability goal
- a measure of the importance of the response
- desirability values

The Response Limits column property defines a desirability function for the response. The Profiler and Contour Profiler use desirability functions to find optimal factor settings. See the Profiler chapter in *Profilers*.

Figure A.3 shows the Response Limits panel in the Column Info window for the response Stretch in the Bounce Data.jmp sample data table, found in the Design Experiment folder.

Figure A.3 Example of the Response Limits Panel

Response Limits

Response Limits are bounds on a response's range of acceptability. The prediction and contour profilers use these values. Click below to key in values.

Goal
Match Target ▼

Importance

	Value	Desirability
Lower	350	.
Middle	450	.
Upper	550	.

☐ Show as graph reference lines

The Response Limits panel consists of the following areas:

Goal Select your response goal from the menu. Available goals are Maximize, Match Target, Minimize, and None. JMP defines a desirability function for the response to match the selected goal. If you specify limits, the desirability function is defined using these limits. If you do not specify limits, JMP bases the desirability function on conservative limit values derived from the distribution of the response. If None is selected as the goal, then all response values are considered equally desirable. See also [“Responses”](#) on page 86 in the [“Custom Designs”](#) chapter.

Importance Enter a relative weighting for each of several responses in computing the overall desirability function. The Importance value can be any positive number. When no Importance value is specified, JMP treats all responses in a given analysis as having equal Importance values. If there is only one response, it receives Importance 1.

Value Specify Lower and Upper limits and a Middle value for your response. JMP uses these values to construct a desirability function for the response. If you do not specify limits, JMP bases the desirability function on conservative limit values. If your goal is Match Target and no Middle value is specified, then the target is defined to be the midpoint of the Lower and Upper limits.

Desirability Specify values that reflect the desirability of your Lower, Middle, and Upper values. Desirability values should be between 0 and 1. If you do not specify Desirability values, JMP assigns values in accordance with the selected Goal.

Show as Graph Reference Lines Shows horizontal reference lines for the Lower, Middle, and Upper values in the Actual by Predicted Plot and the Prediction Profiler. This option applies only if limits are specified.

Response Limits Example

The Coffee Data.jmp sample data table (located in the Design Experiment folder) contains the results of an experiment that was performed to optimize the Strength of coffee. For a complete description of the experimental design and analysis, see “The Coffee Strength Experiment” on page 42 in the “Starting Out with DOE” chapter.

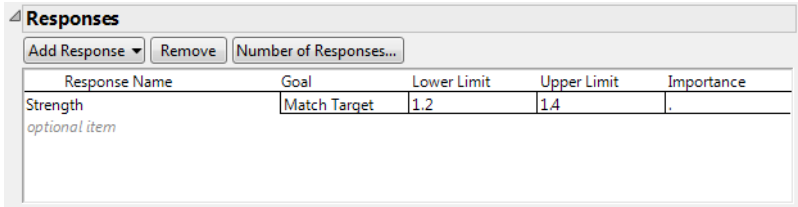
Your goal is to find factor settings that enable you to brew coffee with a target strength of 1.3, which is considered to be the most desirable value. Values less than 1.2 and greater than 1.4 are completely undesirable. The desirability of values between 1.2 and 1.4 decreases as their distance from 1.3 increases.

1. Select **Help > Sample Data Library** and open Design Experiment/Coffee Data.jmp.
2. In the Table panel, click the green triangle next to the **DOE Dialog** script.

The DOE Dialog script re-creates the Custom Design dialog that was used to create the experimental design in Coffee Data.jmp.

3. Open the Responses outline.

Figure A.4 Responses Outline in Custom Design Window

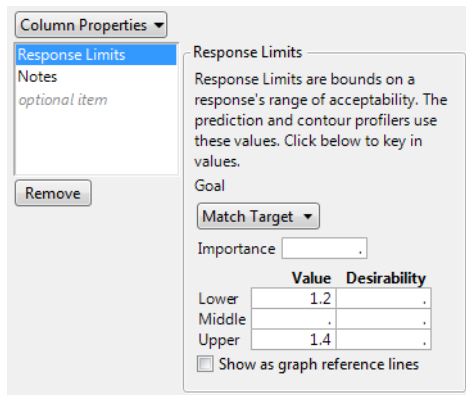


Responses				
Add Response Remove Number of Responses...				
Response Name	Goal	Lower Limit	Upper Limit	Importance
Strength <i>optional item</i>	Match Target	1.2	1.4	.

When you designed this experiment, you specified a response Goal of Match Target with a Lower Limit of 1.2 and an Upper Limit of 1.4. Since there is only one response, you did not specify a value for Importance, because it is 1 by default. When you constructed the design table, JMP assigned the Response Limits column property to Strength.

4. Close the Custom Design window.
5. In the Coffee Data.jmp sample data table, select the Strength column and select **Cols > Column Info**.
6. Select Response Limits in the Column Properties list.

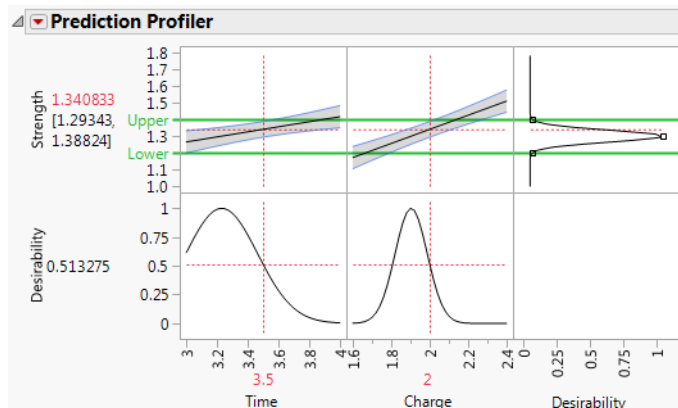
Figure A.5 Response Limits Column Property for Strength



Notice the following:

- The Goal is set to **Match Target**.
 - Importance is missing. When Importance is missing, JMP treats all responses in a given analysis as having equal Importance values. So JMP assigns Strength an Importance value of one.
 - The Lower limit is 1.2.
 - The Upper limit is 1.4.
 - No Middle value is specified.
Because no Middle value is specified, the target is defined to be the midpoint of the Lower and Upper limits, which is 1.3.
 - No Desirability values are specified.
7. Select the **Show as Graph Reference Lines** option.
This option shows horizontal reference lines for the Lower, Middle, and Upper values in the Actual by Predicted Plot and the Prediction Profiler.
 8. Click **OK**.
 9. In the Coffee Data.jmp data table, click the green triangle next to the **Reduced Model** script.
 10. Click **Run**.

The Prediction Profiler appears at the bottom of the report.

Figure A.6 Profiler Showing Desirability Function for Strength

The desirability function for Strength appears in the plot at the right above Desirability. This plot appears because the data table contains a Response Limits column property for Strength. The Prediction Profiler also shows reference lines for the Lower and Upper limits for Strength.

11. Hold down the Ctrl key and click the Strength plot for Desirability.

Figure A.7 Response Goal Window for Strength

The figure shows a Response Goal window for the Strength response. It has a 'Match Target' dropdown menu. Below it, there are input fields for Strength values and Desirability. The High value is 1.4 with Desirability 0.0183. The Middle value is 1.3 with Desirability 1. The Low value is 1.2 with Desirability 0.0183. The Importance is set to 1. Buttons for OK, Cancel, and Help are at the bottom.

Strength	Values	Desirability
High:	1.4	0.0183
Middle:	1.3	1
Low:	1.2	0.0183
Importance:	1	

Notice the following:

- JMP determines the Middle value to be the midpoint of the High and Low limits that you specified in the Response Limits column property.
- Because the Goal is set to Match Target, JMP sets the Desirability for the Middle value to 1.
- JMP sets the Desirability for the High and Low values to very small numbers, 0.0183.
- The Desirability plot in Figure A.6 shows how JMP uses the Desirability values shown in Figure A.7. The Desirability function for Strength is essentially 0 beyond the Low and High values and it increases to 1 gradually as Strength approaches the target of 1.3. The Importance value is set to 1 since there is only one response in the model.

12. Click **Cancel** to exit the window.
13. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

The settings for Time and Charge are updated to show settings for the factors that maximize the desirability function for Strength. However, many other settings also maximize the desirability function. See the Contour Profiler chapter in *Profilers* for information about how to identify other settings that maximize the desirability function.

Editing Response Limits

In the Vinyl Data.jmp sample data table, a Response Limits column property is already assigned to the response thickness. The property has a goal of maximizing thickness. Suppose that instead of maximizing thickness, you want the sheets of vinyl to have a thickness between 6 and 10, with a target thickness of 8.5.

1. Select **Help > Sample Data Library** and open Design Experiment/Vinyl Data.jmp.
2. Select the thickness column and select **Cols > Column Info**.

The Response Limits property appears in the Column Properties list as the only property assigned to thickness. The Response Limits panel appears to the right of the list.

3. Click **Maximize** and select **Match Target**.
4. Type 1 for the **Importance** value.
5. Under Value, type 6 for the Lower value, 8.5 for the Middle value, and 10 for the Upper value.

This is an example of asymmetric response limits. Values of thickness as small as 6 or as large as 10 are acceptable. However, the target for thickness is 8.5.

6. Select **Show as Graph Reference Lines**.

This option shows horizontal reference lines for the Lower, Middle, and Upper values in the Actual by Predicted Plot and the Prediction Profiler.

Figure A.8 Completed Response Limits Panel

'thickness' in Table 'Vinyl Data'

Column Name:

☐ Lock

Data Type:

Modeling Type:

Format: Width:

☐ Use thousands separator (,)

Column Properties ▾

Response Limits
optional item

Response Limits

Response Limits are bounds on a response's range of acceptability. The prediction and contour profilers use these values. Click below to key in values.

Goal:

Importance:

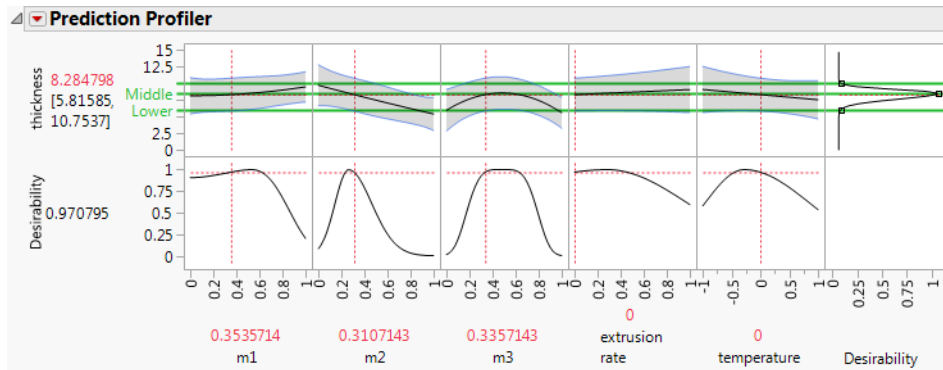
	Value	Desirability
Lower	6	.
Middle	8.5	.
Upper	10	.

☒ Show as graph reference lines

OK
Cancel
Apply
Help

7. Click **OK**.
8. In the Vinyl Data.jmp data table, click the green triangle next to the **Model** script.
Note that m1, m2, and m3 are mixture factors. Also, the design involves a random Whole Plots factor. Because of this, the default Method is REML (Recommended).
9. Click **Run**.
10. Click the Response thickness red triangle and select **Row Diagnostics > Plot Actual by Predicted**.
The reference lines for the Lower, Middle, and Upper limits appear on the Actual by Predicted Plot.
11. Click the Response thickness red triangle and select **Factor Profiling > Profiler**.

Figure A.9 Prediction Profiler Showing Asymmetric Desirability Function



The plot at the right above Desirability shows the desirability function that JMP has constructed for thickness. The desirability is 1 at the Middle limit of 8.5. The desirability is essentially 0 for thickness values below 6 and above 10.

12. Hold down the Ctrl key and click the thickness plot for Desirability.

Figure A.10 Response Goal Window for Thickness

The figure shows the Response Goal window for the 'thickness' response. The 'Match Target' goal is selected. The settings are as follows:

thickness	Values	Desirability
High:	10	0.0183
Middle:	8.5	1
Low:	6	0.0183
Importance:	1	

Buttons for OK, Cancel, and Help are at the bottom.

This window shows your settings for the High, Middle, and Low Values. It also shows the Desirability values that JMP assigns, based on your goal of Match Target. The Desirability function shown in Figure A.9 is a continuous curve that matches the Desirability settings in Figure A.10 at the High, Middle, and Low Values. At other values, the Desirability function assigns desirabilities that are consistent with the selected goal.

13. Click **Cancel**.
14. Click the Prediction Profiler red triangle and select **Optimization and Desirability > Maximize Desirability**.

The settings for the factors are updated to show values that maximize the desirability function for thickness. Keep in mind that many other settings also maximize the desirability function. The predicted response at these optimal settings is 8.5. Recall that you set 8.5 as the target setting, with limits of 6 and 10.

15. Close the Vinyl Data.jmp sample data table without saving the changes.

Design Role

Factors in designed experiments, as well as in more general models, can behave in various ways. JMP uses the design role column property to describe these behaviors. These are the possible design roles:

- Continuous
- Discrete Numeric
- Categorical
- Blocking
- Covariate
- Mixture
- Constant
- Uncontrolled
- Random Block
- Signal
- Noise

In many of the JMP DOE platforms, you can specify factors with different design roles. In some platforms, your design requirements cause JMP to define factors. For example, Whole Plots and Subplots are factors that JMP creates when you specify very-hard-to-change and hard-to-change factors. In platforms where various design roles can occur, when JMP creates the design table for your design, each factor is assigned the Design Role column property.

- For descriptions of the design roles other than Random Block, Signal, and Noise, see [“Factor Types”](#) on page 89 in the “Custom Designs” chapter.
- For a description of the Random Block design role, see [“Changes and Random Blocks”](#) on page 92 in the “Custom Designs” chapter.
- For descriptions of the Signal and Noise design roles, see [“Factors”](#) on page 428 in the “Taguchi Designs” chapter.

Design Role Example

The experiment in the Odor Control Original.jmp sample data table studies the effect of three factors on odor. You designed the 15-run experiment and it was conducted. However, when the results were reported to you, you learned that the experiment was conducted over three days. The first 5 runs were conducted on Day 1, the second 5 runs on Day 2, and the remaining 5 runs on Day 3.

Since variations in temperature and humidity might have an effect on the response, you want to include Day as a random blocking factor. It is easy to add a column for Day to the design table. But you want to use the Evaluate Design platform to compare the design with the unexpected block to your original design. You also want the ability to use the Augment Design platform in case you need to augment the design. To use the Evaluate Design and Augment Design platforms, you need to add the Design Role column property to your new Day column.

1. Select **Help > Sample Data Library** and open **Odor Control Original.jmp**.
2. Select the first column, **Run**.
3. Select **Cols > New Columns**.
4. Next to **Column Name**, type **Day**.
5. From the **Initialize Data** list, select **Sequence Data**.
6. Enter the following:
 - 1 for **From**
 - 3 for **To**
 - 1 for **Step**
 - 5 for **Repeat each value N times**
7. Next to **Number of columns to add**, type 1.

Figure A.11 Completed New Columns Window

Add columns to 'Odor Control Original'

Column Name: Day

☐ Lock

Data Type: Numeric

Modeling Type: Continuous

Format: Best Width: 12

☐ Use thousands separator (,)

Initialize Data: Sequence Data

From: 1

To: 3

Step: 1

Repeat each value N times: 5

Number of columns to add: 1

After selected column: [dropdown]

Column Properties [dropdown]

OK Cancel Apply Next Help

8. Click **OK**.

The Day column is added as the second column in the data table.

9. Select the Day column.
10. Select **Cols > Column Info**.
11. From the Column Properties list, select **Design Role**.
12. In the Design Role panel, click **Continuous** and select **Random Block**.
13. Click **OK**.

In the columns panel, an asterisk appears next to Day.

14. Click the asterisk next to Day to verify that the Design Role column property has been assigned.
15. Close the Odor Control Original.jmp sample data table without saving the changes.

Coding

The Coding column property applies only to columns with a numeric data type. It applies a linear transformation to the data in the column. In the Coding column property window, you specify a Low Value and a High Value. The Low Value and High Value in your original data are transformed to -1 and +1. JMP uses the transformed data values whenever the column is entered as a model effect in the Fit Model platform.

The coding property is useful for the following reasons:

- Coded predictors lead to parameter estimates that are more easily interpreted and compared.
- Coded predictors help reduce multicollinearity in models with interaction and higher-order terms.

When any DOE platform other than Accelerated Life Test Design creates a design, JMP defines a Coding column property for each non-mixture factor with a numeric data type. Figure A.12 shows the Coding column property panel for the column Feed Rate in the Reactor 20 Custom.jmp sample data table, found in the Design Experiment folder.

Figure A.12 Coding Property Panel for Feed Rate

'Feed Rate' in Table 'Reactor 20 Custom'

Column Name:

☐ Lock

Data Type:

Modeling Type:

Format: Width:

☐ Use thousands separator (,)

Column Properties:

- Coding**
- Design Role
- Factor Changes
optional item

Coding

Coding declares that when you fit it in a model, the range Low to High is transformed into the range -1 to 1. This property makes tests and parameters estimates more meaningful.

Low Value:

High Value:

Low and High Values

When the Coding property is applied as part of design construction in a DOE platform, the Low Value is initially set to the minimum setting of the factor and the High Value is initially set to the maximum setting.

When you apply the Coding property to a column that does not contain that property, JMP inserts the minimum value as the Low Value and the maximum value as the High Value. You can change these values as needed.

Caution: After the Coding column property is assigned to a column, JMP does not automatically update it when you make changes to the values in the column. If you change the values in a column that has a Coding column property, review the High Value and Low Value to ensure that they are still appropriate.

The Coding column property centers each value in a column by subtracting the midpoint of the High Value and Low Value. It then divides by half the range. Suppose that H is the High Value and L is the Low Value. Then every X in the column is transformed to the following:

$$\frac{X - (H + L)/2}{(H - L)/2}$$

For each factor, the transformed values have a midpoint equal to 0 and range from -1 to +1.

Coding Column Property and Center Polynomials

The Center Polynomials option is located in the Fit Model launch window, within the Model Specification red triangle menu. Center Polynomials centers a continuous column involved in a polynomial term by subtracting the mean of each value in the column. See the Model Specification chapter in *Fitting Linear Models*.

If the Coding column property is assigned to a column, then the Center Polynomials option has no effect on that column. In a polynomial term involving that column, the values are centered and scaled as specified by their Coding property. Suppose that another column in the model does not have the Coding property and that you select Center Polynomials. Then that column is centered by its mean in any polynomial term where it appears.

Coding Example

The Reactor 20 Custom.jmp sample data table contains data from a 20-run design that was constructed using the Custom Design platform. The experiment investigates the effects of five factors on a yield response (Percent Reacted) for a chemical process.

1. Select **Help > Sample Data Library** and open Design Experiment/Reactor 20 Custom.jmp.
2. In the Table panel, click the green triangle next to the **DOE Dialog** script.
3. Open the **Factors** outline.

Figure A.13 Factors Outline for Design Used in Reactor 20 Custom.jmp

Name	Role	Changes	Values
Feed Rate	Continuous	Easy	10 15
Catalyst	Continuous	Easy	1 2
Stir Rate	Continuous	Easy	100 120
Temperature	Continuous	Easy	140 180
Concentration	Continuous	Easy	3 6

Notice that the settings for Temperature range from 140 to 180. When the design was generated, the Coding column property was assigned to Temperature. The Low Value is set to 140 and the High Value is set to 180.

4. Close the Custom Design window.
5. In the Reactor 20 Custom.jmp sample data table, click the asterisk next to Temperature in the columns panel and select **Coding**.

The Column Info window appears and shows the Coding column property panel. You can see that JMP added the column property, specifying the Low Value and High Value, when it constructed the design table. In fact, by repeating this step, you can verify that JMP added the Coding property for all five factors.

Figure A.14 Coding Panel for Temperature

'Temperature' in Table 'Reactor 20 Custom'

Column Name:

☐ Lock

Data Type:

Modeling Type:

Format: Width:

☐ Use thousands separator (,)

Column Properties:

- Coding
- Design Role
- Factor Changes
- optional item

Remove

Coding

Coding declares that when you fit it in a model, the range Low to High is transformed into the range -1 to 1. This property makes tests and parameters estimates more meaningful.

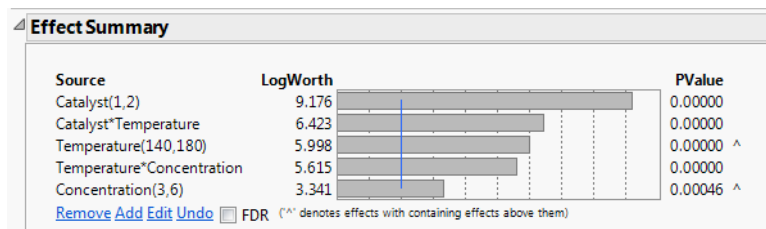
Low Value:

High Value:

OK Cancel Apply Help

- Click **Cancel** to close the Column Info window.
- In the Reactor 20 Custom.jmp sample data table, click the green triangle next to the **Reduced Model** script.
This script fits a model that contains only the five effects determined to be significant based on an analysis of the full model.
- Click **Run**.

Figure A.15 Effect Summary Report for Reduced Model



In the Source list, the High and Low values used in the Coding column property appear in parentheses to the right of the main effects, Catalyst, Temperature, and Concentration. The ranges imposed by the Coding property are not shown for the interaction effects.

Tip: Notice the “^” symbols to the right of the PValues for Temperature and Concentration. These symbols indicate that these main effects are components of interaction effects with smaller *p*-values. If an interaction effect is included in the model, then the principle of effect heredity requires that all component effects are also in the model. See [“Effect Heredity”](#) on page 62 in the “Starting Out with DOE” chapter.

9. Click the Response Percent Reacted red triangle and select **Estimates > Show Prediction Expression**.

Look at the Prediction Expression outline to see how coding affects the prediction formula.

Figure A.16 Prediction Expression for Reduced Model

The screenshot shows a window titled "Prediction Expression" containing the following formula:

$$\begin{aligned}
 &65.46506734 \\
 &+ 9.9423400673 \cdot \left(\frac{(\text{Catalyst} - 1.5)}{0.5} \right) \\
 &+ 5.5576599327 \cdot \left(\frac{(\text{Temperature} - 160)}{20} \right) \\
 &+ -3.076599327 \cdot \left(\frac{(\text{Concentration} - 4.5)}{1.5} \right) \\
 &+ \left(\frac{(\text{Catalyst} - 1.5)}{0.5} \right) \cdot \left(\frac{(\text{Temperature} - 160)}{20} \right) \cdot 6.0349326599 \\
 &+ \left(\frac{(\text{Temperature} - 160)}{20} \right) \cdot \left(\frac{(\text{Concentration} - 4.5)}{1.5} \right) \cdot -5.150673401
 \end{aligned}$$

Each factor is transformed as specified by the Coding column property. For example, for Temperature, notice the following:

- The Low Value in the Coding property was set to 140. The Temperature value of 140 is transformed to -1.
- The High Value in the Coding property was set to 180. The Temperature value of 180 is transformed to +1.
- The midpoint of the Low and High values is 160. The Temperature value of 160 is transformed to 0.

The transformed values help you compare the effects. The estimated coefficient for Catalyst is 9.942 and the estimated coefficient for Concentration is -3.077. It follows that the predicted effect of Catalyst on Percent Reacted is more than three times as large as the effect of Concentration on Percent Reacted. Also, the coefficients indicate that predicted Percent Reacted increases as Catalyst increases and decreases as Concentration increases.

The transformed values help you interpret the coefficients:

- When all factors are at their midpoints, their transformed values are 0. The predicted Percent Reacted is the intercept, which is 65.465.

- When Catalyst and Concentration are at their midpoints, a 20 unit increase in Temperature increases the Percent Reacted by 5.558 units.
- Suppose that Concentration is at its midpoint, so that its transformed value is 0:
 - When Catalyst is at its midpoint, a 20 unit increase in Temperature increases the Percent Reacted by 5.558 units.
 - When Catalyst is at its high setting, a 20 unit increase in Temperature increases the Percent Reacted by $5.558 + 6.035 = 11.593$ units.

It follows that the coefficient of the interaction term, 6.035, is the increase in the slope of the model for predicted Percent Reacted for a 0.5 unit change in Catalyst.

Assigning Coding

The experimental data in the Tiretread.jmp sample data table results from an experiment to study the effects of SILICA, SILANE, and SULFUR on four measures of tire tread performance. In this example, you will consider only one of the responses, ABRASION.

You will first fit a model using the uncoded factors. Then you will assign the coding property to the factors and rerun the model to obtain meaningful parameter estimates.

1. Select **Help > Sample Data Library** and open Tiretread.jmp.
2. Select **Analyze > Fit Model**.
3. Select ABRASION and click **Y**.
4. Select SILICA, SILANE, and SULFUR and click **Macros > Response Surface**.
5. Check **Keep dialog open**.
6. Click **Run**.
7. Click the Response ABRASION red triangle and select **Estimates > Show Prediction Expression**.

Figure A.17 Prediction Expression for Model with Uncoded Factors

Prediction Expression

```

-40.0394111
+ 32.987289493 • SILICA
+ 1.7880765099 • SILANE
+ 21.813076926 • SULFUR
+ (SILICA - 1.2) • ((SILICA - 1.2) • -16.03840362 )
+ (SILICA - 1.2) • ((SILANE - 50) • 1.025 )
+ (SILANE - 50) • ((SILANE - 50) • -0.034471056 )
+ (SILICA - 1.2) • ((SULFUR - 2.3) • 28.5 )
+ (SILANE - 50) • ((SULFUR - 2.3) • 1.575 )
+ (SULFUR - 2.3) • ((SULFUR - 2.3) • -6.288485272 )

```

The coefficients do not help you compare effect sizes. The sizes of the coefficients do not reflect the impact of the effects on ABRASION over the range of their settings. Also, the coefficients are not easily interpreted. For example, the coefficients do not facilitate your understanding of the predicted response when SILICA is at the midpoint of its range.

Apply the Coding column property to the three factors to see how coding makes the coefficients more meaningful.

8. In the Tiretread.jmp data table, select SILICA, SILANE, and SULFUR in the Columns panel. Right-click the highlighted column titles and select **Standardize Attributes**.
9. Select **Column Properties > Coding** in the Standardize Properties panel.
10. Click **OK**.

An asterisk appears in the Columns panel next to SILICA, SILANE, and SULFUR indicating that these have been assigned a column property.

11. In the Fit Model window, click **Run**.
12. Click the Response ABRASION red triangle and select **Estimates > Show Prediction Expression**.

Figure A.18 Prediction Expression for Model with Coded Factors

Prediction Expression

$$\begin{aligned}
 &139.11923872 \\
 &+ 26.934121871 \cdot \left(\frac{(\text{SILICA} - 1.2)}{0.8165} \right) \\
 &+ 29.199289407 \cdot \left(\frac{(\text{SILANE} - 50)}{16.33} \right) \\
 &+ 17.81037731 \cdot \left(\frac{(\text{SULFUR} - 2.3)}{0.8165} \right) \\
 &+ \left(\frac{(\text{SILICA} - 1.2)}{0.8165} \right) \cdot \left(\left(\frac{(\text{SILICA} - 1.2)}{0.8165} \right) \cdot -10.69235863 \right) \\
 &+ \left(\frac{(\text{SILICA} - 1.2)}{0.8165} \right) \cdot \left(\left(\frac{(\text{SILANE} - 50)}{16.33} \right) \cdot 13.666781125 \right) \\
 &+ \left(\frac{(\text{SILANE} - 50)}{16.33} \right) \cdot \left(\left(\frac{(\text{SILANE} - 50)}{16.33} \right) \cdot -9.192358626 \right) \\
 &+ \left(\frac{(\text{SILICA} - 1.2)}{0.8165} \right) \cdot \left(\left(\frac{(\text{SULFUR} - 2.3)}{0.8165} \right) \cdot 19.000159125 \right) \\
 &+ \left(\frac{(\text{SILANE} - 50)}{16.33} \right) \cdot \left(\left(\frac{(\text{SULFUR} - 2.3)}{0.8165} \right) \cdot 21.000175875 \right) \\
 &+ \left(\frac{(\text{SULFUR} - 2.3)}{0.8165} \right) \cdot \left(\left(\frac{(\text{SULFUR} - 2.3)}{0.8165} \right) \cdot -4.192358626 \right)
 \end{aligned}$$

The coefficients for the coded factors enable you to compare effect sizes. SILANE has the largest effect on ABRASION over the range of design settings. The effects of SILICA and the SILANE*SULFUR interaction are large as well.

The coefficients for the coded factors are also more easily interpreted. For example, when all factors are at the center of their ranges, the predicted value of ABRASION is the intercept, 139.12.

13. Close the Tiretread.jmp sample data table without saving the changes.

Mixture

The Mixture column property is useful when a column in a data table represents a component of a mixture. The components of a mixture are constrained to sum to a constant. Because of this, they differ from non-mixture factors. The Mixture column property serves two purposes:

- It identifies a column as a mixture component.
If you add a column with the Mixture column property as a model effect in the Analyze > Fit Model window, JMP automatically generates a no-intercept model.
- It defines the coding for a mixture component.
Coding for mixture components differs from that for non-mixture factors. However, as with non-mixture factors, a benefit of coding for mixture factors is that it helps you interpret parameter estimates. See [“PseudoComponent Coding”](#) on page 751.

Figure A.19 shows the Mixture column property panel for the factor m1 in the Vinyl Data.jmp sample data table, found in the Design Experiment folder.

Figure A.19 Mixture Column Property Panel

'm1' in Table 'Vinyl Data'

Column Name: m1

☐ Lock

Data Type: Numeric

Modeling Type: Continuous

Format: Best Width: 10

☐ Use thousands separator (,)

Column Properties:

- Mixture (selected)
- Design Role
- Factor Changes
- optional item

Remove

Mixture

Mixture is specified if the column participates in a mixture where it and other columns add up to a constant

Lower Limit: 0

Upper Limit: 1

Sum of Terms: 1

☒ L PseudoComponent Coding

☐ U PseudoComponent Coding

(If both checked, it chooses most appropriate)

OK, Cancel, Apply, Help

In the Mixture column property panel, you can specify the following:

Lower Limit Specifies the low value used in PseudoComponent Coding. When the Mixture property is applied as part of design construction in a DOE platform, the Lower Limit is set to the minimum setting of the factor. When you apply the Mixture property to a column that does not contain that property, JMP inserts the minimum value as the Lower Limit. You can change this value as needed.

Upper Limit Specifies the high value used in PseudoComponent Coding. When the Mixture property is applied as part of design construction in a DOE platform, the Upper Limit is set to the maximum setting of the factor. When you apply the Mixture property to a column that does not contain that property, JMP inserts the maximum value as the Upper Limit. You can change this value as needed.

Sum of Terms Specifies the sum of the mixture components. When you apply the Mixture property to a column that does not contain that property, JMP inserts a default value of 1 for the Sum of Terms.

L PseudoComponent Coding Transforms data values so that the Lower Limit corresponds to 0.

U PseudoComponent Coding Transforms data values so that the Upper Limit corresponds to 0.

PseudoComponent Coding

A pseudo-component is a linear transformation. Let S denote the sum of the mixture components. Suppose that i columns $X_1, X_2, \dots, X_q$ have been assigned the Mixture column property. Suppose that the columns and effects constructed from these columns are entered as effects in the Fit Model window.

Define the following:

$$L = \sum_{i=1}^q L_i, \text{ where } L_i \text{ is the Lower Limit for } X_i$$

$$U = \sum_{i=1}^q U_i, \text{ where } U_i \text{ is the Upper Limit for } X_i$$

Let x_i denote a value of the column X_i . The L PseudoComponent at x_i is defined as follows:

$$x_i^L = (X_i - L_i) / (S - L)$$

The U PseudoComponent at x_i is defined as follows:

$$x_i^U = (U_i - X_i) / (U - S)$$

If you select both **L PseudoComponent Coding** and **U PseudoComponent Coding**, the Fit Model platform uses the L coding if $(S - L) < (U - S)$. Otherwise, the U coding is used.

In Fit Model, mixture factors are transformed using pseudo-components before computing parameter estimates. This helps make parameter estimates more meaningful. In reports dealing with parameter estimates, the mixture main effects are given by the coding transformation. Other reports, such as the profilers, are based on the uncoded values.

Mixture Example

The data in the Donev Mixture Data.jmp sample data table, found in the Design Experiment folder, are based on an example from Atkinson and Donev (1992). The design includes three mixture factors and one non-mixture factor. The response and factors are as follows:

- The response is the electromagnetic Damping of an acrylonitrile powder.
- The three mixture ingredients are copper sulphate (CuSO_4), sodium thiosulphate ($\text{Na}_2\text{S}_2\text{O}_3$), and Glyoxal.
- The non-mixture environmental factor of interest is the Wavelength of light.

Though Wavelength is theoretically continuous, the researchers were interested only in predictions at three discrete wavelengths. As a result, Wavelength is treated as a categorical factor with three levels.

For more information about using Custom Design to construct a design for this situation, see “Mixture Experiments” on page 176 in the “Examples of Custom Designs” chapter.

1. Select **Help > Sample Data Library** and open Design Experiment/Donev Mixture Data.jmp.
2. Click the asterisk next to CuSO_4 in the columns panel and select **Mixture**.

Figure A.20 Mixture Column Property Panel for CuSO_4

'CuSO4' in Table 'Donev Mixture Data'

Column Name:

☐ Lock

Data Type:

Modeling Type:

Format: Width:

☐ Use thousands separator (,)

Column Properties:

Mixture

Mixture is specified if the column participates in a mixture where it and other columns add up to a constant

Lower Limit:

Upper Limit:

Sum of Terms:

☒ L PseudoComponent Coding

☐ U PseudoComponent Coding

(If both checked, it chooses most appropriate)

Buttons: OK, Cancel, Apply, Help

Notice the following:

- The Lower Limit is 0.2, the minimum design setting for CuSO₄.
 - The Upper Limit is 0.8, the maximum design setting for CuSO₄.
 - The Sum of Terms is set to 1. This is the sum of the three mixture factors.
 - The **L PseudoComponent Coding** option is selected. See “[PseudoComponent Coding](#)” on page 751.
3. Click **Cancel**.
 4. Click the asterisk next to Glyoxal in the columns panel and select **Mixture**.
For this factor, note the following:
 - The Lower Limit is 0, the minimum design setting for Glyoxal.
 - The Upper Limit is 0.6, the maximum design setting for Glyoxal.
 5. Click **Cancel**.
 6. In the Donev Mixture Data.jmp data table, click the green triangle next to the **Model** script.
The model contains the main effects of the mixture factors and two-way interactions for all four factors.
 7. Click **Run**.

In the Parameter Estimates report, the mixture factors appear in their pseudo-component coded form. When the mixture factors appear in interactions, they are not denoted in coded form. Nevertheless, the model fitting is based on the pseudo-components. The first three terms in the Parameter Estimates report (Figure A.21), show the coded form for the mixture factors.

Figure A.21 Parameter Estimates Report

Parameter Estimates				
Term	Estimate	Std Error	t Ratio	Prob> t
(CuSO ₄ -0.2)/0.6	6.1910821	0.918805	6.74	0.0005*
(Na ₂ S ₂ O ₃ -0.2)/0.6	4.0089179	0.918805	4.36	0.0048*
Glyoxal/0.6	8.1666667	0.921638	8.86	0.0001*
CuSO ₄ *Na ₂ S ₂ O ₃	11.293949	4.728922	2.39	0.0542
CuSO ₄ *Glyoxal	4.3511692	4.512775	0.96	0.3722
CuSO ₄ *Wavelength[L1]	-3.847343	1.113399	-3.46	0.0135*
CuSO ₄ *Wavelength[L2]	1.8781509	1.079113	1.74	0.1324
Na ₂ S ₂ O ₃ *Glyoxal	-18.4845	4.512775	-4.10	0.0064*
Na ₂ S ₂ O ₃ *Wavelength[L1]	-0.275689	1.064004	-0.26	0.8042
Na ₂ S ₂ O ₃ *Wavelength[L2]	0.4452853	1.103681	0.40	0.7006
Glyoxal*Wavelength[L1]	0.198283	1.090627	0.18	0.8617
Glyoxal*Wavelength[L2]	0.1905384	1.090647	0.17	0.8671

8. Click the Response Damping red triangle and select **Estimates > Show Prediction Expression**.

The Prediction Expression report shows the model that was fit. Note that the mixture factors are transformed using the L PseudoComponent coding.

Figure A.22 Prediction Expression for Damping Model

Prediction Expression

$$\begin{aligned}
& 6.1910820765 \cdot \left(\frac{(\text{CuSO4} - 0.2)}{0.6} \right) \\
& + 4.0089179235 \cdot \left(\frac{(\text{Na2S2O3} - 0.2)}{0.6} \right) \\
& + 8.1666666667 \cdot \left(\frac{\text{Glyoxal}}{0.6} \right) \\
& + \left(\frac{(\text{CuSO4} - 0.2)}{0.6} \right) \cdot \left(\frac{(\text{Na2S2O3} - 0.2)}{0.6} \right) \cdot 11.293949004 \\
& + \left(\frac{(\text{CuSO4} - 0.2)}{0.6} \right) \cdot \left(\frac{\text{Glyoxal}}{0.6} \right) \cdot 4.3511691804 \\
& + \left(\frac{(\text{CuSO4} - 0.2)}{0.6} \right) \cdot \text{Match}(\text{Wavelength}) \begin{pmatrix} \text{"L1"} \Rightarrow -3.847342876 \\ \text{"L2"} \Rightarrow 1.8781509412 \\ \text{"L3"} \Rightarrow 1.9691919345 \\ \text{else} \Rightarrow . \end{pmatrix} \\
& + \left(\frac{(\text{Na2S2O3} - 0.2)}{0.6} \right) \cdot \left(\frac{\text{Glyoxal}}{0.6} \right) \cdot -18.48450251 \\
& + \left(\frac{(\text{Na2S2O3} - 0.2)}{0.6} \right) \cdot \text{Match}(\text{Wavelength}) \begin{pmatrix} \text{"L1"} \Rightarrow -0.275688526 \\ \text{"L2"} \Rightarrow 0.44528528 \\ \text{"L3"} \Rightarrow -0.169596754 \\ \text{else} \Rightarrow . \end{pmatrix} \\
& + \left(\frac{\text{Glyoxal}}{0.6} \right) \cdot \text{Match}(\text{Wavelength}) \begin{pmatrix} \text{"L1"} \Rightarrow 0.1982830114 \\ \text{"L2"} \Rightarrow 0.1905384076 \\ \text{"L3"} \Rightarrow -0.388821419 \\ \text{else} \Rightarrow . \end{pmatrix}
\end{aligned}$$

Suppose that you are interested in predictions at Wavelength L2. Suppose also that Na2S2O3 and Glyoxal are set to their low values, 0.2 and 0 respectively, and that CuSO4 is set to its high value, 0.8. In this case, the predicted Damping equals the parameter estimate for CuSO4 (6.191) plus the parameter estimate for CuSO4*Wavelength[L2] (1.878). You can verify this in the Prediction Profiler.

- Click the Response Damping red triangle and select **Save Columns > Save Coding Table**.

Figure A.23 First Three Columns of Coding Table Showing Coded Mixture Factors

	(CuSO ₄ -0.2)/0.6	(Na ₂ S ₂ O ₃ -0.2)/0.6	Glyoxal/0.6
1	1	0	0
2	0.4	0.6	0
3	0	1	0
4	0.5	0	0.5
5	0	0.5	0.5
6	0	0	1
7	1	0	0
8	0.6	0.4	0
9	0	1	0
10	0.5	0	0.5
11	0	0.5	0.5
12	0	0	1
13	1	0	0
14	0.4	0.6	0
15	0	1	0
16	0.5	0	0.5
17	0	0.5	0.5
18	0	0	1

For this particular design, the L Pseudo Component coding transforms the mixture factors to range between 0 and 1. Note that this does not happen in general

Factor Changes

The Factor Changes column property indicates how difficult it is to change factor settings in a designed experiment. The possible specifications for Factor Changes are Easy, Hard, and Very Hard. For example, Figure A.24 shows the Factor Changes column property panel for the factor A1 in the Battery Data.jmp sample data table, located in the Design Experiment folder.

Figure A.24 Factor Changes Column Property Panel

'A1' in Table 'Battery Data'

Column Name: A1

☐ Lock

Data Type: Numeric

Modeling Type: Continuous

Format: Best Width: 12

☐ Use thousands separator (,)

Column Properties

- Coding
- Design Role
- Factor Changes
- Notes
- optional item

Remove

Factor Changes: Very Hard

OK, Cancel, Apply, Help

- When a design contains factors that are hard-to-change and very-hard-to-change, it must also include a subplot and a whole plot factor:
 - The levels of the whole plot factor define the groups of runs for which the levels of the very-hard-to-change factors are held constant.
 - The levels of the subplot factor define the groups of runs for which the levels of the hard-to-change factors are held constant.
- When a design contains only factors that are hard-to-change, but no factors that are very-hard-to-change, it should include a whole plot factor. The levels of the whole plot factor define the groups of runs for which the levels of the hard-to-change factors are held constant. See [“Changes and Random Blocks”](#) on page 92 in the “Custom Designs” chapter.

Augment and Evaluate Design

For the Evaluate Design and Augment Design platforms, the Factor Changes column property identifies factors with Changes specified as Hard or Very Hard. However, these platforms also require that the whole plot and subplot factors be entered as model effects in the launch windows. This is because the whole plot and subplot factors are part of the design structure.

Custom Design

The Custom Design platform enables you to create designs where all factor changes are Easy. You can also construct split-plot, split-split plot, or two-way split-plot (strip-plot) designs. When constructing these designs, you need to identify the factors whose values are hard-to-change or very-hard-to-change. In the Factors outline, you can identify factors as having Changes that are Easy, Hard, or Very Hard. When the Custom Design platform constructs the design table, the Factor Changes property is assigned to every factor that appears in the Factors outline.

The Custom Design platform is the only platform that constructs designs for factors with Changes that are Hard or Very Hard. Other DOE platforms also assign the Factor Changes column property to factors that they construct, but the value of the column property is set to Easy for their factors.

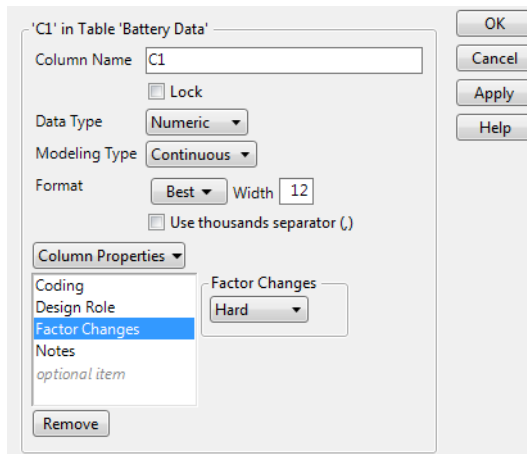
If you Load Factors in the Custom Design window using a table of factors, you can assign the Factor Changes column property to columns in that table. When you Load Factors using that table, your Factor Changes specifications appear in the Factors outline.

Factor Changes Example

The Battery Data.jmp sample data table, found in the Design Experiment folder, contains data from an experiment that studies the open current voltage of batteries (OCV). The design is a two-way split-plot design. For further background, see the [“Examples of Custom Designs”](#) chapter on page 200.

1. Select **Help > Sample Data Library** and open Design Experiment/Battery Data.jmp.
2. Click the asterisk to the right of the factor C1 in the columns panel.
3. Select **Factor Changes**.

Figure A.25 Factor Changes Panel for C1



The value of Factor Changes for C1 is Hard. Figure A.24 shows that the value of Factor Changes for A1 is Very Hard.

4. Click **OK**.
5. In the data table, click the green triangle next to the **DOE Dialog** script.
6. Open the **Factors** outline.

Figure A.26 Factors Outline for Battery Experiment

Factors				
Add Factor ▼ Remove Add N Factors 1				
Name	Role	Changes	Values	
A1	Continuous	Very Hard	-1	1
A2	Continuous	Very Hard	-1	1
A3	Continuous	Very Hard	-1	1
A4	Continuous	Very Hard	-1	1
C1	Continuous	Hard	-1	1
C2	Continuous	Hard	-1	1

The factors A1, A2, A3, and A4 have Changes set to Very Hard, and the factors C1 and C2 have Changes set to Hard. When the Custom Design platform constructs the design table, it saves these specifications to the appropriate columns as Factor Changes column properties.

In the Design outline, notice the Whole Plots and Subplots factors.

Figure A.27 Design Outline Partial View

Design									
Run	Whole Plots	Subplots	A1	A2	A3	A4	C1	C2	
1	1	1	-1	-1	1	-1	-1	-1	
2	1	2	-1	-1	1	-1	1	-1	
3	1	3	-1	-1	1	-1	-1	1	
4	2	4	-1	1	1	1	1	1	
5	2	5	-1	1	1	1	-1	1	
6	2	6	-1	1	1	1	1	-1	
7	3	7	1	1	1	-1	-1	1	
8	3	8	1	1	1	-1	1	1	
9	3	9	1	1	1	-1	-1	-1	
10	4	10	1	-1	1	1	1	1	
11	4	11	1	-1	1	1	1	-1	
12	4	12	1	-1	1	1	-1	-1	

To account for the factor changes that are Hard and Very Hard, two factors are created by the Custom Design platform. The Whole Plots factor groups the runs where the Very Hard factor levels are constant and the Subplots factor groups the runs where the Hard factors levels are constant. These factors need to be included as model effects when you enter columns with the Factor Changes column property in the Evaluate Design and Augment Design platforms.

Value Order

The Value Order column property assigns an order to the values in a column. That order is then used in plots and analyses. You can specify the order in which you want values to appear in reports.

Note: For certain values that have a natural ordering, such as days of the week, JMP automatically orders these in the appropriate way in reports. See The Column Info Window chapter in *Using JMP*.

Figure A.28 shows the Value Order panel for the Type column in the Car Physical Data.jmp sample data table. Reports that involve the values of Type place these levels in the order Sporty, Small, Compact, Medium, and Large. Use the buttons to the right of the Value Order list to specify your desired ordering for the values.

Figure A.28 Value Order Column Property for Type

'Type' in table 'Car Physical Data'

Column Name:

☐ Lock

Data Type:

Modeling Type:

Column Properties:

Remove

Value Order

Specify the order of data items and/or pick rules that determine the order.

☒ Custom Order

Custom Order

Sporty
Small
Compact
Medium
Large

Sorted Order

Custom Order(5)

Sorted Order(0)

Rules for sorted items:

☐ Row Order Levels
☒ Common Order
☒ Numerical Order

☒ Use Locale Comparisons for Characters

OK
Cancel
Apply
Help

In designs created using most DOE platforms, categorical factors, including the constructed factors Whole Plots and Subplots, and blocking factors are assigned the Value Order property. This property orders the levels according to the order in which they appear in the Factors outline. The levels of constructed factors are consecutive integers and the Value Order property specifies this natural ordering. You can modify the Value Order specification for any factor to meet your needs.

The Value Order property is not assigned by the Covering Array or Taguchi Arrays platforms. The Covering Array platform assigns the Value Labels column property. See ["Value Labels"](#) on page 763.

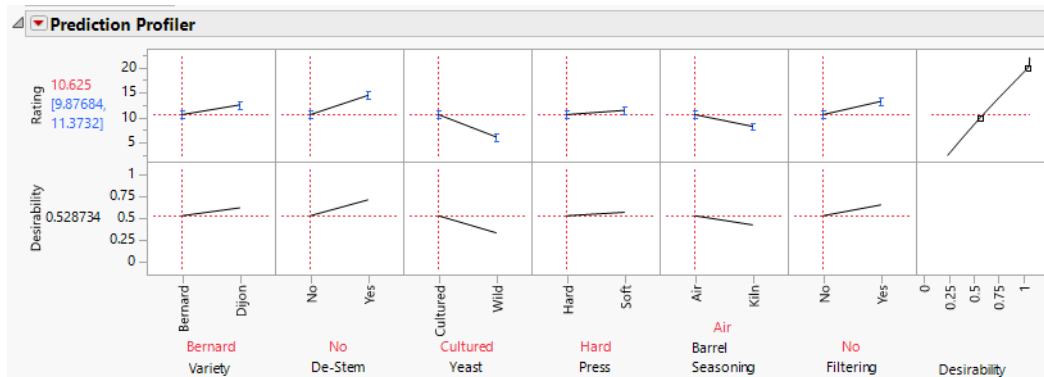
Value Order Example

Suppose that you want the values for a factor to appear in a different order in the Prediction Profiler. Consider an example of a wine tasting experiment, constructed using Custom Design. Wine is rated by five experts, each listed as a Rater in the Wine Data.jmp sample data table. Rater is a fixed blocking factor. Nine factors are studied. Rating is the response.

1. Select **Help > Sample Data Library** and open Design Experiment/Wine Data.jmp.
2. In the Table panel, click the green triangle next to the **Reduced Model** script.
3. Click **Run**.

The Prediction Profiler appears at the bottom of the report.

Figure A.29 Profiler with Original Value Order



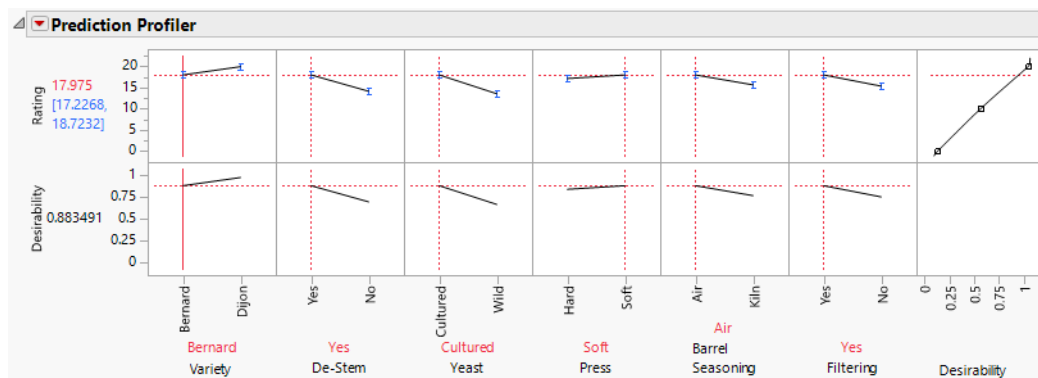
Notice that the values for De-Stem and Filtering appear in the order No followed by Yes. You want to reverse these, so that the Yes level appears first.

Tip: To arrange the profiler click the Prediction Profiler red triangle and select **Appearance > Arrange in rows**. For this example enter 7 and click **OK**.

4. Close the Response Rating report.
5. In the data table, click the asterisk next to De-Stem in the columns panel and select Value Order.
6. Click **Reverse**.
7. Click **OK**.
8. Click the asterisk next to Filtering in the columns panel and select Value Order.
9. Click **Reverse**.
10. Click **OK**.
11. Again, click the green triangle next to the **Reduced Model** script.

12. Click **Run**.

Figure A.30 Profiler with New Value Orders



The levels for De-Stem and Filtering now appear in the order Yes followed by No.

13. Close the Wine Data.jmp sample data table without saving the changes.

Assigning Value Order

Consider the Candy Bars.jmp sample data table. Of the 18 brands lists under Brand, Hershey and M&M/Mars have the largest numbers of types of candy as listed in the Name column. You want these two brands to appear first in reports.

1. Select **Help > Sample Data Library** and open Candy Bars.jmp.
2. Select the Brand column.
3. Select **Cols > Column Info**.
4. Under Column Properties, select Value Order.
5. In the Value Order list, select Hershey.
6. Click the **up arrow** five times.
Hershey is now first in the Value Order list.
7. In the Value Order list, select M&M/Mars.
8. Click the **up arrow** seven times.
M&M/Mars is now second in the Value Order list.
9. Click **OK**.

An asterisk indicating the Value Order column property appears next to Brand in the columns panel. JMP now lists Hershey and M&M/Mars first in reports involving Brand.

10. Select **Analyze > Distribution**.

11. Select **Calories** and click **Y, Columns**.
12. Select **Brand** and click **By**.
13. Click **OK**.
14. Press **Ctrl**, click the **Calories** red triangle, and select **Display Options > Horizontal Layout**.
Note that the Distribution reports for **Hershey** and **M&M/Mars** appear first among the 18 brands.
15. Close the **Candy Bars.jmp** sample data table without saving the changes.

Value Labels

The Value Labels column property represents values in a column with specified labels. These labels are displayed in the data table and are used in plots and reports. In the data table, you can view the original values by double-clicking within a cell. For more information about how to assign and work with the Value Labels column property, see The Column Info Window chapter in *Using JMP*.

The Covering Arrays platform is the only DOE platform that assigns the Value Labels column property. The Covering Arrays platform saves factors to the data table with a Nominal modeling type. The underlying values are consecutive integers ranging from 1 to the number of levels that you specify in the Covering Array Factors outline. The Values that you specify in the Factors outline are the Value Labels that are assigned to the underlying integers.

Value Labels Example

You want to test an internet-based software application to detect issues arising from components in the operating environment. The four components of interest consist of a browser, the operating system, the computer's RAM, and the connection speed. To minimize testing time, you restrict yourself to testing two-way interactions.

Construct a Strength 2 covering array to test the required combinations of factor levels.

1. Select **DOE > Special Purpose > Covering Array**.
2. Select **Help > Sample Data Library** and open **Design Experiment/Software Factors.jmp**.
3. Click the Covering Array red triangle and select **Load Factors**.

The factors and their levels appear in the Factors outline.

Figure A.31 Factors Outline for Software Factors

Factors

Load Design Add Factor Remove Add N Factors 1 Strength: t = 2

Name	Role	Values
Web Browser	Categorical	Safari IE Firefox Chrome Other
Operating System	Categorical	Macintosh Windows
RAM	Categorical	16 MB 4 MB 8 MB
Connection Speed	Categorical	0-1 Mbps 1-5 Mbps >5 Mbps

Notice that the Role of the four factors is described as Categorical.

4. Click **Continue**.
- The Restrict Factor Level Combinations outline opens. Since all combinations of settings are possible, leave this set to **None**.
5. Click **Make Design**.
6. Click **Make Table**.
7. In the columns panel, click the asterisk next to Web Browser and select **Value Labels**.

Figure A.32 Column Info Window for Factor A

'Web Browser' in Table 'Covering Array'

Column Name Web Browser

☐ Lock

Data Type Numeric

Modeling Type Nominal

Format Best Width 12

☐ Use thousands separator (,)

Column Properties

Value Labels
optional item

Remove

Value Labels

If a column has value labels, and Use Value Labels is checked, the labels are displayed wherever the column data are displayed.

1 = Safari
2 = IE
3 = Firefox
4 = Chrome
5 = Other
optional item

☐ Allow Ranges

Value

Label

☒ Use Value Labels

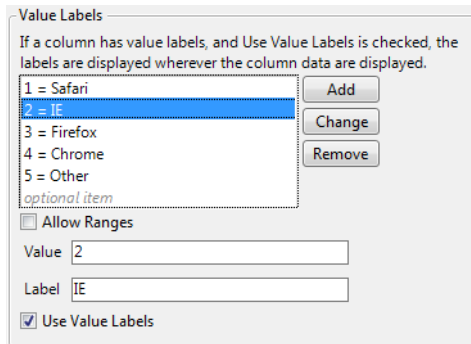
OK
Cancel
Apply
Help

Notice that Web Browser has a Numeric data type and a Nominal modeling type. The underlying values of Web Browser are the integers 1, 2, 3, 4, and 5. These values are assigned value labels corresponding to the values that you entered when you constructed the design.

Change the value label for 2 from IE to Internet Explorer.

8. Select **2 = IE** in the Value Labels list.

Figure A.33 Value Labels Panel with Selection



9. Type **Internet Explorer** next to Label.
10. Click **Change**.

The change appears in the data table.

Note: To use the numeric values and not the labels, deselect **Use Value Labels**.

RunsPerBlock

When you use the DOE platforms to construct a design containing a blocking factor, the factor is assigned the Design Role column property with the value Blocking. JMP also assigns the RunsPerBlock column property to each Blocking factor. The RunsPerBlock property indicates the maximum allowable number of runs in each block. This property is used by the Evaluate Design and Augment Design platforms to indicate the blocking structure for the factor. See [“Blocking”](#) on page 90 in the “Custom Designs” chapter.

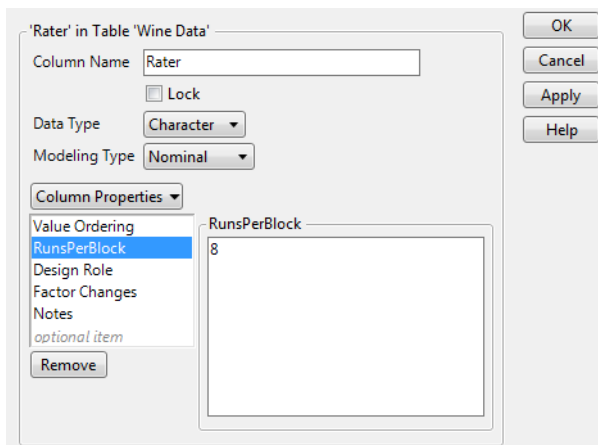
Note: The RunsPerBlock column property is assigned by JMP as part of design construction. You cannot directly assign this column property.

RunsPerBlock Example

Consider the wine tasting experiment described in “[Example of a Custom Design](#)” on page 68 in the “Custom Designs” chapter. Wine samples are tasted by five raters (Rater) and each rater tastes eight samples.

1. Select **Help > Sample Data Library** and open Design Experiment/Wine Data.jmp.
2. Click the asterisk next to Rater in the columns panel and select **Design Role**.
Notice that the Design Role is set to Blocking.
3. Click **Cancel**.
4. Click the asterisk next to Rater in the columns panel and select **RunsPerBlock**.

Figure A.34 RunsPerBlock Column Property Panel for Rater



Notice that the value of RunsPerBlock is 8. The design constructed by the DOE Dialog script has 40 runs. Since there are five raters, JMP constructs a design with $40/5 = 8$ runs for each rater.

ConstraintState

In the Custom and Mixture Design platforms, you can Save Constraints that you specify for a design. When you select Save Constraints, the coefficients of each linear constraint appear in a column in a data table. The value that bounds the inequality is given in the last row of the table.

Each constraint column is assigned the ConstraintState column property. This property specifies the direction of the inequality that defines the constraint. When you select Load Constraints from a design platform, the ConstraintState column property tells JMP the direction of the inequality.

Note: The ConstraintState column property is assigned by JMP as part of design construction. You cannot directly assign this column property.

ConstraintState Example

The sample data table Piepel.jmp, located in the Design Experiment folder, contains a mixture design with three continuous factors. The design is based on an experiment presented in Snee (1979) and Piepel (1988), where there are boundary constraints on each factor and three additional linear constraints. In the following example, you do the following:

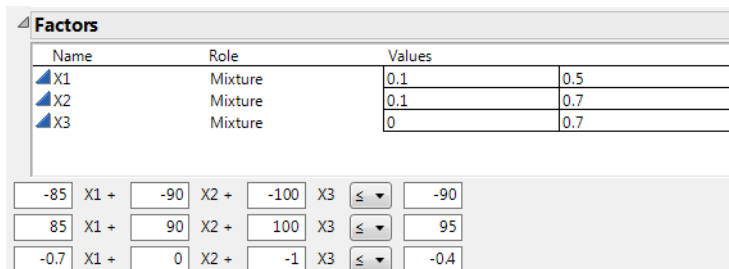
1. Change one of the three additional constraints
2. Save the constraints to a table
3. Observe how the ConstraintState column property describes the direction of the inequality in the constraint

For further development of this example, see [“An Extreme Vertices Example with Linear Constraints”](#) on page 407 in the “Mixture Designs” chapter.

1. Select **Help > Sample Data Library** and open Design Experiment/Piepel.jmp.
2. In the Table panel, click the green triangle next to the **DOE Dialog** script.

Notice the three linear constraints below the Factors outline. To make the constraints more interpretable, you want to reformulate the first constraint in terms of a “greater than or equal to” inequality.

Figure A.35 Linear Constraints beneath Factors Outline



Name	Role	Values
X1	Mixture	0.1 0.5
X2	Mixture	0.1 0.7
X3	Mixture	0 0.7

-85	X1 +	-90	X2 +	-100	X3	≤	-90
85	X1 +	90	X2 +	100	X3	≤	95
-0.7	X1 +	0	X2 +	-1	X3	≤	-0.4

3. In the first constraint, do the following:
 - Type 85 next to X1.

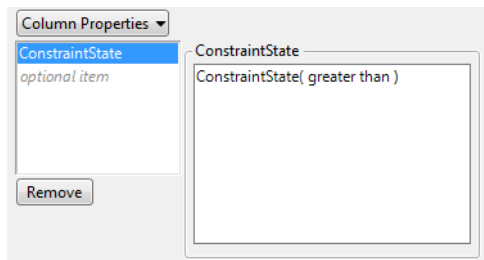
- Type 90 next to X2.
 - Type 100 next to X3.
 - Select $\geq$ from the inequality menu.
 - Type 90 to the right of the inequality sign.
4. Click the Mixture Design red triangle and select **Save Constraints**.
 A table containing information about the constraints appears.

Figure A.36 Constraint Table

	Constraint 1	Constraint 2	Constraint 3
1	-85	85	-0.7
2	-90	90	0
3	-100	100	-1
4	-90	95	-0.4

Each column contains the coefficients of the factors X1, X2, and X3 in rows 1 through 3. Row 4 contains the value that appeared to the right of the inequality sign.

5. Click the asterisk next to Constraint 1.
6. Click **ConstraintState**.

Figure A.37 ConstraintState Column Property Panel


The ConstraintState panel for X1 indicates that the direction of the inequality is “greater than” indicating greater than or equal to $\geq$.

7. Click **Cancel**.

Appendix **B**

Technical Details

This appendix contains information on the Alias Matrix, power calculations and the Relative Prediction Variance.

- [“The Alias Matrix”](#) on page 771
- [“Power Calculations”](#) on page 772
- [“Relative Prediction Variance”](#) on page 775

Contents

The Alias Matrix 771

 Designs with Hard or Very Hard Factor Changes 771

 Designs with If Possible Effects 772

Power Calculations 772

 Power for a Single Parameter 772

 Power for a Categorical Effect 774

Relative Prediction Variance 775

The Alias Matrix

The Alias Matrix entries represent the degree of bias imparted to model parameters by the effects that you specified in the Alias Terms outline. The Alias Matrix is also used in defining alias optimality. See [“Alias Optimality”](#) on page 131 in the “Custom Designs” chapter.

Calculations for the Alias Matrix are based on the model matrix. See [“Model Matrix”](#) on page 113 in the “Custom Designs” chapter.

Let $\mathbf{X}_1$ be the model matrix corresponding to the terms in the Model outline. Denote the matrix of model terms for the effects specified in the Alias Terms outline by $\mathbf{X}_2$.

The assumed model is given as follows:

$$\mathbf{Y} = \mathbf{X}_1\beta_1 + \varepsilon$$

Suppose that some of the alias terms are active and that the true model is given as follows:

$$\mathbf{Y} = \mathbf{X}_1\beta_1 + \mathbf{X}_2\beta_2 + \varepsilon$$

The least squares estimator of β_1 is given by:

$$\hat{\beta}_1 = (\mathbf{X}_1'\mathbf{X}_1)^{-1}\mathbf{X}_1'\mathbf{Y}$$

Under the usual regression assumptions, the expected value of $\hat{\beta}_1$ is given by:

$$E[\hat{\beta}_1] = \beta_1 + \mathbf{A}\beta_2$$

where $\mathbf{A} = (\mathbf{X}_1'\mathbf{X}_1)^{-1}\mathbf{X}_1'\mathbf{X}_2$.

The matrix $\mathbf{A}$ is called the *alias matrix*.

Designs with Hard or Very Hard Factor Changes

For designs with hard-to-change or very-hard-to-change factors, the alias matrix is given as follows:

$$\mathbf{A} = (\mathbf{X}_1'\mathbf{V}^{-1}\mathbf{X}_1)^{-1}\mathbf{X}_1'(\mathbf{V}^{-1}\mathbf{X}_2)$$

where $\mathbf{V}$ is the block diagonal covariance matrix of the responses.

Designs with If Possible Effects

For designs with If Possible effects (Bayesian D- and I-optimal designs), the alias matrix is given as follows:

$$\mathbf{A} = (\mathbf{X}_1' \mathbf{X}_1 + \mathbf{K}^2)^{-1} \mathbf{X}_1' \mathbf{X}_2$$

where $\mathbf{K}$ is a diagonal matrix with these values:

- $k = 0$ for Necessary terms
- $k = 1$ for If Possible main effects, powers, and interactions involving a categorical factor with more than two levels
- $k = 4$ for all other If Possible terms

In the Bayesian case, the alias matrix gives the aliasing of effects corresponding to a ridge regression with a prior variance of $\mathbf{K}^2$. For additional detail on Bayesian designs, see [“Bayesian D-Optimality”](#) on page 128 in the “Custom Designs” chapter and [“Bayesian I-Optimality”](#) on page 130 in the “Custom Designs” chapter.

In the Custom Design platform, you can control the weights used for If Possible terms by selecting **Advanced Options > Prior Parameter Variance** from the red triangle menu. There you can set prior variances for all model terms by specifying the diagonal elements of $\mathbf{K}$. The option updates to show the default weights when you click Make Design.

Power Calculations

The Power Analysis report gives power calculations for single parameter values and, when the design includes a categorical factor with three or more levels, for whole effects. This section describes the calculations in the two cases:

- [“Power for a Single Parameter”](#) on page 772
- [“Power for a Categorical Effect”](#) on page 774

Power for a Single Parameter

This section describes how power for the test of a single parameter is computed. Use the following notation:

- X** The model matrix. See the Standard Least Squares Report and Options chapter in *Fitting Linear Models* for information on the coding for nominal effects. Also, See [“Model Matrix”](#) on page 113 in the “Custom Designs” chapter.

Note: You can view the model matrix by running Fit Model. Then select **Save Columns > Save Coding Table** from the red triangle menu for the main report.

β_i The parameter corresponding to the term of interest.

$\hat{\beta}_i$ The least squares estimator of β_i

β_i^A The Anticipated Coefficient value. The difference you want to detect is $2\beta_i^A$.

The variance of $\hat{\beta}_i$ is given by the i^{th} diagonal entry of $\sigma^2(\mathbf{X}'\mathbf{X})^{-1}$, where σ^2 is the error variance. Denote the i^{th} diagonal entry of $(\mathbf{X}'\mathbf{X})^{-1}$ by $(\mathbf{X}'\mathbf{X})_{ii}^{-1}$.

The error variance, σ^2 , is estimated by the MSE, and has $n - p - 1$ degrees of freedom, where n is the number of observations and p is the number of terms other than the intercept in the model. If $n - p - 1 = 0$, then JMP sets the degrees of freedom for the error to 1. This allows the power to be estimated for parameters in a saturated design.

The test of $H_0: \beta_i = 0$ is given by:

$$\frac{\hat{\beta}_i}{\sqrt{MSE(\mathbf{X}'\mathbf{X})_{ii}^{-1}}}$$

or equivalently by:

$$F_0 = \frac{\hat{\beta}_i^2}{MSE(\mathbf{X}'\mathbf{X})_{ii}^{-1}}$$

Under the null hypothesis, the test statistic F_0 has an F distribution on 1 and $n - p - 1$ degrees of freedom.

If the true value of β_i is β_i^A , then F_0 has a noncentral F distribution with noncentrality parameter given by:

$$\lambda = \frac{(\beta_i^A)^2}{\sigma^2(\mathbf{X}'\mathbf{X})_{ii}^{-1}}$$

To compute the power of the test, first solve for the α -level critical value F_c :

$$\alpha = 1 - FDist(F_c, 1, n - p - 1)$$

Then calculate the power as follows:

$$Power = 1 - FDist(F_c, 1, n - p - 1, \lambda)$$

Power for a Categorical Effect

This section describes how power for the test for a whole categorical effect is computed. Use the following notation:

X Model matrix. See [“The Alias Matrix”](#) on page 771 in the “Technical Details” appendix.

β Vector of parameters.

$\hat{\beta}$ Least squares estimate of β .

β^A Vector of Anticipated Coefficient values.

L Matrix that defines the test for the categorical effect. The matrix **L** identifies the values of the parameters in β corresponding to the categorical effect and sets them equal to 0. The null hypothesis for the test of the categorical effect is given by:
 $H_0: L\beta = \mathbf{0}$

r Rank of **L**. Alternatively, r is the number of levels of the categorical effect minus one.

Note: You can view the design matrix by running Fit Model. Then select **Save Columns > Save Coding Table** from the red triangle menu for the main report.

The covariance matrix of $\hat{\beta}$ is given by $\sigma^2(\mathbf{X}'\mathbf{X})^{-1}$, where σ^2 is the error variance.

The error variance, σ^2 , is estimated by the MSE, and has $n - p - 1$ degrees of freedom, where n is the number of observations and p is the number of terms other than the intercept in the model. If $n - p - 1 = 0$, then JMP sets the degrees of freedom for the error to 1. This allows the power to be estimated for parameters in a saturated design.

The test of $H_0: L\beta = \mathbf{0}$ is given by:

$$F_0 = \left((\mathbf{L}\hat{\beta})' [\mathbf{L}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{L}']^{-1} (\mathbf{L}\hat{\beta}) \right) / (rMSE)$$

Under the null hypothesis, the test statistic F_0 has an F distribution on r and $n - p - 1$ degrees of freedom.

If the true value of β is β^A , then F_0 has a noncentral F distribution with noncentrality parameter given by:

$$\lambda = \left((\mathbf{L}\beta^A)' [\mathbf{L}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{L}']^{-1} (\mathbf{L}\beta^A) \right) / \sigma^2$$

To compute the power of the test, first solve for the α -level critical value F_c :

$$\alpha = 1 - FDist(F_c, r, n - p - 1)$$

Then calculate the power as follows:

$$Power = 1 - FDist(F_c, r, n - p - 1, \lambda)$$

Relative Prediction Variance

Consider the following notation:

X Model matrix. See “[Model Matrix](#)” on page 113 in the “Custom Designs” chapter. Custom designs provide a script that shows the model matrix. See “[Save X Matrix](#)” on page 112 in the “Custom Designs” chapter.

σ^2 Error variance.

$\hat{\beta}$ Vector of least squares estimates of the parameters.

$\mathbf{x}_i'$ The i^{th} row of **X**.

Using this notation, the predicted response for the i^{th} row of **X** is given by:

$$\hat{Y} = \mathbf{x}_i' \hat{\beta}$$

The relative prediction variance at the settings defined by $\mathbf{x}_i$ is given by:

$$\frac{\mathbf{x}_i' var(\hat{Y}) \mathbf{x}_i}{\sigma^2} = \frac{\mathbf{x}_i' var(\mathbf{x}_i' \hat{\beta}) \mathbf{x}_i}{\sigma^2} = \mathbf{x}_i' (\mathbf{X}' \mathbf{X})^{-1} \mathbf{x}_i$$

Appendix **C**

References

- Agresti, A., and Coull, B. A. (1998). "Approximate Is Better than 'Exact' for Interval Estimation of Binomial Proportions." *American Statistician* 52:119–126.
- Agresti, A., and Caffo, B. N. (2000). "Simple and Effective Confidence Intervals for Proportions and Differences of Proportions Result from Adding Two Successes and Two Failures." *American Statistician* 54:280–288.
- Atkinson, A. C., and Donev, A. N. (1992). *Optimum Experimental Designs*. New York: Oxford University Press.
- Barker, C. (2011). "Power and Sample Size Calculations in JMP." Technical Report, Cary NC: SAS Institute Inc. Accessed May 6, 2018.
<https://www.jmp.com/content/dam/jmp/documents/en/technical-reports/powerAndSampleSize.pdf>.
- Borror, C. M., and Montgomery, D. C. (2000). "Mixed Resolution Designs as Alternatives to Taguchi Inner-Outer Array Designs for Robust Design Problems." *Quality and Reliability Engineering International* 16:117–128.
- Box, G. E. P. (1988). "Signal-to-Noise Ratio, Performance Criteria, and Transformations." *Technometrics* 30:1–40.
- Box, G. E. P., and Behnken, D. W. (1960). "Some New Three-Level Designs for the Study of Quantitative Variables." *Technometrics* 2:455–475.
- Box, G. E. P., and Draper, N. R. (1987). *Empirical Model-Building and Response Surfaces*. New York: John Wiley & Sons.
- Box, G. E. P., Hunter, W. G., and Hunter, J. S. (1978). *Statistics for Experimenters*. New York: John Wiley & Sons.
- Box, G. E. P., Hunter, W. G., and Hunter, J. S. (2005). *Statistics for Experimenters: Design, Innovation, and Discovery*. 2nd ed. New York: John Wiley & Sons.
- Box, G. E. P., and Meyer, R. D. (1986). "An Analysis of Unreplicated Fractional Factorials." *Technometrics* 28:11–18.
- Box, G. E. P., and Wilson, K. B. (1951). "On the Experimental Attainment of Optimum Conditions." *Journal of the Royal Statistical Society, Series B* 13:1–45.
- Byrne, D. M., and Taguchi, G. (1986). "The Taguchi Approach to Parameter Designs." *Quality Congress Transactions* 177:168–177.
- Clopper, C. J., and Pearson, E. S. (1934). "The Use of Confidence or Fiducial Limits Illustrated in the Case of the Binomial." *Biometrika* 26:404–413.

- Cochran, W. G., and Cox, G. M. (1957). *Experimental Designs*. 2nd ed. New York: John Wiley & Sons.
- Cohen, M. B., Gibbons, P. B., Mugridge, W. B., and Colbourn, C. J. (2003). "Constructing Test Suites for Interaction Testing." In *Proceedings 25th International Conference on Software Engineering*, 38–48.
- Cohen, M. B., Dwyer, M. B., and Shi, J. (2007). "Interaction Testing of Highly-Configurable Systems in the Presence of Constraints." In *Proceedings of the 2007 International Symposium on Software Testing and Analysis*, 129–139.
- Colbourn, C. (2004). "Combinatorial Aspects of Covering Arrays." *Le Matematiche* 58:121–167.
- Colbourn, C., and Dinitz, J., eds. (2010). *Handbook of Combinatorial Designs*. 2nd ed. Boca Raton, FL: CRC Press.
- Cornell, J. A. (1990). *Experiments with Mixtures*. 2nd ed. New York: John Wiley & Sons.
- Dalal, S. R., and Mallows, C. L. (1998). "Factor-Covering Designs for Testing Software." *Technometrics* 40:234–243.
- DuMouchel, W., and Jones, B. (1994). "A Simple Bayesian Modification of *D*-Optimal Designs to Reduce Dependence on an Assumed Model." *Technometrics* 36:37–47.
- Errore, A., Jones, B., Li, W., and Nachtsheim, C. (2017). "Using Definitive Screening Designs to Identify Active First- and Second-Order Factor Effects." *Journal of Quality Technology* 49:244–264.
- Fries, A., and Hunter, W. G. (1984). "Minimum Aberration 2^{k-p} Designs." *Technometrics* 22:601–608.
- Gotwalt, C. M., Jones, B. A., and Steinberg, D. M., (2009). "Fast Computation of Designs Robust to Parameter Uncertainty for Nonlinear Settings." *Technometrics* 51:88–95.
- Goos, P. (2002). *The Optimal Design of Blocked and Split Plot Experiments*. New York: Springer.
- Goos, P., and Jones, B. (2011). *Optimal Design of Experiments: A Case Study Approach*. New York: John Wiley & Sons.
- Hamada, M., and Wu, C. F. J., (1992). "Analysis of Designed Experiments with Complex Aliasing." *Technometrics* 24:130–137.
- Hartman, A., and Raskin, L. (2004). "Problems and Algorithms for Covering Arrays." *Discrete Mathematics* 284:149–156.
- Horn, R., and Johnson, C. (2012). *Matrix Analysis*. 2nd ed. Cambridge: Cambridge University Press.
- Huber, J., and Zwerina, K. (1996). "The Importance of Utility Balance in Efficient Choice Designs." *Journal of Marketing Research* 33:307–317.
- Jennrich, R. I. (1969). "Asymptotic Properties of Non-Linear Least Squares Estimators." *Annals of Mathematical Statistics* 40:633–643.
- Johnson, M. E., and Nachtsheim, C. J. (1983). "Some Guidelines for Constructing Exact *D*-Optimal Designs on Convex Design Spaces." *Technometrics* 25:271–277.

- Jones, B. (1991). "An Interactive Graph For Exploring Multidimensional Response Surfaces." In *Proceedings of the Annual Joint Statistical Meetings of the American Statistical Association*. Alexandria, VA: American Statistical Association.
- Jones, B., and Goos, P. (2007). "A Candidate-set-free Algorithm for Generating D-optimal Designs." *Applied Statistics* 56:1–18. Accessed January 23, 2017.
<https://www.ncbi.nlm.nih.gov/pmc/articles/PMC3001117/>.
- Jones, B., and Goos, P. (2015). "Optimal Design of Blocked Experiments in the Presence of Supplementary Information about the Blocks." *Technometrics* 47:01–317.
- Jones, B., and Johnson, R. T. (2009). "Design and Analysis for the Gaussian Process Model." *Quality and Reliability Engineering International* 25:515–524. Accessed January 23, 2017.
https://www.stat.osu.edu/~comp_exp/jour.club/Jones_Johnson_QREI09.pdf.
- Jones, B., Lekivetz, R., Majumdar, D., Nachtsheim, C. J., and Stallings, J. W. (2019). "Construction, Properties, and Analysis of Group-Orthogonal Supersaturated Designs." *Technometrics*. Accessed September 24, 2019.
<https://doi.org/10.1080/00401706.2019.1654926>
- Jones, B., Lin, D. K. J., and Nachtsheim, C. J. (2008). "Bayesian D-Optimal Supersaturated Designs." *Journal of Statistical Planning and Inference* 138:86–92.
- Jones, B., and Nachtsheim, C. J. (2011a). "A Class of Three-Level Designs for Definitive Screening in the Presence of Second-Order Effects." *Journal of Quality Technology* 43:1–15. Accessed January 23, 2017.
https://www.jmp.com/en_us/whitepapers/jmp/class-three-level-definitive-screening.html.
- Jones, B., and Nachtsheim, C. J. (2011b). "Efficient Designs with Minimal Aliasing." *Technometrics* 53:62–71.
- Jones, B., and Nachtsheim, C. J. (2013). "Definitive Screening Designs with Added Two-Level Categorical Factors." *Journal of Quality Technology* 45:121–129.
- Jones, B., and Nachtsheim, C. J. (2016a). "Blocking Schemes for Definitive Screening Designs." *Technometrics* 58:74–83.
- Jones, B., and Nachtsheim, C. J. (2016b). "Effective Model Selection for Definitive Screening Designs." *Technometrics*. Accessed September 28, 2017.
<https://www.tandfonline.com/doi/full/10.1080/00401706.2016.1234979>.
- Joseph, V. R., Gul, E., and Ba, S. (2015). "Maximum Projection Designs for Computer Experiments." *Biometrika* 102:371–380.
- Kessels, R., Jones, B., Goos, P. (2011). "Bayesian Optimal Designs for Discrete Choice Experiments with Partial Profiles." *Journal of Choice Modelling* 4:52–74.
- Kowalski, S. M., Cornell, J. A., and Vining, G. G. (2002). "Split Plot Designs and Estimation Methods for Mixture Experiments with Process Variables." *Technometrics* 44:72–79.
- Lenth, R. V. (1989). "Quick and Easy Analysis of Unreplicated Fractional Factorials." *Technometrics* 31:469–473.
- Lekivetz, R. (2014). "What is an Alias Matrix?" Accessed January 23, 2017.
<https://blogs.sas.com/content/jmp/2014/06/05/what-is-an-alias-matrix/>.

- Lekivetz, R., Sitter, R., Bingham, D., Hamada, M. S., Moore, L. M., and Wendelberger, J. R. (2015). "On Algorithms for Obtaining Orthogonal and Near-Orthogonal Arrays for Main-Effects Screening." *Journal of Quality Technology* 47:2–13.
- Lin, D. K. J. (1993). "A New Class of Supersaturated Design." *Technometrics* 35:28–31.
- Martirosyan, S. (2003). "Perfect Hash Families, Identifiable Parent Property Codes and Covering Arrays." Ph.D. diss., Universitat Duisburg-Essen.
- Meyer, R. K., and Nachtsheim, C. J. (1995). The Coordinate Exchange Algorithm for Constructing Exact Optimal Designs." *Technometrics* 37:60–69.
- Meyer, R. D., Steinberg, D. M., and Box, G. (1996). Follow-up Designs to Resolve Confounding in Multifactor Experiments. *Technometrics* 38:307.
- Miller, A., and Sitter, R. R. (2005). "Using Folded-Over Nonorthogonal Designs." *Technometrics* 47:502–513.
- Montgomery, D. C. (2013). *Design and Analysis of Experiments*. 8th ed. New York: John Wiley & Sons.
- Morgan, J. (2009). "Constrained Covering Arrays: Resolving Invalid Level Combination Constraints." Technical Report, Cary NC: SAS Institute Inc., also in *Supplemental Proceedings of the 20th International Symposium on Software Reliability Engineering*. Accessed May 6, 2018. <https://www.jmp.com/content/dam/jmp/documents/en/technical-reports/ConstrainedCovArray09.pdf>.
- Morgan, J. (2014). "Computing Coverage and Diversity Metrics for Constrained Covering Arrays," Technical Report, Cary NC: SAS Institute Inc. Accessed May 6, 2018. <https://www.jmp.com/content/dam/jmp/documents/en/technical-reports/CoverageDiversity2014.pdf>.
- Morgan, J. P., and Stallings, J. (2017). "Optimal experimental design that targets meaningful information." *WIREs Comput Stat* 2017, 9:e1393. doi: 10.1002/wics.1393.
- Myers, R. H., Montgomery, D. C., and Anderson-Cook, C. M. (2009). *Response Surface Methodology, Process and Product Optimization Using Designed Experiments*. 3rd ed. New York: John Wiley & Sons.
- Mylona, K., Goos, P., and Jones, B. (2014). "Optimal Design of Blocked and Split-plot Experiments for Fixed Effects and Variance Component Estimation." *Technometrics* 56:132–144.
- Nayeri, P., Colbourn, C., and Konjevod, G. (2013). "Randomized Postoptimization of Covering Arrays." *European Journal of Combinatorics* 34:91–103.
- Piepel, G. F. (1988). "Programs for Generating Extreme Vertices and Centroids of Linearly Constrained Experimental Regions." *Journal of Quality Technology* 20:125–139.
- Sall, J. (2008). "Choice Experimental Designs Are Different." Accessed January 22, 2016. <https://blogs.sas.com/content/jmp/2008/12/16/choice-experimental-designs-are-different/>.
- SAS Institute Inc. (2013). "The %MktBIBD Macro." Accessed September 27, 2017. <https://support.sas.com/rnd/app/macros/mktbibd/mktbibd.htm>.
- Schoen, E. D. (1999). "Designing Fractional Two-level Experiments with Nested Error Structures." *Journal of Applied Statistics* 26: 495–508.

- Shewry, M. C., and Wynn, H. P. (1987). "Maximum Entropy Sampling." *Journal of Applied Statistics* 14:165–170.
- Snee, R. D., and Marquardt, D. W. (1974). "Extreme Vertices Designs for Linear Mixture Models." *Technometrics* 16:391–408.
- Snee, R. D. (1975). "Experimental Designs for Quadratic Models in Constrained Mixture Spaces." *Technometrics* 17:149–159.
- Snee, R. D. (1979). "Experimental Designs for Mixture Systems with Multicomponent Constraints." *Communications in Statistics -Theory and Methods* 8:303–326.
- Snee, R. D. (1985). "Computer Aided Design of Experiments - Some Practical Experiences." *Journal of Quality Technology* 17:222–236.
- Vivacqua, C. A., and Bisgaard, S. (2004). "Strip-Block Experiments for Process Improvement and Robustness." *Quality Engineering* 16:495–500.
- Worley, B. A. (1987). *Deterministic Uncertainty Analysis*. Technical Report ORN-0628, Oak Ridge National Laboratory.
- Wu, C. F. J. (1981). "Asymptotic Theory of Nonlinear Least Squares Estimation." *Annals of Mathematical Statistics* 9:501–513.
- Wu, C. F. J., and Hamada, M. S. (2009). *Experiments: Planning, Analysis, and Optimization*. 2nd ed. New York: John Wiley & Sons.
- Xiao, L., Lin, D. K. J., and Bai, F. (2012). "Constructing Definitive Screening Designs Using Conference Matrices." *Journal of Quality Technology* 44:2–8.
- Ye, K. Q., and Hamada, M. (2000). "Critical Values of the Lenth Method for Unreplicated Factorial Designs." *Journal of Quality Technology* 32:57–66.
- Yilmaz, C., Fouche, S., Cohen, M. B., Porter, A., Demiroz, G., and Koc, U. (2014). "Moving Forward with Combinatorial Interaction Testing." *Computer* 47:37–45.
- Zahran, A., Anderson-Cook, C. M., and Myers, R. H. (2003). "Fraction of Design Space to Assess the Prediction Capability of Response Surface Designs." *Journal of Quality Technology* 35:377–386.
- Zhang, Z., and Zhang, J. (2011). "Characterizing Failure-Causing Parameter Interactions by Adaptive Testing." In *Proceedings of the 2011 International Symposium on Software Testing and Analysis*, 331–341.

Technology License Notices

- Scintilla - Copyright © 1998-2017 by Neil Hodgson <neilh@scintilla.org>.

All Rights Reserved.

Permission to use, copy, modify, and distribute this software and its documentation for any purpose and without fee is hereby granted, provided that the above copyright notice appear in all copies and that both that copyright notice and this permission notice appear in supporting documentation.

NEIL HODGSON DISCLAIMS ALL WARRANTIES WITH REGARD TO THIS SOFTWARE, INCLUDING ALL IMPLIED WARRANTIES OF MERCHANTABILITY AND FITNESS, IN NO EVENT SHALL NEIL HODGSON BE LIABLE FOR ANY SPECIAL, INDIRECT OR CONSEQUENTIAL DAMAGES OR ANY DAMAGES WHATSOEVER RESULTING FROM LOSS OF USE, DATA OR PROFITS, WHETHER IN AN ACTION OF CONTRACT, NEGLIGENCE OR OTHER TORTIOUS ACTION, ARISING OUT OF OR IN CONNECTION WITH THE USE OR PERFORMANCE OF THIS SOFTWARE.

- Progress® Telerik® UI for WPF: Copyright © 2008-2019 Progress Software Corporation. All rights reserved. Usage of the included Progress® Telerik® UI for WPF outside of JMP is not permitted.
- ZLIB Compression Library - Copyright © 1995-2005, Jean-Loup Gailly and Mark Adler.
- Made with Natural Earth. Free vector and raster map data @ naturalearthdata.com.
- Packages - Copyright © 2009-2010, Stéphane Sudre (s.sudre.free.fr). All rights reserved.

Redistribution and use in source and binary forms, with or without modification, are permitted provided that the following conditions are met:

Redistributions of source code must retain the above copyright notice, this list of conditions and the following disclaimer.

Redistributions in binary form must reproduce the above copyright notice, this list of conditions and the following disclaimer in the documentation and/or other materials provided with the distribution.

Neither the name of the WhiteBox nor the names of its contributors may be used to endorse or promote products derived from this software without specific prior written permission.

THIS SOFTWARE IS PROVIDED BY THE COPYRIGHT HOLDERS AND CONTRIBUTORS "AS IS" AND ANY EXPRESS OR IMPLIED WARRANTIES, INCLUDING, BUT NOT LIMITED TO, THE IMPLIED WARRANTIES OF MERCHANTABILITY AND FITNESS FOR A PARTICULAR PURPOSE ARE DISCLAIMED. IN NO EVENT SHALL THE COPYRIGHT OWNER OR CONTRIBUTORS BE LIABLE FOR ANY DIRECT, INDIRECT, INCIDENTAL, SPECIAL, EXEMPLARY, OR CONSEQUENTIAL DAMAGES

(INCLUDING, BUT NOT LIMITED TO, PROCUREMENT OF SUBSTITUTE GOODS OR SERVICES; LOSS OF USE, DATA, OR PROFITS; OR BUSINESS INTERRUPTION) HOWEVER CAUSED AND ON ANY THEORY OF LIABILITY, WHETHER IN CONTRACT, STRICT LIABILITY, OR TORT (INCLUDING NEGLIGENCE OR OTHERWISE) ARISING IN ANY WAY OUT OF THE USE OF THIS SOFTWARE, EVEN IF ADVISED OF THE POSSIBILITY OF SUCH DAMAGE.

- iODBC software - Copyright © 1995-2006, OpenLink Software Inc and Ke Jin (www.iodbc.org). All rights reserved.

Redistribution and use in source and binary forms, with or without modification, are permitted provided that the following conditions are met:

- Redistributions of source code must retain the above copyright notice, this list of conditions and the following disclaimer.
- Redistributions in binary form must reproduce the above copyright notice, this list of conditions and the following disclaimer in the documentation and/or other materials provided with the distribution.
- Neither the name of OpenLink Software Inc. nor the names of its contributors may be used to endorse or promote products derived from this software without specific prior written permission.

THIS SOFTWARE IS PROVIDED BY THE COPYRIGHT HOLDERS AND CONTRIBUTORS “AS IS” AND ANY EXPRESS OR IMPLIED WARRANTIES, INCLUDING, BUT NOT LIMITED TO, THE IMPLIED WARRANTIES OF MERCHANTABILITY AND FITNESS FOR A PARTICULAR PURPOSE ARE DISCLAIMED. IN NO EVENT SHALL OPENLINK OR CONTRIBUTORS BE LIABLE FOR ANY DIRECT, INDIRECT, INCIDENTAL, SPECIAL, EXEMPLARY, OR CONSEQUENTIAL DAMAGES (INCLUDING, BUT NOT LIMITED TO, PROCUREMENT OF SUBSTITUTE GOODS OR SERVICES; LOSS OF USE, DATA, OR PROFITS; OR BUSINESS INTERRUPTION) HOWEVER CAUSED AND ON ANY THEORY OF LIABILITY, WHETHER IN CONTRACT, STRICT LIABILITY, OR TORT (INCLUDING NEGLIGENCE OR OTHERWISE) ARISING IN ANY WAY OUT OF THE USE OF THIS SOFTWARE, EVEN IF ADVISED OF THE POSSIBILITY OF SUCH DAMAGE.

- This program, “bzip2”, the associated library “libbzip2”, and all documentation, are Copyright © 1996-2019 Julian R Seward. All rights reserved.

Redistribution and use in source and binary forms, with or without modification, are permitted provided that the following conditions are met:

1. Redistributions of source code must retain the above copyright notice, this list of conditions and the following disclaimer.
2. The origin of this software must not be misrepresented; you must not claim that you wrote the original software. If you use this software in a product, an acknowledgment in the product documentation would be appreciated but is not required.
3. Altered source versions must be plainly marked as such, and must not be misrepresented as being the original software.

4. The name of the author may not be used to endorse or promote products derived from this software without specific prior written permission.

THIS SOFTWARE IS PROVIDED BY THE AUTHOR "AS IS" AND ANY EXPRESS OR IMPLIED WARRANTIES, INCLUDING, BUT NOT LIMITED TO, THE IMPLIED WARRANTIES OF MERCHANTABILITY AND FITNESS FOR A PARTICULAR PURPOSE ARE DISCLAIMED. IN NO EVENT SHALL THE AUTHOR BE LIABLE FOR ANY DIRECT, INDIRECT, INCIDENTAL, SPECIAL, EXEMPLARY, OR CONSEQUENTIAL DAMAGES (INCLUDING, BUT NOT LIMITED TO, PROCUREMENT OF SUBSTITUTE GOODS OR SERVICES; LOSS OF USE, DATA, OR PROFITS; OR BUSINESS INTERRUPTION) HOWEVER CAUSED AND ON ANY THEORY OF LIABILITY, WHETHER IN CONTRACT, STRICT LIABILITY, OR TORT (INCLUDING NEGLIGENCE OR OTHERWISE) ARISING IN ANY WAY OUT OF THE USE OF THIS SOFTWARE, EVEN IF ADVISED OF THE POSSIBILITY OF SUCH DAMAGE.

Julian Seward, jseward@acm.org

bzip2/libbzip2 version 1.0.8 of 13 July 2019

- R software is Copyright © 1999-2012, R Foundation for Statistical Computing.
- MATLAB software is Copyright © 1984-2012, The MathWorks, Inc. Protected by U.S. and international patents. See www.mathworks.com/patents. MATLAB and Simulink are registered trademarks of The MathWorks, Inc. See www.mathworks.com/trademarks for a list of additional trademarks. Other product or brand names may be trademarks or registered trademarks of their respective holders.
- libopc is Copyright © 2011, Florian Reuter. All rights reserved.

Redistribution and use in source and binary forms, with or without modification, are permitted provided that the following conditions are met:

- Redistributions of source code must retain the above copyright notice, this list of conditions and the following disclaimer.
- Redistributions in binary form must reproduce the above copyright notice, this list of conditions and the following disclaimer in the documentation and / or other materials provided with the distribution.
- Neither the name of Florian Reuter nor the names of its contributors may be used to endorse or promote products derived from this software without specific prior written permission.

THIS SOFTWARE IS PROVIDED BY THE COPYRIGHT HOLDERS AND CONTRIBUTORS "AS IS" AND ANY EXPRESS OR IMPLIED WARRANTIES, INCLUDING, BUT NOT LIMITED TO, THE IMPLIED WARRANTIES OF MERCHANTABILITY AND FITNESS FOR A PARTICULAR PURPOSE ARE DISCLAIMED. IN NO EVENT SHALL THE COPYRIGHT OWNER OR CONTRIBUTORS BE LIABLE FOR ANY DIRECT, INDIRECT, INCIDENTAL, SPECIAL, EXEMPLARY, OR CONSEQUENTIAL DAMAGES (INCLUDING, BUT NOT LIMITED TO, PROCUREMENT OF SUBSTITUTE GOODS OR SERVICES; LOSS OF

USE, DATA, OR PROFITS; OR BUSINESS INTERRUPTION) HOWEVER CAUSED AND ON ANY THEORY OF LIABILITY, WHETHER IN CONTRACT, STRICT LIABILITY, OR TORT (INCLUDING NEGLIGENCE OR OTHERWISE) ARISING IN ANY WAY OUT OF THE USE OF THIS SOFTWARE, EVEN IF ADVISED OF THE POSSIBILITY OF SUCH DAMAGE.

- libxml2 - Except where otherwise noted in the source code (e.g. the files hash.c, list.c and the trio files, which are covered by a similar license but with different Copyright notices) all the files are:

Copyright © 1998 - 2003 Daniel Veillard. All Rights Reserved.

Permission is hereby granted, free of charge, to any person obtaining a copy of this software and associated documentation files (the "Software"), to deal in the Software without restriction, including without limitation the rights to use, copy, modify, merge, publish, distribute, sublicense, and/or sell copies of the Software, and to permit persons to whom the Software is furnished to do so, subject to the following conditions:

The above copyright notice and this permission notice shall be included in all copies or substantial portions of the Software.

THE SOFTWARE IS PROVIDED "AS IS", WITHOUT WARRANTY OF ANY KIND, EXPRESS OR IMPLIED, INCLUDING BUT NOT LIMITED TO THE WARRANTIES OF MERCHANTABILITY, FITNESS FOR A PARTICULAR PURPOSE AND NONINFRINGEMENT. IN NO EVENT SHALL DANIEL VEILLARD BE LIABLE FOR ANY CLAIM, DAMAGES OR OTHER LIABILITY, WHETHER IN AN ACTION OF CONTRACT, TORT OR OTHERWISE, ARISING FROM, OUT OF OR IN CONNECTION WITH THE SOFTWARE OR THE USE OR OTHER DEALINGS IN THE SOFTWARE.

Except as contained in this notice, the name of Daniel Veillard shall not be used in advertising or otherwise to promote the sale, use or other dealings in this Software without prior written authorization from him.

- Regarding the decompression algorithm used for UNIX files:

Copyright © 1985, 1986, 1992, 1993

The Regents of the University of California. All rights reserved.

THIS SOFTWARE IS PROVIDED BY THE REGENTS AND CONTRIBUTORS "AS IS" AND ANY EXPRESS OR IMPLIED WARRANTIES, INCLUDING, BUT NOT LIMITED TO, THE IMPLIED WARRANTIES OF MERCHANTABILITY AND FITNESS FOR A PARTICULAR PURPOSE ARE DISCLAIMED. IN NO EVENT SHALL THE REGENTS OR CONTRIBUTORS BE LIABLE FOR ANY DIRECT, INDIRECT, INCIDENTAL, SPECIAL, EXEMPLARY, OR CONSEQUENTIAL DAMAGES (INCLUDING, BUT NOT LIMITED TO, PROCUREMENT OF SUBSTITUTE GOODS OR SERVICES; LOSS OF USE, DATA, OR PROFITS; OR BUSINESS INTERRUPTION) HOWEVER CAUSED AND ON ANY THEORY OF LIABILITY, WHETHER IN CONTRACT, STRICT LIABILITY, OR TORT (INCLUDING NEGLIGENCE OR OTHERWISE) ARISING IN ANY WAY OUT OF THE USE OF THIS SOFTWARE, EVEN IF ADVISED OF THE POSSIBILITY OF SUCH DAMAGE.

1. Redistributions of source code must retain the above copyright notice, this list of conditions and the following disclaimer.

2. Redistributions in binary form must reproduce the above copyright notice, this list of conditions and the following disclaimer in the documentation and/or other materials provided with the distribution.

3. Neither the name of the University nor the names of its contributors may be used to endorse or promote products derived from this software without specific prior written permission.

- Snowball - Copyright © 2001, Dr Martin Porter, Copyright © 2002, Richard Boulton.
All rights reserved.

Redistribution and use in source and binary forms, with or without modification, are permitted provided that the following conditions are met:

1. Redistributions of source code must retain the above copyright notice, this list of conditions and the following disclaimer.

2. Redistributions in binary form must reproduce the above copyright notice, this list of conditions and the following disclaimer in the documentation and/or other materials provided with the distribution.

3. Neither the name of the copyright holder nor the names of its contributors may be used to endorse or promote products derived from this software without specific prior written permission.

THIS SOFTWARE IS PROVIDED BY THE COPYRIGHT HOLDERS AND CONTRIBUTORS "AS IS" AND ANY EXPRESS OR IMPLIED WARRANTIES, INCLUDING, BUT NOT LIMITED TO, THE IMPLIED WARRANTIES OF MERCHANTABILITY AND FITNESS FOR A PARTICULAR PURPOSE ARE DISCLAIMED. IN NO EVENT SHALL THE COPYRIGHT HOLDER OR CONTRIBUTORS BE LIABLE FOR ANY DIRECT, INDIRECT, INCIDENTAL, SPECIAL, EXEMPLARY, OR CONSEQUENTIAL DAMAGES (INCLUDING, BUT NOT LIMITED TO, PROCUREMENT OF SUBSTITUTE GOODS OR SERVICES; LOSS OF USE, DATA, OR PROFITS; OR BUSINESS INTERRUPTION) HOWEVER CAUSED AND ON ANY THEORY OF LIABILITY, WHETHER IN CONTRACT, STRICT LIABILITY, OR TORT (INCLUDING NEGLIGENCE OR OTHERWISE) ARISING IN ANY WAY OUT OF THE USE OF THIS SOFTWARE, EVEN IF ADVISED OF THE POSSIBILITY OF SUCH DAMAGE.

- Pako - Copyright © 2014–2017 by Vitaly Puzrin and Andrei Tuputcyn.

Permission is hereby granted, free of charge, to any person obtaining a copy of this software and associated documentation files (the "Software"), to deal in the Software without restriction, including without limitation the rights to use, copy, modify, merge, publish, distribute, sublicense, and/or sell copies of the Software, and to permit persons to whom the Software is furnished to do so, subject to the following conditions:

The above copyright notice and this permission notice shall be included in all copies or substantial portions of the Software.

THE SOFTWARE IS PROVIDED “AS IS”, WITHOUT WARRANTY OF ANY KIND, EXPRESS OR IMPLIED, INCLUDING BUT NOT LIMITED TO THE WARRANTIES OF MERCHANTABILITY, FITNESS FOR A PARTICULAR PURPOSE AND NONINFRINGEMENT. IN NO EVENT SHALL THE AUTHORS OR COPYRIGHT HOLDERS BE LIABLE FOR ANY CLAIM, DAMAGES OR OTHER LIABILITY, WHETHER IN AN ACTION OF CONTRACT, TORT OR OTHERWISE, ARISING FROM, OUT OF OR IN CONNECTION WITH THE SOFTWARE OR THE USE OR OTHER DEALINGS IN THE SOFTWARE.

- HDF5 (Hierarchical Data Format 5) Software Library and Utilities Copyright 2006 –2015 by The HDF Group. NCSA HDF5 (Hierarchical Data Format 5) Software Library and Utilities Copyright 1998-2006 by the Board of Trustees of the University of Illinois. All rights reserved. DISCLAIMER: THIS SOFTWARE IS PROVIDED BY THE HDF GROUP AND THE CONTRIBUTORS “AS IS” WITH NO WARRANTY OF ANY KIND, EITHER EXPRESSED OR IMPLIED. In no event shall The HDF Group or the Contributors be liable for any damages suffered by the users arising out of the use of this software, even if advised of the possibility of such damage.
- agl-aglfn technology is Copyright © 2002, 2010, 2015 by Adobe Systems Incorporated. All Rights Reserved.

Redistribution and use in source and binary forms, with or without modification, are permitted provided that the following conditions are met:

- Redistributions of source code must retain the above copyright notice, this list of conditions and the following disclaimer.
- Redistributions in binary form must reproduce the above copyright notice, this list of conditions and the following disclaimer in the documentation and/or other materials provided with the distribution.
- Neither the name of Adobe Systems Incorporated nor the names of its contributors may be used to endorse or promote products derived from this software without specific prior written permission.

THIS SOFTWARE IS PROVIDED BY THE COPYRIGHT HOLDERS AND CONTRIBUTORS “AS IS” AND ANY EXPRESS OR IMPLIED WARRANTIES, INCLUDING, BUT NOT LIMITED TO, THE IMPLIED WARRANTIES OF MERCHANTABILITY AND FITNESS FOR A PARTICULAR PURPOSE ARE DISCLAIMED. IN NO EVENT SHALL THE COPYRIGHT HOLDER OR CONTRIBUTORS BE LIABLE FOR ANY DIRECT, INDIRECT, INCIDENTAL, SPECIAL, EXEMPLARY, OR CONSEQUENTIAL DAMAGES (INCLUDING, BUT NOT LIMITED TO, PROCUREMENT OF SUBSTITUTE GOODS OR SERVICES; LOSS OF USE, DATA, OR PROFITS; OR BUSINESS INTERRUPTION) HOWEVER CAUSED AND ON ANY THEORY OF LIABILITY, WHETHER IN CONTRACT, STRICT LIABILITY, OR TORT (INCLUDING NEGLIGENCE OR OTHERWISE) ARISING IN ANY WAY OUT OF THE USE OF THIS SOFTWARE, EVEN IF ADVISED OF THE POSSIBILITY OF SUCH DAMAGE.

- dmlc/xgboost is Copyright © 2019 SAS Institute.

Licensed under the Apache License, Version 2.0 (the “License”); you may not use this file except in compliance with the License. You may obtain a copy of the License at

<http://www.apache.org/licenses/LICENSE-2.0>

Unless required by applicable law or agreed to in writing, software distributed under the License is distributed on an “AS IS” BASIS, WITHOUT WARRANTIES OR CONDITIONS OF ANY KIND, either express or implied. See the License for the specific language governing permissions and limitations under the License.

- libzip is Copyright (C) 1999-2019 Dieter Baron and Thomas Klausner.

This file is part of libzip, a library to manipulate ZIP archives. The authors can be contacted at <libzip@nih.at>.

Redistribution and use in source and binary forms, with or without modification, are permitted provided that the following conditions are met:

1. Redistributions of source code must retain the above copyright notice, this list of conditions and the following disclaimer.
2. Redistributions in binary form must reproduce the above copyright notice, this list of conditions and the following disclaimer in the documentation and/or other materials provided with the distribution.
3. The names of the authors may not be used to endorse or promote products derived from this software without specific prior written permission.

THIS SOFTWARE IS PROVIDED BY THE AUTHORS “AS IS” AND ANY EXPRESS OR IMPLIED WARRANTIES, INCLUDING, BUT NOT LIMITED TO, THE IMPLIED WARRANTIES OF MERCHANTABILITY AND FITNESS FOR A PARTICULAR PURPOSE ARE DISCLAIMED. IN NO EVENT SHALL THE AUTHORS BE LIABLE FOR ANY DIRECT, INDIRECT, INCIDENTAL, SPECIAL, EXEMPLARY, OR CONSEQUENTIAL DAMAGES (INCLUDING, BUT NOT LIMITED TO, PROCUREMENT OF SUBSTITUTE GOODS OR SERVICES; LOSS OF USE, DATA, OR PROFITS; OR BUSINESS INTERRUPTION) HOWEVER CAUSED AND ON ANY THEORY OF LIABILITY, WHETHER IN CONTRACT, STRICT LIABILITY, OR TORT (INCLUDING NEGLIGENCE OR OTHERWISE) ARISING IN ANY WAY OUT OF THE USE OF THIS SOFTWARE, EVEN IF ADVISED OF THE POSSIBILITY OF SUCH DAMAGE.

